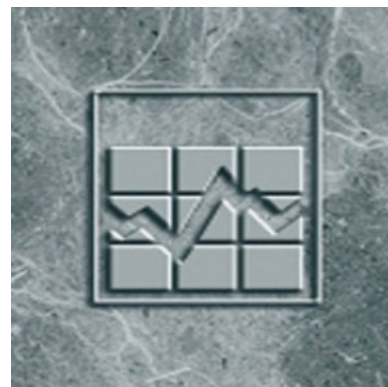## Research Paper

Household expenditures research paper series

Survey of Household Spending 2008: Data Quality Indicators

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Survey of Household Spending 2008: Data Quality Indicators

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Highlights

Sampling errors

➢ The coefficients of variation (CVs) of the estimates of average total expenditure per household for the provinces vary between 1.3% and 2.1%. The CV at the national level is 0.9%.

➢ The coefficients of variation (CVs) of the average estimates for the different summary-level expenditure categories are in most cases less than or equal to 2.7% at the national level and generally less than 5.0% at the provincial level. The results for the dwelling and household equipment characteristics are similar. Since the sample size is smaller in Prince Edward Island, the CVs tend to be higher than those of the other provinces.

Nonresponse

➢ The final response rate is 63.4%. The provincial response rates range from 57.2% in Alberta to 71.3% in Newfoundland and Labrador.

➢ The non-response rate is 36.6%. Nonresponse is due to refusals (22.6%), households that could not be contacted (10.6%), and households whose data were considered unusable (3.4%).

➢ The nonresponse rate tends to increase with the urbanization level. The non-response rate is 30.8% in rural areas and 37.9% in urban centres with a population of 1 million or more.

➢ According to an analysis of nonresponse rates in strata consisting of the high-income geographic areas created under the sample design, the rate in high-income strata (40.4%) is higher than the rate observed in other strata (36.2%). The refusal rate for high-income strata (27.0%) is higher than the rate observed for other strata (22.2%).

Coverage errors

➢ Undercoverage of households is 11.7% at the national level.

➢ Undercoverage of persons is 15.0% at the national level.

Processing errors related to imputation

(i) Expenditure variables

➢ It was necessary to impute at least one expenditure variable[1] for 56.0% of households nationally. Contributing strongly to this rate is the section of the questionnaire dealing with expenditures related to residential communications services (telephone, cell phone and Internet access), cable television services, satellite distribution services and security services. For these services, respondent households have been allowed since 2004 to provide only the total expenditure for a package (bundled services), indicating which services are included. Approximately 43% of usable households required imputation for at least one of these six services.

➢ The overall imputation rate excluding these six services is 25.7% at the national level. For these expenditures, most households required imputation of one or two of the 246 expenditure variables. At the provincial level, imputation rates vary from a low of 17.6% for Newfoundland and Labrador to a high of 32.5% for Quebec.

➢ Imputed values account for 41.8% of the estimate of total expenditure on cable television services and 49.0% of the estimate of total expenditure on Internet access services. Imputation rates for households reporting expenditures on one or the other of these two services are respectively 42.1% and 58.7%. These higher rates are mainly due to the fact that among households that reported paying for a package, a large proportion of packages included these two services.

➢ About 10% of individuals required imputation of clothing variables, but in the vast majority of such cases, the respondents provided the totals and only the components were imputed.

➢ About 13% of individuals aged 15 and over required imputation of at least one variable in the Personal Taxes, Security and Money Gifts section.

(ii) Income variables

➢ Imputation of at least one income variable was required for 7.0% of individuals aged 15 and over.

(iii) Categorical variables

➢ For 9.9% of households, it was necessary to impute at least one categorical variable. Approximately 78% of them had only one variable imputed. Provincially, imputation rates range from a low of 7.3% for Quebec to a high of 14.6% for Manitoba. Categorical variables that required imputation can be found in the following sections of the questionnaire: Dwelling Characteristics; Facilities Associated with the Dwelling; Tenure; and Tobacco and Miscellaneous, for variables pertaining to purchases through direct sales.

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1. Excluding the Clothing section and the Personal Taxes, Security and Money Gifts section, which are collected at the individual level and not the household level.

Introduction

The Survey of Household Spending (SHS) is an annual survey that collects data on household income and expenditure using personal interviews. The 2008 SHS sample consists of 15,445 households[2] distributed throughout the ten provinces. This is a decrease of about 25% in the usual SHS sample size, and as a result, the estimates will have higher coefficients of variation. Collection takes place in January, February and March, and income and spending figures are obtained for the period from January 1 to December 31 of the previous year. Following a redesign that took place in 1997, this survey replaces the periodic Family Expenditure Survey and the Household Facilities and Equipment Survey (with modifications to questionnaires and samples).

Like all surveys, the SHS is subject to errors, despite all the precautions taken at the different stages of the survey to control them. While there is no comprehensive measure of the quality of the data generated by a survey, some quality measures produced at the different stages of the survey can provide users with the information needed in order to interpret the data properly.

This report describes the quality indicators produced for the 2008 Survey of Household Spending. It covers the usual quality indicators that generally help users interpret data, such as coefficients of variation, response and nonresponse rates, slippage rates and imputation rates.

Quality indicators have been classified according to the main types of error encountered in a survey. Section 1 deals with sampling errors—that is, errors due to the fact that the inferences about the population drawn from the survey are based on information collected from a sample of the population, rather than the entire population. The remaining sections cover non-sampling errors. Nonresponse and coverage errors are discussed in sections 2 and 3. Response errors and processing errors are dealt with in sections 4 and 5 respectively.

This report focuses on data quality. For a detailed description of the survey's methodology, see reference [1].

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2. The initial sample is made up of 18,609 dwellings. From these dwellings, it is necessary to identify and exclude ineligible dwellings (see section 2.1) to obtain the 15,445 households from which data on expenditure and income are collected.

1. Sampling errors

Sampling errors occur when inferences made about a population following a survey are based on information collected from a sample of the population, and not from the entire population. In addition to the sample design and the estimation method used in the Survey of Household Spending, the sample size and the variability of each characteristic are factors that determine sampling error. Characteristics that are rare or are distributed very unevenly in the population will have greater sampling error than characteristics that are observed more frequently or are more homogeneous in the population.

