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Forecast for southern British Columbia coho salmon in 1999

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## Abstract

This Working Paper documents forecasts of marine survival, abundance and distribution for the coho salmon of southern British Columbia (Thompson River, lower Fraser, Strait of Georgia, and west Vancouver Island).

Marine survival: Our recommendations for the marine survival forecast for the five hatchery indicators and one wild coho indicator are given in the following Table. Survivals are expected to remain poor for all Strait of Georgia wild and hatchery stocks and are forecast to be either unchanged or lower in 1999 compared to 1998 (following Table). Survival of Black Creek coho, the single wild indicator on the Strait of Georgia for which there is a forecast, is one of the survivals expected to be lower in 1999 compared to 1998. Nevertheless, survival rates appear to be substantially greater toward the north end of Georgia Strait compared to the lower mainland and the Fraser. CWT escapement data are not yet available for Thompson coho and consequently there is no forecast of marine survival for this aggregate. There are no indications in the magnitude of the escapement that survivals improved in 1998 and the forecast of abundance remains dismal. The forecast survival for wVI coho is slightly lower than in 1997 and 1998.

| indicator | Model | $\hat{s}_{1999}$ | $(50 \% \mathrm{CI})$ | change <br> relative to <br> 1998 |
| :--- | :--- | :---: | :--- | :---: |
| Big Qualicum | LLY | 0.003 | $(0.0013-0.008)$ | same |
| Quinsam | LLY | 0.021 | $(0.013-0.034)$ | same |
| Chilliwack | RAT3 | 0.017 | $(0.010-0.027)$ | lower |
| Inch Creek | LLY | 0.005 | $(0.003-0.010)$ | same |
| Black Creek | 3YRA | 0.042 | $(0.031-0.056)$ | lower |
| Robertson Creek | sibling | 0.029 | $(0.020-0.041)$ | lower |
| regression |  |  |  |  |

Abundance forecast: Without fisheries information, forecasting abundance is highly problematic, and because we are using time-series models, the forecast is dependent on the highly uncertain estimate of abundance in 1998. With those caveats the RAT3 forecast of the StG-Fr aggregate is $2.0 \times 10^{5}\left(50 \% \mathrm{CI}: 1.5 \times 10^{5}-2.8 \times 10^{5}\right)$. This forecast portends a very worrisome further deterioration in the status of Strait of Georgia coho.
The LLY forecast for the wVI aggregate is $4.5 \times 10^{5}\left(50 \% \mathrm{CI}: 3.1 \times 10^{5}-6.5 \times 10^{5}\right)$. This forecast is $77 \%$ of the overall average abundance of $5.9 \times 10^{5}$.

The abundance forecast for Thompson coho is for continued severe depression. Brood year escapements in the Thompson were very low and there is no indication that marine survival will improve either in the southern Strait of Georgia or the west coast of Vancouver Island. We conclude that it is unlikely that stock size will increase appreciably for either the North or South Thompson aggregate in 1999.

Distribution forecast: The predicted proportion of catch inside the Strait of Georgia ( $p_{\text {inside }}$ ) should there be no fishing restrictions is 0.33 ( $50 \% \mathrm{CI} 0.25-0.42$ ), which can be characterized as a moderately strong outside distribution. The confidence interval suggests that an extreme outside year ( $p_{\text {inside }}<0.2$ ) is about as likely as a return to a "normal" distribution ( $p_{\text {inside }}>0.4$ ).

## Résumé

Le présent document traite des prévisions de la survie en mer, de l'abondance et de la répartition du saumon coho du sud de la Colombie-Britannique (rivière Thompson, bas Fraser, détroit de Géorgie et ouest de l'île de Vancouver).
Survie en mer : Nos recommandations relatives à la prévision de la survie en mer des cinq stocks d'élevage et du stock sauvage de saumon coho servant d'indicateurs sont présentées dans le tableau ci-après. Les taux de survie devraient demeurer faibles pour tous les stocks sauvages et d'élevage du détroit de Géorgie et l'on prévoit qu'ils demeureront inchangés ou diminueront en 1999, comparativement à 1998 (tableau). La survie du coho de Black Creek, le seul stock sauvage indicateur du détroit de Géorgie pour lequel nous disposons d'une prévision, est l'un de ceux dont le taux de survie prévu sera inférieur en 1999, par rapport à 1998. Le taux de survie semble cependant passablement plus élevé dans la partie nord du détroit de Géorgie, comparativement au lower mainland et au Fraser. Nous ne disposons pas encore des données sur les échappées obtenues par étiquettes à fils codés pour le coho de la Thompson et nous n'avons donc pas de prévision pour la survie en mer de ces stocks. L'importance des échappées n'indique pas qu'il y ait eu amélioration du taux de survie par rapport à 1998 et l'abondance prévue demeure extrêmement faible. Le taux de survie prévu pour le coho de l'ouest de l'île de Vancouver est légèrement inférieur à ceux de 1997 et 1998.

| Indicateur | Modèle | $\hat{s}_{1999}$ | (IC de 50\%) | Écart par <br> rapport à 1998 |
| :--- | :--- | :--- | :--- | :---: |
| Big Qualicum | LLY | 0,003 | $(0,0013-0,008)$ | inchangé |
| Quinsam | LLY | 0,021 | $(0,013-0,034)$ | inchangé |
| Chilliwack | RAT3 | 0,017 | $(0,010-0,027)$ | inférieur |
| Inch Creek | LLY | 0,005 | $(0,003-0,010)$ | inchangé |
| Black Creek | 3YRA | 0,042 | $(0,031-0,056)$ | inférieur |
| Robertson Creek | régression | (cl. | 0,029 | $(0,020-0,041)$ |
| jumelles) |  | inférieur |  |  |

Prévision de l'abondance : En l'absence de renseignements obtenus des pêches, la prévision de l'abondance s'avère très difficile et comme nous utilisons des modèles fondés sur des séries chronologiques, la prévision est dépendante des estimations fortement incertaines obtenues pour 1998. Si l'on fait abstraction de ces lacunes, la prévision RAT3 du groupe StG-Fr est de $2,0 \times 10^{5}$ (IC de $50 \%$ : $1,5 \times 10^{5}-2,8 \times 10^{5}$ ). Cette prévision indique une détérioration supplémentaire préoccupante de la situation du coho de détroit de Géorgie.
La prévision LLY du groupe wVI est de $4,5 \times 10^{5}$ (IC de $50 \%: 3,1 \times 10^{5}-6,5 \times 10^{5}$ ). Cette prévision correspond à $77 \%$ de l'abondance moyenne générale de $5,9 \times 10^{5}$.
La prévision d'abondance du coho de la Thompson fait état du maintien d'un appauvrissement sévère. Les échappées dans la Thompson ont été très faibles et rien n'indique que la survie en mer s'améliorera que soit dans le sud du détroit de Géorgie ou
sur la côte ouest de l'île de Vancouver. Nous concluons qu'il est peu probable que les effectifs des stocks de la North ou de la South Tompson augmentent de façon appréciable en 1999.

Prévision de la répartition : La proportion prévue des captures à l'intérieur du détroit de Géorgie ( $p_{i n}$ ) en l'absence d'une limitation de la pêche est de 0,33 (IC de $50 \%: 0,25-$ 0,42 ), ce qui peut être qualifié de répartition modérément forte en faveur de l'extérieur. L'intervalle de confiance porte à croire qu'une répartition extrême vers l'extérieur ( $p_{i n}<$ 0.2 ) est pratiquement aussi probable qu'un retour à une répartition «normale » $\left(p_{i n}>0.4\right)$.
Table of Contents
page

1. Introduction ..... 10
2. Data Sources ..... 10
3. Forecasting Models and Retrospective Analysis of Predictive Power. ..... 11
3.1 Forecasting models ..... 11
3.2 Retrospective analyses ..... 12
4. 1998 Marine Survival Estimates ..... 13
4.1 Retrospective Analysis of Survival Predictors ..... 14
4.2 Biologically based forecast for wVI coho ..... 14
4.3 Marine Survival Rate Forecast. ..... 15
5. Forecast of distribution ..... 15
6. Forecasts of abundance ..... 16
6.1 Estimation of abundance in 1998 ..... 17
6.2 Forecast abundance in 1999. ..... 17
6.3 Upper Fraser and Thompson coho ..... 18
7. Conclusions ..... 19
7.1 Marine survival ..... 19
7.2 Abundance forecast ..... 20
7.3 Distribution forecast ..... 20
8. References ..... 21

## Tables

page
Table 1. Release and recovery summaries for the six indicator streams used to generate
forecasts.......................................................................................................... 23
Table 2. Retrospective performance statistics for predictors of $\hat{s}_{\text {smolt }}$. Within each
hatchery indicator the methods were compared over the same set of years. In the top
section of the Table the number of observations in the comparison was determined by
the requirement that the sibling regression data-set held at least nine observations. If
in that comparison a time-series model outperformed the sibling regression then the
comparison was restricted to the time-series models. These comparisons are shown in
the bottom section of the Table. The 'best' model for each hatchery indicator is
shaded. ................................................................................................................... 26
Table 3. Forecast of age 3 return $\left(\hat{R}_{3}^{1999}\right)$ and survival $\left(\hat{s}_{\text {molt }}\right)$ for 1996 brood year for the four Strait of Georgia indicators and Robertson Creek using sibling regressions. Data used are found in Table 1. The slope and intercept are for the sibling regression model (Equation 6)
Table 4. Forecasts of age 3 survival ( $\hat{s}_{\text {smolt }}$ ) with confidence levels for the 1996 brood
year, for the four Strait of Georgia hatchery indicators and one wild indicator and the wVI hatchery indicator. The predictions of the best performing models are shaded. 28
Table 5. Data used for the biologically based survival forecast for Robertson Creek coho. The euphausiid biomass is the average June to October biomass of Thysanoessa spinifera in Barkley Sound in the smolt year (BY+2). The marine survival data are from Table 1
Table 6. Data used in forecasting catch distribution of Strait of Georgia coho salmon. The salinity time series is the average sea-surface salinity measured at Sisters Island and Chrome Island lighthouses in February of BY+3.
Table 7. Correlations between $p_{\text {inside }}$ for the four hatchery stocks on the Strait of Georgia. $N=15$ for all correlations.
Table 8. Forecast of $p_{\text {inside }}$ for 1999 for Strait of Georgia hatchery indicators using the salinity model. Data used are in.
Table 9. Total stock size $\left(N_{h}\right)$ for the four inside hatchery indicators and their proportions $(p)$ of $N_{A}$, the StG-Fr aggregate of wild stocks. Total stock size is the expanded recoveries of age 3 fish as recorded in MRP plus estimates of terminal FW harvest........................................................................................................................ 32
Table 10. Estimates of $\hat{N}$ for the StG-Fr aggregate. ........................................................ 32
Table 11. Total stock size $\left(N_{h}\right)$ for the single outside hatchery indicators and its proportion $(p)$ of $N_{A}$, the wVI aggregate of wild stocks............................................. 33
Table 12. Forecasts of abundance for StG-Fr and wVI aggregates in 1999 ( $\hat{A}_{1999}$ ), with retrospective analysis and confidence limits. The recommended models are shaded. ...................................................................................... 33
Table 13. Coho escapement time series for the eight North Thompson index streams.... 34
Table 14. Coho escapement time series for eight index streams in the South Thompson River drainage from 1975 to 1998.
Table 15. Coho escapement time series for eight additional index streams in the South Thompson River drainage from 1975 to 1998.

