Catalogue no. 82-003-X

Health Reports



Volume 23, Number 3



Statistics Canada

Statistique Canada



How to obtain more information

Specific inquiries about this product and related statistics or services should be directed to: Health Analysis Division, Statistics Canada, Ottawa, Ontario, K1A 0T6 (telephone: 613-951-1765).

For information about this product or the wide range of data available from Statistics Canada, visit our website at **www.statcan.gc.ca** or contact us by e-mail at **infostats@statcan.gc.ca** or by phone from 8:30 am to 4:30 pm Monday to Friday at:

Toll-free telephone (Canada and the United States):

Statistics Canada national contact centre: Fax line	1-613-951-8116 1-613-951-0581
Depository Services Program fax line	1-800-565-7757
Depository Services Program inquiries line	1-800-635-7943
Fax line	1-877-287-4369
National telecommunications device for the hearing impaired	1-800-363-7629
Inquiries line	1-800-263-1136

Accessing and ordering information

This product, catalogue no. 82-003-XIE, is available for free in electronic format. To obtain a single issue, visit our website at **www.statcan.gc.ca** and select **Publications**.

This product, catalogue no. 82-003-XPE, is also available as a standard printed publication at a price of CAN\$24.00 per issue and CAN\$68.00 for a one-year subscription.

The following additional shipping charges apply for delivery outside Canada:

	Single issue	Annual subscription
United States	CAN\$6.00	CAN\$24.00
Other countries	CAN\$10.00	CAN\$40.00

All prices exclude sales taxes.

The printed version of this publication can be ordered by

Phone (Canada and United States)
 Fax (Canada and United States)
 E-mail
 1-800-267-6677
 1-877-287-4369
 infostats@statcan.gc.ca

Mail Statistics Canada
 Finance Division
 R.H. Coats Bldg., 6th Floor

100 Tunney's Pasture Driveway Ottawa (Ontario) K1A 0T6

· In person from authorized agents and bookstores.

When notifying us of a change in your address, please provide both old and new addresses.

Standards of service to the public

Statistics Canada is committed to serving its clients in a prompt, reliable and courteous manner. To this end, the Agency has developed standards of service which its employees observe in serving its clients. To obtain a copy of these service standards, please contact Statistics Canada toll free at 1-800-263-1136. The service standards are also published on www.statcan.gc.ca under About us > Providing services to Canadians.

A Canadian peer-reviewed journal of population health and health services research

Published by authority of the Minister responsible for Statistics Canada

© Minister of Industry, 2012

All rights reserved. Use of this publication is governed by the Statistics Canada Open Licence Agreement (http://www.statcan.gc.ca/reference/copyright-droit-auteur-eng.htm).

September 2012

Catalogue no. 82-003-XPE, Vol. 23, No. 3 ISSN 0840-6529

Catalogue no. 82-003-XIE, Vol. 23, No. 3 ISSN 1209-1367

Frequency: Quarterly

Ottawa

Note of Appreciation

Canada owes the success of its statistical system to a long-standing partnership between Statistics Canada, the citizens of Canada, its businesses, governments and other institutions. Accurate and timely statistical information could not be produced without their continued cooperation and goodwill.

Editor-in-Chief Didier Garriguet

Senior Editor Mary Sue Devereaux

Managing Editor Janice Felman

Assistant Editor Anne Marie Baxter

Production Manager Robert Pellarin

Creative Services Rasha Bradic

Administration Amber Doy-Yat

Associate Editors

Isabelle Amano Statistics Canada

David Buckeridge McGill University

Elizabeth Lin
The Clarke Institute of Psychiatry

Doug Manuel
Ottawa Health Research Institute
and Statistics Canada

Nazeem Muhajarine University of Saskatchewan

Georgia Roberts
Statistics Canada

Geoff Rowe Statistics Canada

Michelle Simard
Statistics Canada

Pamela White Statistics Canada

Author information: We seek submissions from researchers based in government or academia. Submissions can come in the form of a traditional research article, a shorter descriptive piece that we call "Health Matters," or a contribution that addresses technical issues related to the analysis of complex health surveys or administrative databases—"Methodological Insights." For detailed author guidelines, please visit the journal's website at: www.statcan.gc.ca/healthreports.

Electronic version: *Health Reports* is available free in PDF or HTML format. The current issue may be obtained at *www.statcan.gc.ca/healthreports*. For previous issues, select "Other issues in the series" from the left sidebar of the *Health Reports* website.

Aussi disponible en français : Rapports sur la santé, nº 82-003-X au catalogue

Symbols

The following standard symbols are used in Statistics Canada publications:

- . not available for any reference period
- .. not available for specific reference period
- ... not applicable
- p preliminary
- r revised
- x suppressed to meet the confidentiality requirements of the Statistics Act
- E use with caution
- F too unreliable to be published

The paper used in this publication meets the minimum requirements of American National Standard for Information Sciences – Permanence of Paper for Printed Library Materials, ANSI Z39.48 – 1984.

About Health Reports

Health Reports publishes original research on diverse topics related to the health of populations and the delivery of health care. The journal archives, for the research and policy communities and for the general public, discoveries from analyses of national/provincial surveys and administrative databases, as well as results of international comparative health research. Health Reports is also a forum for sharing methodological information by those using health surveys or administrative databases. Health Reports is produced by the Health Analysis Division at Statistics Canada. Articles appear monthly in electronic format and quarterly in print, and are indexed in Index Medicus and MEDLINE.

For more information about *Health Reports*, contact Janice Felman, Health Analysis Division, Statistics Canada, 24th Floor, R.H. Coats Building, Ottawa, Ontario, Canada K1A 0T6. Telephone: (613) 951-6446; fax: (613) 951-3959; email: HealthReports@statcan.gc.ca

Editorial Board

David L. Streiner, Scientific Editor University of Toronto

Bill Avison
University of Western Ontario

Adam Baxter-Jones
University of Saskatchewan

James Dunn

University of Toronto and Centre for Research on Inner City Health

Bob Evans University of British Columbia

David Feeny
Kaiser Permanente

Rick Glazier

Institute for Clinical Evaluative Sciences and University of Toronto

Judy Guernsey
Dalhousie University

Howard Morrison
Public Health Agency of Canada

Cameron Mustard

Institute for Work and Health, University of Toronto

Tom Noseworthy
University of Calgary

Patricia O'Campo University of Toronto and Centre for Research on Inner City Health

Jennifer O'Loughlin University of Montreal

Nancy Ross McGill University

Paul Veugelers
University of Alberta

Michael Wolfson University of Ottawa

Xiaoyi Yan Health Canada

In this issue

Research articles



by Lisa N. Oliver and Dafna E. Kohen

Rates of hospitalization for most causes of unintentional injury are higher for children and youth in areas with a high percentage of Aboriginal identity residents, compared with children and youth in low-percentage Aboriginal identity areas.



☐ Mortality rates among children and teenagers living in Inuit Nunangat, 1994 to 200817

by Lisa N. Oliver, Paul A. Peters and Dafna E. Kohen

In 1994-1998 and in 2004-2008, age-standardized mortality rates were higher for children and teenagers in Inuit Nunangat, compared with those in the rest of Canada.



Cause-specific mortality by education in Canada: A 16-year follow-up study23

by Michael Tjepkema, Russell Wilkins and Andrea Long

All-cause age-specific mortality rates were highest among people with less than secondary school graduation and lowest for university degree-holders.

Health matters



by Annie Turner and Leanne Findlay

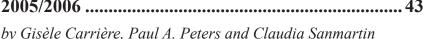
More than one-third of Canadians aged 45 or older reported caring for a senior with a short- or long-terms health condition or limitation.



Karen C. Roberts, Margot Shields, Margaret de Groh, Alfred Aziz and Jo-Anne Gilbert

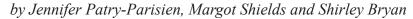
According to the World Health Organization approach, close to a third of 5- to 17-year-olds in Canada are overweight or obese, compared with about a quarter according to the classification cut-offs established by the International Obesity Task Force.

Methodological insights



Hospital records that lack information on country of birth can be analysed with aggregate census data to compare hospitalization rates for areas with greater or lesser percentages of foreign-born residents.





Waist circumference measures based on the World Health Organization and National Institutes of Health protocols differ significantly.



ELECTRONIC PUBLICATIONS AVAILABLE AT

www.statcan.gc.ca

Unintentional injury hospitalizations among children and youth in areas with a high percentage of Aboriginal identity residents: 2001/2002 to 2005/2006

by Lisa N. Oliver and Dafna E. Kohen

Abstract

Background

Because administrative data typically do not contain Aboriginal identifiers, national unintentional injury hospitalization rates among Aboriginal children have not been reported. This study examines rates of unintentional injury hospitalization for children in areas with a high-percentage Aboriginal identity population.

Data and Methods

Data are from the Hospital Morbidity Database (2001/2002 to 2005/2006). Rates of unintentional injury hospitalization were calculated for 0- to 19-year-olds in census Dissemination Areas (DAs) where at least 33% of residents reported an Aboriginal identity. DAs were classified as high-percentage First Nations, Métis or Inuit identity based on the predominant group.

Results

Unintentional injury hospitalization rates of children and youth in high-percentage Aboriginal identity areas were at least double the rate for their contemporaries in low-percentage Aboriginal identity areas. Falls and land transportation were the most common causes of unintentional injury hospitalization, regardless of Aboriginal identity status, but disparities between rates for high- and low-percentage Aboriginal identity areas were often greatest for less frequent causes, such as fire, natural/environmental, and drowning/ suffocation.

Interpretation

The geographic areas where children live were associated with hospitalization rates for injury.

Keywords

Child health, drowning, hospital records, Inuit, Métis, poisoning, trauma

Authors

Lisa N. Oliver (lisa.oliver@statcan.gc.ca) is with the Research Data Centre at Simon Fraser University and Dafna E. Kohen (1-613-951-3346; dafna.kohen@statcan.gc.ca) is with the Health Analysis Division at Statistics Canada, Ottawa, Ontario, K1A 0T6. Inintentional injury is the leading cause of death¹ and morbidity² among Canadian children.

Not only are injuries associated with increased health care costs, hospitalizations and physician care,³ but injuries sustained in childhood also have consequences that can last throughout the lifecourse.⁴ For these reasons, childhood injuries have been identified as a public health issue. Among Aboriginal children, in particular, injury rates have been reported to be relatively high.⁵⁻⁷

Canadian studies of injury in Aboriginal populations have largely focused on adults.^{8,9} Most studies of hospitalization due to injury have been restricted by the lack of Aboriginal identifiers on hospitalization records. To address this problem, some researchers have adopted a geographic approach and examined hospitalizations in areas that have a high percentage of Aboriginal residents.¹⁰⁻¹² These studies reported higher rates of injury hospitalization for people living in such areas.

While there are limitations to a geographic approach, the present analysis examines hospitalizations for unintentional injury among children and youth in communities where at least 33% of residents reported Aboriginal identity. The purposes are to: (1) calculate rates

of unintentional injury hospitalization by cause for areas with a relatively high percentage of First Nations, Métis, and Inuit identity residents, and (2) compare those rates with rates for children and youth in areas with a low percentage of Aboriginal identity residents.

Methods

The data are from the 2001/2002 to 2005/2006 Hospital Morbidity Database (HMDB), which contains discharge records for all hospital separations in Canada. For each separation, information on the patient's age, sex, residential postal code, and diagnoses is available.

This analysis pertains only to acutecare facilities. The data represent the number of hospital separations, not

Unintentional injury hospitalizations among children and youth in areas with a high percentage of Aboriginal identity residents: 2001/2002 to 2005/2006 • Research article

the number of individuals (a single individual may have been hospitalized more than once). Multiple diagnoses may be listed on discharge records; the presence of at least one diagnosis of unintentional injury made the record eligible for inclusion in this analysis.

Unintentional injuries are those for which there was no intent to harm; adverse effects due to drugs or medical care are excluded. The International Classification of Diseases (ICD) was used to classify unintentional injuries based on the external cause of the injury. The version of the ICD codes submitted by each province and recorded on the HMDB were used for this analysis. Causes of injury hospitalizations were grouped based on the International Collaborative Effort on Injury Statistics¹³: falls, land transportation, motor vehicle traffic (subset of land transportation), being struck, being cut/pierced, natural/ environmental, poisoning, fire (includes hot substances), drowning/suffocation, and other.

Residential postal codes on individual hospital separation records were linked to census Dissemination Areas (DAs) via the Postal Code Conversion File. 14 Because discharge records for Quebec contain only the first three digits of the six-digit postal code, they were excluded from this study.

DAs where at least 33% of residents reported an Aboriginal identity to the 2001 Census were considered to be "high-percentage Aboriginal identity" areas.10,15 These DAs were further classified as First Nations, Métis or Inuit identity areas, based on the predominant Aboriginal identity group in the DA. On average, predominant First Nations DAs had 76% First Nations identity; predominant Métis DAs, 38% Métis identity; and predominant Inuit DAs, 79% Inuit identity. All other DAs were designated low-percentage Aboriginal identity areas. Excluding Quebec, there were 1,862 predominant First Nations identity DAs, 135 predominant Métis identity DAs, 69 predominant Inuit identity DAs, and 38,774 low-percentage Aboriginal identity DAs.

The denominator for hospitalization rates was derived from the 2001 and 2006 Censuses. The denominator was the sum of the interpolated populations (aged 0 to 19) for each of the five years of hospitalization data (2001/2002 to 2005/2006, excluding Quebec) and was based on the midpoint (October) of the fiscal year (April to March). For the 2002/2003 fiscal year, Nunavut did not submit hospital separation data, and the population count for this year was excluded.

Because of small populations, global non-response, or incompletely enumerated Indian Reserves, a small number of DAs lacked the detailed age and sex data needed to provide a complete denominator. To retain these DAs in the sample, age and sex were estimated from total population counts

or population estimates of incompletely enumerated Indian Reserves.

Valid cases of unintentional injury hospitalization for this analysis totalled 117,605. Because of invalid or missing postal codes, 3,320 unintentional injury hospitalizations were excluded, and another 327 were excluded owing to insufficient census information at the DA level.