1.1 Measures of sampling error

The standard error is a commonly used measure of sampling error. The standard error is the degree of variation of the estimate due to the fact that one sample was selected rather than another, among all possible samples of the same size under the same sample design. Since the SHS uses a complex sample design and estimation method, the standard error is estimated using a resampling method known as the bootstrap technique. Prior to the 2003 reference year, the jackknife resampling method was used to produce standard error estimates for the SHS. The decision was made to use the bootstrap method for the 2003 and subsequent reference years, mainly because Income Statistics Division (ISD) was planning to publish median expenditure estimates and needed the coefficients of variation of those estimates. The bootstrap method is suitable for variance estimation of non-smooth statistics such as quantiles. For more details on this method, see reference [2].

The coefficient of variation (CV) is also a frequently used measure of the reliability of an estimate. It simply expresses the standard error as a percentage of the estimate. Thus, if an estimate Y is obtained for a certain characteristic and SE is the estimated standard error, then the CV will be (SE/Y) x 100.

Finally, either the standard error or the coefficient of variation may be used to derive another measure of the precision of estimates, namely the confidence interval. This measure indicates the level of confidence one can have that the true value of an observed characteristic for the population lies within the interval. An interval with a confidence level of 95% corresponds to the estimate obtained from the sample ± 2 standard errors: (Y ± 2 SE).[3] This means that if the sampling were repeated a large number of times, each sample would provide a different interval and 95% of the intervals would contain the true value of the characteristic. Similarly, if the sampling were repeated, the interval Y ± SE would contain the true value in 68% of cases.

1.2 Coefficients of variation

Estimated coefficients of variation are calculated for estimates of many characteristics collected in the SHS. The CVs of detailed average household expenditure and the CVs of dwelling characteristics and household facilities and equipment are available on request for the national and provincial levels (Income Statistics Division, 1-888-297-7355: income@statcan.gc.ca)

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3. The confidence interval is calculated directly from the CV in similar fashion, namely Y ± 2 (CV x Y)/100.

It should be noted that the estimated CVs do not take account of the fact that some of the data were imputed and thus may underestimate the true CVs. For most variables, the imputation rates are low (see section 5) and the provided CVs represent good estimates of the true CVs. However, to assess the reliability of detailed expenditures with a high imputation rate, the CV and the imputation rate should both be considered.

Table 1.1 presents an overview of the CVs of estimates of household averages for a few of the summary-level expenditure categories and for income at the provincial and national levels.

Table 1.1
Coefficients of variation for the estimate of average household expenditures for several summary-level expenditure categories and for the estimate of average income, Canada and the provinces

	Can.	N.L.	P.E.I.	N.S.	N.B.	Que.	Ont.	Man.	Sask.	Alta.	B.C.
						percent					
Total expenditure	0.9	1.3	2.1	1.7	1.6	1.6	1.8	1.7	1.6	1.5	1.5
Total current consumption	0.8	1.5	2.0	1.8	1.6	1.4	1.6	1.6	1.7	1.7	1.6
Food	0.7	1.5	2.0	1.5	1.6	1.4	1.4	1.8	1.8	1.7	1.6
Shelter	1.3	2.7	2.7	2.2	2.7	2.3	2.5	2.2	2.4	2.6	2.3
Household operation	1.3	2.1	2.8	2.4	2.1	2.3	2.5	2.5	2.3	3.3	2.4
Furnishings	2.2	3.8	4.8	4.9	4.9	3.6	4.3	5.4	5.6	5.3	5.1
Clothing	1.6	3.0	4.0	3.4	3.3	2.6	3.0	3.6	3.8	3.4	4.2
Transportation	1.6	3.8	4.7	3.5	4.1	3.0	3.4	4.1	4.2	4.4	3.3
Health care	2.7	3.5	3.9	3.4	4.3	3.2	7.2	3.7	3.7	3.7	3.7
Personal care	1.2	2.5	3.5	3.1	2.8	2.5	2.3	2.8	2.8	3.1	2.7
Recreation	1.8	3.6	6.3	4.1	4.4	4.0	2.9	4.5	4.2	5.9	4.5
Reading & printed material	2.1	6.8	6.7	4.2	4.5	4.4	4.1	4.4	4.3	5.0	5.3
Education	4.6	13.2	13.4	9.9	12.8	6.6	7.8	7.7	10.6	13.8	8.1
Tobacco, alcoholic beverages	2.1	4.5	6.7	4.8	5.3	3.8	4.3	4.7	5.0	6.2	4.5
Games of chance (net)	4.9	7.5	10.3	10.0	15.2	12.7	9.0	7.0	9.2	10.6	13.3
Miscellaneous expenditures	3.3	4.1	7.1	9.2	8.7	7.7	6.5	6.1	7.1	8.1	5.0
Personal income tax	2.1	2.4	4.4	3.3	3.8	3.2	4.4	3.5	3.1	3.0	3.6
Personal insurance and pension contributions	1.3	2.6	4.2	2.6	2.9	2.3	2.7	3.6	3.2	2.2	2.6
Gifts and contributions	4.6	7.9	18.0	10.4	8.0	18.2	5.8	11.6	12.3	13.1	7.4
Income	0.9	1.1	2.1	1.7	1.5	1.4	2.0	1.6	1.3	1.4	1.6

Table 1.2 presents an overview of the CVs for some dwelling characteristics and household equipment estimates at the provincial and national levels.