Table 16. Exploitation rate, escapement and total return for North and South Thompson average-streams. The marine survival and exploitation rates are averages of the Thompson indicator streams...................................................................................... 37
Table 17. Forecasts of average-stream total returns to the North and South Thompson for 1999 with their associated confidence intervals. Both forecasts were based on the 3YRA model............................................................................................................... 37

## Figures

page
Figure 1. Marine survivals vs. return year for seven coho indicators in southern British
Columbia. The forecast survivals for 1999 are shown with associated $50 \%$ CIs. The Thompson values are a composite of all available smolt release data. A 1999 forecast survival is not available for the Thompson. 38
Figure 2. Return and survival forecast for Robertson Creek coho in 1999 using the sibling regression model. The lower panel is the sibling relationship. The upper panel is the probability distribution for the predicted age 3 return. Returns can be converted to survival using the middle scale............................................................................... 39
Figure 3. Confidence intervals around the time-series forecasts of marine survivals for four hatchery indicators and Black Creek.
Figure 4. Marine survival at Robertson Creek vs. euphausiid biomass in Barkley Sound.
The outlier (1994 return year) was excluded from the fitted line shown. .................. 41
Figure 5. Predicting pinside for 1999 using average Chrome Island and the Sisters February salinities. The lower panel is the predictive relationship. The upper panel is the probability distribution for the point predictions. A February salinity of 28.07 was used
Figure 6. Abundance estimates for the Strait of Georgia+Fraser aggregate and the West Coast Vancouver Island aggregate of southern British Columbia coho. Values shown for 1998 are for estimates derived from each hatchery indicator and for all hatcheries combined (the overall $p_{h}$ ) for the period 1990 to 1997. For the StG-Fr aggregate the preferred estimate of 1998 abundance is shaded. The forecast abundances for 1999 with associated $50 \%$ CI are shown for both aggregates. ............................................ 43
Figure 7. The ratio of the return to each hatchery to the estimated abundance of the StGFr aggregate. The overall proportion was calculated by summing the hatchery indicator abundances before calculating the ratio.44

Figure 8. Probability plots for the abundance forecasts for $\mathrm{StG}-\mathrm{Fr}$ and wVI aggregate abundance in 1999 using the recommended models. ................................................. 45
Figure 9. Total returns to North and South Thompson aggregates from 1975 to 1998.
The forecasts for 1999 with their $50 \%$ CI are shown.

## 1. Introduction

Forecasts of the marine survival rate, the ocean distribution and the ocean abundance of southern British Columbia coho in 1999 are presented in this Working Paper. The methods we used in developing the forecasts marine survival rate and ocean distribution are similar to those used in previous working papers (Holtby and Kadowaki 1998, Kadowaki et al. 1996; Kadowaki 1997).

## 2. Data Sources

Catch and escapement data for coded-wire tagged coho from the Big Qualicum River, Quinsam River, Chilliwack River, Inch Creek and Robertson Creek hatchery stocks and catches of Black Creek coho were obtained from the Mark Recovery Program (MRP) data base maintained at the Pacific Biological Station in Nanaimo, B.C. CWT'd smolt releases in 1997 from the Big Qualicum River, Chilliwack River, and Inch Creek facilities included fish whose left ventral and adipose fins had been removed. These doubly marked fish were not included in any calculations of survival. All hatchery coho were included in estimates of abundance. Escapement data for Black Creek was obtained from program sources in the Stock Assessment Division. Coho could not be retained in Canadian waters in 1998 as part of the conservation measures undertaken to protect Thompson coho. There were some exceptions in terminal areas where surpluses were identified. However, many fisheries proceeded. Coho that were caught were released with minimal harm.

CWT recovery data for 1998 are preliminary and may change once catch and escapement estimates are finalized. A full treatment of the 1998 season will be contained in a future PSARC paper. For forecasting purposes, an exploitation rate of $5 \%$ in Canadian fisheries was applied to all stocks. The exploitation rate in southern US fisheries was assumed to be half of the historic average rate. The exploitation rate in Alaskan fisheries was calculated using recovery data in MRP. Thus, the estimates of total exploitation rate must be viewed as preliminary.

Freshwater sport recoveries of CWT'd coho from the Chilliwack and Inch Creek hatcheries were added to the escapement rather than treated as catch to better represent the exploitation rate on wild stocks, which are not exposed to this intense terminal fishery. Unfortunately the age composition of the terminal sport catch was not estimated. To estimate recoveries at Chilliwack we first assumed that the no-pin rate observed at the hatchery (70 of 1223) was the same for LV/adipose- and adipose-clipped fish. The hatchery escapement of adipose clipped CWT'd age 3 coho was therefore 932. Adiposeclipped age- 3 coho were $4.7 \%$ of the marked coho in the return to the hatchery and we assumed that the same proportion would apply to the age- 3 coho taken in the sport fishery. Marked jacks (all would have been adipose clips) were $0.4 \%$ of the hatchery escapement. Assuming that jacks were equally likely to be caught in the sport fishery as age- 3 coho means that of the 11,886 coho estimated caught in the terminal fishery
approximately 557 were adipose-clipped age- 3 coho. That estimate is reduced to 554 if we assume that jacks are three-times as likely to be caught as age- 3 fish. Our estimate of the escapement of adipose-clipped only coho is 1485 .

Estimating the number of Inch Creek hatchery fish taken in the terminal sport fishery is not possible with certainty. Terminal harvest of two adults and two jacks was permitted regardless of marking and no records were taken of the age composition. The fishery also involved returns from releases to the Stave River and the Nicomen Slough. We assumed that the return was proportional to release and the $90 \%$ of the fish taken were age- 3 adults. We estimated that 912 of the reported harvest of 1062 were age- 3 Inch Creek coho and that the escapement of that stock was 3,712 .

Salinity data for the two lighthouses in the Strait of Georgia were obtained from C. Perkins, Institute of Ocean Sciences, Sidney, BC.

## 3. Forecasting Models and Retrospective Analysis of Predictive Power.

### 3.1 Forecasting models

In this paper we forecast marine survival rates $(s)$, catch distribution ( $p_{\text {inside }}$ ) and stock size or abundance (A). All of these variables are forecast using four quasi-time series models. In each model the variable being forecast $\left(v_{t}\right)$ is first transformed so that

$$
\begin{equation*}
Z_{t}=\mathfrak{I}\left(v_{t}\right) \tag{1}
\end{equation*}
$$

The Log transformation was used for abundance. The Logit transformation $\square_{\text {was applied }}$ to proportions such as $s$ or $p_{\text {inside. }}$. The four models can then be described as follows:

| mnemonic | model |
| :--- | :--- |
| LLY ("Like Last Year") | $Z_{t+1}=Z_{t}+\varepsilon_{t}$ |
| 3YRA (3-year average) | $Z_{t+1}=\frac{\sum_{k=t-2, t} Z_{k}}{3}+\varepsilon_{t}$ |
| RAT1 (1 year trend) | $Z_{t+1}=\frac{Z_{t}^{2}}{Z_{t-1}}+\varepsilon_{t}$ |
| ${ }^{1} Z_{t}=\log v_{t} / 1-v_{t}$ |  |


| mnemonic | model | Equation |
| :---: | :---: | :---: |
| RAT3 (average 3-year trend) | $Z_{t+1}=\frac{\sum_{k=t-2, t} Z_{k} / Z_{k-1}}{3} Z_{t}+\varepsilon_{t}$ |  |

For each model we assume that the error term is normally distributed $\boldsymbol{\varepsilon} \sim N\left(0, \sigma^{2}\right) \boldsymbol{\}}$ nd is independent of time. For the purpose of estimating uncertainty in the forecast value $\left(Z_{t+1}\right)$, an estimate of $\sigma^{2}$ was obtained for the distribution of observed minus predicted for years $1 . . . t$.
The differences between the four models are summarized in the following Table:

|  |  | years used in prediction |  |
| :--- | :---: | :---: | :---: |
|  |  | 1 |  |
|  |  |  | $3(\approx 1$ cycle $)$ |
| project | NO | LLY | 3 YRA |
| trends? | YES | RAT1 | RAT3 |
|  |  |  |  |

Marine survival rates were also predicted using a "sibling-regression" model, where the total return of age- 3 fish $\left(R_{3}^{B Y+3}\right)$ is predicted from the observed age- 2 male escapement ( $R_{2}^{B Y+2}$, 'jacks'):

$$
\begin{equation*}
\log _{e} R_{3}^{B Y+3}=b \log _{e} R_{2}^{B Y+2}+a \tag{6}
\end{equation*}
$$

Survival ( $s_{\text {smolt }}$ ) was then calculated by dividing the age- 3 return by the number of smolts released ( $N_{\text {smolt }}$ ).
Catch distribution or the proportion of the catch caught in waters inside the Strait of Georgia ( $p_{\text {inside }}$ ) was estimated using the model:

$$
\begin{equation*}
\operatorname{Logit}\left(p_{\text {inside }}\right)=b S+a \tag{7}
\end{equation*}
$$

where $S$ is the average February surface salinity at Chrome Island and Sisters in BY $+3^{\underline{2} \text {. }}$ Confidence limits around forecasts in the case of the latter two models were determined using linear regression analysis.

### 3.2 Retrospective analyses

To compare the performance of the forecast models we computed both the Root Mean Square Error (RMSE):

[^0]\[

$$
\begin{equation*}
R M S E=\sqrt{\boldsymbol{\omega}_{\text {bsereved }, t+1}-v_{\text {preaicead }, t+1} \mathbf{I}^{2}} \tag{8}
\end{equation*}
$$

\]

and the Mean Absolute Deviation (MAD):

$$
\begin{equation*}
M A D=\overline{\boldsymbol{K}_{\text {bsereved, } 141}-v_{\text {prediced, }, t+1}} \overline{\boldsymbol{T}} \tag{9}
\end{equation*}
$$

Note that this calculation is performed in the variable space and not in the transformed (equation 1) space.