Hospitalization rates were agestandardized to the Aboriginal identity population based on the 2001 Census. Age-standardized hospitalization rates (ASHRs) per 10,000 person-years at risk and rate ratios (RRs) for those in high-percentage First Nations, Métis and Inuit identity areas, compared with those in low-percentage Aboriginal identity areas, were calculated. ASHRs and RRs were calculated by sex, age group (0 to 9 and 10 to 19), and cause of injury. Confidence intervals were based

Table 1 Number and percentage distribution of hospitalizations for unintentional injury and crude rate, by sex, age group, Dissemination Area reporting Aboriginal identity, and cause of injury, population aged 0 to 19, Canada (excluding Quebec), 2001/2002 to 2005/2006

	Hospitaliz	ations	Crude rate per 10,000
	Number	%	person-years at risk
Total	117,605	100.0	39.4
Sex			
Male	77,960	66.3	50.9
Female	39,645	33.7	27.2
Age (years)			
0 to 9	46,954	39.9	33.9
10 to 19	70,651	60.1	44.1
Type of Dissemination Area			
High % First Nations	6,712	5.7	86.2
High % Métis	828	0.7	89.2
High % Inuit	546	0.5	83.3
Low % Aboriginal	109,519	93.1	37.8
Cause of injury [†]			
Falls	43,713	37.2	14.6
Land transportation	29,076	24.7	9.7
Motor vehicle traffic	13,842	11.8	4.6
Struck	13,400	11.4	4.5
Poisoning	6,647	5.7	2.2
Cut/Pierce	3,499	3.0	1.2
Fire	3,010	2.6	1.0
Natural/Environmental	2,920	2.5	1.0
Drowning/Suffocation	1,683	1.4	0.6
Other	14,524	12.3	4.9

[†] because multiple injuries were recorded, causes add to more than total

Notes: Dissemination Areas where at least 33% of the population reported Aboriginal identity are classified as high-percentage
Aboriginal identity. Classification as high-percentage First Nations, Métis or Inuit identity is based on the predominant group
Source: 2001/2002 to 2005/2006 Hospital Morbidity Database.

on a Poisson distribution. A t-test was used to determine if unintentional injury hospitalization rates for high-percentage First Nations, Métis, and Inuit identity areas differed significantly from rates for low-percentage Aboriginal identity areas (p<0.05). All statistical analyses were performed using SAS version 9.1.

Results

Hospitalization rates

From 2001/2002 through 2005/2006, Canadian hospitals (excluding Quebec) recorded 117,605 separations of children and youth aged 0 to 19 for unintentional injury (Table 1). Two-thirds of these hospitalizations were of males, and 60% were of 10- to 19-year-olds.

Overall, the crude unintentional injury hospitalization rate for the population aged 0 to 19 was 39.4 per 10,000 person-years at risk. Rates were considerably higher among children and youth in DAs where at least a third of the population reported Aboriginal identity: 86.2 per 10,000 person-years at risk in high-percentage First Nations identity areas; 89.2 in high-percentage Métis identity areas; and 83.3 in high-percentage Inuit identity areas.

When the rates were age-standardized to account for the different age distributions of each group, patterns were similar (Table 2). The age-standardized unintentional injury hospitalization rate (ASHR) for 0- to 19-year-olds was 37.1 per 10,000 person-years at risk in low-percentage Aboriginal identity areas, compared with 85.9 in high-percentage First Nations identity areas, 88.2 in high-percentage Métis identity areas, and 83.0 in high-percentage Inuit identity areas.

Rate ratios (RRs) were calculated to compare unintentional injury hospitalization rates in the three types of high-percentage Aboriginal identity areas with the rates for the low-percentage Aboriginal identity areas. RRs in the high-percentage Aboriginal identity areas ranged from 2.2 to 2.4; that is, ASHRs in these areas were more than twice those in low-percentage Aboriginal identity areas.

Table 2 Number of hospitalizations for unintentional injury, age-standardized rate, and rate ratio, by cause of injury and Dissemination Area reporting Aboriginal identity, population aged 0 to 19, Canada (excluding Quebec), 2001/2002 to 2005/2006

Cause of injury and type of		Age-standardized rate (per 10,000 person-years	95% confide inter	ence	Rate	95% confidence interval		
Dissemination Area	Number	at risk)	from	to	ratio	from	to	
Total								
High % First Nations	6,712	85.9*	83.9		2.3	2.3	2.4	
High % Métis	828	88.2*	82.4		2.4	2.2	2.6	
High % Inuit	546	83.0*	76.3		2.2	2.1	2.4	
Low % Aboriginal	109,519	37.1	36.8	37.3	1.0			
Falls	0.000	00.7*	07.0	20.0	0.0	1.0	0.4	
High % First Nations	2,223	28.7*	27.6		2.0 2.0	1.9	2.1	
High % Métis High % Inuit	267 150	29.0* 23.1*	25.7 19.7		1.6	1.8 1.4	1.9	
Low % Aboriginal	41,073	14.4	14.3		1.0	1.4		
· ·	41,070	17.7	14.5	14.0	1.0			
Land transportation High % First Nations	1,694	21.5*	20.5	22 6	2.5	2.4	2.6	
High % Métis	251	26.1*	23.0		3.0	2.7	3.4	
High % Inuit	178	27.0*	23.3		3.1	2.7	3.6	
Low % Aboriginal	26,953	8.6	8.5	8.7	1.0			
Motor vehicle traffic								
High % First Nations	797	10.1*	9.4	10.8	2.5	2.3	2.7	
High % Métis	88	9.0*	7.3	11.1	2.3	1.8	2.8	
High % Inuit	59	8.9*		11.5	2.2	1.7	2.9	
Low % Aboriginal	12,898	4.0	3.9	4.1	1.0			
Struck	_							
High % First Nations	552	7.0*	6.5	7.7	1.7	1.6	1.9	
High % Métis	81	8.5*	6.8	10.6	2.0	1.6	2.6	
High % Inuit	40	6.1*	4.5	8.3	1.5	1.1	2.0	
Low % Aboriginal	12,727	4.1	4.1	4.2	1.0			
Poisoning	558	7.1*	6.5	7.7	3.3	3.1	3.6	
High % First Nations High % Métis	44	4.9*	3.6	6.6	2.3	1.7	3.1	
High % Inuit	31	4.6*	3.2	6.5	2.2	1.5	3.1	
Low % Aboriginal	6,014	2.1	2.1	2.2	1.0			
Cut/Pierce	5,5							
High % First Nations	251	3.2*	2.8	3.6	3.0	2.7	3.5	
High % Métis	33	3.5*	2.5	5.0	3.4	2.4	4.8	
High % Inuit	17	2.6*	1.6	4.2	2.5	1.5	4.0	
Low % Aboriginal	3,198	1.0	1.0	1.1	1.0			
Fire								
High % First Nations	305	3.9*	3.5	4.4	4.1	3.6	4.6	
High % Métis	22	2.4*	1.6	3.7	2.5	1.7	3.9	
High % Inuit	10	1.5	0.8	2.8	1.6	8.0	2.9	
Low % Aboriginal	2,673	1.0	0.9	1.0	1.0			
Natural/Environmental High % First Nations	265	3.4*	3.0	3.8	3.7	3.2	4.2	
High % Métis	19	2.0*	1.3	3.2	2.2	1.4	3.5	
High % Inuit	35	5.3*	3.8	7.4	5.8	4.1	8.1	
Low % Aboriginal	2,601	0.9	0.9	1.0	1.0		0.1	
Drowning/Suffocation	_,	2.0		-		***		
High % First Nations	124	1.6*	1.3	1.9	2.8	2.3	3.3	
High % Métis	15	1.7*	1.0	2.8	2.9	1.8	4.9	
High % Inuit	11	1.7*	0.9	3.0	3.0	1.6	5.3	
Low % Aboriginal	1,533	0.6	0.5	0.6	1.0			
Other								
High % First Nations	784	10.0*		10.7	2.2	2.0	2.4	
High % Métis	102			13.0	2.3	1.9	2.8	
High % Inuit	12 561			14.5	2.5	2.0	3.2	
Low % Aboriginal	13,561	4.6	4.5	4.7	1.0			

^{*} significantly different from low-percentage Aboriginal identity Dissemination Areas ... not applicable

Notes: Dissemination Areas where at least 33% of the population reported Aboriginal identity are classified as high-percentage
Aboriginal identity. Classification as high-percentage First Nations, Métis or Inuit identity is based on the predominant group
Source: 2001/2002 to 2005/2006 Hospital Morbidity Database.

Unintentional injury hospitalizations among children and youth in areas with a high percentage of Aboriginal identity residents: 2001/2002 to 2005/2006 • Research article

Causes of injury

Almost without exception, ASHRs by cause of unintentional injury were significantly higher for children and youth in high-percentage First Nations, Métis, and Inuit identity areas than for their counterparts in low-percentage Aboriginal identity areas.

Falls ranked first as a cause of unintentional injury hospitalization among children and youth, whether they lived in high- or low-percentage Aboriginal identity DAs. This category includes falls on one level (for example, on ice) and falling off an object (for example, furniture). In high-percentage Aboriginal identity areas, ASHRs for injury due to falls ranged from 23.1 to 29.0 per 10,000 person-years at risk, compared with 14.4 in low-percentage Aboriginal identity areas.

Land transportation, the second most prevalent cause of unintentional injury hospitalizations among children and youth, involves motorized and non-motorized vehicles, on and off public highways. ASHRs exceeded 20.0 per 10,000 person-years at risk in high-percentage Aboriginal identity areas, compared with 8.6 in low-percentage Aboriginal identity areas.

ASHRs for unintentional injury due to each of the other causes were much lower, never surpassing 8.5 per 10,000 person-years at risk for children and youth in high-percentage Aboriginal identity areas, or 4.1 per 10,000 person-years at risk for those in low-percentage Aboriginal identity areas.

The higher ASHRs for unintentional injuries among children and youth in high-percentage Aboriginal identity areas are reflected in rate ratios (RRs). While overall RRs for unintentional injury hospitalization among those in highpercentage Aboriginal identity areas were a little more than 2.0, for some causes and for some high-percentage Aboriginal identity group areas, RRs approached or exceeded 3.0. This was the case for injuries due to land transportation (high-percentage Métis and identity areas) drowning/suffocation (all high-percentage Aboriginal identity

areas), cut/pierce (high-percentage First Nations and Métis identity areas), and poisoning (high-percentage First Nations identity areas). Moreover, RRs for hospitalizations due to injury from fire and natural environmental causes were 4.1 and 3.7, respectively, for children and youth in high-percentage First Nations identity areas, and close to 6.0 for injuries related to natural/environmental causes in high-percentage Inuit areas.

Hospitalization rates vary by age group

ASHRs for total unintentional injuries varied by age group and were generally higher among 10- to 19-year-olds than among children aged 0 to 9. For example, in high-percentage Métis identity areas, the ASHR at ages 10 to 19 was 100.4 per 10,000 person-years at risk, compared with 76.9 at ages 0 to 9 (Table 3). In low-percentage Aboriginal identity areas, the corresponding rates were 42.2 and 32.2. However, this general pattern masks considerable age differences by cause of injury.

For unintentional injuries due to falls, poisoning, fire, natural/environmental causes and drowning/suffocation, ASHRs were higher at ages 0 to 9 than at ages 10 to 19. By contrast, ASHRs for unintentional injuries due to land transportation, being struck, and being cut/pierced were higher at ages 10 to 19.

At ages 0 to 9, RRs compared with children in low-percentage Aboriginal identity areas were particularly high (approximately 3.0) for injuries due to land transportation (high-percentage First Nations and Métis identity areas), poisoning (high-percentage First Nations identity areas), fire (high-percentage First Nations identity areas), natural environment (high-percentage First Nations and Inuit identity areas), and drowning/suffocation (high-percentage First Nations and Métis identity areas).

At ages 10 to 19, RRs in highpercentage First Nations identity areas were 3.0 or more for hospitalization for unintentional injuries due to poisoning, being cut/pierced, fire, and natural/ environmental causes. RRs were

What is already known on this subject?

- Unintentional injury is the leading cause of death and morbidity among Canadian children.
- Studies of injury in Aboriginal populations have tended to focus on adults.
- Little is known about injury hospitalization rates among Aboriginal children.

What does this study add?

- For most causes of unintentional injury, rates of hospitalization are higher for children and youth in areas with a high-percentage (33% or more) of Aboriginal identity residents, compared with those in lowpercentage Aboriginal identity areas.
- Unintentional injury hospitalization rates for children and youth differed among high-percentage First Nations, Métis and Inuit identity areas.
- While hospitalization rates were higher among males, for many causes of unintentional injury, the disparity between high- and lowpercentage Aboriginal identity areas was greater among females.

also 3.0 or more for hospitalization for unintentional injury due to land transportation in high-percentage Métis and Inuit identity areas. As well, in high-percentage Inuit identity areas, the RR for hospitalization for unintentional injuries due to natural/environmental causes was 7.8.