Table 1.2
Coefficients of variation for some dwelling characteristics and household equipment, Canada and the provinces

	Can.	N.L.	P.E.I.	N.S.	N.B.	Que.	Ont.	Man.	Sask.	Alta.	B.C.
						percent					
Owner	1.0	1.9	3.0	2.1	2.0	2.2	2.1	2.6	2.4	2.3	2.0
Renter	2.0	6.4	7.3	5.3	7.1	2.9	4.4	5.6	6.0	6.2	4.3
Washing machine	0.7	1.1	1.8	1.9	1.3	1.3	1.6	2.0	1.9	1.9	1.7
Clothes dryer	0.8	1.2	1.8	2.1	1.5	1.4	1.7	1.8	1.8	2.0	1.6
Dishwasher	1.3	3.6	4.2	3.8	3.9	2.6	2.8	3.6	2.6	2.4	2.2
Freezer	1.2	2.9	2.9	2.5	2.2	2.7	2.6	2.0	2.1	2.4	2.9
Microwave oven	0.4	0.6	0.8	0.6	0.7	0.7	0.8	0.9	0.8	0.6	1.0
Cellular phone	0.8	2.2	2.6	2.2	2.4	2.0	1.6	1.9	1.8	1.6	1.5
CD player	0.7	1.8	2.2	1.8	1.9	1.4	1.4	1.7	1.7	1.7	1.5
Cable TV	1.3	3.2	4.1	2.8	3.8	2.7	2.6	2.4	4.0	3.4	2.2
Satellite dish	3.0	6.3	5.6	6.7	4.9	6.0	5.9	5.7	6.1	7.5	8.2
DVD player	0.6	1.6	1.8	1.4	1.7	1.2	1.2	1.3	1.3	1.4	1.1
Home computer	0.7	2.3	2.5	1.9	2.0	1.5	1.4	1.8	1.7	1.5	1.3
Regular tel. connection to a computer (modem)	6.1	19.8	21.2	13.0	15.4	9.0	12.0	15.3	15.1	17.2	22.9
High-speed tel. connection to a computer	2.5	5.1	6.0	5.9	4.3	6.5	5.1	5.1	4.4	6.7	5.3
Cable connection to a computer	2.2	8.6	17.8	6.1	11.4	4.8	4.1	7.5	10.2	5.2	4.1
Use of Internet (home)	0.8	2.4	2.6	2.0	2.3	1.9	1.5	2.1	1.9	1.7	1.5
Owned vehicles (one)	1.7	4.4	5.9	4.0	4.6	3.3	3.3	4.0	4.8	4.9	3.7
Owned vehicles (2 or more)	1.8	4.8	5.4	4.7	4.6	4.5	3.8	3.6	3.3	3.7	2.9

1.3 Model for deriving an approximation of the CV

Estimates for different domains of interest (for example, by income quintile) for the summary-level expenditure categories are available in the publication *Spending Patterns in Canada* (see reference [4]). Estimates for different domains of interest for detailed expenditure categories are available on request from Income Statistics Division. (For more details on tables available on request from Income Statistics Division, see reference [3] or [4].) For operational reasons, it is impossible to produce CVs for all levels of aggregation that may interest users.

1.3.1 Approximation of the CV for domain estimates

It is, however, possible to calculate an approximation of the CV by using a relationship between the number of households in the sample that reported expenditures for a given category and the CV for an aggregate level. This relationship, based on the CV's tendency to increase in proportion to a decrease in the square root of the number of households reporting an expenditure, is illustrated below.

Formula for approximating the CV for a domain (a subgroup of the population)

If CV(Y) represents the CV for the estimate of the household average of a certain characteristic for the entire population, then an approximation of the CV of the estimate of that characteristic can be calculated for a domain (which can be regarded as a subgroup of the population, such as a household type, an income quintile or an urbanization level) with the following equation:

$$CV \ (Y_d) \ = \ CV \ (Y) \times \sqrt{\frac{nP}{n_d \, P_d}}$$

where

n: number of households in the sample
P: estimate of the proportion of households reporting a value > 0 for this characteristic in the population
n_d: number of households in the sample in domain d
P_d: estimate of the proportion of households reporting a value > 0 for this characteristic in domain d

Generally, approximations for the different domains are calculated using the CV, size n and proportion P at the national level. If an approximation of a CV is desired for a domain that is entirely contained within a single province (for example, a metropolitan area), then it is preferable to use these values at the provincial level, since provincial CVs are published for the 2008 SHS (reference [3]). It should be noted that a CV obtained using this approach is only an approximation of the actual value.

1.3.2 Approximation of the CV using the microdata file

Microdata file users can obtain an approximation of the CV of the estimates using another method that will generally provide better results than the method described in the previous section for the CVs of detailed expenditure categories. This approach is described in detail in the documentation provided with the 2008 microdata file. This method of approximation can be used only with the microdata file, since it requires having data and weights for each household.

The 1997 SHS data quality report contains the results of the performance evaluation of these two CV approximation methods.

1.4 Suppression of unreliable data in estimate tables

Since the coefficient of variation is an indicator of data reliability, we would like to use it to determine whether or not the estimates should be published. Estimates whose estimated CV is greater than 33% are not considered sufficiently reliable to be published. However, estimated CV are not calculated for many of the published estimates. The suppression rule for expenditure estimates is therefore based on the number of households reporting a value greater than zero.[4]

It can be shown that CVs are usually less than 33% when the number of households reporting an expenditure is greater than 30. Since this is an approximate rule, some estimates may be published even though the CV is greater than 33%, and some estimates will not be published even though the CV is less than 33%. The 1997 SHS data quality report contains the results of the evaluation of the risk of error due to the suppression rule.

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4. In practice, we use the estimate of the proportion of households reporting an expenditure, which is multiplied by the sample size.

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2.   Nonresponse

Errors due to nonresponse occur when some potential respondents do not provide the necessary information or the provided information proves to be unusable. When the respondent has failed to respond to only some questions, this is referred to as partial nonresponse. In such a case, the missing data are imputed. Errors associated with imputation are described in section 5, which deals with processing errors. In the present section, nonresponse includes collection nonresponse, which is mainly due to inability to contact the household or to refusal by household members to participate in all or part of the survey, as well as data that are collected from households but prove to be unusable.

The main impact of nonresponse on data quality is that it can introduce a bias in the estimates if the characteristics of respondents and nonrespondents differ and the difference has an impact on the characteristics studied. Nonresponse rates are easy to calculate, but they provide only an indication of data quality, since they do not measure the bias associated with the estimates. The magnitude of nonresponse can be considered an indicator of the risks of bias in the estimates.

2.1   Response, nonresponse and vacancy rates

Since the units selected in the SHS are dwellings, interviewers must first identify ineligible dwellings, that is, dwellings occupied by persons who are not part of the target population, as well as dwellings that no longer exist (demolished, mobile home moved or dwelling converted to business) and vacant dwellings (unoccupied, seasonal or under construction).

Then, for eligible dwellings, the proportion of households that did not respond to the survey is evaluated. This is called the collection nonresponse rate. Included are households that refused to participate in the survey and households for which respondents could not be contacted, either because they were absent or because of special circumstances (language problem, illness, death).