For models involving regressions (either equation 6 or 7), values of $R M S E$ and $M A D$ were compiled by fitting the model using only the first nine observations and forecasting the tenth year. The tenth year was then added to the data set, the model refit and the eleventh year forecast, and so on. In this way we assessed the predictive power of the model and not its fit to the data. For comparison of the performance of all models used, the $R M S E$ and $M A D$ values were computed over the largest set of years common to all of the forecasts. For the Robertson Creek indicator only BY1992 was a clear outlier to the sibling relationship and it was excluded from the calculation of the sibling regression, but not from the calculations of RMSE and MAD.

## 4. 1998 Marine Survival Estimates

Marine survival rates for the five hatchery indicators and Black Creek, a wild indicator, are presented in Table 1 and Figure 1. Survival rates at Big Qualicum, Inch Creek and Chilliwack were lower in 1998 than in 1997. Survival rates at Quinsam and Black Creek, the northern most indicators in the Strait of Georgia, and at Robertson Creek increased in 1998 compared to 1997 Figure 1. Preliminary indications of escapement to streams around the Strait of Georgia are variable but suggest that survivals of wild coho were likely poor in the southern Strait and the lower Fraser. We think it unlikely that the survival of wild coho is as poor as it is at hatcheries but other than Black Creek we have no data that would allow a test of this.

Survival forecasts for 1998 were prepared for four of the hatcheries we consider in this Working Paper (Holtby and Kadowaki 1998). The survival rate was lower than predicted at Big Qualicum but higher than predicted at Chilliwack, Quinsam and Robertson Creek (see following Table). The observed survival was within the $50 \%$ CI for three of the indicators and only slightly above at the remaining one.

| indicator | forecasting model | $\hat{S}_{1998}$ | $50 \%$ CI for $s$ | $S_{1998}$ |
| :--- | :--- | :---: | :---: | :---: |
| Big Qualicum | RAT3 | 0.007 | $0.002-0.02$ | 0.003 |
| Chilliwack | RAT3 | 0.019 | $0.012-0.03$ | 0.021 |
| Quinsam | LLY | 0.012 | $0.008-0.018$ | 0.021 |
| Robertson Creek | sibling regression | 0.030 | $0.021-0.043$ | 0.038 |

### 4.1 Retrospective Analysis of Survival Predictors

Comparison between the four time series models and the sibling regression model is complicated by the varying number of years in each data set. For Chilliwack and Inch Creek hatcheries the size of the data set was very limited because of the requirement to include at least 9 years of data in the sibling regressions. We first did a comparison of the sibling regression model and the time-series models. If the best performer in that comparison was a time series model then the time-series models were compared further using the longer time series.

Performance measures are summarized in Table 2 for the six indicator stocks. When the comparison between models was restricted to the time period where there were at least nine cases in the sibling regression models then either the RAT3 or RAT1 model performed best in the Strait of Georgia hatchery indicators, although the LLY model also performed well. The Sibling Regression model performed best at Robertson Creek. The relative performance of the LLY model improved when the time period was expanded to include all available data for the Strait of Georgia hatchery indicators (Table 2. The 3YRA model was the best performer for survival at Black Creek where survival has been nearly constant for the past three years (Table 1). The recommended model for each indicator is shown in the following Table.

| indicator | recommended model |
| :--- | :--- |
| Quinsam | LLY |
| Big Qualicum | LLY |
| Chilliwack | RAT3 |
| Inch | LLY |
| Black Creek (wild) | 3YRA |
| Robertson Creek | Sibling Regression |

### 4.2 Biologically based forecast for wVI coho

Marine survival of Carnation Creek coho appears related to early-ocean growth rates (Holtby et al. 1990), which are probably dependent on the amount of available food. Although juvenile coho feed on many species of zooplankton in their first few months in the ocean, euphausiids are the most important food item (Healey 1978; Petersen et al. 1982; Brodeur 1989; Brodeur and Pearcy 1990; Morris and Healey 1990; Brodeur et al. 1992). Euphausiid populations within Barkley Sound have undergone marked declines in recent years (RWT, unpubl. data) which prompted us to examine the relationship between the abundance of Thysanoessa spinifera in Barkley Sound in the smolt year with marine survival of Robertson Creek coho.

Collection and processing protocols for euphausiids are fully described in Tanasichuk (1998). The measure of abundance used here is the average biomass $/ \mathrm{m}^{2}$ (dry weight) during June through October of the smolt year (BY+2) of animals ranging in total length from 9 to 12 mm . This is the size range of susceptibility to juvenile coho (Petersen et al.1982). A total of seven observations were available (Table 5).

After appropriate transformations of the Robertson Creek marine survival data (Table 1) and the euphausiid biomass data (Table 5) we found a suggestive relationship between survival and biomass Figure 4. The 1991 brood year is a clear outlier. Mackerel incursions into Barkley Sound in 1993 are the suspected cause of near-zero survival in
coho and chinook. Excluding that year there is a marginally significant relationship between euphausiid biomass $\left(B M_{e}\right)$ and coho survival $(s)$ :

$$
\begin{aligned}
& \operatorname{Logit}(s)=0.459 \log \left(B M_{e}\right)-5.758 \\
& \quad\left(N=6 ; \text { adj. } \cdot r^{2}=0.44 ; P<0.1\right)
\end{aligned}
$$

We included this predictor here because it leads to a different and considerably lower forecast for 1999 survival than the sibling regression and time-series models.

### 4.3 Marine Survival Rate Forecast

Survival forecasts and associated confidence intervals are shown for the sibling regressions in Table 3 and for the time-series models in Table 4 The survival forecasts made by the best performing model and associated $50 \%$ confidence intervals are presented in the following Table.

| indicator | model | $\hat{s}_{1999}$ | $(50 \% \mathrm{CI})$ |
| :--- | :--- | :--- | :--- |
| Big Qualicum | LLY | 0.0032 | $(0.0013-0.0079)$ |
| Quinsam | LLY | 0.021 | $(0.013-0.034)$ |
| Chilliwack | RAT3 | 0.017 | $(0.010-0.027)$ |
| Inch Creek | LLY | 0.005 | $(0.0027-0.010)$ |
| Black Creek | 3YRA | 0.042 | $(0.031-0.056)$ |
| Robertson Creek | sibling regression | 0.029 | $(0.020-0.041)$ |

The survival outlook for the hatcheries on the Strait of Georgia and in the lower Fraser is mixed but generally poor. Survivals at Big Qualicum and Inch Creek are expected to be considerably lower than $1 \%$. The survival of Black Creek coho is expected to remain around $4 \%$, where it has been for the past three years. Survival at the single outside indicator is expected to be lower than in recent years.
The biologically based predictor for wVI, introduced in Section 4.2. gives a survival forecast of 0.0153 ( $50 \% \mathrm{CI}: 0.011-0.021$ ). While this predictor is based on a short timeseries, it is sufficiently lower than the conventional forecasts to warn against complacency for wVI coho originating in Barkley Sound.

## 5. Forecast of distribution

Variable proportions of the catch of coho originating the coho producing systems around the Strait of Georgia are taken in the sport, troll and net fisheries that have operated within the Strait (Kadowaki 1997; Simpson et al. 1997). Distribution is expressed as the proportion of the catch of hatchery indicator stocks taken in fisheries wholly within the Strait of Georgia ( $p_{\text {inside }}$ ). There was no catch of coho in British Columbia during 1998. The estimated coho mortality in sBC amounted to an exploitation rate between $3 \%$ and $5 \%$. Consequently, there is no estimate of $p_{\text {inside }}$ for 1998 and the time series models that were developed in 1998 cannot be applied (Holtby and Kadowaki 1998). However, we note that the salinity model outperformed the time-series models by a large margin.

Surface salinity at two stations located in the central Strait of Georgia (Sisters and Chrome Island) are correlated with $p_{\text {inside }}$. Salinity in February of the year of return (brood year +3 ) is the best predictor of $p_{\text {inside. }}$. In Kadowaki (1997), the mean of the Chrome Island and Sisters Island February salinities was used to generate the distribution forecast, while in Kadowaki et al. (1996) and Holtby and Kadowaki (1998) the salinity at Chrome Island was used. We have reverted to the average of Chrome Island and Sisters. The differences between the predictions are small and of no practical significance.

Data used in the forecast are given in Table 6. Because of the high correlation between the catch distributions of the hatchery indicators (Table 7) we have forecast only the average distribution. At the time of writing, salinity was available only to the third week of February. With the record amount of rainfall that the region has seen in the past month it is possible that the eventual February salinity will be lower than the value available to us. Where GSsal is the average of the average February salinity at Chrome Island and Sisters:

$$
\begin{aligned}
& \operatorname{logit}\left(\hat{p}_{\text {inside }}\right)=1.002 \text { GSsal }-28.9 \\
&\left(N=23 ; \quad \text { adj. adj. } \quad r^{2}=0.69 ;\right. \\
&P<0.001)
\end{aligned}
$$

Figure 5 shows the fitted relationship and a probability plot of the confidence interval. Confidence levels are tabulated in Table 8. A predicted value of 0.329 could be characterized as a moderately strong outside distribution. The confidence interval suggests that an extreme outside year ( $p_{\text {inside }}<0.2$ ) is about as likely as a return to a "normal" distribution ( $p_{\text {inside }}>0.4$ ). Although there is a tendency to overestimate $p_{\text {inside }}$ when its value is small Figure 5), the preliminary salinity value is well above the low values ( $27 \%$ o) associated with strong outside distributions and marked over-prediction.

## 6. Forecasts of abundance

In southern British Columbia, all fisheries were managed to eliminate coho mortality wherever possible_and to minimize it where not. Fisheries that were permitted were assigned mortality ${ }^{3}$ ceilings based on forecasts of abundance of Strait of Georgia-Fraser (StG-Fr) and west Vancouver Island (wVI) stock aggregates. The StG-Fr aggregate includes stocks originating in streams draining into the Strait of Georgia and Johnstone Straits, including the Fraser and Thompson Rivers. The wVI aggregate is comprised of stocks on the West Coast of Vancouver Island. Holtby and Kadowaki (1998) forecast abundance for these aggregates using fishery mortality (catch), estimates of the stock composition of the catch and estimated mean exploitation rates. A similar method was used to forecast the abundance of coho in the WCVI troll fishing area (Kadowaki et al. 1996; Kadowaki 1997). A similar reconstruction could not be done for 1998 fisheries because there was no catch.

Our method for estimating abundance of the aggregate $(A)$ in 1998 depends directly on past estimates of abundance. Estimates of total stock size $\left(N_{t}\right)$ for individual hatcheries

[^1]were made for the five indicator hatcheries. The ratio $p_{i j}$, was then calculated for each hatchery $i$ in every year $j$ possible:
$$
p_{i j}=\frac{N_{i j}}{A_{j}}
$$

The abundance in 1998 was then estimated for each hatchery $i$ and for the sum of all hatcheries as:

$$
A_{i}=\frac{N_{h}}{p_{i}}
$$

where $p_{i}$ is an average taken over either the entire time series or a recent period. This method assumes that past estimates of $A$ and $N_{i}$ were accurate.