Higher rate ratios for females

Regardless of whether an area was designated high- or low-percentage Aboriginal identity, ASHRs for unintentional injury were generally higher among males than females

Table 3
Age-standardized hospitalization rate and rate ratio for unintentional injury, by age group, cause of injury and Dissemination Area reporting Aboriginal identity, population aged 0 to 19, Canada (excluding Quebec), 2001/2002 to 2005/2006

		Age	es 0 to 9					Age	s 10 to 19			
Cause of injury and type of	Age-standardized rate (per 10,000 person-years	95% confid inter	ence val	Rate	95% confide interv	ence val	Age-standardized rate (per 10,000 person-years	95° confid inter	lence val	Rate	95% confide inter	ence val
Dissemination Area	at risk)	from	to	ratio	from	to	at risk)	from	to	ratio	from	to
Total High % First Nations High % Métis High % Inuit Low % Aboriginal	83.8* 76.9* 66.7* 32.2	81.0 69.2 58.5 31.9	86.7 85.4 76.1 32.5	2.6 2.4 2.1 1.0	2.5 2.1 1.8	2.7 2.7 2.4 	88.2* 100.4* 100.5* 42.2	91.8	91.2 109.8 112.0 42.6	2.1 2.4 2.4 1.0	2.0 2.2 2.1	2.2 2.6 2.7
Falls High % First Nations High % Métis High % Inuit Low % Aboriginal	33.9* 33.1* 24.8* 15.5	32.1 28.2 20.0 15.3	35.7 38.9 30.8 15.7	2.2 2.1 1.6 1.0	2.1 1.8 1.3	2.3 2.5 2.0	23.2* 24.6* 21.2* 13.3	21.7 20.5 16.7 13.1	29.5	1.7 1.9 1.6 1.0	1.6 1.5 1.3	1.9 2.2 2.0
Land transportation High % First Nations High % Métis High % Inuit Low % Aboriginal	11.9* 12.8* 10.1* 4.0	10.8 9.9 7.2 3.9	13.0 16.6 14.2 4.1	3.0 3.2 2.5 1.0	2.7 2.5 1.8	3.3 4.2 3.6	31.9* 40.3* 45.0* 13.5	30.2 35.0 38.3 13.3	46.4 53.0	2.4 3.0 3.3 1.0	2.2 2.6 2.8	2.5 3.4 3.9
Motor vehicle traffic High % First Nations High % Métis High % Inuit Low % Aboriginal	4.7* 3.8* 3.4* 1.6	4.1 2.3 1.9 1.5	5.5 6.0 6.1 1.7	2.9 2.3 2.1 1.0	2.5 1.4 1.2	3.4 3.8 3.8	15.8* 14.7* 14.9* 6.6	14.6 11.7 11.2 6.4	17.1 18.6 19.8 6.7	2.4 2.2 2.3 1.0	2.2 1.8 1.7	2.6 2.8 3.0
Struck High % First Nations High % Métis High % Inuit Low % Aboriginal	4.9* 4.9* 3.6 2.1	4.2 3.2 2.1 2.1	5.6 7.4 6.4 2.2	2.3 2.3 1.7 1.0	2.0 1.5 1.0	2.6 3.5 3.0	9.4* 12.4* 8.7 6.3	8.5 9.6 6.0 6.2	10.4 16.0 12.6 6.4	1.5 2.0 1.4 1.0	1.3 1.5 1.0	1.7 2.5 2.0
Poisoning High % First Nations High % Métis High % Inuit Low % Aboriginal	9.1* 7.3* 5.4* 2.7	8.2 5.2 3.5 2.6	10.1 10.3 8.5 2.8	3.4 2.7 2.0 1.0	3.1 1.9 1.3	3.8 3.9 3.2	4.9* 2.3 3.7* 1.5	4.2 1.3 2.1 1.5	5.6 4.1 6.5 1.6	3.2 1.5 2.4 1.0	2.7 0.8 1.4	3.7 2.7 4.3
Cut/Pierce High % First Nations High % Métis High % Inuit Low % Aboriginal	2.0* 3.3* 2.1* 0.7	1.6 2.0 1.0 0.7	2.4 5.5 4.4 0.8	2.7 4.5 2.8 1.0	2.1 2.7 1.3	3.4 7.5 6.0	4.5* 3.8* 3.1* 1.4	3.9 2.4 1.7 1.3	5.2 6.0 5.8 1.4	3.3 2.7 2.3 1.0	2.8 1.7 1.2	3.8 4.4 4.2
Fire High % First Nations High % Métis High % Inuit Low % Aboriginal	5.0* 3.5* 2.3 1.3	4.4 2.2 1.2 1.2	5.8 5.8 4.6 1.4	3.9 2.7 1.8 1.0	3.4 1.7 0.9	4.5 4.5 3.6	2.7* 1.3 x 0.6	2.2 0.6 0.6	3.3 2.8 0.6	4.5 2.1 x 1.0	3.7 0.9 	5.6 4.7
Natural/Environmental High % First Nations High % Métis High % Inuit Low % Aboriginal	4.2* 2.2* 5.4* 1.2	3.6 1.2 3.4 1.1	4.9 4.1 8.5 1.2	3.7 1.9 4.7 1.0	3.1 1.0 2.9	4.3 3.6 7.4	2.5* 1.9* 5.3* 0.7	2.1 1.0 3.3 0.6	3.1 3.6 8.5 0.7	3.7 2.7 7.8 1.0	3.0 1.4 4.8	4.6 5.3 12.6
Drowning/Suffocation High % First Nations High % Métis High % Inuit Low % Aboriginal	2.8* 2.7* 2.1* 0.9	2.3 1.5 1.0 0.9	3.3 4.7 4.4 1.0	3.1 2.9 2.3 1.0	2.5 1.7 1.1	3.7 5.2 4.9	0.3 x x 0.2	0.2 0.2	0.6 0.2	1.4 x x 1.0	0.8	2.6
Other High % First Nations High % Métis High % Inuit Low % Aboriginal	10.6* 7.5* 10.8* 4.1	9.6 5.4 7.9 4.0	11.6 10.5 15.0 4.2	2.6 1.8 2.6 1.0	2.3 1.3 1.9	2.8 2.6 3.6	9.4* 14.2* 12.5* 5.1	8.5 11.2 9.1 5.0	10.5 18.0 17.0 5.2	1.9 2.8 2.5 1.0	1.7 2.2 1.8	2.1 3.6 3.4

^{*} significantly different from low-percentage Aboriginal identity Dissemination Areas

Notes: Dissemination Areas where at least 33% of the population reported Aboriginal identity are classified as high-percentage Aboriginal identity. Classification as high-percentage First Nations, Métis or Inuit identity is based on the predominant group.

Source: 2001/2002 to 2005/2006 Hospital Morbidity Database.

^{...} not applicable

 $[\]boldsymbol{x}$ suppressed to meet confidentiality requirements of Statistics \boldsymbol{Act}

Unintentional injury hospitalizations among children and youth in areas with a high percentage of Aboriginal identity residents: 2001/2002 to 2005/2006 • Research article

Table 4 Age-standardized hospitalization rate and rate ratio for unintentional injury, by sex, cause of injury and Dissemination Area reporting Aboriginal identity, population aged 0 to 19, Canada (excluding Quebec), 2001/2002 to 2005/2006

		N	lales					Fe	emales			
Cause of injury and type of	Age-standardized rate (per 10,000 person-years	95% confide interv	ence	Rate	95% confide interv	ence	Age-standardized rate (per 10,000 person-years	95° confid inter	ence	Rate	95% confide inter	ence
Disseminatin Area	at risk)	from	to	ratio	from	to	at risk)	from	to	ratio	from	to
Total High % First Nations High % Métis High % Inuit Low % Aboriginal	103.6* 107.7* 103.6* 47.6	100.5 98.8 93.2 47.2	117.4 115.1	2.2 2.3 2.2 1.0	2.1 2.1 2.0	2.2 2.5 2.4	67.4* 68.3* 61.6* 26.0	64.8 61.1 53.6 25.7	70.0 76.4 70.8 26.2	2.6 2.6 2.4 1.0	2.5 2.4 2.1	2.7 2.9 2.7
Falls High % First Nations High % Métis High % Inuit Low % Aboriginal	34.1* 35.3* 29.7* 17.9	32.3 30.4 24.3 17.6	35.9 41.1 36.2 18.1	1.9 2.0 1.7 1.0	1.8 1.7 1.4	2.0 2.3 2.0	23.1* 22.6* 16.3* 10.8	21.6 18.5 12.4 10.6	24.7 27.5 21.4 10.9	2.1 2.1 1.5 1.0	2.0 1.7 1.2	2.3 2.6 2.0
Land transportation High % First Nations High % Métis High % Inuit Low % Aboriginal	26.1* 33.5* 32.5* 11.5	24.5 28.8 26.9 11.3	27.7 39.1 39.2 11.6	2.3 2.9 2.8 1.0	2.1 2.5 2.3	2.4 3.4 3.4	16.8* 18.4* 21.3* 5.6	15.5 14.9 16.8 5.5	18.1 22.7 27.0 5.7	3.0 3.3 3.8 1.0	2.8 2.7 3.0	3.3 4.1 4.8
Motor vehicle traffic High % First Nations High % Métis High % Inuit Low % Aboriginal	10.9* 9.8* 9.5* 4.8	9.9 7.4 6.7 4.7	11.9 13.0 13.4 4.9	2.3 2.0 2.0 1.0	2.1 1.5 1.4	2.5 2.7 2.8	9.2* 8.3* 8.3* 3.2	8.3 6.1 5.7 3.1	10.2 11.3 12.2 3.2	2.9 2.6 2.6 1.0	2.6 1.9 1.8	3.3 3.6 3.9
Struck High % First Nations High % Métis High % Inuit Low % Aboriginal	10.4* 11.7* 10.1* 6.3	9.4 9.0 7.2 6.2	11.4 15.2 14.1 6.4	1.6 1.9 1.6 1.0	1.5 1.4 1.1	1.8 2.4 2.2	3.6* 5.2* 1.9 1.9	3.0 3.5 0.8 1.8	4.2 7.8 4.2 1.9	1.9 2.8 1.0 1.0	1.6 1.9 0.5	2.3 4.2 2.3
Poisoning High % First Nations High % Métis High % Inuit Low % Aboriginal	6.7* 4.2* 4.9* 2.2	5.9 2.7 3.1 2.2	7.5 6.6 7.9 2.3	3.0 1.9 2.2 1.0	2.6 1.2 1.4	3.4 2.9 3.5	7.5* 5.6* 4.2* 2.0	6.7 3.8 2.5 1.9	8.4 8.2 7.2 2.1	3.8 2.8 2.1 1.0	3.4 1.9 1.3	4.3 4.2 3.6
Cut/Pierce High % First Nations High % Métis High % Inuit Low % Aboriginal	4.7* 5.0* 3.9* 1.5	4.1 3.4 2.3 1.5	5.5 7.5 6.7 1.6	3.1 3.3 2.6 1.0	2.7 2.2 1.5	3.6 4.9 4.4	1.6* 2.0* x 0.5	1.2 1.0 0.5	2.0 3.9 0.6	2.9 3.7 x 1.0	2.2 1.9 	3.7 7.2
Fire High % First Nations High % Métis High % Inuit Low % Aboriginal	5.1* 2.6* 1.7 1.2	4.5 1.5 0.8 1.1	5.9 4.5 3.8 1.3	4.3 2.1 1.4 1.0	3.7 1.2 0.6	5.0 3.8 3.2	2.6* 2.3* x 0.7	2.2 1.3 0.7	3.2 4.3 0.8	3.7 3.3 x 1.0	3.0 1.7 	4.5 6.1
Natural/Environmental High % First Nations High % Métis High % Inuit Low % Aboriginal	3.9* 2.3* 5.0* 1.0	3.4 1.3 3.1 1.0	4.6 4.1 8.1 1.1	3.8 2.2 4.9 1.0	3.3 1.2 3.0	4.5 4.0 7.9	2.8* 1.8* 5.6* 0.8	2.3 0.9 3.5 0.8	3.4 3.6 8.9 0.9	3.4 2.2 6.9 1.0	2.8 1.1 4.3	4.2 4.4 10.9
Drowning/Suffocation High % First Nations High % Métis High % Inuit Low % Aboriginal	1.7* 1.3 1.8* 0.7	1.3 0.6 0.8 0.7	2.1 2.8 4.0 0.8	2.4 1.8 2.6 1.0	1.8 0.8 1.1	3.0 4.0 5.7	1.5* 2.1* 1.6* 0.4	1.1 1.1 0.6 0.4	1.9 4.0 3.7 0.5	3.5 4.9 3.6 1.0	2.7 2.5 1.5	4.6 9.5 8.8
Other High % First Nations High % Métis High % Inuit Low % Aboriginal	11.8* 12.5* 14.5* 5.6	10.7 9.7 11.0 5.5	12.9 16.0 19.2 5.8	2.1 2.2 2.6 1.0	1.9 1.7 1.9	2.3 2.8 3.4	8.2* 9.0* 8.6* 3.5	7.4 6.6 5.9 3.4	9.2 12.2 12.5 3.6	2.4 2.6 2.5 1.0	2.1 1.9 1.7	2.7 3.5 3.6

^{*} significantly different from low-percentage Aboriginal identity Dissemination Areas ... not applicable

... not applicable
x suppressed to meet confidentiality requirements of Statistics Act
Notes: Dissemination Areas where at least 33% of the population reported Aboriginal identity are classified as high-percentage Aboriginal identity. Classification as high-percentage First Nations, Métis or Inuit identity is based on the predominant group.

Source: 2001/2002 to 2005/2006 Hospital Morbidity Database.

(Table 4). Rate ratios, however, tended to be higher among females, indicating a greater difference compared with low-percentage Aboriginal areas. For instance, the RRs in high-percentage First Nations identity areas were 2.6 for females and 2.2 for males.

Rate ratios among females in high-percentage Aboriginal identity areas were particularly elevated (approximately 3.0 or more) for unintentional injury hospitalizations due to land transportation and drowning/suffocation. This was also the case for unintentional injury hospitalizations due to poisoning, cut/pierce, and fire in high-percentage First Nations and Métis identity areas, and due to natural/environmental causes in high-percentage First Nations and Inuit identity areas.

Discussion

This study reveals associations between the geographic areas where children and youth live and hospitalization for unintentional injury. Those in areas where at least 33% of residents reported Aboriginal identity were hospitalized at approximately twice the rate of their counterparts in low-percentage Aboriginal identity areas. studies based on provincial data11,12,16 had similar results. For instance, First Nations children in Western Canada were reported to have injury hospitalization rates 1.2 to 2.9 times higher than those of the general population.⁶ Similarly, First Nations children in Alberta were 1.4 times more likely to be hospitalized for an injury than were non-Aboriginal children.16

In the present analysis, ASHRs for injuries related to drowning/suffocation in high-percentage Aboriginal identity areas were approximately three times those in low-percentage Aboriginal identity areas. This is consistent with other research indicating that Aboriginal people are at increased risk of injuries due to drowning.¹⁷ Even so, the overall ASHR for injuries due to drowning/suffocation was low, compared with causes such as falls and land transportation.