Also for eligible dwellings, the rate of unusable data is determined. Unusable data refers to the number of households whose questionnaires were at least partially completed but were rejected during data processing. When many income or expenditure questions have been left unanswered, the questionnaire is classified as incomplete and is not used.

Note that all rates provided in this section are unweighted. For the 2008 Survey of Household Spending, the final response rate is 63.4%. Table 2.1-1 shows the final response rate as well as the sample size (eligible households) broken down by refusals, units not contacted, unusable data and usable data. This rate is provided for the national and provincial levels.

Table 2.1-1
Sample size and response rate, Canada and the provinces

	Eligible households	No contact	Refusals	Unusables	Usables	Final response rate (at estimation stage)[1]
	number					percent
Canada	**15,445**	**1,641**	**3,495**	**522**	**9,787**	**63.4**
Newfoundland and Labrador	1,356	98	254	37	967	71.3
Prince Edward Island	786	44	189	17	536	68.2
Nova Scotia	1,466	144	351	61	910	62.1
New Brunswick	1,307	138	262	45	862	66.0
Quebec	1,948	143	505	48	1,252	64.3
Ontario	2,338	283	614	89	1,352	57.8
Manitoba	1,480	222	303	37	918	62.0
Saskatchewan	1,453	156	285	26	986	67.9
Alberta	1,522	197	369	86	870	57.2
British Columbia	1,789	216	363	76	1,134	63.4

1. Usable/eligible households x 100

Table 2.1-2 shows the final nonresponse rate; the collection nonresponse rate, broken down by refusals and units not contacted; and the rate of households with unusable data owing to incomplete questionnaires. The vacancy rate is also included. These rates are provided for the national and provincial levels

Note that the vacancy rates shown in the tables in section 2 include vacant dwellings (unoccupied, seasonal or under construction) and dwellings that no longer exist (demolished, mobile home moved or dwelling converted to business).

Table 2.1-2
Nonresponse and vacancy rates, Canada and the provinces

	Vacancy rate	Collection nonresponse rate			Unusable data rate			Final nonresponse rate (at estimation stage)
		Total	No contact	Refusal	Total	Incomplete	Out-of-balance	
	percent							
Canada	14.8	**33.3**	10.6	22.6	**3.4**	0.7	2.6	36.6
Newfoundland and Labrador	16.2	**26.0**	7.2	18.7	**2.7**	0.7	2.1	28.7
Prince Edward Island	20.2	**29.6**	5.6	24.0	**2.2**	0.0	2.2	31.8
Nova Scotia	18.1	**33.8**	9.8	23.9	**4.2**	1.1	3.1	37.9
New Brunswick	21.4	**30.6**	10.6	20.0	**3.4**	0.6	2.8	34.0
Quebec	14.1	**33.3**	7.3	25.9	**2.5**	0.2	2.3	35.7
Ontario	12.4	**38.4**	12.1	26.3	**3.8**	1.0	2.8	42.2
Manitoba	9.8	**35.5**	15.0	20.5	**2.5**	0.9	1.6	38.0
Saskatchewan	13.1	**30.4**	10.7	19.6	**1.8**	0.6	1.2	32.1
Alberta	11.4	**37.2**	12.9	24.2	**5.7**	1.2	4.5	42.8
British Columbia	14.8	**32.4**	12.1	20.3	**4.2**	0.7	3.5	36.6

2.2 Nonresponse by urbanization level

Nonresponse varies with urbanization level. The various national rates are shown by urbanization level in Table 2.2.[5]

Table 2.2
Nonresponse and vacancy rates by urbanization level

	Vacancy rate	Collection nonresponse rate			Unusable data rate			Final nonresponse rate (at estimation stage)
		Total	No contact	Refusal	Total	Incomplete	Out-of-balance	
					percent			
Urban								
1,000,000 or more	9.4	35.1	10.3	24.9	2.8	0.3	2.4	37.9
500,000 to 999,999	8.0	37.9	14.9	23.0	4.1	0.9	3.2	41.9
250,000 to 499,999	10.2	39.6	13.5	26.1	3.8	1.6	2.2	43.4
100,000 to 49,999	12.2	35.4	11.6	23.8	3.0	0.7	2.3	38.4
30,000 to 99,999	12.9	29.3	7.6	21.7	3.8	0.8	3.0	33.1
Less than 30,000	14.0	30.4	9.4	21.1	2.5	0.6	1.8	32.9
Rural	29.5	26.7	7.9	18.8	4.1	0.8	3.3	30.8
Total	14.8	33.3	10.6	22.6	3.4	0.7	2.6	36.6

2.3 Nonresponse by income stratum

Since income information is not available for nonrespondents, it is impossible to compare income-specific nonresponse rates. However, the LFS sample design, used for the SHS, was constructed in such a way that in census metropolitan areas, there are strata consisting of geographic areas with a high concentration of high-income households. While the number of high-income strata remains relatively small (51 out of a total of 1,060 strata), comparison of response rates in this group in relation to the other strata provides relevant information on the potential effect of nonresponse (see Table 2.3).

Note that in addition to regular strata, the "Other" strata category includes the following four types of strata: strata with a high vacancy rate, high-cost strata, strata with a concentration of immigrants, and strata with a concentration of Aboriginals. Since the portion of the SHS sample allocated to the latter four strata was smaller, the results for them are not broken out in Table 2.3.

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5. Tables on nonresponse rates by urbanization level and province are available on request from Household Survey Methods Division.

Table 2.3
Nonresponse and vacancy rates in high-income strata and other strata

	Vacancy rate	Collection nonresponse rate			Unusable data rate			Final nonresponse rate (at estimation stage)
		Total	No contact	Refusal	**Total**	Incomplete	Out-of-balance	
				percent				
High-income	7.8	**37.8**	10.7	27.0	**2.6**	0.6	2.1	**40.4**
Others	15.4	**32.8**	10.6	22.2	**3.5**	0.8	2.7	**36.2**
Total	14.8	**33.3**	10.6	22.6	**3.4**	0.7	2.6	**36.6**

2.4 Adjustment for nonresponse

To compensate for nonresponse, the weights in the SHS are inflated by the inverse of the weighted response rate within certain predefined groups. Following the overhaul of the LFS sample design, the nonresponse adjustment groups were redefined. As was the case for previous years, these groups are defined on the basis of the different urbanization levels in each province and of subprovincial geographic areas for Quebec, Ontario and British Columbia. Also, specific nonresponse adjustment groups were created for high-income strata. As was seen in the previous section, the number of high-income strata increased with the introduction of the new LFS sample design. Thus, it is now possible to form nonresponse adjustment groups for high-income strata in all provinces except Prince Edward Island. Such a group cannot be formed in Prince Edward Island because that province has no high-income strata.