### 6.1 Estimation of abundance in 1998

The ratio of hatchery stock size to the estimated total abundance of $\mathrm{StG}-\mathrm{Fr}$ coho was more variable prior to 1990 than between 1990 and 1997 (Table 9, Figure 7, except for Inch Creek hatchery where the ratio has been roughly constant since 1985. Estimates of the 1998 StG-Fr abundance made using the average (1990-1997) ratios for individual hatcheries and for all four hatcheries summed range from $5.2 \times 10^{4}$ to $7.0 \times 10^{5}$ ( Table 10. We have chosen the value estimated from the overall proportion $\left(3.15 \times 10^{5}\right)$ as the 1998 estimate of StG-Fr abundance. That abundance is greater than the 1998 forecast $\left(2.4 \times 10^{5}\right.$; Holtby and Kadowaki 1998) but within the $50 \%$ CI $\left(1.8 \times 10^{5}-3.3 \times 10^{5}\right)$. Preliminary estimates of escapement to the region suggest to us that the estimate generated from the return to Big Qualicum is too low, while the estimate generated from the returns to Chilliwack may be too high. Of the three remaining estimates the smallest was generated from the overall return. This estimate was selected because it is the most conservative. Strong returns to many wVI streams suggest that the single estimate of wVI abundance in $1998\left(4.5 \times 10^{5}\right)$ is at least plausible. That estimate is considerably greater than the forecast for 1998 of $1.8 \times 10^{5}\left(50 \% \mathrm{CI}: 1.3 \times 10^{5}-2.5 \times 10^{5}\right)$.

### 6.2 Forecast abundance in 1999

The four time series models were used to forecast abundance in 1999 for StG-Fr and wVI aggregates. In the period beginning in 1993 when abundance of the StG-Fr aggregate was clearly trending downward (Figure 6), the "best" model continued to be the RAT3 model (Table 12). With this model the forecast $\mathrm{StG}-\mathrm{Fr}$ abundance is $2.0 \times 10^{5}$ ( $50 \% \mathrm{CI}: 1.5 \times 10^{5}-$ $2.8 \times 10^{5}$; Table 12). A probability distribution of this forecast is shown in Figure 8. For the wVI aggregate the LLY model was the best performer over the period 1993-1998. The forecast abundance using this model is $4.5 \times 10^{5}$ Table 12. A probability distribution of this forecast is shown in Figure 8.

### 6.3 Upper Fraser and Thompson coho

Although coho returning to the upper Fraser/Thompson are part of the $\mathrm{StG}-\mathrm{Fr}$ stock aggregate, they are considered separately because of the role they played in determining salmon fisheries management in southern BC during 1998.

We restricted ourselves to the North and South Thompson drainages. Escapement data from other areas of the upper Fraser are of lower quality and the North and South Thompson aggregates were identified during the 1998 PSARC reviews as the main southern coho stocks of concern. From the available time series of visual escapement data we selected streams that were not enhanced and for which there were at least 19 annual escapement estimates (out of a possible maximum of 24 ). The survey effort expended in many systems during 1998 exceeded the effort given in previous years. We therefore adjusted our escapement estimates for 1998 to reflect historical survey efforts ${ }^{4}$. Escapement estimates for 8 streams in the North Thompson drainage (Table 13) and 16 streams in the South Thompson drainage Table 14, Table 15 extend from 1975 to present. The exploitation rate and marine survival estimates are averages of all available data for Thompson indicators ${ }^{5}$. Survival rate estimates are for smolt releases only. An exploitation rate of 0.05 for Canadian fisheries was assumed for 1998. Thompson coho are exploited in southern US fisheries. For 1998 we assumed the exploitation rate was half of the historical average of 0.11 . No CWT recoveries in the escapement were available at the time of writing. All data from 1998 are preliminary and are subject to revision. For the period 1975 to 1986 we used the average of the first five years of the measured exploitation rate. Data from prior to 1998 are from Irvine et al. (1999).

The escapement time series were manipulated to produce an "average-stream" escapement within each drainage. First, the escapement $(E)$ in each stream $i$ was scaled to the maximum escapement recorded in that stream across all years $t$ :

$$
\begin{equation*}
p_{i, t}=\frac{E_{i, t}}{\max \mathbf{Q} \mathbf{f}} \tag{10}
\end{equation*}
$$

Then the $p_{i, t}$ were averaged across streams $i$ within each year $t$ to give a time series $p_{\bar{i}, t}$ or $p_{\max }$. The average stream escapement was constructed by multiplying $p_{\max }$ by the average across the $i$ streams of $\max \left(E_{i}\right)$. Total returns could then be estimated using the exploitation rate time series. The resulting escapement and return time-series are shown in Table 16 The time-series models (equations 2 to 5) were used to forecast the 1999 return. The forecast from the best performing of these models over the period 1986 to 1998 was chosen. The North and South Thompson were forecast separately.
Escapement in 1998 relative to the brood year of 1995 was varied but in aggregate was considerably improved in the North Thompson but somewhat worse in the South Thompson (Table 13 to Table 15, Figure 9). The apparent failure of the South Thompson aggregate to respond more strongly to the near cessation in harvest is very worrisome, particularly since escapement in 1996 was the lowest yet observed in both Thompson drainages.

[^2]The temporal pattern in total stock size is far from encouraging Table 16 Figure 9. The substantial increase in escapement to the North Thompson appears to have been a transfer of catch to escapement with total stock size remaining roughly constant compared to 1997 and the brood year 1995. Total stock size of the South Thompson average-stream in 1998 was unchanged from 1997 but less than half that of the brood year.
In retrospective analyses the averaging models (LLY and 3YRA) considerably outperformed the ratio models in forecasts of total return, and the 3YRA model was the model of choice. Forecasts for the North and South Thompson aggregate are detailed in Table 17. The forecast total return to both areas is around $20 \%$ of the average total return and in both areas slightly more than the brood return. The forecast average-stream return for the North Thompson is 108 ( $50 \%$ CI: 64-181). Over the period 1975 to 1998 the average escapement was 188 and the average maximum escapement was 420 . The forecast average-stream total return for the South Thompson average-stream is 77 ( $50 \%$ CI:48-124). Over the period 1975 to 1998 the average escapement was 141 and the maximum escapement 320 . In both areas the forecast total return is around $55 \%$ of the past average escapement.
We have no analytical escapement target or standard for Thompson coho. If the period 1983 and 1991 were used to define a desirable escapement, then the forecasts for 1999 represent approximately $15 \%$ of such a target for the South Thompson and $20 \%$ for the North Thompson.

Measured marine survivals of Thompson indicators are difficult to interpret because there are no continuous time-series. However, marine survival appears to have declined (Figure 11), and escapement to the Thompson would suggest that there have been wide-spread declines in the survival of Thompson wild coho, although the effects of FW habitat loss cannot be discounted. Marine survivals for the Strait of Georgia indicators are not expected to improve in 1999. Brood year escapements in the Thompson were very low. These two factors make the conclusion that the outlook for Thompson coho in 1999 is critically poor a necessary one.

## 7. Conclusions

### 7.1 Marine survival

Our recommendations for the marine survival forecast for the five hatchery indicators and one wild coho indicator are given in the following Table. The survival outlook for the inside hatcheries is mixed but generally poor, and there is no indication that an already bleak situation will improve in 1999. Survivals at Inch Creek and Big Qualicum are expected to be considerably lower than $1 \%$. Survivals for Black Creek coho, the single wild indicator on the Strait of Georgia for which there is a forecast, and for Robertson Creek, the single west coast Vancouver Island Vancouver Island indicator, are expected to be lower than in recent years. The forecast of lower survival for west coast Vancouver Island is reinforced by the lower forecast survival rate of 0.015 given by a biologically based model.

| indicator | model | $\hat{s}_{1999}$ | $(50 \% \mathrm{CI})$ | change relative <br> to 1998 |
| :--- | :--- | :--- | :--- | :---: |
| Big Qualicum | LLY | 0.003 | $(0.0013-0.008)$ | same |
| Quinsam | LLY | 0.021 | $(0.013-0.034)$ | same |
| Chilliwack | RAT3 | 0.017 | $(0.010-0.027)$ | lower |
| Inch Creek | LLY | 0.005 | $(0.003-0.010)$ | same |
| Black Creek | 3YRA | 0.042 | $(0.031-0.056)$ | lower |
| Robertson Creek | sibling regression | 0.029 | $(0.020-0.041)$ | lower |

### 7.2 Abundance forecast

Without fisheries information, forecasting abundance is highly problematic, and because we are using time-series models the forecast is dependent on the highly uncertain estimate of abundance in 1998. With those caveats the RAT3 forecast of the StG-Fr aggregate is $2.0 \times 10^{5}\left(50 \% \mathrm{CI}: 1.5 \times 10^{5}-2.8 \times 10^{5}\right)$. This forecast portends a very worrisome further deterioration in the status of Strait of Georgia wild coho.
The LLY forecast for the wVI aggregate is $4.5 \times 10^{5}\left(50 \% \mathrm{CI}: 3.1 \times 10^{5}-6.5 \times 10^{5}\right)$. This forecast is $77 \%$ of the overall average abundance of $5.9 \times 10^{5}$.
The abundance forecast for Thompson coho is for continued severe depression. Brood year escapements in the Thompson were very low and there is no indication that marine survival will improve either in the southern Strait of Georgia or on the west coast of Vancouver Island. We conclude that it is unlikely that stock size will increase appreciably for either the North or South Thompson aggregate in 1999.

### 7.3 Distribution forecast

The predicted proportion of catch inside the Strait of Georgia ( $p_{\text {inside }}$ ) should there be no fishing restrictions is $0.33(50 \% \mathrm{CI} 0.25-0.42)$, which can be characterized as a moderately strong outside distribution. The confidence interval suggests that an extreme outside year ( $p_{\text {inside }}<0.2$ ) is about as likely as a return to a "normal" distribution ( $p_{\text {inside }}>0.4$ ). Qualitative information (e.g. research trawl catches and observations by sport fishermen) suggests a strong outside distribution comparable to those observed in the past few years.