Unintentional injury hospitalization rates by cause differed among high-percentage First Nations, Métis and Inuit identity areas. ASHRs for injuries due to natural/environmental causes were highest in high-percentage Inuit identity areas, possibly reflecting conditions and activities specific to northern areas. ASHRs for injuries due to fires were almost four times higher in high-percentage First Nations identity areas than in low-percentage Aboriginal identity areas, a finding consistent with earlier research.

The RRs comparing unintentional injury hospitalizations in high- and lowpercentage Aboriginal identity areas were often greater for females than males, a pattern reported in previous work.11 Thus, although males were more likely to be hospitalized, the difference between residents of high- and low-Aboriginal identity areas was greater among females. In particular, RRs for females were greater for hospitalization for injuries due to poisoning and land transportation in high-percentage First Nations identity areas. Also, RRs for young children in high-percentage First Nations identity areas were greater than those for 10- to 19-year-olds for most causes of unintentional injury hospitalization.

Strengths and limitations

A strength of this study is the use of five years of national population-based data to examine rates and types of unintentional injury hospitalization for children and youth in high- and low-percentage Aboriginal identity areas, and the provision of breakdowns by injury type, age group and sex.

However, some limitations warrant discussion. Because Aboriginal identifiers were not available on the HMDB, a geographical proxy was used to designate census DAs as high- or low-percentage Aboriginal identity. Therefore, this is an ecological study reporting results for geographic areas; the associations observed do not necessarily apply at the individual level. As well,

the geographic location where the injury occurred was not available.

A threshold of 33% was used to designate a DA as high-percentage Aboriginal identity.¹⁰ For highpercentage Inuit identity areas, the 33% cut-off results in the selection of DAs with an average of 93% Aboriginal identity residents younger than age 20 (Appendix Table A). For high-percentage First Nations and Métis identity areas, the 33% cut-off results in the selection of DAs with 86% and 67% Aboriginal identity residents younger than age 20.19 (A higher cut-point would have resulted in a substantial loss of high-percentage Métis identity areas.) Consequently, the findings are not representative of the First Nations, Métis, or Inuit identity populations in Canada.

High-percentage First Nations, Métis, and Inuit identity areas defined in this study differ in urban/rural location, population size, and socioeconomic characteristics—all factors that have been associated with injury rates. 19,20 For example, 100% of the population living in high-percentage Inuit identity DAs were in weak or non-Metropolitan-Influenced compared with 8% of the population in low-percentage Aboriginal identity DAs (Appendix Table A). Similarly, 27% of the population in high-percentage Inuit identity DAs lived in crowded dwellings, compared with 3% of the population in low-percentage Aboriginal identity areas. Also, information about individual and family characteristics such as income, education, and individual behaviours that may influence injury risk was not available on hospital records.

Previous research has found that injuries sustained among First Nations populations tend to be more severe. 8,12 Although the injuries included in this analysis were serious enough to result in hospitalization, the severity of those injuries was not assessed. And of course, injuries so severe that they resulted in death before hospital admission were not included in this analysis.

The results do not represent the entire country. Incomplete postal code

Unintentional injury hospitalizations among children and youth in areas with a high percentage of Aboriginal identity residents: 2001/2002 to 2005/2006 • Research article

information for Quebec meant that the province had to be excluded from the study.

Finally, counts for some causes of injury for some Aboriginal identity groups were small.

Conclusion

Elevated rates of unintentional injury hospitalization among children and youth in high-percentage Aboriginal identity areas, compared with those in low-percentage Aboriginal identity areas, prevailed for all causes of injury examined in this analysis. While falls and land transportation injury hospitalizations were the most common, regardless of Aboriginal identity status, disparities between hospitalization rates for high- and low-percentage Aboriginal

identity areas were often greatest for less frequent causes such as fire, natural/environmental causes, and drowning/suffocation. The extent of the difference in unintentional injury hospitalizations between high- and low-percentage Aboriginal identity areas varied, depending on whether the comparison was with high-percentage First Nations, Métis or Inuit identity areas.

References

- Public Health Agency of Canada. Leading causes of death, Canada, 2004, males and females combined: counts (crude death rate per 100,000). Ottawa: Public Health Agency of Canada, 2009.
- Public Health Agency of Canada. Leading causes of hospitalizations, Canada, 2004, males and females combined: counts (crude rate per 100,000). Ottawa: Public Health Agency of Canada, 2009.
- SMARTRISK. The Economic Burden of Injury in Canada. Toronto: SMARTRISK, 2009.
- Holbrook TL, Hoyt DB, Coimbra R, et al.
 Trauma in adolescents causes long-term marked deficits in quality of life: adolescent children do not recover preinjury quality of life or function up to two years postinjury compared to national norms. *Journal of Trauma* 2007; 62(3): 577-83.
- Smilie J, Adomako P. Indigenous Children's Health Report: Health Assessment in Action. Toronto: Centre for Research on Inner City Health, 2009.
- Health Canada. A Statistical Profile on the Health of First Nations in Canada, 2000. Ottawa: Health Canada, 2008.
- Adelson N. The embodiment of inequity. Health disparities in Aboriginal Canada. Canadian Journal of Public Health 2005; 96 (Suppl. 2): S45-61.

- Karmali A, Laupland K, Robert Harrop A, et al. Epidemiology of severe trauma among status Aboriginal Canadians: a population-based study. Canadian Medical Association Journal 2005; 172(8): 1007-11.
- Tjepkema M, Wilkins R, Senécal S, et al. Mortality of Métis and Registered Indian adults in Canada: An 11-year follow-up study. Health Reports 2009; 20: 31-51.
- Carrière G, Garner R, Sanmartin C. Acute-care Hospitalizations and Aboriginal Identity in Canada, 2001/2002 (Statistics Canada, Catalogue 82-622-x) Ottawa: Statistics Canada, 2010. Available at: http://www.statcan.gc.ca/pub/82-622-x/82-622-x2010005-eng.pdf.
- Fantus D, Shah B, Qui Feng, Hux J, Rochon P. Injury in First Nations communities in Ontario. *Canadian Journal of Public Health* 2009; 100(4): 258-62.
- Alaghehbandan R, Sikdar KC, MacDonald D, et al. Unintentional injuries among children and adolescents in aboriginal and non-aboriginal communities, Newfoundland and Labrador, Canada. *International Journal* of Circumpolar Health 2010; 69(1): 61-71.
- 13. Public Health Agency of Canada. *Injury Surveillance On-Line: ICD10-ICD9 Transition Matrix. Report.* 2008. Available at: http://dsol-smed.phac-aspc.gc.ca/dsol-smed/is-ib/chirpp/ICD10-ICD9Transition-MatrixISOL.pdf.

- 4. Wilkins R. PCCF + Version 4G User's Guide: Automated Geographic Coding Based on the Statistics Canada Postal Code Conversion Files Including Postal Codes to October 2005 (Statistics Canada, Catalogue 82F0086-XDB) Ottawa: Statistics Canada, 2006.
- Wilkins R, Uppal S, Finès, P, et al. Life expectancy in the Inuit-inhabited areas of Canada, 1989 to 2003. *Health Reports* 2008; 19(1): 1-13.
- Spady DW, Saunders DL, Schopflocher DP, Svenson LW. Patterns of injury in children: A population-based approach. *Pediatrics* 2004; 113(3): 522-9.
- 17. Health Canada. *Unintentional and Intentional Injury Profile for Aboriginal People in Canada*. Ottawa: Health Canada, 2001.
- Zukewich N, O'Donnell V. Aboriginal Children's Survey, 2006: Family, Community and Child Care (Statistics Canada, Catalogue 89-634-XWE2008001) Ottawa: Statistics Canada, 2008.
- Oliver LN, Kohen DE. Neighbourhood variation in hospitalization for unintentional injury among children and teenagers. *Health Reports* 2010; 21(4): 9-17.
- Oliver L, Kohen D. Neighbourhood income gradients in hospitalisations due to motor vehicle traffic incidents among Canadian children. *Injury Prevention* 2009; 15(3): 163-9.

Appendix

Table A
Demographic and socio-economic census characteristics, by Aboriginal identity group in
Dissemination Area, Canada (excluding Quebec), 2001

Characteristics	Low % Aboriginal	High % First Nations	High % Métis	High % Inuit
Total population	19,137,200	338,500	49,000	31,800
Aboriginal identity population (%)	3	77	57	87
Population aged 0 to 19 (%)	30	45	41	48
Population aged 0 to 19 of Aboriginal identity (%)	4	86	67	93
Population in weak or non-Metropolitan-Influenced Zone (%)	8	66	70	100
Population without secondary graduation (%)	30	53	45	52
Population not in labour force (%)	30	42	31	33
Population living in crowded dwellings (%)	3	21	10	27
Population in dwellings in need of major repair (%)	8	32	20	23
Average household income per person (\$)	26,381	12,878	16,737	21,123

Notes: Dissemination Areas where at least 33% of the population reported Aboriginal identity are classified as high-percentage Aboriginal identity. Classification as high-percentage First Nations, Métis or Inuit is based on the predominant group.

Source: 2001 Census of Canada.

ELECTRONIC PUBLICATIONS AVAILABLE AT

www.statcan.gc.ca

Mortality rates among children and teenagers living in Inuit Nunangat, 1994 to 2008

by Lisa N. Oliver, Paul A. Peters and Dafna E. Kohen

Abstract

Background

Because Vital Statistics data do not include information on Inuit identity in all jurisdictions, mortality rates cannot be calculated specifically for Inuit. However, Inuit in Canada are geographically concentrated—78% live in Inuit Nunangat, and 82% of the area's total population identify as Inuit. While there are limitations, geographic approaches can be employed to calculate mortality for the population of that area.

Data and methods

The Vital Statistics Database (1994 to 2008) and population estimates were used to calculate age-standardized mortality rates (ASMRs) in five-year intervals around the 1996 and 2006 Census years. Mortality rates were calculated for 1- to 19-year-olds living in Inuit Nunangat and those living elsewhere in Canada.

Results

The ASMR in 2004-2008 for 1- to 19-year-olds in Inuit Nunangat was 188.0 deaths per 100,000 person-years at risk, five times the rate (35.3) elsewhere in Canada. The disparity had not narrowed over the previous decade. In Inuit Nunangat, injuries were responsible for 64% of deaths of children and teenagers, compared with 36% in the rest of Canada.

Interpretation

The persistently high mortality rates for children and teenagers living in Inuit Nunangat, compared with the rest of Canada, are important in understanding the health and socio-economic situation of residents of this region.

Keywords

Aboriginal, age-standardized mortality rates, child health, death rates, suicide, vital statistics, wounds and injuries

Authors

Lisa N. Oliver (lisa.oliver@statcan.gc.ca) is with the Research Data Centre at Simon Fraser University and Paul A. Peters (1-613-951-0616; paul.a.peters@statcan.gc.ca) and Dafna E. Kohen (1-613-951-3346; dafna.kohen@statcan.gc.ca) are with the Health Analysis Division at Statistics Canada, Ottawa, Ontario K1A 0T6.

A number of recent studies have examined life expectancy, mortality, hospitalization, and other health indicators for the four Inuit Nunangat land claim regions. Life expectancy at birth for residents of that area is 6 to 11 years less than for people in the rest of Canada, and the infant mortality rate is higher. To date, child and youth mortality rates for residents of Inuit Nunangat have not been calculated.

Because Vital Statistics data do not include information on Inuit identity in all jurisdictions, mortality rates cannot be calculated specifically for Inuit. However, Inuit in Canada are geographically concentrated—78% live in Inuit Nunangat, and 82% of the area's total population identify as Inuit (Table 1). While there are limitations, geographic approaches can be employed to calculate mortality for the population of that area.

Inuit Nunangat is the Inuktitut term for "Inuit homeland," an expanse comprising more than one-third of Canada's land mass, which extends from northern Labrador to the Northwest Territories. It consists of the four Inuit land claim regions: Nunatsiavut (Northern coastal Labrador), Nunavik (Northern Quebec), the territory of Nunavut, and the

Inuvialuit Settlement Region (Northwest Territories and Yukon).

The population of Inuit Nunangat is fast-growing and young. Between 1996 and 2006, the population of the area rose by 14%, compared with an 8% increase in the rest of Canada (Table 1). The primary driver of the increase in Inuit Nunangat was high growth among Inuit, at 18%. The result was a much younger population age structure than that of most other populations in Canada.⁷ In 2006, 42% of the population in Inuit Nunangat were aged 1 to 19, compared with 24% of the population elsewhere in Canada.

This study examines disparities⁸ in mortality between 1- to 19-year-old residents of Inuit Nunangat and the rest of Canada from 1994 to 2008. Mortality rates are calculated by cause of death.

Table 1
Total, non-Aboriginal and Inuit population, by residence, Canada, 2006

				Inuit su	Inuit subregions			
	Outside Inuit Nunangat	Inuit Nunangat	Inuvialuit Settlement Region	Nunavut	Nunavik	Nunatsiavut		
Total population	30,071,225	48,015	5,705	29,325	10,570	2,415		
% aged 1 to 19	24	42	34	43	45	36		
% change since 1996	8	14	-2	19	21	-14		
Non-Aboriginal	30,060,225	7,065	1,520	4,410	920	215		
% of total population	96	15	26	15	9	9		
% aged 1 to 19	23	15	18	16	11	12		
% change since 1996	8	4	-4	11	-2	-28		
Inuit % Inuit % aged 1 to 19 % change since 1996	11,000	39,475	3,115	24,635	9,565	2,160		
	0	82	54	84	89	89		
	39	47	40	47	48	39		
	62	18	-3	20	25	3		

Source: 2006 Census of Population.

Methods

A geographic approach was employed to estimate mortality rates for children and teenagers in Inuit Nunangat and in the rest of Canada. A similar technique has been used to examine life expectancy, mortality, hospitalization, cancer incidence, and crime in this region. 1,5,9 Deaths of children younger than age 1 are considered to be infant mortality, and so are not included in this analysis.

A large majority (92%) of the population aged 1 to 19 in Inuit Nunangat identified as Inuit on the 2006 Census (data not shown). At more than 96%, Nunavik and Nunatsiavut had the highest percentages reporting Inuit identity; 65% of the population aged 1 to 19 in the Inuvialuit Region reported Inuit identity.