The weighted rates differ from the rates presented in this section, since the former take the sampling weight of each household into account. An algebraic description of the nonresponse adjustment is provided in Appendix A.

The adjustment of weights for nonresponse takes account of differences in nonresponse by urbanization level (as illustrated in section 2.2) and geographic area or by groups of high-income strata. It will reduce the bias insofar as the characteristics of respondents and non-respondents are similar for a given urbanization level and geographic area or for a given group of high-income strata.

It should be noted that a nonresponse adjustment group can be combined with another group if the number of households in the group is too small or the adjustment factor is too high.

3. Coverage errors

The target population was defined in the design of the survey. It is useful to go over this definition, since a good understanding of the target population is necessary in order to properly interpret the survey data. It is important to note that the SHS uses the sampling frame of the Labour Force Survey (LFS).

Target population

> The target population consists of individuals living in private households. It therefore excludes residents of institutions such as prisons, chronic care hospitals or senior citizens' homes, as well as members of religious orders and other groups living communally, members of the Armed Forces living in military compounds, and individuals residing permanently in hotels or rooming houses. Also excluded are foreign countries' official representatives residing in Canada and their families as well as individuals residing on Indian reserves or public lands (except for the territories). With these exclusions, the survey covers nearly 98% of the population in the ten provinces. The territories are excluded from the target population for the 2008 SHS, as they are covered only once every two years.

We did not collect data from persons temporarily living away from their families (for example, students at university) because the information would be obtained from their families if selected in the sample.

Coverage errors result from inadequate representation of the target population based on the units in the sampling frame. Some units of the target population may be omitted from the sampling frame, in which case there is undercoverage. Other units that are not in the target population may be included by error, or some units may be included more than once. These units are responsible for overcoverage.

3.1   Undercoverage and overcoverage: slippage rates

In the SHS, the sample is selected using a list of dwellings in each selected cluster. Factors contributing to undercoverage are: the omission of dwellings in the creation of the list, new dwellings that are added between the creation of the list and the interviewer's visit (mainly in developing areas), and the erroneous classification of vacant dwellings. The inclusion of dwellings that are not within the boundaries of the cluster is a source of overcoverage. Similarly, errors can occur during data collection, due to improper identification of persons as members of the selected household. These errors also contribute to undercoverage or overcoverage.

Also, as described in section 2.4, reweighting methods are implemented to take account of nonresponse. However, when these adjustments are made, it is impossible to correct the survey weights to ensure that all subgroups within the population are well represented.

A good representation of the target population is essential to the production of realistic expenditure estimates. The sample must adequately represent the individuals in the target population as well as the distribution of households according to size.

There is generally net undercoverage of the number of persons and the number of households in the SHS. This undercoverage is corrected by an adjustment of weights using auxiliary or reference data based on postcensal demographic estimates. The slippage rate (see Appendix A) is a measure of the percentage of difference between the

estimates from these auxiliary data and the survey estimates calculated using weights not adjusted with these data.[6] Slippage therefore represents the combined effect of undercoverage and imbalance in certain population subgroups created by survey nonresponse that could not be corrected at the reweighting stage.

Slippage rates by age group for the national and provincial levels are shown in Table 3.1, while slippage rates by household size, used in adjusting weights, are shown in Table 3.2. A positive rate indicates overcoverage of the number of persons or households in the survey.

Table 3.1
Slippage rates by age group, Canada and the provinces

	0 to 6	7 to 17	18 to 24	25 to 34	35 to 44	45 to 54	55 to 64	65 and older	Total
	percent								
Newfoundland and Labrador	-5.5	-0.2	-40.2	-31.4	-15.3	-13.1	-8.1	-7.7	**-14.4**
Prince Edward Island	9.1	-10.4	-32.8	-21.5	-12.1	-13.5	-3.0	0.6	**-10.8**
Nova Scotia	-0.3	4.6	-32.6	-31.2	-10.4	-4.7	-11.8	-3.7	**-10.7**
New Brunswick	3.5	-10.6	-29.9	-18.9	-15.8	-21.2	-22.1	3.3	**-14.6**
Quebec	-14.2	-6.7	-3.7	-21.0	-10.7	-1.5	-13.3	-0.7	**-8.7**
Ontario	0.8	-18.5	-24.9	-26.0	-15.5	-15.6	-13.5	-15.5	**-16.8**
Manitoba	-16.6	-12.2	-24.1	-26.4	-10.2	-20.0	-6.7	-8.0	**-15.4**
Saskatchewan	-2.7	0.7	-24.9	-16.7	-2.6	-23.3	5.9	5.1	**-7.5**
Alberta	-21.6	-19.9	-26.4	-31.5	-21.4	-13.0	-16.1	-6.1	**-20.0**
British Columbia	-10.0	-8.4	-36.2	-27.1	-21.2	-20.4	-19.7	-10.0	**-19.2**
Canada	**-7.6**	**-12.9**	**-22.6**	**-25.5**	**-15.2**	**-12.7**	**-13.8**	**-8.3**	**-15.0**

Table 3.2
Slippage rates by household size, Canada and the provinces

	All households	One-person households	Two-person households	Households of three or more persons
	percent			
Canada	**-11.7**	**-9.7**	**-9.8**	**-14.9**
Newfoundland and Labrador	**-9.4**	-2.3	-5.0	-18.5
Prince Edward Island	**-8.5**	-12.6	-4.5	-9.8
Nova Scotia	**-11.3**	-14.9	-13.1	-6.1
New Brunswick	**-12.1**	-11.1	-10.6	-14.5
Quebec	**-8.1**	-11.5	-5.0	-8.0
Ontario	**-11.4**	-1.4	-12.7	-16.5
Manitoba	**-12.0**	-7.8	-6.6	-20.7
Saskatchewan	**-6.0**	-10.8	-2.6	-5.4
Alberta	**-15.6**	-8.9	-14.5	-20.6
British Columbia	**-17.6**	-25.2	-9.7	-18.5

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6. The subweight, which is the sampling weight adjusted for nonresponse, is used (see Appendix A).