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Table 1. Release and recovery summaries for the six indicator streams used to generate forecasts.

| brood year | number of codedwire tagged smolts | estimated return |  | marine survival age 3 |
| :---: | :---: | :---: | :---: | :---: |
|  |  | age 3 | age 2 (jacks) |  |
| Big Qualicum |  |  |  |  |
| 1972 | 112427 | 40122 | 1398 | 0.357 |
| 1973 | 57425 | 16546 | 931 | 0.288 |
| 1974 | 75512 | 12368 | 1482 | 0.164 |
| 1975 | 210520 | 28019 | 5860 | 0.133 |
| 1976 | 150348 | 28420 | 1504 | 0.189 |
| 1977 | 101224 | 21430 | 621 | 0.212 |
| 1978 | 107328 | 12181 | 543 | 0.113 |
| 1979 | 55435 | 5705 | 733 | 0.103 |
| 1980 | 51984 | 5791 | 271 | 0.111 |
| 2981 | 49274 | 3882 | 643 | 0.079 |
| 1982 | 42453 | 2127 | 181 | 0.050 |
| 1983 | 191620 | 1207 | 184 | 0.006 |
| 1984 | 152273 | 598 | 71 | 0.004 |
| 1985 | 119424 | 1393 | 440 | 0.012 |
| 1986 | 77760 | 1079 | 257 | 0.014 |
| 1987 | 102747 | 3776 | 739 | 0.037 |
| 1988 | 64833 | 3259 | 277 | 0.050 |
| 1989 | 36474 | 2134 | 187 | 0.059 |
| 1990 | 37362 | 2492 | 363 | 0.067 |
| 1991 | 38235 | 2618 | 188 | 0.068 |
| 1992 | 37957 | 1129 | 48 | 0.030 |
| 1993 | 38917 | 6198 | 237 | 0.016 |
| 1994 | 37616 | 525 | 87 | 0.014 |
| 1995 | 38827 | 124 | 41 | 0.003 |
| Chilliwack |  |  |  |  |
| 1980 | 54665 | 6544 | 891 | 0.120 |
| 1981 | 28502 | 4097 | 626 | 0.144 |
| 1982 | 100841 | 18866 | 771 | 0.187 |
| 1983 | 72194 | 7172 | 198 | 0.099 |
| 1984 | 129770 | 21880 | 555 | 0.169 |
| 1985 | 59935 | 10863 | 845 | 0.181 |
| 1986 | 68658 | 8646 | 350 | 0.126 |
| 1987 | 39250 | 4164 | 271 | 0.106 |
| 1988 | 39801 | 3604 | 233 | 0.091 |
| 1989 | 395 | 2239 | 151 | 0.057 |
| 1990 | 39797 | 2361 | 152 | 0.059 |
| 1991 | 79613 | 3598 | 134 | 0.045 |
| 1992 | 39654 | 1481 | 153 | 0.037 |
| 1993 | 39808 | 1577 | 207 | 0.040 |
| 1994 | 36256 | 870 | 75 | 0.024 |
| 1995 | 74456 | 1563 | 117 | 0.021 |
| Inch Creek |  |  |  |  |
| 1983 | 38711 | 2560 | 26 | 0.066 |
| 1984 | 38774 | 3440 | 197 | 0.089 |
| 1985 | 19723 | 4007 | 148 | 0.203 |
| 1986 | 19504 | 2116 | 22 | 0.108 |
| 1987 | 27458 | 2206 | 127 | 0.080 |
| 1988 | 38019 | 2690 | 36 | 0.071 |


| brood year | number of codedwire tagged smolts | estimated return |  | marine survival age 3 |
| :---: | :---: | :---: | :---: | :---: |
|  |  | age 3 | age 2 (jacks) |  |
| 1989 | 29367 | 2851 | 37 | 0.097 |
| 1990 | 31629 | 2607 | 91 | 0.082 |
| 1991 | 21172 | 1279 | 112 | 0.060 |
| 1992 | 20303 | 1116 | 10 | 0.055 |
| 1993 | 21540 | 834 | 90 | 0.039 |
| 1994 | 21174 | 226 | 5 | 0.011 |
| 1995 | 38707 | 201 | 12 | 0.005 |
| Quinsam |  |  |  |  |
| 1975 | 73442 | 7129 | 2204 | 0.097 |
| 1976 | 139968 | 9303 | 3242 | 0.066 |
| 1977 | 168286 | 16778 | 2177 | 0.100 |
| 1978 | 226186 | 12602 | 2311 | 0.056 |
| 1979 | 280127 | 13387 | 3117 | 0.048 |
| 1980 | 76237 | 4973 | 501 | 0.065 |
| 1981 | 279799 | 15019 | 1343 | 0.054 |
| 1982 | 317306 | 27648 | 3443 | 0.087 |
| 1983 | 220929 | 17963 | 1530 | 0.081 |
| 1984 | 77380 | 6135 | 968 | 0.079 |
| 1985 | 42176 | 3352 | 924 | 0.079 |
| 1986 | 192294 | 14824 | 2765 | 0.077 |
| 1987 | 39362 | 3067 | 791 | 0.078 |
| 1988 | 39466 | 1650 | 299 | 0.042 |
| 1989 | 394 | 2317 | 251 | 0.059 |
| 1990 | 39411 | 1365 | 233 | 0.035 |
| 1991 | 42470 | 966 | 315 | 0.023 |
| 1992 | 43742 | 1098 | 353 | 0.025 |
| 1993 | 38947 | 377 | 129 | 0.010 |
| 1994 | 80125 | 953 | 128 | 0.012 |
| 1995 | 38827 | 831 | 643 | 0.021 |
| Black Creek (wild indicator) |  |  |  |  |
| 1983 | 24134 | 3012 |  | 0.125 |
| 1984 | 31648 | 3602 |  | 0.114 |
| 1985 | 35640 | 4510 |  | 0.127 |
| 1986 | 74997 | 8500 |  | 0.113 |
| 1987 | 29203 | 3618 |  | 0.124 |
| 1988 | 118382 | 9004 |  | 0.076 |
| 1989 | 52351 | 6319 |  | 0.121 |
| 1990 | 49860 | 3161 |  | 0.063 |
| 1991 | 54996 | 3131 |  | 0.057 |
| 1992 | 75970 | 3416 |  | 0.045 |
| 1993 | 18152 | 611 |  | 0.034 |
| 1994 | 13736 | 599 |  | 0.044 |
| 1995 | 69996 | 3346 |  | 0.048 |
| Robertson Creek |  |  |  |  |
| 1972 | 44536 | 2954 | 1624 | 0.066 |
| 1973 | 44071 | 3411 | 1234 | 0.077 |
| 1974 | 55672 | 4007 | 1054 | 0.072 |
| 1975 | 51460 | 2507 | 1628 | 0.049 |
| 1976 | 43047 | 3776 | 486 | 0.088 |
| 1977 | 51019 | 2369 | 433 | 0.046 |
| 1978 | 51916 | 1167 | 307 | 0.022 |


|  |  | estimated return |  |  |
| :---: | :---: | ---: | ---: | :---: |
| brood <br> year | number of coded- <br> wire tagged smolts | age 3 | age 2 (jacks) | marine <br> survival <br> age 3 |
| 1979 |  |  |  | 0.020 |
| 1980 | 48776 | 9195 | 110 | 0.057 |
| 1981 | 144742 | 8661 | 1038 | 0.069 |
| 1982 | 125895 | 1932 | 44 | 0.020 |
| 1983 | 94740 | 2038 | 85 | 0.039 |
| 1984 | 52092 | 1335 | 54 | 0.029 |
| 1985 | 46061 | 764 | 86 | 0.018 |
| 1986 | 41474 | 2514 | 412 | 0.049 |
| 1987 | 50967 | 5525 | 615 | 0.090 |
| 1988 | 61191 | 2569 | 139 | 0.059 |
| 1989 | 43524 | 1926 | 57 | 0.046 |
| 1990 | 41773 | 964 | 140 | 0.024 |
| 1991 | 40221 | 19 | 0 | $<0.0005$ |
| 1992 | 38419 | 679 | 2 | 0.013 |
| 1993 | 36873 | 1312 | 23 | 0.016 |
| 1994 | 42248 | 1497 | 228 | 0.031 |
| 1995 | 43005 |  | 54 | 0.038 |

Table 2. Retrospective performance statistics for predictors of $\hat{s}_{\text {smolt }}$. Within each hatchery indicator the methods were compared over the same set of years. In the top section of the Table the number of observations in the comparison was determined by the requirement that the sibling regression data-set held at least nine observations. If in that comparison a time-series model outperformed the sibling regression then the comparison was restricted to the time-series models. These comparisons are shown in the bottom section of the Table. The 'best' model for each hatchery indicator is shaded.

| predictor | RMSE |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | MAD |  |  |  |  |  |
|  | Big Qualicum | Chilliwack | Quinsam | Inch Creek | $\begin{aligned} & \hline \begin{array}{l} \text { Black Creek } \\ \text { (wild) } \end{array} \\ & \hline \end{aligned}$ | Robertson Creek |
| $N$ | 15 | 7 | 12 | 5 | no data | 15 |
| sibling regression | 0.0532 | 0.0362 | 0.0432 | 0.0641 |  | 0.0193 |
|  | 0.0341 | 0.0312 | 0.0305 | 0.0594 |  | 0.0154 |
| LLY | 0.0208 | 0.0156 | 0.0148 | 0.0173 |  | 0.0236 |
|  | 0.0157 | 0.0115 | 0.0102 | 0.0139 |  | 0.0200 |
| 3YRA | 0.0315 | 0.0247 | 0.0154 | 0.0262 |  | 0.0274 |
|  | 0.0269 | 0.0210 | 0.0116 | 0.0244 |  | 0.0227 |
| RAT1 | 0.0189 | 0.0160 | 0.0214 | 0.0143 |  | 0.0382 |
|  | 0.0142 | 0.0138 | 0.0147 | 0.0109 |  | 0.0298 |
| RAT3 | 0.0173 | 0.0102 | 0.0151 | 0.0162 |  | 0.0299 |
|  | 0.0133 | 0.0087 | 0.0109 | 0.0151 |  | 0.0228 |
| $N$ | 20 | 12 | 17 | 10 | 9 | 20 |
| LLY | 0.0317 | 0.0294 | 0.0158 | 0.0259 | 0.0302 | 0.0261 |
|  | 0.0216 | 0.0210 | 0.0116 | 0.0201 | 0.0224 | 0.0222 |
| 3YRA | 0.0361 | 0.0305 | 0.0165 | 0.0275 | 0.0264 | 0.0275 |
|  | 0.0304 | 0.0270 | 0.0124 | 0.0234 | 0.0216 | 0.0238 |
| RAT1 | 0.0406 | 0.0489 | 0.0260 | 0.0481 | 0.0514 | 0.0418 |
|  | 0.0267 | 0.0325 | 0.0201 | 0.0289 | 0.0360 | 0.0324 |
| RAT3 | 0.0453 | 0.0294 | 0.0170 | 0.0405 | 0.0337 | 0.0318 |
|  | 0.0279 | 0.0192 | 0.0127 | 0.0283 | 0.0254 | 0.0259 |