Each death record in the Vital Statistics Deaths Database contains the decedent's sex, age and usual place of residence (Census Subdivision - CSD). A single underlying cause of death is recorded, based on the International Classification of Diseases 9th (1994 to 1999) and 10th (2000 onwards) revisions (ICD-9 and ICD-10).

The analyses in this study are based on records for people who were aged 1 to 19 when they died. It excludes deaths of non-residents of Canada, deaths of residents of Canada whose province or territory of residence was unknown, and death records lacking the decedent's

age or sex. The approach used for this analysis required that death records include geographic identifiers. For the Inuvialuit Region and Nunatsiavut, the complete CSD code was necessary; for Nunavik, either the CSD or Census Division code was needed; and for Nunavut, only the territory code.

The underlying causes of death, based on ICD-9 or ICD-10 codes, were grouped according to the Global Burden of Disease (GBD) classification.¹⁰ The GBD framework differs from the ICD chapters, which categorize diseases by body systems. According to the GBD classification, deaths fall into three main groups: I - communicable, maternal, perinatal, and nutritional conditions; II - non-communicable diseases: and III - injuries. Groups I and II are further classified into specific diseases or conditions. Group III is subdivided into unintentional or intentional injuries and then classified by type. Because self-inflicted injury (suicide) accounted for nearly the entire intentional injury category in this study, for reasons of confidentiality, homicide could not be reported separately. Additional information on the GBD classification, the ICD-9 and ICD-10 codes, and the use of the GBD in Inuit Nunangat is available elsewhere.10,11

Small-area population estimates were used to provide detailed counts of those aged 1 to 19 for each CSD in Inuit

Nunangat from 1996 through 2006.¹² These estimates are more accurate than those available on the census. The closest population estimates available were used as a proxy. Person-years at risk were calculated by aggregating the five years surrounding each census year (1996 and 2006).

Age-standardized mortality (ASMRs) (deaths per 100,000 personyears at risk) were calculated using the method of Chiang¹³ in five-year intervals around the 1996 and 2006 Census years (1994 to 1998 and 2004 to 2008) in order to obtain the minimum counts needed to produce rates. The total number of deaths in each five-year interval was divided by the person-years at risk. Mortality rates were age-standardized to the estimated 2001 population age structure of Inuit Nunangat. The Spiegelman method¹⁴ was used to calculate 95% confidence intervals. Rates were also calculated for each of the four regions of Inuit Nunangat. Rate ratios were calculated in order to compare Inuit Nunangat with the rest of Canada.

Results

Overall mortality rates

In 2004-2008, the ASMR at ages 1 to 19 in Inuit Nunangat was 188.0 deaths per 100,000 person-years at risk; this compared with 35.3 deaths per 100,000 in the rest of Canada (Table 2). For both populations, these ASMRs marked a decline from 1994-1998, when the rates had been 210.1 deaths per 100,000 person-years at risk in Inuit Nunangat, and 44.7 in the rest of Canada. The simultaneous decreases in ASMRs meant that the rate ratio comparing children and teenagers in Inuit Nunangat with the rest of Canada did not change significantly over the two periods.

The high overall ASMRs for Inuit Nunangat prevailed across subregions. For instance, in 2004-2008, rates were 152.5 deaths per 100,000 person-years at risk in Nunavut, 307.8 in Nunavik, and 269.1 in Nunatsiavut. Because of small numbers, rates for the Inuvialuit Region were not calculated.

Table 2
Age-standardized mortality rates (ASMR) per 100,000 person-years at risk, by sex, population aged 1 to 19, Canada,
Inuit Nunangat, and Inuit subregions 1994 to 1998 and 2004 to 2008

	C	anada		Inuit N		ngat Nunavut Nunavik		Nunangat Nunavut			Nunavut		Nui	natsiav	ut
		95 confid inter	lence		95 confid inte	dence		95 confic inte	lence		95 confid inte			confi	5% dence erval
Sex/Years	ASMR	from	to	ASMR	from	to	ASMR	from	to	ASMR	from	to	ASMR	from	to
Both sexes															
1994 to 1998	44.7	44.0	45.4	210.1	182.6	241.7	211.2	176.3	252.9	262.4	200.1	344.1	248.7	147.2	420.1
2004 to 2008	35.3	34.7	35.9	188.0	163.9	215.7	152.5	125.3	185.6	307.8	245.3	386.3	269.1	177.9	407.0
Males															
1994 to 1998	53.2	52.2	54.2	258.2	215.9	308.8	233.7	183.4	297.7	377.1	274.0	519.1	324.0	168.5	623.0
2004 to 2008	41.7	40.7	42.6	244.5	206.6	289.5	210.4	166.7	265.7	366.9	274.6	490.3	348.8	209.0	582.1
Females															
1994 to 1998	35.7	34.9	36.6	161.2	128.6	202.0	188.2	143.7	246.6	146.4	87.8	244.0	173.9	72.3	418.8
2004 to 2008	28.6	27.8	29.4	129.5	102.2	164.2	91.8	63.7	132.3	247.1	171.5	356.1	187.2	92.1	380.5

[†] excludes residents of Inuit Nunangat

Note: Because of small numbers, rates were not calculated for the Inuvialuit Settlement Region.

Source: Custom population estimates, Demography Division; Vital Statistics Database.

Cause of death

Communicable diseases

Communicable diseases are conditions such as tuberculosis and respiratory infections. In 2004-2008, the ASMR due to communicable diseases for 1- to 19-year-olds was 35.6 deaths per 100,000 person-years at risk in Inuit Nunangat, compared with 9.9 elsewhere in Canada (Table 3). The rate ratio was almost unchanged since 1994-1998, remaining around 3.6 (Table 4).

Non-communicable diseases

Non-communicable diseases are conditions such as cancer, congenital anomalies and neurologic diseases. In 2004-2008, the ASMR for deaths due to non-communicable diseases among children and teenagers was 22.4 deaths per 100,000 person-years at risk in Inuit Nunangat and 12.0 elsewhere in Canada. Since 1994-1998, the ASMRs for non-communicable diseases had fallen in both Inuit Nunangat and the rest of Canada, so the rate ratio was relatively constant over the period at about 2.0.

Injuries

Injuries were the largest contributor to mortality of children and teenagers, accounting for a much larger share of deaths in Inuit Nunangat than in the rest of Canada: 64% versus 36% in 2004-2008 (data not shown). The ASMR for all injuries combined was 115.3 deaths per 100,000 person-years at risk in Inuit Nunangat, compared with 10.9 elsewhere in Canada.

Deaths due to *unintentional* injuries are those in which there was no intent to harm (for example, accidental motor vehicle collisions, unintentional drownings). Deaths due to *intentional* injuries refer to suicide (self-inflicted) and homicide.

In 2004-2008, in Inuit Nunangat, the ASMR per 100,000 person-years at risk was 40.4 deaths for *unintentional* injuries and 74.9 deaths for *intentional* injuries. Rates were much lower elsewhere in Canada, and the rate was higher for *unintentional* (7.8) than for *intentional* injuries (3.1). Since 1994-1998, the rate ratios had not changed significantly.

For children and teenagers in both Inuit Nunangat and the rest of Canada, the majority of deaths due to *intentional* injuries were self-inflicted, that is, suicides; ASMRs for homicide could not be reported because of small numbers. Suicides accounted for a much larger share of all deaths of young people in Inuit Nunagat than elsewhere in Canada: 40% versus 8%. 15-18

The suicide rate of girls and young women in Inuit Nunangat was

approximately 40 deaths per 100,000 person-years at risk from 1994-1998 to 2004-2008, compared with around 2 deaths per 100,000 person-years at risk in the rest of Canada. The rate ratios show suicide rates for girls and young women in Inuit Nunangat to be more than 20 times those in the rest of Canada.

Among boys and young men in Inuit Nunangat, the suicide rate was 77.2 deaths per 100,000 person-years at risk in 1994-1998 and 101.6 in 2004-2008, rates that were not statistically different. By contrast, the suicide rate among boys and young men in the rest of Canada fell significantly from 6.1 to 4.2 deaths per 100,000 person-years at risk. As a result, the suicide rate ratio rose from 15 to 35.

From 1994-1998 to 2004-2008, the percentage of suicides due to hanging/suffocation rose among children and teenagers in Inuit Nunangat (from 70% to 85%) and also in the rest of Canada (from 55% to 72%). Rate ratios for suicides from hanging/suffocation were 38 times higher in Inuit Nunangat than elsewhere in Canada, and those for suicides due to firearms, 51 times higher.

Discussion

Mortality rates among 1- to 19-yearolds in Inuit Nunangat declined since 1994-1998, but so have rates in the rest

Table 3 Age-standardized mortality rates (ASMR) per 100,000 person-years at risk, by sex and cause of death, population aged 1 to 19, Canada † and Inuit Nunangat, 1994 to 1998 and 2004 to 2008

			1994 t	o 1998					2004 t	o 2008		
	C	anada		Inuit	Inuit Nunangat			anada		Inuit	Nunang	gat
	95% confidence interval		95% confidence interval				95% confidence interval			95% confidence interval		
Sex/Cause of death	ASMR	from	to	ASMR	from	to	ASMR	from	to	ASMR	from	to
Both sexes - All causes	44.7	44.0	45.4	210.1	182.6	241.7	35.3	34.7	35.9	188.0	163.9	215.7
Group I: Communicable diseases	9.4	9.1	9.7	34.3	24.5	48.0	9.9	9.6	10.3	35.6	25.6	49.3
Infectious and parasitic	1.0	0.9	1.1	5.2	2.1	12.4	0.8	0.7	0.9	9.1	4.8	17.6
Group II: Non-communicable diseases	15.7	15.3	16.1	36.3	26.0	50.6	12.0	11.7	12.4	22.4	14.8	33.7
Congenital anomalies	6.9	6.6	7.2	16.4	10.0	26.9	4.7	4.5	5.0	11.5	6.5	20.2
Group III: Injuries	15.5	15.1	15.9	109.6	89.9	133.7	10.9	10.6	11.2	115.3	97.1	136.9
Unintentional	11.2	10.9	11.6	46.8	34.7	63.2	7.8	7.5	8.0	40.4	30.0	54.4
Road traffic	6.6	6.3	6.8	9.6	5.0	18.5	4.3	4.1	4.5	7.7	3.8	15.5
Drownings	1.0	0.9	1.1	Χ			0.7	0.6	0.8	7.4	3.7	14.8
Intentional	4.2	4.0	4.4	62.8	48.2	81.8	3.1	3.0	3.3	74.9	60.7	92.5
Self-inflicted	3.3	3.1	3.5	58.5	44.5	77.0	2.2	2.1	2.3	72.1	58.2	89.3
Firearm	0.8	0.7	0.9	16.1	9.5	27.1	0.2	0.1	0.2	9.5	5.2	17.1
Hanging/Suffocation	1.8	1.7	2.0	41.3	29.8	57.3	1.6	1.5	1.7	60.9	48.2	76.9
Males - All causes	53.2	52.2	54.2	258.2	215.9	308.8	41.7	40.7	42.6	244.5	206.6	289.5
Group I: Communicable diseases	10.4	9.9	10.8	40.1	25.8	62.2	10.9	10.4	11.4	40.8	26.6	62.7
Infectious and parasitic	1.1	0.9	1.2	Χ			0.8	0.7	1.0	13.9	6.6	29.2
Group II: Non-communicable diseases	16.9	16.3	17.5	41.3	26.6	64.2	13.2	12.7	13.7	18.8	10.1	35.0
Congenital anomalies	7.5	7.2	7.9	22.7	12.5	41.0	4.9	4.6	5.2	12.8	6.1	27.0
Group III: Injuries	20.9	20.3	21.5	147.2	115.6		14.5	14.0	15.0	163.5	133.5	
Unintentional	14.8	14.3	15.3	68.1	47.9	97.0	10.3	9.9	10.7	58.2	41.1	82.4
Road traffic	8.3	7.9	8.7	16.8	8.4		5.5	5.2	5.8	11.6	5.2	25.8
Drownings	1.5	1.3	1.7	X			1.1	1.0	1.3	12.8	6.1	27.0
Intentional	6.1	5.8	6.5	79.1		110.2	4.2	4.0	4.5	105.3		135.1
Self-inflicted	5.0	4.7	5.3	77.2		108.0	2.9	2.7	3.1	103.5		131.0
Firearm	1.5	1.3	1.6	22.7	12.2		0.4	0.3	0.4	18.7	10.3	33.8
Hanging/Suffocation	2.6	2.4	2.9	54.5		81.4	2.0	1.8	2.2	81.3		108.0
Females - All causes	35.7	34.9	36.6	161.2	128.6	202.0	28.6	27.8	29.4	129.5	102.2	164.2
Group I: Communicable diseases	8.3	7.9	8.8	28.4	16.8	48.1	8.8	8.4	9.3	30.1	18.1	50.0
Infectious and parasitic	0.9	0.8	1.1	X			0.7	0.6	0.9	X		
Group II: Non-communicable diseases	14.3	13.8	14.9	31.2	18.8	51.9	10.9	10.4	11.4	26.1	15.1	44.9
Congenital anomalies	6.2	5.8	6.6	10.2	4.2		4.5	4.2	4.9	10.1	4.2	24.3
Group III: Injuries	9.8	9.3	10.2	71.5		101.1	7.1	6.7	4.9 7.4	65.6	47.5	90.6
Unintentional	9.o 7.5	7.2	7.9	25.3	14.4	44.7	5.1	4.8	7.4 5.4	22.0	12.5	38.8
	7.5 4.7	4.5	7.9 5.1	25.3 X			3.1	2.9	3.3			
Road traffic				X					3.3 0.4	X		•••
Drownings	0.5	0.4	0.7			 74 F	0.3	0.2		X		C4.5
Intentional	2.2	2.0	2.4	46.1	29.7	71.5	2.0	1.8	2.2	43.6	29.4	64.5
Self-inflicted	1.5	1.3	1.6	39.5	24.5	63.5	1.5	1.3	1.6	41.6	27.9	62.1
Firearm	0.1	0.1	0.2	Х			X			Х		
Hanging/Suffocation	1.0	0.9	1.1	27.9	15.8	49.1	1.2	1.1	1.3	39.9	26.5	60.1

[†] excludes residents of Inuit Nunangat

Source: Custom population estimates, Demography Division; Vital Statistics Database.

of Canada. Consequently, the rate ratio remained approximately five times higher throughout the decade.