3.2   Adjustment at the population and household levels

To correct the problem of the sample's representativeness, shown in Table 3.1, and to reduce the resulting bias, the survey data are adjusted during weighting using demographic estimates for the age groups defined in this table, for each province. For more details on the adjustment methodology, see references [1] and [5]. This adjustment reduces the bias but does not eliminate it entirely if the characteristics of the individuals omitted from the survey differ from those of individuals included for a given age group in a province.

It should also be noted that the effectiveness of the adjustment based on demographic estimates depends largely on the quality of those estimates and their accuracy in representing the target population of the survey. Demographic estimates are not error-free. They are postcensal estimates based on the population counts from the 2001 Census adjusted for net undercoverage, and they take into account recent statistics on migration, births, deaths, etc. These demographic estimates are adjusted to account for certain exclusions specific to household surveys, such as persons living in institutions. Conceptually, they differ slightly from the SHS target population in that they include persons living in non-institutional collective dwellings, such as members of groups living communally and individuals permanently residing in hotels or rooming houses. However, this difference is considered negligible, since such individuals represent less than 0.4% of the Canadian population.

To remedy the problem of the sample's representativeness in terms of the number of households by size as illustrated in Table 3.2, the survey data are adjusted using auxiliary data. By adjusting the SHS weights to reflect postcensal estimates of the number of households by size, the goal is to compensate for the bias resulting from inadequate representation of households. However, the bias will not necessarily be eliminated if characteristics of households not interviewed (i.e., missed or non-responding households) differ from those of responding households for a given household size. As in the case of demographic estimates of population, the effectiveness of the adjustment will depend on the quality of the auxiliary data on the number of households.

In addition to demographic estimates for age groups by province, two other auxiliary data sets are used during weighting to adjust the survey data and thereby improve their representativeness. The first data set is used to control for the number of children and adults in certain major cities. For the second set, counts for major categories of income from wages and salaries are used when adjusting weights to ensure a degree of consistency in income distributions between the SHS and outside sources.

4. Response errors

Response errors represent a lack of accuracy in responses to questions. They can be attributed to different factors, including a questionnaire that requires improvements, misinterpretation of questions by interviewers or respondents, and errors in respondents' statements.

In the SHS, there can be various reasons for errors in respondents' statements. First, there are recall errors that occur when a respondent forgets expenditures made during the period covered by the survey (which corresponds to the calendar year), or when a respondent provides an erroneous value because of the time interval that has elapsed between the time of purchase and the date of the interview. Recall errors are probably the survey's largest source of response error, since the reference period is long (12 months) and a great variety of information is requested.

To reduce the magnitude of this type of error, the respondent is also encouraged to consult various documents (bills, bank statements, etc.) so as to provide more accurate data. To determine expenditures for small items purchased at regular intervals, interviewers generally suggest that respondents estimate the frequency of the purchases and the price generally paid in order to derive expenditures for a 12-month period.

A second source of error in respondents' reporting is telescopic error, which consists of including in the reference period events that occurred before or after it. In the SHS, the use of the calendar year is considered to provide a good marker for the start of the reference period. Furthermore, since the reference period is a long one, telescopic error has less impact.

Responses by proxy can also contribute to response error. The household member who made an expenditure is generally best able to report it accurately. This is definitely the case with, say, personal purchases. Expenditures reported by an intermediary are more likely to be tainted by response error, and this type of error tends to have a greater effect on certain types of expenditures.

Among other sources of response error, the extent of the respondent's co-operation should not be overlooked. For personal reasons, the respondent may choose not to mention particular expenditures or decide to twist the facts.

In the SHS, another factor is the response burden, owing to the length of the interview and the great variety of items to be reported, as well as the pace of the interview. This can lead to respondent fatigue and affect the quality of the responses obtained. The interview time varies greatly from one household to another, depending on household size, income and various other characteristics.

While response errors are a major source of error in a recall interview, they are the aspect of data quality that is the hardest to measure. Generally, it is necessary to conduct quite costly special studies in an attempt to measure them. Efforts are made to combat response errors by using survey techniques designed to reduce them.

5. Processing errors

Errors can arise in all types of data handling. The main stages of data processing are response coding, data entry, editing, imputation of partial nonresponse and weighting. In the SHS, different procedures are applied at each stage in order to minimize processing errors, and the survey estimates are compared with other data sources prior to release. Errors related to the adjustments made at the weighting stage are described in sections 2 and 3. The other types of processing errors are covered in this section.

Because of the shift to a computer-assisted collection method in 2006, data processing and quality control procedures were altered. Automated edits incorporated into the questionnaire replaced the previously conducted balance edit checks and edits in regional offices. For the 2008 SHS, interviewers entered responses on a portable computer and conducted an initial edit simultaneously. Thus, interval controls, which showed minimums and maximums for certain purchases, were applied if the interviewer entered an unusual amount. Other edits targeted inconsistent responses, such as a household renting its dwelling but paying no rent.

The processing of SHS data also involves imputation for partial nonresponse. Partial nonresponse occurs when the respondent refuses to answer or does not know the answer to certain questions. The imputation approach differs depending on whether the data are categorical or continuous. Categorical data take on only specific values (as in yes/no questions or type of dwelling questions), while continuous data can take any numerical value (such as for income and expenditure data).

Income and expenditure data are imputed using the nearest neighbour technique. The imputation is done on one group of variables at a time, with the groups formed on the basis of the relationships among the variables. A group generally corresponds to a section of the questionnaire. For each group, the missing values for a recipient (a household that has some missing data for at least one of these variables) are imputed from data on the most similar record among all donors (households that have no missing values for these variables). For each recipient, the closest donor is chosen as the one that minimizes a particular distance function. This function is based on matching variables that are chosen because they are correlated with the variables to be imputed. For example, the total income of a household is chosen as a matching variable for all sections pertaining to expenditures. It must also be ensured that, after receiving the donor values, the recipient household satisfies certain consistency rules. In general, the imputation is done at the household level, but in some groups (e.g., income and clothing expenditures), the imputation is done at the person level since the original data are collected at that level for these variables.