Table 3. Forecast of age 3 return $\left(\hat{R}_{3}^{1999}\right)$ and survival $\left(\hat{s}_{\text {smolt }}\right)$ for 1996 brood year for the four Strait of Georgia indicators and Robertson Creek using sibling regressions. Data used are found in Table 1. The slope and intercept are for the sibling regression model (Equation 6).

|  | Big Qualicum | Chilliwack | Quinsam | Inch Creek | Robertson Creek ${ }^{\S}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $a$ (intercept) | 1.743 | 2.840 | 1.361 | 5.260 | 5.553 |
| $b$ (slope) | 1.084 | 0.979 | 1.039 | 0.549 | 0.387 |
| $N$ | 24 | 16 | 21 | 13 | 22 |
| $r_{a d j}^{2}$ | 0.74 | 0.63 | 0.79 | 0.41 | 0.51 |
| $R_{2}^{1998}$ | 143 | 43 | 90 | 7 | 46 |
| smolts released | 40331 | 37282 | 39813 | 41918 | 39578 |
| $\hat{R}_{3}^{1999}$ | 1237 | 680 | 418 | 560 | 1136 |
| $\hat{S}_{\text {smolt }}$ | 0.031 | 0.018 | 0.011 | 0.013 | 0.029 |
| CI:1\% lower ${ }^{\text {* }}$ | 0.004 | 0.003 | 0.005 | 0.001 | 0.008 |
| CI:5\% lower | 0.008 | 0.005 | 0.006 | 0.003 | 0.012 |
| CI:10\% lower | 0.011 | 0.007 | 0.007 | 0.004 | 0.014 |
| CI:25\% lower | 0.018 | 0.011 | 0.008 | 0.008 | 0.020 |
| CI:75\% lower | 0.052 | 0.029 | 0.013 | 0.024 | 0.041 |

§: 1992 brood year was excluded
$\ddagger$ : In this case $1 \%$ of the observed values are expected to be less than the stated value.

Table 4. Forecasts of age 3 survival ( $\hat{s}_{\text {smolt }}$ ) with confidence levels for the 1996 brood year, for the four Strait of Georgia hatchery indicators and one wild indicator and the wVI hatchery indicator. The predictions of the best performing models are shaded.

|  |  | Strait of Georgia indicator |  |  |  |  | wVIindicator |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Quinsam River | Big Qualicum River | Chilliwack River | Inch Creek | Black Creek (wild) |  |
| LLY | CI:75\% | 0.0340 | 0.0079 | 0.0337 | 0.0101 | 0.0673 | 0.0847 |
|  | $\hat{s}_{\text {smolt }}$ | 0.0214 | 0.0032 | 0.0210 | 0.0052 | 0.0478 | 0.0378 |
|  | CI:25\% | 0.0134 | 0.0013 | 0.0130 | 0.0027 | 0.0338 | 0.0164 |
|  | CI:10\% | 0.0087 | 0.0006 | 0.0082 | 0.0014 | 0.0241 | 0.0076 |
|  | CI:5\% | 0.0066 | 0.0003 | 0.0061 | 0.0010 | 0.0193 | 0.0046 |
|  | CI:1\% | 0.0038 | 0.0001 | 0.0033 | 0.0004 | 0.0119 | 0.0017 |
| 3YRA | CI:75\% | 0.0199 | 0.0190 | 0.0415 | 0.0146 | 0.0563 | 0.0448 |
|  | $\hat{s}_{\text {smolt }}$ | 0.0137 | 0.0090 | 0.0272 | 0.0102 | 0.0415 | 0.0263 |
|  | CI:25\% | 0.0094 | 0.0042 | 0.0178 | 0.0072 | 0.0305 | 0.0153 |
|  | CI:10\% | 0.0066 | 0.0021 | 0.0117 | 0.0051 | 0.0225 | 0.0092 |
|  | CI:5\% | 0.0053 | 0.0013 | 0.0089 | 0.0041 | 0.0184 | 0.0067 |
|  | CI:1\% | 0.0033 | 0.0005 | 0.0048 | 0.0025 | 0.0116 | 0.0035 |
| RAT1 | CI:75\% | 0.0616 | 0.0014 | 0.0339 | 0.0088 | 0.0838 | 0.2424 |
|  | $\hat{s}_{\text {smolt }}$ | 0.0351 | 0.0004 | 0.0183 | 0.0045 | 0.0518 | 0.0468 |
|  | CI:25\% | 0.0198 | 0.0001 | 0.0098 | 0.0023 | 0.0316 | 0.0075 |
|  | CI:10\% | 0.0116 | 0.0000 | 0.0054 | 0.0012 | 0.0195 | 0.0013 |
|  | CI:5\% | 0.0082 | 0.0000 | 0.0037 | 0.0008 | 0.0142 | 0.0004 |
|  | CI:1\% | 0.0040 | 0.0000 | 0.0017 | 0.0003 | 0.0071 | 0.0001 |
| RAT3 | CI:75\% | 0.0325 | 0.0033 | 0.0273 | 0.0033 | 0.0732 | 0.1640 |
|  | $\hat{s}_{\text {smolt }}$ | 0.0193 | 0.0011 | 0.0168 | 0.0019 | 0.0484 | 0.0501 |
|  | $\mathrm{CI}: 25 \%$ | 0.0114 | 0.0004 | 0.0103 | 0.0010 | 0.0317 | 0.0140 |
|  | CI:10\% | 0.0070 | 0.0001 | 0.0064 | 0.0006 | 0.0213 | 0.0042 |
|  | CI:5\% | 0.0051 | 0.0001 | 0.0046 | 0.0004 | 0.0165 | 0.0020 |
|  | CI:1\% | 0.0026 | 0.0000 | 0.0023 | 0.0002 | 0.0098 | 0.0004 |

Table 5. Data used for the biologically based survival forecast for Robertson Creek coho. The euphausiid biomass is the average June to October biomass of Thysanoessa spinifera in Barkley Sound in the smolt year (BY+2). The marine survival data are from Table 1.

| return year <br> $(\mathrm{BY}+3)$ | euphausiid biomass <br> $\left(\mathrm{mg}\right.$ dry mass $\left./ \mathrm{m}^{2}\right)$ | Robertson Creek <br> marine survival |
| :---: | :---: | :---: |
| 1992 | 183 | 0.046 |
| 1993 | 127 | 0.024 |
| 1994 | 40 | 0.0048 |
| 1995 | 42 | 0.013 |
| 1996 | 42 | 0.016 |
| 1997 | 291 | 0.031 |
| 1998 | 76 | 0.036 |
| 1999 | 32 |  |

Table 6. Data used in forecasting catch distribution of Strait of Georgia coho salmon. The salinity time series is the average sea-surface salinity measured at Sisters Island and Chrome Island lighthouses in February of BY+3.

|  | proportion of catch in the Strait of Georgia |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Brood year | Big <br> Qualicum | Quinsam | Chilliwack | Inch Creek | Average <br> $\left(P_{\text {inside }}\right)$ | February <br> salinity $(\%$ o $)$ |
| 1972 | 0.6564 |  |  | 0.6564 | 29.55 |  |
| 1973 | 0.4518 |  |  | 0.4518 | 28.75 |  |
| 1974 | 0.5034 |  |  | 0.5034 | 29.30 |  |
| 1975 | 0.5171 | 0.5266 |  | 0.5218 | 28.54 |  |
| 1976 | 0.7854 | 0.5687 |  | 0.6771 | 29.50 |  |
| 1977 | 0.6204 | 0.4266 |  |  | 0.5235 | 28.75 |
| 1978 | 0.4134 | 0.2504 |  | 0.3319 | 27.65 |  |
| 1979 | 0.4358 | 0.3172 |  | 0.3765 | 28.34 |  |
| 1980 | 0.3637 | 0.2756 | 0.3667 | 0.4071 | 0.3533 | 28.04 |
| 1981 | 0.3449 | 0.3835 | 0.3802 | 0.4591 | 0.3919 | 28.25 |
| 1982 | 0.5699 | 0.5865 | 0.6432 | 0.7445 | 0.6361 | 29.20 |
| 1983 | 0.4686 | 0.4480 | 0.4785 | 0.6519 | 0.5118 | 29.10 |
| 1984 | 0.5669 | 0.6916 | 0.5492 | 0.7226 | 0.6325 | 28.95 |
| 1985 | 0.6751 | 0.5578 | 0.7338 | 0.8053 | 0.6930 | 29.75 |
| 1986 | 0.5269 | 0.4985 | 0.3711 | 0.6255 | 0.5055 | 28.95 |
| 1987 | 0.5910 | 0.6524 | 0.4974 | 0.6084 | 0.5873 | 29.05 |
| 1988 | 0.0938 | 0.1081 | 0.0290 | 0.2027 | 0.1084 | 27.00 |
| 1989 | 0.5717 | 0.5978 | 0.4896 | 0.5760 | 0.5588 | 27.40 |
| 1990 | 0.8265 | 0.8055 | 0.6726 | 0.8714 | 0.7940 | 29.34 |
| 1991 | 0.3603 | 0.4698 | 0.1759 | 0.4057 | 0.3529 | 28.75 |
| 1992 | 0.0215 | 0.0917 | 0.0021 | 0.1775 | 0.0732 | 27.70 |
| 1993 | 0.0439 | 0.1594 | 0.0327 | 0.1289 | 0.0912 | 26.90 |
| 1994 | 0.1010 | 0.0710 | 0.0001 | 0.4610 | 0.1583 | 27.14 |

Table 7. Correlations between $p_{\text {inside }}$ for the four hatchery stocks on the Strait of Georgia. $N=15$ for all correlations.

|  | Inch Creek | Big Qualicum | Chilliwack |
| ---: | :---: | :---: | :---: |
| Big Qualicum | 0.934 | 1.000 |  |
| Chilliwack | 0.914 | 0.950 | 1.000 |
| Quinsam | 0.864 | 0.958 | 0.889 |

Table 8. Forecast of $p_{\text {inside }}$ for 1999 for Strait of Georgia hatchery indicators using the salinity model. Data used are in.

|  | overall <br> $\left(p_{\text {inside }}\right)$ |
| :--- | :---: |
| $a^{\dagger}$ | -28.9 |
| $b$ | 1.002 |
| $N$ | 23 |
|  |  |
| $\hat{p}_{\text {inside }}$ | 0.329 |
| $\mathrm{CI}: 1 \%$ lower |  |
| $\mathrm{CI}: 5 \%$ lower | 0.100 |
| $\mathrm{CI}: 10 \%$ lower | 0.151 |
| $\mathrm{CI}: 25 \%$ lower | 0.184 |
| $\mathrm{CI}: 75 \%$ lower | 0.246 |

$\dagger$ : The fitted model was $\operatorname{Logit}\left(p_{\text {inside }}\right)=b S+a$ where $S$ is the average February surface salinity at Chrome Island and the Sisters in BY+3.
$\ddagger$ : In this case $1 \%$ of the observed values are expected to be less than the stated value.