The greatest disparity was for injuries, with rate ratios in 2004-2008 about 10 times higher among children and

teenagers in Inuit Nunangat than in the rest of Canada.

In 2004-2008, children and teenagers in Inuit Nunangat were more than 30 times as likely to die from suicide as were those in the rest of Canada. Similarly high suicide rates have been reported for

the total population in Inuit regions. 19,20 Half of all deaths of young people in Inuit Nunangat were suicides, compared with approximately 10% in the rest of Canada.

While rate ratios were highest for injuries, disparities also emerged in

^{...} not applicable

X suppressed to meet confidentiality requirements of the Statistics Act

Table 4 Rate ratios for age-standardized mortality rates, by sex and cause of death, population aged 1 to 19, Inuit Nunangat compared with Canada, † 1994 to 1998 and 2004 to 2008

	199	4 to 19	98	20	2004 to 2008				
	Rate	95% confide interv	ence	Rate	95% confide inter	ence			
Sex/Cause of death	ratio	from	to	ratio	from	to			
Both sexes - All causes	4.7	4.1	5.4	5.3	4.6	6.1			
Group I: Communicable diseases	3.7	2.6	5.1	3.6	2.6	5.0			
Infectious and parasitic	5.2	2.1	12.5	11.7	6.0	22.8			
Group II: Non-communicable diseases	2.3	1.7	3.2	1.9	1.2	2.8			
Congenital anomalies	2.4	1.5	3.9	2.4	1.4	4.3			
Group III: Injuries	7.1	5.8	8.7	10.6	8.9	12.6			
Unintentional	4.2	3.1	5.6	5.2	3.9	7.0			
Road traffic	1.5	0.8	2.8	1.8	0.9	3.6			
Drownings	X			10.1	5.0	20.5			
Intentional	14.9	11.4	19.4	24.1	19.4	29.9			
Self-inflicted	17.8	13.5	23.6	32.8		40.9			
Firearm	19.6	11.5	33.4	51.3	27.4	96.1			
Hanging/Suffocation	22.5		31.5	38.0		48.5			
Males - All causes	4.9	4.1	5.8	5.9	5.0	7.0			
Group I: Communicable diseases	3.9	2.5	6.0	3.7	2.4	5.7			
Infectious and parasitic	Χ			16.9	7.9	36.0			
Group II: Non-communicable diseases	2.4	1.6	3.8	1.4	0.8	2.7			
Congenital anomalies	3.0	1.7	5.5	2.6	1.2	5.5			
Group III: Injuries	7.0	5.5	9.0	11.3	9.2	13.8			
Unintentional	4.6	3.2	6.6	5.6	4.0	8.0			
Road traffic	2.0	1.0	4.1	2.1	0.9	4.7			
Drownings	Χ			11.5	5.4	24.4			
Intentional	12.9	9.2	18.1	25.0	19.3	32.4			
Self-inflicted	15.4	10.9	21.7	35.0	26.9	45.6			
Firearm	15.5	8.3	29.2	52.4	28.0	98.2			
Hanging/Suffocation	20.6		31.1	40.9	30.4	55.0			
Females - All causes	4.5	3.6	5.7	4.5	3.6	5.7			
Group I: Communicable diseases	3.4	2.0	5.8	3.4	2.0	5.7			
Infectious and parasitic	X			X					
Group II: Non-communicable diseases	2.2	1.3	3.6	2.4	1.4	4.1			
Congenital anomalies	1.6	0.7	4.0	2.2	0.9	5.4			
Group III: Injuries	7.3	5.2	10.4	9.3	6.7	12.9			
Unintentional	3.4	1.9	6.0	4.3	2.4	7.6			
Road traffic	X			X					
Drownings	X			X					
Intentional	20.8	13.3	32.5	22.2	14.8	33.3			
Self-inflicted	27.2	16.7	44.3	28.4	18.8	43.1			
Firearm									
Hanging/Suffocation	28.3	15.8	50.7	33.4	21.8	51.2			

[†] excludes residents of Inuit Nunangat

Source: Custom population estimates, Demography Division; Vital Statistics Database.

mortality rates due to communicable diseases. In 2004-2008, children and teenagers in Inuit Nunangat were 3.6 times more likely to die from communicable diseases than were those elsewhere in Canada consistent with other evidence.²¹ As well, throughout the

decade, children and teenagers in Inuit Nunangat were approximately twice as likely to die due to non-communicable diseases, compared with those in the rest of Canada.

Overall, whether they lived in Inuit Nunangat or in the rest of Canada, girls

What is already known on this subject?

- Life expectancy at birth for residents of Inuit Nunangat is 6 to 11 years less than that of people in the rest of Canada.
- In 2006, the overall mortality rate of residents of Inuit Nunangat was double that of Canada as a whole.

What does this study add?

- This study provides mortality rates by detailed cause of death for children and teenagers aged 1 to 19 living in Inuit Nunangat, compared with those living elsewhere in Canada for two five-year periods: 1996 (1994-1998) and 2006 (2004-2008).
- Age-standardized mortality rates were higher for children and teenagers in Inuit Nunangat compared with the rest of Canada in both 1994-1998 and 2004-2008.
- Injuries accounted for the largest component of mortality among children and teenagers in Inuit Nunangat.
- In 2004-2008, age-standardized suicide rates were up to 30 times higher among children and teenagers in Inuit Nunangat than in the rest of Canada.

and young women had lower ASMRs than did boys and young men.

Limitations

While the use of national death records to calculate mortality rates for Inuit Nunangat is a strength of this study, this data source has inherent limitations. Because Vital Statistics data do not contain Inuit identifiers, this analysis used a geographic approach to produce mortality rates for Inuit regions rather than for the Inuit population per se. Thus, the rates presented here are not representative of all Inuit in Canada.

^{...} not applicable

X suppressed to meet confidentiality requirements of Statistics Act

As well, the Vital Statistics data record only the primary cause of death, although it is possible that some deaths had multiple underlying causes.

The mortality rates for youth living in Inuit Nunangat are based on small populations and very small numbers of deaths. Confidence intervals are wide and frequently overlap. As a result, what may seem to be substantial differences can be based on few events, and apparently large changes are not significant. In fact, low numbers prevented separate calculations of ASMRs by cause of death for the four Inuit Nunangat land claims regions, and the number of deaths in the Inuvialuit Settlement Region was so low that separate ASMRs could not be calculated.

Conclusion

The findings show higher mortality rates for children and teenagers in Inuit Nunangat, compared with the rest of Canada. The overall mortality rate in 2004-2008 was about five times higher, a disparity that has persisted since the mid-1990s. Self-inflicted injuries are the largest contributor to mortality among young people living in Inuit Nunangat. ■

References

- Garner R, Carrière G, Sanmartin C. The health of Inuit, Métis and First Nations adults Living off-reserve in Canada: The impact of socio-economic status on inequalities in health. Health Research Working Paper Series (Statistics Canada, Catalogue 82-622-X, No. 004) Ottawa: Statistics Canada, 2010.
- Luo ZC, Senécal S, Simonet F, et al. Birth outcomes in the Inuit-inhabited areas of Canada. Canadian Medical Association Journal 2010; 182: 235-42. doi:10.1503/ cmaj.082042.
- Peters PA. Shifting transitions: health inequalities in Inuit Nunangat in perspective. *Journal of Rural and Community Development* 2012; 7(1): 35-58.
- Statistics Canada. *Inuit in Canada: Selected Findings of the 2006 Census.* Ottawa: Statistics Canada, 2008. Available at: http://www.statcan.gc.ca/pub/11-008-x/2008002/article/10712-eng.htm. Accessed July 27, 2011.
- Wilkins R, Uppal S, Finès P, et al. Life expectancy in the Inuit-inhabited areas of Canada, 1989 to 2003. *Health Reports* 2008; 19(1): 7-19.
- Statistics Canada. Table 102-0704. Mortality, by selected causes of death (ICD-10) and sex, five-year average, Canada and Inuit regions, every 5 years. CANSIM (database), 2010. Accessed February 6, 2012.
- Bélanger A, Gilbert S. The fertility of immigrant women and their Canadian-born daughters. Report on the Demography Situation in Canada (Statistics Canada, Catalogue 91-209-XIE) Ottawa: Statistics Canada, 2003.

- Kindig D. Understanding population health terminology. *The Milbank Quarterly* 2007; 85: 139-61.
- Charron S, Penney C, Senecal S. Police-reported crime in Inuit-Nunangat. Crime and Justice Research Paper Series (Statistics Canada, Catalogue 85-561-MWE00) Ottawa: Statistics Canada, 2010.
- Lopez AD, Mathers CD, Ezzati M, et al. Global Burden of Disease and Risk Factors. Oxford University Press and TheWorld Bank: New York. 2006.
- Peters PA. Causes and contributions to differences in life expectancy for Inuit Nunangat and Canada, 1994-2003. International Journal of Circumpolar Health 2010; 69: 66-77.
- Statistics Canada. Census Inuit Tables (Statistics Canada, Catalogue 89-636-x 2008, No. 1) Ottawa: Statistics Canada, 2008. Available at: http://www.statcan. gc.ca/pub/89-636-x/89-636-x2008001-eng. htm. Accessed July 27, 2011.
- Chiang CL. The Life Table and Its Applications. Malabar: Robert E. Kreiger Publishing, 1984.
- Spiegelman M. Introduction to Demography. Revised Edition. Cambridge Massachusetts: Harvard University Press, 1968.
- Kirmayer LJ. Suicide among Canadian Aboriginal peoples. *Transcultural Psychiatry* 1994; 31(1): 3-58.

- Kirmayer LJ, Malus M, Boothroyd LJ. Suicide attempts among Inuit youth: a community survey of prevalence and risk factors. *Acta Psychiatrica Scandinavica* 1996; 94: 8-17.
- Kirmayer LJ, Boothroyd L, Hodgins S. Attempted suicide among Inuit youth: psychosocial correlates and implications for prevention. *Canadian Journal of Psychiatry* 1998; 43: 816-22.
- Boothroyd L, Kirmayer L, Spreng S, et al. Completed suicides among the Inuit of northern Quebec, 1982-1996: a case control study. *Canadian Medical Association Journal* 2001; 165(6): 749-55.
- Macaluay A, Orr P, Macdonald S, et al. Mortality in the Kivalliq region of Nunavut, 1987-1996. *International Journal of Circumpolar Health* 2004; 63(Suppl 2): 80-5.
- Inuit Tapiriit Kanatami. Health Indicators of Inuit Nunangat within the Canadian Context 1994-1998 and 1999-2003. Ottawa: Inuit Tapiriit Kanatami, 2010.
- 21. Bjerregaard P, Young TK, Dewailly E, Ebbesson SOE. Indigenous health in the Arctic: an overview of the circumpolar Inuit population. *Scandinavian Journal of Public Health* 2004; 32: 390-5

Cause-specific mortality by education in Canada: A 16-year follow-up study

by Michael Tjepkema, Russell Wilkins and Andrea Long

Abstract

Background

People with lower levels of education tend to have higher rates of disease and death, compared with people who have higher levels of education. However, because death registrations in Canada do not contain information on the education of the deceased, unlinked vital statistics cannot be used to examine mortality differentials by education.

Methods

This study examines cause-specific mortality rates by education in a broadly representative sample of Canadians aged 25 or older. The data are from the 1991 to 2006 Canadian census mortality follow-up study, which included about 2.7 million people and 426,979 deaths. Age-standardized mortality rates (ASMRs) were calculated by education for different causes of death. Rate ratios, rate differences and excess mortality were also calculated.

Results

All-cause ASMRs were highest among people with less than secondary graduation and lowest for university degree-holders. If all cohort members had the mortality rates of those with a university degree, the overall ASMRs would have been 27% lower for men and 22% lower for women. The causes contributing most to that "excess" mortality were ischemic heart disease, lung cancer, chronic obstructive pulmonary disease, stroke, diabetes, injuries (men), and respiratory infections (women). Causes associated with smoking and alcohol abuse had the steepest gradients.

Interpretation

A mortality gradient by education was evident for many causes of death.

Keywords

Age-standardized mortality rates, rate differences, rate ratios, socio-economic inequalities

Authors

Michael Tjepkema (1-613-951-3896; michael. tjepkema@statcan.gc.ca) is with the Health Analysis Division at Statistics Canada, Ottawa, Ontario K1A 0T6. Russell Wilkins is with the Health Analysis Division and the University of Ottawa. Andrea Long is with the Public Health Agency of Canada, Ottawa, Ontario.

The social, economic and environmental conditions that people experience throughout their lives are the most important influences on their health. Known as the *social determinants of health*, these factors include income, occupation, living conditions, and importantly, education.

The level of education that a person achieves is influenced by circumstances that include family income during childhood and intergenerational effects such as the mother's education.^{1,2-5} Differences in educational attainment may be associated with different health trajectories. People with lower levels of education tend to have high rates of disease and mortality, compared with those with higher levels of attainment.⁶

Education can affect health through multiple pathways.^{6,7} Educational attainment may be an indicator of intraand inter-personal skills that are needed to produce and maintain good health.8,9 Health literacy—the ability to access and use health information to make decisions that contribute to the maintenance of basic health—is considered a critical link between education and health outcomes.¹⁰ People with higher levels of education may be more receptive to prevention messages and better able to change their behaviours and to use the health care system effectively.11 For instance, people with higher levels of education may be less likely to engage in health risk behaviours, such as smoking. 12 Education is also closely connected with other social determinants of health. Higher attainment can increase opportunities for employment and income security, 1 and research consistently documents better health in higher-income groups. 13-15

The association between education and mortality is well established in western and eastern European countries. 16,17 In Canada, however, death registrations do not contain information about the education of the deceased. As a result, unlinked vital statistics cannot be used to examine differences in mortality by level of education. To overcome this obstacle, several smaller-scale record linkage-based mortality studies have been conducted. 18-23 These studies demonstrated socio-economic differentials in mortality in Canada, but their applicability may be limited by the scope of the universe covered (geographically, or by age, sex and/or occupation), small sample size, lack of information about causes of death, or a combination of factors.