Note that since 2001, the imputation of all expenditure and income data has been done using the Canadian Census Edit and Imputation System (CANCEIS) of Statistics Canada. This new system is based on methodology that is slightly different from that in the system used previously. The new system allows a better use of categorical variables as matching fields when selecting a donor. Moreover, this system lends itself to the imputation of both continuous and categorical data. The new system was tested prior to its implementation and the results were similar to those provided by the old system. As of 2003, categorical data, which are found mainly in the dwelling characteristics and facilities sections of the questionnaire, have been imputed with the CANCEIS system.

The categorical data were previously imputed with the help of a "hot deck" imputation technique that randomly chooses a donor from a group of respondent households with similar characteristics.

The bias caused by imputation of partial nonresponse is difficult to evaluate. It depends on the differences between respondents and nonrespondents as well as the ability of the imputation method to produce unbiased estimates. However, the imputation rates give an indication of the importance of partial nonresponse. They are presented in the following section.

5.1 Proportion of households or individuals requiring imputation, at the national and provincial levels

A first indication of the magnitude of partial nonresponse is the proportion of households requiring imputation and the number of variables imputed per household. The questionnaire can be divided into two major groups of variables: those collected at the household level and those collected at the individual level (such as income and clothing expenditure). For the latter, it is important to note that the respondent may provide only the total income or total clothing expenditures if he/she is unable to provide the breakdowns by source of income or type of expenditure. This increases the level of imputation for the components of income and clothing expenditure, but it does not affect total income, total clothing expenditure or total expenditure.

The percentage of households requiring imputation for household expenditure (excluding clothing expenditures and expenditures in the section on Personal Taxes, Security and Money Gifts) is presented in the next subsection. The subsection after that presents the percentage of persons requiring imputation for a clothing expenditure variable, the percentage of persons requiring imputation for an income variable and the percentage of persons requiring imputation for a variable in the section on Personal Taxes, Security and Money Gifts. The last subsection presents the results for the percentage of households requiring imputation of at least one categorical variable. After data imputation by the system, some corrections might have been needed on both imputed and non-imputed variables to ensure data consistency. In reality, these changes constitute only a very small percentage. The results are provided for the national and provincial levels. This gives an indication of which provinces are most affected by imputation.

5.1.1 Household expenditure imputation for the provinces

Table 5.1-1 shows the percentage of usable households requiring imputation of at least one expenditure variable. Usable households are all households living in eligible dwellings, excluding households that could not be contacted, refused to participate in the survey or provided incomplete data (see definitions in section 2.1). The table is broken down by the number of imputed variables (out of 246) for a household.

Note that regular mortgage payments and mortgage insurance premiums are included under shelter costs and thus under total expenditure. Starting with 2002, these two variables were added to the calculation of imputation rates shown in Table 5.1-1. The impact of this change is a higher overall imputation rate.

Starting in 2004, a change was made to the questionnaire regarding expenditures on residential communication services (telephone, cell phone and Internet access), cable television services, satellite distribution services and security systems. Because of the growing use of packages (bundled services), a household may be billed for combined services and may therefore be unable to provide expenditures for individual services. In such a case, the respondent household may provide only the total expenditure for these services while indicating which services are included in the package. Expenditures for individual services are then imputed in two stages. First, we impute households for which only a few services are missing, followed by households for which only the total expenditure for the package is available. For the latter households, the imputed expenditures for services (those included in the package) are adjusted proportionally so that their sum corresponds to the total expenditure on the package as provided by the respondent household. Since this change has had a major impact on the overall imputation rate for expenditures, the imputation rates in Table 5.1-1 are shown separately with and without the costs of residential communications services, rental of cable television services, rental of satellite distribution services and rental of security services. Also, since this change has had an impact on the level of imputation of expenditures for these six services, Table 5.1-2 is provided, showing the imputation rate and a measure of the impact of imputation for each of these services.

Table 5.1-1
Households requiring expenditure imputation, Canada and the provinces

	Imputed variables out of 246 excluding expenditures related to communications services in the home and rental services for cable television, satellite distribution and security systems				Imputed variables out of 252 including expenditures related to communications services in the home and rental services for cable television, satellite distribution and security systems			
	1	2	3 or more	Total	1	2	3 or more	Total
	percent							
Canada	**14.0**	**6.1**	**5.6**	**25.7**	**7.7**	**18.8**	**29.5**	**56.0**
Newfoundland and Labrador	10.3	4.7	2.6	**17.6**	4.6	19.8	25.0	**49.3**
Prince Edward Island	13.2	4.9	5.4	**23.5**	5.6	17.0	40.5	**63.1**
Nova Scotia	16.6	5.7	7.0	**29.3**	6.2	18.6	43.3	**68.0**
New Brunswick	16.6	8.0	6.8	**31.4**	10.2	18.1	26.8	**55.1**
Quebec	18.5	6.8	7.3	**32.5**	10.1	15.1	34.0	**59.2**
Ontario	10.4	5.4	4.7	**20.4**	7.2	12.4	27.1	**46.7**
Manitoba	16.0	5.9	7.5	**29.4**	10.6	13.3	27.0	**50.9**
Saskatchewan	14.6	7.7	3.1	**25.5**	6.6	21.8	31.1	**59.5**
Alberta	10.1	5.2	6.4	**21.7**	6.8	23.0	22.6	**52.4**
British Columbia	14.0	6.0	5.0	**25.0**	7.8	30.2	22.7	**60.6**

Note: Includes regular mortgage payments and mortgage insurance premiums. Excludes clothing expenditures and expenditures in the section on Personal Taxes, Security and Money gifts.

Since there is a higher level of imputation for expenditures related to residential communications services, cable television services, satellite distribution services and security systems, it is important to measure the effect of imputation on the estimates of totals for these six variables. This measure, along with the imputation rate, can be used to see how the amount of imputation done for these variables changes over time. Owing

to the growing popularity of packages (bundled services), the imputation level should increase over time. To measure the impact of imputation, the weighted total of the imputed data is divided by the total estimate (sum of weighted values). This measure represents the proportion of the total value of the estimate that is obtained from imputed data.