Table 9. Total stock size $\left(N_{h}\right)$ for the four inside hatchery indicators and their proportions $(p)$ of $N_{A}$, the StG-Fr aggregate of wild stocks. Total stock size is the expanded recoveries of age 3 fish as recorded in MRP plus estimates of terminal FW harvest.

| return <br> year | $N_{A}$ | Big Qualicum |  | Quinsam |  | Inch |  | Chilliwack |  | all combined |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $N$ | $p$ | $N$ | $p$ | $N$ | $p$ | $N$ | p | $N$ | $p$ |
| 1984 | 2415519 | 92117 | 0.038 | 49579 | 0.021 | 656 | 0.000 | 26545 | 0.011 | 168897 | 0.070 |
| 1985 | 1542008 | 62904 | 0.041 | 98521 | 0.064 | 14282 | 0.009 | 83161 | 0.054 | 258868 | 0.168 |
| 1986 | 2019138 | 35599 | 0.018 | 131623 | 0.065 | 13392 | 0.007 | 210735 | 0.104 | 391349 | 0.194 |
| 1987 | 1801199 | 19818 | 0.011 | 98329 | 0.055 | 24890 | 0.014 | 314431 | 0.175 | 457468 | 0.254 |
| 1988 | 2376256 | 24361 | 0.010 | 102257 | 0.043 | 48526 | 0.020 | 326694 | 0.137 | 501838 | 0.211 |
| 1989 | 1288356 | 15000 | 0.012 | 137412 | 0.107 | 12688 | 0.010 | 211979 | 0.165 | 377079 | 0.293 |
| 1990 | 2053328 | 38486 | 0.019 | 82410 | 0.040 | 8629 | 0.004 | 184251 | 0.090 | 313776 | 0.153 |
| 1991 | 1555158 | 66694 | 0.043 | 48992 | 0.032 | 18719 | 0.012 | 165546 | 0.106 | 299951 | 0.193 |
| 1992 | 1974723 | 62121 | 0.031 | 69180 | 0.035 | 17135 | 0.009 | 108570 | 0.055 | 257006 | 0.130 |
| 1993 | 1881718 | 76190 | 0.040 | 42295 | 0.022 | 18736 | 0.010 | 113798 | 0.060 | 251019 | 0.133 |
| 1994 | 1393793 | 80011 | 0.057 | 27694 | 0.020 | 24275 | 0.017 | 130166 | 0.093 | 262146 | 0.188 |
| 1995 | 1287089 | 34047 | 0.026 | 28231 | 0.022 | 17396 | 0.014 | 73158 | 0.057 | 152832 | 0.119 |
| 1996 | 797267 | 21191 | 0.027 | 11552 | 0.014 | 17865 | 0.022 | 71288 | 0.089 | 121896 | 0.153 |
| 1997 | 363607 | 18198 | 0.050 | 13830 | 0.038 | 3036 | 0.008 | 42161 | 0.116 | 77225 | 0.212 |
| 1998 |  | 1913 |  | 12318 |  | 3912 |  | 32026 |  | 44558 |  |

Table 10. Estimates of $\hat{N}_{A}$ for the StG-Fr aggregate.

|  | whole period |  | recent years <br> $(1990-1997)$ |  |
| :--- | :---: | ---: | :---: | ---: |
| hatchery | $p$ | $\hat{N}_{A}$ | $p$ | $\hat{N}_{A}$ |
| Big Qualicum | 0.0394 | 48609 | 0.0368 | 52051 |
| Quinsam | 0.0428 | 287567 | 0.0279 | 440943 |
| Chilliwack | 0.0648 | 493960 | 0.0460 | 696196 |
| Inch Creek | 0.0120 | 341962 | 0.0121 | 341157 |
| average $\hat{N}_{A}$ |  | 288972 |  | 382587 |
|  |  |  |  |  |
| overall | 0.185 | 272762 | 0.160 | 314513 |

Table 11. Total stock size $\left(N_{h}\right)$ for the single outside hatchery indicators and its proportion $(p)$ of $N_{A}$, the wVI aggregate of wild stocks.

| return year | $N_{A}$ | $N_{h}$ | $p$ |
| :---: | ---: | ---: | :--- |
| 1984 | 660059 | 61352 | 0.0929 |
| 1985 | $-{ }^{\dagger}$ | 25958 | - |
| 1986 | 608314 | 64290 | 0.1057 |
| 1987 | 1295715 | 42344 | 0.0327 |
| 1988 | 616236 | 21297 | 0.0346 |
| 1989 | 601214 | 56441 | 0.0939 |
| 1990 | 977262 | 48960 | 0.0501 |
| 1991 | 548391 | 69217 | 0.1262 |
| 1992 | 504734 | 71415 | 0.1415 |
| 1993 | 320019 | 33022 | 0.1032 |
| 1994 | 456507 | 373 | 0.0008 |
| 1995 | 501723 | 9742 | 0.0194 |
| 1996 | 382846 | 14397 | 0.0376 |
| 1997 | 176225 | 3953 | 0.0224 |
| 1998 | 450730 | 32310 | 0.0715 |

$\dagger$ The estimation procedure for $N_{w}$ could not estimate stock composition in 1985.

Table 12. Forecasts of abundance for StG-Fr and wVI aggregates in $1999\left(\hat{A}_{1999}\right)$, with retrospective analysis and confidence limits. The recommended models are shaded.

|  | StG-Fr aggregate abundance ( $\times 10^{5}$ ) |  |  |  | wVI aggregate abundance ( $\times 10^{5}$ ) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | LLY | 3YRA | RAT1 | RAT3 | LLY | 3YRA | RAT1 | RAT3 |
| $\hat{A}_{1999}$ | 3.1 | 4.5 | 2.7 | 2.0 | 4.5 | 3.1 | 12 | 4.4 |
| $1 \% \mathrm{CI}^{\dagger}$ | 1.2 | 1.3 | 0.51 | 0.59 | 1.1 | 0.90 | 1.0 | 0.65 |
| 5\%CI | 1.6 | 2.0 | 0.91 | 0.92 | 1.7 | 1.4 | 2.5 | 1.3 |
| 10\%CI | 1.9 | 2.4 | 1.2 | 1.1 | 2.2 | 1.7 | 3.6 | 1.8 |
| 25\%CI | 2.4 | 3.3 | 1.8 | 1.5 | 3.1 | 2.3 | 6.6 | 2.8 |
| 75\%CI | 4.1 | 6.1 | 4.2 | 2.8 | 6.5 | 4.3 | 23 | 7.1 |
| MAD ( $\times 10^{5}$ ) |  |  |  |  |  |  |  |  |
| 1988 to present | 4.6 | 4.6 | 7.7 | 5.4 | 2.3 | 1.7 | 4.9 | 2.7 |
| 1993 to present | 2.8 | 4.6 | 3.3 | 2.6 | 1.6 | 1.5 | 2.0 | 1.9 |
| $\underline{\text { RMSE }\left(\times 10^{5}\right)}$ |  |  |  |  |  |  |  |  |
| 1988 to present | 5.4 | 5.1 | 9.7 | 6.6 | 3.0 | 1.9 | 7.9 | 3.8 |
| 1993 to present | 3.4 | 5.2 | 3.7 | 3.0 | 1.8 | 1.9 | 2.2 | 2.1 |

${ }^{\dagger}$ stated \% of observations will be less than tabulated value

Table 13. Coho escapement time series for the eight North Thompson index streams.

| year | Barrierre River | Cook <br> Creek | E. Barrierre River | Fennel Creek | N. Thompson River | Raft <br> River | Reg Christie Creek | Tumtum Creek |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1975 | 300 | - | 60 | 90 | 1500 | 500 | - | 6 |
| 1976 | 360 | - | 30 | 95 | 1500 | 250 | 50 | - |
| 1977 | 420 | - | 18 | 380 | 400 | 350 | 8 | 10 |
| 1978 | 400 | 60 | 110 | 300 | 300 | 250 | 20 | 10 |
| 1979 | 400 | 60 | 120 | 600 | 125 | 120 | 5 | 4 |
| 1980 | 60 | 10 | 25 | 40 | 100 | 90 | 10 | 4 |
| 1981 | 350 | 45 | 60 | 100 | 300 | 110 | 15 | - |
| 1982 | 450 | 50 | 75 | 450 | 90 | 200 | 15 | 2 |
| 1983 | 250 | 100 | 100 | 280 | 125 | 250 | 5 | 50 |
| 1984 | 500 | - | 250 | 700 | 700 | 960 | 25 | 25 |
| 1985 | 425 | 0 | 140 | 450 | 100 | - | 0 | 25 |
| 1986 | - | 65 | 250 | 1250 | 500 | 800 | 25 | 80 |
| 1987 | 500 | 200 | 100 | 580 | 500 | 400 | 0 | 0 |
| 1988 | 600 | 25 | 225 | 800 | 600 | 650 | 15 | 0 |
| 1989 | 175 | 70 | 160 | 60 | 680 | 170 | 22 | 25 |
| 1990 | - | 100 | - | 200 | 774 | 50 | 200 | 26 |
| 1991 | 0 | 10 | 0 | - | 667 | 200 | 24 | 0 |
| 1992 | 100 | 20 | 0 | - | 740 | 100 | 70 | 45 |
| 1993 | 100 | 4 | 50 | - | 350 | 50 | 1 | 12 |
| 1994 | - | 0 | 0 | 50 | 358 | 301 | 0 | 2 |
| 1995 | 85 | 0 | 50 | 0 | 150 | 40 | 0 | 2 |
| 1996 | - | 0 | 10 | - | 92 | 15 | 20 | 0 |
| 1997 | 0 | 0 | - | 0 | 200 | - | 30 | 40 |
| $\underline{\underline{1998}}$ | 85 | 0 | 47 | 75 | 101 | 144 | 75 | 14 |

Table 14. Coho escapement time series for eight index streams in the South Thompson River drainage from 1975 to 1998.