Recently, a broadly representative sample of Canadian adults aged 25 or older was linked to almost 16 years of mortality data.²⁴⁻²⁶ This cohort has been used to examine gradients in allcause mortality (and life expectancy) by various socio-economic indicators. All-cause mortality rates were shown to be lower in each successively higher category of socio-economic status, whether defined by income, education or occupation. But the cohort has not been used to assess educational gradients in cause-specific death rates. The objective of this study, therefore, is to examine cause-specific mortality rates by level of education to determine if the association between education and mortality differs by cause of death.

Methods

The data are from the 1991-to-2006 Canadian census mortality follow-up study, which tracked mortality in a 15% sample of the adult population.²⁴⁻²⁶ Respondents to the 1991 Census were eligible to be included in the study cohort if they were: (1) 25 or older and a usual resident of Canada on the day of the census (June 4, 1991); (2) not a long-term resident of an institution such as a prison, hospital or nursing home; and (3) enumerated using the long-form questionnaire that was administered to one in five private households, and to all residents of non-institutional collective dwellings and Indian reserves. Approximately 3.6 million individuals met these criteria.

The electronic 1991 census database does not contain names, which are needed to determine mortality. To obtain names, census records were first linked to tax-filer data from 1990 and 1991 using probabilistic matching, based on dates of birth and postal codes of the individual and his/her spouse or common-law partner (if any). About three-quarters of those who were in-scope were successfully linked to non-financial tax-filer data. The cohort was then linked to the Canadian Mortality Database (June 4, 1991 to December 31, 2006) using probabilistic methods.²⁷ Even

Table 1 Educational attainment, by sex and age group, non-institutional cohort members aged 25 or older, Canada, 1991 (baseline)

Sex and age group	Total number	Less than secondary graduation	Secondary graduation	Post- secondary diploma	University degree
			%		
Men					
25 or older	1,358,200	35.0	37.6	12.4	15.1
25 to 44	725,500	24.1	42.7	15.7	17.5
45 to 64	433,400	41.4	34.6	9.9	14.0
65 to 74	135,700	58.0	26.9	6.2	8.9
75 or older	63,600	65.3	22.3	4.9	7.5
Women					
25 or older	1,376,600	34.8	35.2	18.4	11.7
25 to 44	765,100	23.0	40.2	21.5	15.3
45 to 64	388,300	42.7	31.5	16.9	8.9
65 to 74	136,300	59.3	25.9	10.5	4.3
75 or older	86,800	64.2	21.9	10.0	3.9

Source: 1991 to 2006 Canadian census mortality and cancer follow-up study.

without a match to a death registration, follow-up status (alive, dead, emigrated, or lost to follow-up) could usually be determined from tax-filer data.²⁵ Overall, the cohort consisted of 2.7 million people, 16% of whom (426,979) died during the follow-up period (Appendix Table A).

Underlying cause of death was coded according to the World Health Organization's International Classification of Diseases, Ninth Revision28 for deaths that occurred in 1991 through 1999, and according to the Tenth Revision29 for deaths that occurred in 2000 through 2006. Deaths were grouped by Global Burden of Disease categories,30 and by behavioural health risk factors, namely, smokingrelated,16 alcohol-related,16 and drugrelated³¹ diseases. Deaths before age 75 that were potentially amenable to medical intervention, such as those due to cerebrovascular disease, hypertension, and breast cancer pneumonia/ influenza, 16,32 were also examined.

Highest level of education at cohort inception (baseline) was grouped into four categories: less than secondary graduation, secondary graduation (or trades certificate), postsecondary certificate or diploma (short of a bachelor's degree), and university degree.

For each cohort member, person-days of follow-up were calculated from the day of the census (June 4, 1991) to the date of death, date of emigration or the last day of the study period (December 31, 2006). Person-days of follow-up were divided by 365.25 to obtain person-years at risk. Age-at-baseline-, sex-, and educational attainment-specific mortality rates by 5-year age groups were used to calculate age-standardized mortality rates (ASMRs), using the cohort population structure (person-years at risk), both sexes together, as the standard population.

Relative inequalities were assessed by rate ratios (RRs) and percent excess mortality. RRs were calculated by dividing the ASMR for those with lower levels of education (less than secondary graduation, secondary graduation, or postsecondary diploma) by the ASMR for those with a university degree. RRs greater than 1.00 indicate an increased mortality risk. Percent excess mortality was calculated by subtracting the ASMR for those with a university degree from the ASMR for the total cohort, then dividing by the total ASMR and multiplying by 100.

Absolute inequalities were assessed by rate differences (RDs) and absolute excess mortality. RDs were calculated by subtracting the ASMR for those with a university degree from the ASMR of those with lower levels of education (less than secondary graduation, secondary graduation, or postsecondary diploma). RDs greater than zero indicate excess mortality. Absolute excess mortality was calculated by subtracting the ASMR of those with a university degree from the ASMR for the total cohort. The difference represents the number of deaths (per 100,000) that hypothetically could have been avoided if all cohort

members had experienced the mortality rate of those with a university degree.

Based on previously described methods,³³ 95% confidence intervals for ASMRs, RRs and RDs were calculated.

Results

The percentages of male and female cohort members, respectively, at each level of education were 35% and 35% for less than secondary graduation, 38% and 35% for secondary graduation, 12%

and 18% for postsecondary diploma, and 15% and 12% for university degree (Table 1). Younger cohort members tended to have higher levels of education than did older members.

The ASMR for all causes of death showed a clear stair-stepped gradient by level of education, with higher mortality rates for those with lower levels of education. Compared with people who had a university degree, the rate ratio (RR) for those with less than secondary graduation was 1.55 for men and 1.44 for women (Tables 2 and 3).

Table 2
Age-standardized mortality rates per 100,000 person-years at risk for selected causes of death, by educational attainment, male cohort members aged 25 or older at baseline, Canada 1991 to 2006

Course of death	Total	University	Post- secondary	Secondary	Less than secondary	DD.	DD	- France	%
Cause of death	Total	degree	diploma	graduation	graduation	RR	RD	Excess	excess
All causes	1,372.8	1,008.9	1,145.7	1,315.2	1,561.9	1.55*	553.0*	363.9	26.5
Communicable diseases	57.1	50.9	49.7	53.0	62.1	1.22*	11.1*	6.2	10.8
HIV/AIDS	5.8	7.4	6.7	5.6	4.7	0.63*	-2.7*	-1.5	-26.5
Respiratory infections	36.4	31.5	30.6	33.9	39.9	1.27*	8.4*	4.9	13.5
Non-communicable diseases	1,188.4	863.8	993.7	1,139.2	1,342.8	1.55*	479.0*	324.6	27.3
Malignant neoplasms	416.5	296.3	353.9	408.3	467.3	1.58*	171.0*	120.2	28.9
Stomach cancer	15.5	9.5	12.4	14.0	18.5	1.94*	9.0*	6.0	38.6
Colon and rectal cancers	44.0	34.3	40.7	43.0	48.0	1.40*	13.7*	9.7	22.0
Liver cancer	9.0	6.6	8.5	8.8	10.0	1.52*	3.4*	2.4	26.9
Pancreatic cancer	19.9	18.6	18.7	20.0	20.9	1.12	2.3	1.3	6.6
Trachea, bronchus, and lung cancers	124.1	55.3	84.3	115.5	154.7	2.80*	99.3*	68.7	55.4
Prostate cancer	51.7	46.2	50.9	51.3	53.0	1.15*	6.9*	5.5	10.6
Diabetes mellitus	38.3	25.0	30.5	34.9	45.6	1.83*	20.6*	13.3	34.7
Neuropsychiatric conditions	65.7	65.4	64.0	63.2	70.4	1.08*	5.0	0.3	0.4
Alcohol use disorders	5.0	2.5	3.4	4.4	7.0	2.82*	4.5*	2.6	50.9
Alzheimer's disease and other dementias	32.9	34.6	32.9	32.5	33.0	0.96	-1.5	-1.7	-5.1
Cardiovascular diseases	483.4	360.3	408.1	459.6	544.3	1.51*	184.0*	123.1	25.5
Ischemic heart disease	293.3	209.5	246.7	279.8	333.5	1.59*	124.0*	83.8	28.6
Cerebrovascular disease	79.3	67.8	65.8	74.6	87.3	1.29*	19.5*	11.5	14.5
Respiratory diseases	86.0	47.2	59.9	76.7	102.5	2.17*	55.3*	38.8	45.1
Chronic obstructive pulmonary disease	65.4	31.1	40.5	56.0	80.3	2.58*	49.2*	34.4	52.5
Digestive diseases	46.5	29.6	34.0	45.5	55.1	1.86*	25.5*	16.9	36.4
Cirrhosis of the liver	15.1	7.7	10.3	15.3	18.8	2.44*	11.1*	7.4	48.9
Injuries	70.2	47.0	54.2	69.3	91.3	1.94*	44.4*	23.2	33.1
Unintentional injuries	45.7	32.1	37.0	45.3	58.2	1.81*	26.1*	13.6	29.8
Road traffic accidents	7.9	5.1	6.4	7.8	10.1	1.98*	5.0*	2.8	35.3
Intentional injuries	24.4	14.8	17.2	24.0	33.1	2.23*	18.3*	9.6	39.2
Suicide	22.7	13.9	16.2	22.6	30.2	2.17*	16.3*	8.8	38.7
Smoking-related diseases	216.2	102.6	143.9	198.3	266.1	2.59*	163.5*	113.5	52.5
Alcohol-related diseases	15.8	7.4	9.6	15.1	21.6	2.90*	14.1*	8.4	53.0
Drug-related diseases	5.4	3.4	3.4	5.5	7.7	2.25*	4.3*	2.0	36.7
Amenable to medical intervention (younger than 75)	141.8	121.8	128.5	138.8	152.4	1.25*	30.6*	20.0	14.1

^{*} significantly different from rate for university degree (p< 0.05)

RR = rate ratio (less than secondary graduation / university degree)

RD = rate difference (less than secondary graduation - university degree)

Excess = (total - university degree)

[%] excess = [100 * (total - university degree)/total]

Note: Reference population (person-years at risk) for age-standardization was taken from internal cohort age distribution (5-year age groups).

Source: 1991 to 2006 Canadian census mortality and cancer follow-up study.

Table 3
Age-standardized mortality rates per 100,000 person-years at risk for selected causes of death, by educational attainment, female cohort members aged 25 or older at baseline, Canada 1991 to 2006

		University		Secondary					%
Cause of death	Total	degree	diploma	graduation	graduation	RR	RD	Excess	excess
All causes	869.4	677.7	736.6	820.9	977.7	1.44*	300.0*	191.7	22.0
Communicable diseases	34.2	27.2	29.0	31.7	39.0	1.43*	11.8*	7.0	20.4
HIV/AIDS	0.5	0.3	0.2	0.5	0.9	3.21*	0.6*	0.2	40.9
Respiratory infections	21.8	17.1	19.1	20.7	23.9	1.40*	6.9*	4.7	21.7
Non-communicable diseases	748.0	581.2	639.4	710.1	835.0	1.44*	253.8*	166.8	22.3
Malignant neoplasms	273.3	231.1	250.1	272.6	294.8	1.28*	63.7*	42.2	15.5
Stomach cancer	6.8	4.8	5.5	6.1	8.0	1.64*	3.1*	2.0	29.4
Colon and rectal cancers	28.0	23.8	26.7	27.9	29.8	1.25*	5.9*	4.2	14.8
Liver cancer	4.3	3.2	4.0	3.8	4.9	1.57*	1.8	1.1	26.1
Pancreatic cancer	14.9	11.0	13.4	15.6	15.5	1.41*	4.5*	3.9	26.2
Trachea, bronchus, and lung cancers	61.7	29.2	45.1	59.3	77.0	2.64*	47.8*	32.4	52.6
Female breast cancer	49.2	57.2	51.1	51.8	46.0	0.80*	-11.2*	-8.1	-16.4
Cervix uteri cancer	3.8	2.3	2.8	3.3	5.3	2.36*	3.1*	1.5	40.5
Ovarian cancer	14.5	15.3	14.9	14.7	14.1	0.92	-1.2	-0.8	-5.9
Diabetes mellitus	24.3	11.6	14.3	18.4	32.1	2.78*	20.5*	12.8	52.5
Neuropsychiatric conditions	55.9	57.3	53.0	55.2	58.5	1.02	1.2	-1.3	-2.4
Alcohol use disorders	1.5	0.8	0.9	1.2	2.3	2.77*	1.5	0.6	43.2
Alzheimer's disease and other dementias	36.7	36.3	35.5	36.5	37.1	1.02	0.8	0.3	0.9
Cardiovascular diseases	280.4	204.3	231.3	258.5	315.6	1.54*	111.3*	76.0	27.1
Ischemic heart disease	142.0	97.1	111.6	129.0	162.5	1.67*	65.4*	44.8	31.6
Cerebrovascular disease	64.0	52.9	57.0	61.6	69.1	1.31*	16.2*	11.1	17.4
Respiratory diseases	42.5	23.6	32.4	37.4	51.2	2.17*	27.6*	18.9	44.4
Chronic obstructive pulmonary disease	30.1	14.2	21.8	25.9	36.7	2.59*	22.5*	15.9	52.9
Digestive diseases	32.3	22.4	24.2	31.4	38.0	1.69*	15.6*	9.9	30.6
Cirrhosis of the liver	7.1	3.8	4.5	6.7	9.6	2.55*	5.9*	3.3	46.8
Injuries	31.2	29.4	30.3	30.2	36.2	1.23*	6.8*	1.9	6.0
Unintentional injuries	24.4	24.0	23.9	23.2	27.9	1.16*	3.9	0.4	1.5
Road traffic accidents	3.5	3.6	3.7	3.2	4.3	1.18	0.6	-0.1	-3.1
Intentional injuries	6.9	5.4	6.4	7.0	8.3	1.55*	3.0*	1.5	21.9
Suicide	6.0	4.5	6.0	6.0	7.1	1.58*	2.6*	1.5	25.6
Smoking-related diseases	100.8	49.7	74.0	94.3	124.2	2.50*	74.5*	51.1	50.7
Alcohol-related diseases	5.7	2.9	3.7	5.6	8.0	2.75*	5.1*	2.8	49.1
Drug-related diseases	4.5	2.6	3.6	4.3	6.9	2.68*	4.3*	1.9	43.0
Amenable to medical intervention (younger than 75)	148.4	133.7	138.5	147.4	156.5	1.17*	22.8*	14.8	9.9

^{*} significantly different from rate for university degree (p< 0.05)

Note: Reference population (person-years at risk) for age-standardization was taken from internal cohort age distribution (5-year age groups).