Table 5.1-2
Impact of imputation of residential communications services, cable television services, satellite distribution services and security systems at the national level

	Impact of imputation	Reporting households requiring imputation of expenditure for residential communications services, cable television services, satellite distribution services or security systems
	percent	
Telephone services	29.6	36.4
Cell phone, pager and handheld text messaging services	11.7	13.7
Cable television services	41.8	42.1
Satellite distribution services	10.5	10.4
Internet access services	49.0	58.7
Security systems	3.5	6.9

5.1.2   Individual expenditure and income imputation for the provinces

Since some respondents provide only totals for clothing expenditure and income variables, a two-step procedure is used to impute these variables (at the individual level). Individuals who require imputation of only certain components are imputed first, followed by those for whom only totals are available but imputation on all components is required. See reference [1] for a more detailed description of this process.

The percentage of usable individuals (members of usable households) requiring imputation of at least one income variable is presented for the provinces in Table 5.2. The table shows the percentage of persons who had exactly one variable imputed, the percentage who had two or more variables (but not all) imputed and the percentage of persons for whom only total income was available (and hence required having all their components imputed). The total percentage of persons requiring some form of income imputation is also provided. The second to last column of Table 5.2 indicates the total percentage of persons requiring some form of imputation for clothing expenditure variables. The last column of Table 5.2 indicates the total percentage of persons requiring some form of imputation for the Personal Taxes, Security and Money Gifts section of the questionnaire.

Note that only household members aged 15 or over on December 31 of the reference year must answer the questions relating to personal income, personal taxes, security

and money gifts. Thus, since the 2003 reference year, the percentage of persons requiring some form of imputation for income variables as well as for the Personal Taxes, Security and Money Gifts section has been calculated using only persons aged 15 or over, rather than all persons as in previous years. This modification resulted in a slightly higher imputation rate for those variables. As in previous years, the percentage of persons requiring imputation for clothing expenditure variables is based on all persons, since those expenditure questions are asked for each household member.

Table 5.2
Persons requiring income imputation, persons requiring clothing expenditure imputation and persons requiring imputation of variables in the Personal Taxes, Security and Money Gifts section, Canada and the provinces

	Persons requiring imputation of income variables				Persons requiring imputation of at least one of the 15 clothing expenditure variables	Persons requiring imputation of at least one of the 15 variables in the Personal Taxes, Security and Money Gifts section
	1 income variable imputed	2 or more income variables imputed (not all)	All income variables imputed (total income known)	Total (any form of income imputation)		
	percent					
Canada	4.2	1.5	0.9	**7.0**	10.2	13.4
Newfoundland and Labrador	1.6	0.9	1.0	**3.8**	3.9	8.6
Prince Edward Island	3.1	3.0	0.1	**6.4**	8.5	13.8
Nova Scotia	4.1	1.9	0.8	**7.2**	8.6	16.4
New Brunswick	4.7	2.2	1.4	**8.4**	12.0	15.9
Quebec	5.0	1.8	0.8	**7.7**	12.8	14.2
Ontario	4.9	1.5	0.8	**7.3**	4.0	15.7
Manitoba	3.8	1.3	0.5	**5.6**	15.1	10.7
Saskatchewan	4.1	0.9	0.9	**6.2**	4.9	9.6
Alberta	3.4	0.8	1.4	**6.9**	12.5	10.0
British Columbia	6.1	1.5	1.4	**9.5**	19.9	17.3

5.1.3   Imputation of categorical variables for the provinces

Table 5.3 shows the percentage of usable households requiring imputation of at least one categorical variable. The table is broken down by the number of imputed variables (out of 41) for a household. Categorical variables that required imputation can be found in the following sections of the questionnaire: Dwelling Characteristics (with the exception of the dwelling type variable); Facilities Associated with the Dwelling; Tenure (with the exception of variables related to a tenure change during the reference year); Tobacco and Miscellaneous for variables pertaining to purchases through direct sales (yes/no questions). Note that other categorical variables from the questionnaire, such as the household composition variables or questionnaire skips, are edited and validated by subject-matter experts from Income Statistics Division. Therefore, the latter variables are not imputed using the nearest neighbour technique.

Table 5.3
Households requiring imputation of categorical variables, Canada and the provinces

	Number of variables imputed (out of 41)			Total
	1	2	3 or more	
	percent			
Canada	7.7	1.4	0.9	9.9
Newfoundland and Labrador	7.9	0.3	0.1	8.3
Prince Edward Island	12.3	1.5	0.4	14.2
Nova Scotia	8.2	0.9	0.5	9.7
New Brunswick	7.0	0.7	0.1	7.8
Quebec	6.2	0.4	0.8	7.3
Ontario	7.8	2.0	1.5	11.3
Manitoba	10.3	2.9	1.3	14.6
Saskatchewan	6.1	2.0	1.1	9.2
Alberta	5.5	1.0	1.5	8.0
British Columbia	8.0	1.8	1.0	10.8

References

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[4] Statistics Canada, Income Statistics Division. 2008. *Spending Patterns in Canada*. Catalogue no. 62-202. Ottawa.

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Appendix A

Algebraic Notation

1. Nonresponse adjustment

The subweight (i.e. the design weight adjusted for nonresponse) for a household k, denoted as w_k^{NR}, is

$$w_k^{NR} = \pi_k^{-1} * \frac{1}{rate_g} \quad with \quad rate_g = \frac{\sum\limits_{k \in s_{g,r}} \pi_k^{-1}}{\sum\limits_{k \in s_{g,r}} \pi_k^{-1} + \sum\limits_{k \in s_{g,nr}} \pi_k^{-1}}$$

where

$s_{g,r}$ is the set of respondent households in nonresponse group g,

$s_{g,nr}$ is the set of nonrespondent households (refusals, units not contacted, unusable data) in nonresponse group g, and

π_k^{-1} is the design weight attributed to household k.

2. Calculation of the slippage rate

The slippage rate for a control group c, denoted as $rate_c$, is

$$rate_c = 100 * \frac{\left(\sum\limits_{k \in s_{c,r}} w_k^{NR}\right) - t_c}{t_c}$$

where

$s_{c,r}$ is the set of respondent households in control group c,

w_k^{NR} is the subweight of household k, and

t_c is the total of the auxiliary data for control group c.