| year | Adams River | Bessette <br> System | Blurton <br> Creek | Bolean Creek Canoe Creek | Hunakwa <br> Creek | Kingfisher <br> Creek | Scotch Creek |  |
| ---: | :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1975 | 200 | 1220 | - | - | 25 | 25 | 25 | 25 |
| 1976 | 10 | 495 | 25 | 50 | 10 | 25 | 10 | 5 |
| 1977 | 338 | 166 | 40 | 0 | 0 | 0 | 62 | 0 |
| 1978 | 150 | 580 | 10 | 50 | 100 | 200 | 10 | 0 |
| 1979 | 100 | 530 | 25 | 50 | 10 | 75 | 25 | 0 |
| 1980 | 200 | 490 | 16 | 20 | 60 | 42 | 0 | - |
| 1981 | 100 | 345 | 15 | 55 | 30 | 25 | 25 | - |
| 1982 | 100 | 320 | 0 | 100 | 80 | 50 | 100 | 20 |
| 1983 | 100 | 610 | 0 | 50 | 20 | 50 | 75 | 25 |
| 1984 | 650 | 1000 | 50 | 10 | 30 | 125 | 25 | - |
| 1985 | 500 | 1350 | 50 | 100 | 100 | 0 | 25 | 0 |
| 1986 | 150 | 1120 | 50 | 50 | 30 | 0 | 80 | 50 |
| 1987 | 150 | 950 | 50 | 0 | 100 | 0 | 120 | 50 |
| 1988 | 500 | 1300 | 70 | 50 | 75 | 150 | 150 | 100 |
| 1989 | 350 | 1190 | 35 | 35 | 100 | 120 | 50 | 50 |
| 1990 | 100 | 475 | 50 | 35 | 50 | 30 | 0 | 50 |
| 1991 | 100 | 50 | 0 | 0 | 30 | 75 | 0 | 40 |
| 1992 | 250 | 950 | 50 | 0 | 20 | 120 | 45 | 50 |
| 1993 | 20 | 144 | 30 | - | - | 60 | 10 |  |
| 1994 | 70 | 284 | 8 | 0 | 6 | 90 | 32 | 10 |
| 1995 | 75 | 475 | 0 | 20 | 10 | 60 | 25 | 20 |
| 1996 | 16 | 62 | 0 | 0 | 0 | 2 | 0 | 16 |
| 1997 | 40 | 43 | 0 | - | 10 | 20 | 50 | 2 |
| 1998 | 116 | 101 | 0 | 0 | 0 | 48 | 29 | 0 |

Table 15. Coho escapement time series for eight additional index streams in the South Thompson River drainage from 1975 to 1998.

| year | Shuswap <br> River <br> (lower) | Shuswap <br> River <br> (mid) | South Pass <br> Creek | Tappen <br> Creek | Trinity <br> Creek | Wap Creek | Upper <br> Adams <br> River | Sinmax <br> Creek |
| :---: | :---: | :---: | :---: | ---: | ---: | ---: | ---: | ---: |
| 1975 | 100 | 150 | - | - | 21 | 150 | 60 | 60 |
| 1976 | 40 | 60 | 20 | 1 | 8 | 20 | 50 | 18 |
| 1977 | 100 | 594 | 40 | 12 | 21 | 516 | 150 | 40 |
| 1978 | 300 | 350 | 50 | 2 | 4 | 300 | 100 | 55 |
| 1979 | 300 | 500 | 60 | 3 | 45 | 400 | 475 | 140 |
| 1980 | 350 | 550 | 20 | 0 | 10 | 250 | 75 | - |
| 1981 | 250 | 250 | 20 | 15 | 10 | 100 | 100 | - |
| 1982 | 300 | 350 | 50 | 5 | 35 | 225 | 200 | 15 |
| 1983 | 200 | 250 | 10 | 0 | 10 | 80 | 300 | 10 |
| 1984 | 300 | 700 | 25 | 20 | 30 | 150 | 200 | - |
| 1985 | 500 | 1200 | 50 | 30 | 20 | 250 | 500 | 75 |
| 1986 | 600 | 650 | 50 | 30 | 60 | 200 | 1100 | 80 |
| 1987 | 350 | 500 | 53 | 30 | 20 | 450 | 500 | 120 |
| 1988 | 400 | 1200 | 75 | 40 | 50 | 250 | 700 | 80 |
| 1989 | 250 | 500 | 50 | 35 | 50 | 250 | 1100 | 40 |
| 1990 | 200 | 200 | 50 | 15 | 50 | 200 | 220 | 0 |
| 1991 | 200 | 300 | 25 | 0 | 10 | 75 | 100 | 0 |
| 1992 | 250 | 800 | 85 | 15 | 50 | 300 | 0 | 0 |
| 1993 | 20 | 20 | 7 | - | 6 | - | 60 | 8 |
| 1994 | 100 | 300 | 2 | 0 | 13 | 180 | 159 | 0 |
| 1995 | 25 | 50 | 25 | 10 | 0 | 50 | 126 | 17 |
| 1996 | 0 | 120 | 0 | 0 | 8 | 33 | 120 | 7 |
| 1997 | 0 | 200 | 2 | - | 0 | 35 | 105 | 9 |
| 1998 | 0 | 115 | 3 | 20 | 0 | 93 | 230 | 4 |

Table 16. Exploitation rate, escapement and total return for North and South Thompson averagestreams. The marine survival and exploitation rates are averages of the Thompson indicator streams.

| year | marine survival rate | $\begin{aligned} & \text { exploitation } \\ & \text { rate } \\ & \hline \end{aligned}$ | South Thompson |  | North Thompson |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | escapement | total return | escapement | total return |
| 1975 |  | 0.664 | 124 | 370 | 253 | 752 |
| 1976 |  | 0.664 | 58 | 174 | 242 | 721 |
| 1977 |  | 0.664 | 121 | 359 | 169 | 502 |
| 1978 |  | 0.664 | 150 | 446 | 184 | 547 |
| 1979 |  | 0.664 | 169 | 503 | 174 | 518 |
| 1980 |  | 0.664 | 111 | 330 | 43 | 127 |
| 1981 |  | 0.664 | 104 | 311 | 136 | 407 |
| 1982 |  | 0.664 | 154 | 460 | 160 | 475 |
| 1983 |  | 0.664 | 92 | 273 | 200 | 594 |
| 1984 |  | 0.664 | 183 | 543 | 387 | 1151 |
| 1985 |  | 0.664 | 243 | 723 | 181 | 538 |
| 1986 |  | 0.664 | 238 | 710 | 416 | 1237 |
| 1987 | 0.036 | 0.534 | 221 | 475 | 272 | 583 |
| 1988 | 0.046 | 0.719 | 323 | 1150 | 301 | 1071 |
| 1989 | 0.076 | 0.653 | 242 | 698 | 188 | 541 |
| 1990 | 0.028 | 0.739 | 146 | 558 | 268 | 1027 |
| 1991 | 0.0037 | 0.673 | 64 | 197 | 74 | 227 |
| 1992 | 0.019 | 0.818 | 184 | 1009 | 160 | 879 |
| 1993 | 0.0028 | 0.88 | 66 | 548 | 74 | 620 |
| 1994 | 0.011 | 0.434 | 59 | 104 | 56 | 98 |
| 1995 | 0.024 | 0.566 | 62 | 144 | 40 | 92 |
| 1996 | 0.037 | 0.764 | 17 | 72 | 23 | 97 |
| 1997 | 0.012 | 0.352 | 29 | 45 | 82 | 127 |
| 1998 |  | 0.107 | 44 | 49 | 91 | 102 |

Table 17. Forecasts of average-stream total returns to the North and South Thompson for 1999 with their associated confidence intervals. Both forecasts were based on the 3YRA model.

|  | North Thompson |  | South Thompson |  |
| ---: | ---: | :---: | :---: | :---: |
| CI | total return | \% of average <br> total return | total return | \% of average <br> total return |
| $99 \%$ | 724 | 133 | 437 | 102 |
| $95 \%$ | 395 | 73 | 252 | 59 |
| $90 \%$ | 293 | 54 | 192 | 45 |
| $75 \%$ | 181 | 33 | 124 | 29 |
| $\mathbf{5 0 \%}$ | $\mathbf{1 0 8}$ | $\mathbf{2 0}$ | $\mathbf{7 7}$ | $\mathbf{1 8}$ |
| $\mathbf{2 5 \%}$ | 64 | 12 | 48 | 11 |
| $10 \%$ | 40 | 7 | 31 | 7 |
| $5 \%$ | 30 | 5 | 24 | 6 |
| $1 \%$ | 16 | 3 | 14 | 3 |



Figure 1. Marine survivals vs. return year for seven coho indicators in southern British Columbia. The forecast survivals for 1999 are shown with associated $50 \%$ CIs. The Thompson values are a composite of all available smolt release data. A 1999 forecast survival is not available for the Thompson.

## Robertson Creek 1998



Figure 2. Return and survival forecast for Robertson Creek coho in 1999 using the sibling regression model. The lower panel is the sibling relationship. The upper panel is the probability distribution for the predicted age 3 return. Returns can be converted to survival using the middle scale.


Figure 3. Confidence intervals around the time-series forecasts of marine survivals for four hatchery indicators and Black Creek.


Figure 4. Marine survival at Robertson Creek vs. euphausiid biomass in Barkley Sound. The outlier (1994 return year) was excluded from the fitted line shown.


Figure 5. Predicting pinside for 1999 using average Chrome Island and the Sisters February salinities. The lower panel is the predictive relationship. The upper panel is the probability distribution for the point predictions. A February salinity of 28.07 was used.


Figure 6. Abundance estimates for the Strait of Georgia+Fraser aggregate and the West Coast Vancouver Island aggregate of southern British Columbia coho. Values shown for 1998 are for estimates derived from each hatchery indicator and for all hatcheries combined (the overall $p_{h}$ ) for the period 1990 to 1997 . For the StG-Fr aggregate the preferred estimate of 1998 abundance is shaded. The forecast abundances for 1999 with associated $50 \% \mathrm{CI}$ are shown for both aggregates.


Figure 7. The ratio of the return to each hatchery to the estimated abundance of the StG-Fr aggregate. The overall proportion was calculated by summing the hatchery indicator abundances before calculating the ratio.

## RAT3 forecast abundance of StG-Fr



LLY abundance of $w V I$


Figure 8. Probability plots for the abundance forecasts for StG-Fr and wVI aggregate abundance in 1999 using the recommended models.


Figure 9. Total returns to North and South Thompson aggregates from 1975 to 1998. The forecasts for 1999 with their $50 \% \mathrm{CI}$ are shown.


[^0]:    ${ }^{2}$ BY is the brood year. The progeny of fish spawning in year 1 are caught and spawn in year 4 or BY +3 .

[^1]:    ${ }^{3}$ Mortality is the product of an assumed mortality per encounter and an encounter rate estimated from observation.

[^2]:    ${ }_{5}^{4}$ A detailed examination of this bias is presented in Irvine et al. (1999)
    ${ }^{5}$ 1987-1993: Eagle River; 1994-1995: Salmon River; 1995-1997 Louis \& Lemieux Creeks.