Source: 1991 to 2006 Canadian census mortality and cancer follow-up study.

The mortality gradient by education differed by cause of death. For men, RRs comparing those with less than secondary graduation to those with a university degree were particularly high for deaths due to trachea, bronchus and lung cancers (RR=2.80), chronic obstructive pulmonary disease (RR=2.58), and cirrhosis of the liver (RR=2.44) (Table 2). By contrast, the gradient was reversed for deaths due to HIV/AIDS (RR=0.63), and not statistically different

for pancreatic cancer (RR=1.12) and dementia (RR=0.96).

For women, the RRs comparing those with less than secondary graduation to university degree-holders were notably high for deaths due to trachea, bronchus and lung cancers (RR=2.64), alcohol use disorders (RR=2.77), chronic obstructive pulmonary disease (RR=2.59), and cirrhosis of the liver (RR=2.55) (Table 3). On the other hand, the gradient was reversed for female breast cancer

(RR=0.80), and not statistically different for ovarian cancer (RR=0.92), dementia (RR=1.02), and road traffic accidents (RR=1.18).

The education-related percentage "excess" (last column in Tables 2 and 3) shows that if every cohort member had experienced the age-specific mortality rates of those with a university degree, the all-cause ASMR would have been 27% lower for men and 22% lower for women, representing 364 and 192

RR = rate ratio (less than secondary graduation / university degree)

RD = rate difference (less than secondary graduation - university degree)

Excess = (total - university degree)

[%] excess = [100 * (total - university degree)/total]

Table 4
Age-standardized mortality rates per 100,000 person-years at risk for selected causes of death, by educational attainment, age group and sex, cohort members aged 25 or older at baseline, Canada 1991 to 2006

Sex, age group at baseline and cause of death	Total	University degree	Post- secondary diploma	Secondary graduation	Less than secondary graduation	RR	RD	Excess	% excess
Men									
25 to 44									
All causes	211.5	117.6	159.4	216.6	305.6	2.60*	188.0*	93.9	44.4
Unintentional injuries	27.7	11.0	18.6	28.5	44.1	4.01*	33.1*	16.7	60.3
Ischemic heart disease	30.6	14.4	22.7	30.7	47.8	3.31*	33.3*	16.2	52.9
Suicide	23.8	12.5	17.7	25.2	33.2	2.66*	20.7*	11.3	47.5
Trachea, bronchus, and lung cancers	13.1	4.9	8.8	14.5	19.9	4.07*	15.0*	8.2	62.7
45 to 64	4.040.0	007.4	4.040.0	4.054.4	4 500 0	4.00*	770.0*	540.0	00.0
All causes	1,319.3	807.4	1,048.6	1,251.4	1,583.6	1.96*	776.2*	512.0	38.8
Ischemic heart disease	274.6	159.6	224.9	261.8	330.9	2.07*	171.4*	115.1	41.9
Trachea, bronchus, and lung cancers	181.7	72.2	121.6	165.8	239.4	3.32*	167.2*	109.5	60.3
Chronic obstructive pulmonary disease	42.4 51.6	12.4	22.1	34.5	59.5	4.79*	47.1*	30.0	70.7 39.0
Cerebrovascular disease	31.0	31.5	35.9	48.3	62.8	2.00*	31.3*	20.1	39.0
65 to 74 All causes	5,304.4	3,923.5	4,320.0	5,058.1	5,775.8	1.47*	1,852.3*	1,380.8	26.0
Ischemic heart disease	1,180.8	818.3	986.6	1,124.4	1,292.8	1.58*	1,032.3 474.6*	362.5	30.7
Trachea, bronchus, and lung cancers	499.5	225.3	327.4	464.1	584.1	2.59*	358.8*	274.2	54.9
Chronic obstructive pulmonary disease	318.9	137.1	166.3	267.9	392.0	2.86*	254.9*	181.8	57.0
Cerebrovascular disease	333.1	266.4	266.1	308.2	364.8	1.37*	98.4*	66.7	20.0
	000.1	200.4	200.1	000.2	004.0	1.07	JO. T	00.1	20.0
75 or older All causes	9,517.4	8,201.0	8,417.8	9,078.8	9,988.4	1.22*	1,787.4*	1,316.4	13.8
Ischemic heart disease	2,187.4	1,844.3	1,896.9	2,083.8	2,304.1	1.25*	459.9*	343.1	15.7
Trachea, bronchus, and lung cancers	568.4	294.6	428.1	516.9	643.7	2.18*	349.1*	273.8	48.2
Chronic obstructive pulmonary disease	594.4	332.8	402.9	523.2	678.8	2.04*	346.0*	261.6	44.0
Diabetes mellitus	247.3	207.9	195.4	227.1	265.2	1.28*	57.3*	39.4	15.9
Women									
25 to 44									
All causes	134.5	87.2	109.6	131.5	193.1	2.22*	105.9*	47.4	35.2
Trachea, bronchus, and lung cancers	13.9	5.2	9.9	14.2	22.6	4.38*	17.4*	8.8	63.0
Ischemic heart disease	6.6	2.5	4.7	6.3	11.5	4.61*	9.0*	4.2	62.6
Unintentional injuries	9.0	5.6	7.9	8.1	14.2	2.52*	8.6*	3.4	37.7
Suicide	6.8	4.2	6.5	6.9	8.5	2.04*	4.3*	2.6	38.5
45 to 64									
All causes	780.9	521.5	616.0	717.0	926.2	1.78*	404.7*	259.5	33.2
Trachea, bronchus, and lung cancers	103.4	42.6	71.9	97.3	131.2	3.08*	88.6*	60.8	58.8
Ischemic heart disease	89.9	33.7	56.1	79.1	118.0	3.50*	84.3*	56.2	62.5
Chronic obstructive pulmonary disease	27.8	6.8	17.4	20.0	39.8	5.83*	33.0*	21.0	75.4
Diabetes mellitus	24.0	8.3	12.9	17.5	34.3	4.11*	26.0*	15.6	65.2
65 to 74									
All causes	3,198.6	2,493.2	2,660.5	2,993.8	3,446.1	1.38*	952.9*	705.4	22.1
Ischemic heart disease	571.8	396.0	408.0	505.6	645.5	1.63*	249.4*	175.8	30.7
Trachea, bronchus, and lung cancers	196.6	100.8	155.7	187.3	216.0	2.14*	115.2*	95.8	48.7
Chronic obstructive pulmonary disease	146.7	62.8	99.6	126.9	170.8	2.72*	108.0*	83.9	57.2
Cerebrovascular disease	250.9	188.0	218.8	241.4	266.1	1.42*	78.1*	62.9	25.1
75 or older	0.400.0	E 000 4	F 740 0	0.454.0	0.004.0	4.40*	4 044 0*	700 4	44 -
All causes	6,420.2	5,682.1	5,740.0	6,151.6	6,694.0	1.18*	1,011.9*	738.1	11.5
Ischemic heart disease	1,350.0	1,105.5	1,165.4	1,244.3	1,439.5	1.30*	334.0*	244.5	18.1
Chronic obstructive pulmonary disease	228.7	144.6	181.9	219.2	246.6	1.71*	102.0*	84.1	36.8
Diabetes mellitus Trachea, bronchus, and lung cancers	171.9	103.2	107.3	131.5	203.2	1.97* 1.54*	100.0* 66.5*	68.7	40.0 30.9
rrachea, pronchus, and lung cancers	179.3	124.0	144.1	177.2	190.4	1.54*	66.5*	55.3	30.9

^{*} significantly different from rate for university degree (p< 0.05)

Note: Reference population (person-years at risk) for age-standardization was taken from internal cohort age distribution (5-year age groups).

Source: 1991 to 2006 Canadian census mortality and cancer follow-up study.

RR = rate ratio (less than secondary graduation / university degree)

RD = rate difference (less than secondary graduation - university degree)

Excess = (total - university degree)

[%] excess = [100 * (total - university degree)/total]

fewer deaths per 100,000, respectively. ASMRs for trachea, bronchus and lung cancers, diabetes mellitus (women), chronic obstructive pulmonary disease, and alcohol use disorders (men) would each have been at least 50% lower.

The causes of death contributing the most to education-related absolute excess mortality (next-to-last column in Tables 2 and 3) were ischemic heart disease, lung cancer, chronic obstructive pulmonary disease, cerebrovascular disease, diabetes, injuries (men), and respiratory infections (women). Together, these seven causes accounted for about two-thirds of the total education-related excess mortality for men (65%) and women (64%) (percentages not shown).

RRs for smoking-, alcohol- and drug-related disease deaths all exceeded 2.00 (Tables 2 and 3). ASMRs for smoking- and alcohol-related diseases would have been about 50% lower if all cohort members had experienced the age-specific mortality rates of cohort members with a university degree.

For deaths potentially amenable to medical intervention, the gradient in mortality by education was less steep (RR=1.25 for men and 1.17 for women). The percent excess was 14% for men and 10% for women.

The gradient in RRs by education was steepest in the youngest age group (25 to 44 at baseline) and less steep in each successively older age group (Table 4). For men, the RRs were 2.60 at ages 25 to 44 and 1.22 at age 75 or older. For women, the RRs were 2.22 at ages 25 to 44 and 1.18 at age 75 or older. Although RRs across levels of educational attainment were highest in younger age groups, absolute differences were greatest in older age groups (among whom most deaths occur).

The causes of death that contributed the most to excess mortality differed by sex and age group. Among cohort members aged 25 to 44 at baseline, unintentional injuries was the largest contributor to excess mortality for men, and cancer of the trachea, bronchus and lung was the largest contributor for women. For both sexes aged 45 to 74,

ischemic heart disease and lung cancer were the two largest contributors to excess mortality. For those aged 75 or older, ischemic heart disease was the largest contributor, followed by chronic obstructive pulmonary disease for women and lung cancer for men.

Discussion

This analysis shows important differences in cause-specific mortality rates by level of education. For most causes of death, the higher the level of educational attainment, the lower the mortality rate, which is broadly with European^{11,16} consistent American^{8,9,34} research. Compared with university degree-holders, people with less than secondary graduation had agestandardized mortality rate ratios of 1.55 for men and 1.44 for women. If the entire cohort had experienced the age-specific mortality rates of those with a university degree, the all-cause ASMR would have been 27% lower for men and 22% lower for women. Extrapolated to the total non-institutional adult population, that equates to an estimated 50,000 fewer deaths per year: 33,000 among men and 17,000 among women. A similar reduction in mortality could have been achieved if all ischemic heart disease and cerebrovascular disease deaths had been eliminated.

Among the causes of death that contribute the most to absolute excess mortality, cardiovascular diseases (including ischemic heart disease and cerebrovascular disease) consistently rank as the most costly in terms of health care system use and lost productivity due to morbidity and premature mortality. Respiratory illnesses such as chronic obstructive pulmonary disease are also among the top five with respect to the direct and indirect economic burden of illness. 35,36

In terms of RRs, causes of death closely associated with health risk behaviours (for instance, smoking and excessive alcohol consumption) tended to have a steeper mortality gradient by education than did causes not as strongly

Why is this study important?

- The reduction of socio-economic inequalities in health outcomes is an explicit objective of health policies in Canada.
- Understanding socio-economic inequalities by cause of death may help achieve this objective.

What is already known on this subject?

 All-cause mortality rates are higher for people with relatively low levels of educational attainment.

What does this study add?

- If all cohort members had experienced the age-specific mortality rates of those with a university degree, the agestandardized mortality rate would have been 27% lower for men, and 22% lower for women.
- For both sexes, the causes of death contributing most to that "excess" mortality were ischemic heart disease, lung cancer, chronic obstructive pulmonary disease, stroke, diabetes, injuries (men) and respiratory infections (women).

associated with those behaviours. This is consistent with research indicating that people in lower socio-economic categories are more likely to engage in health risk behaviours. ^{15,37-39}

The current study revealed greater relative inequalities at younger than at older ages. It has been suggested that a selection effect may be operating at older ages: individuals with lower socio-economic status may die earlier, so that only the healthiest survive into old age, leading to reduced socio-economic inequalities in mortality.¹¹ It is also possible that the decrease in relative risks

may be partly explained by attenuation of the association between risk factors and chronic diseases in old age.⁴⁰ However, absolute differences in mortality rates were larger in older age groups (65 to 74 and 75 or older) than in younger age groups. Thus, reduction in mortality inequalities at older ages would have the greatest impact on reducing the number of education-related excess deaths.

Strengths and limitations

The large sample on which this study is based is broadly representative of all Canadian adults, and allowed for analysis of mortality differences (in relative and absolute terms) by education within detailed cause of death groupings.

This study was not intended to assess the relative importance of direct and indirect effects of education on mortality: for example, to what extent could educational differences in mortality be explained by associated differences in income?

The data did not include information on risk factors (such as smoking), and thus, might overestimate the effect of education on mortality. Nevertheless, other research concludes that socioeconomic differences in various health outcomes (including mortality) largely persist even after controlling for behavioural risk factors. 7.41,42 Further research to determine the degree to which individual behaviours and risk

factors explain (or fail to explain) the higher mortality rates experienced by persons of lower socio-economic status in Canada would require long-term mortality follow-up from health surveys that collect data on behavioural risk factors and on indicators of socio-economic status.

This analysis was based on the credentials cohort members had attained by June 4, 1991. Since then, Canadians have become more educated, so the percentage of the population in the lowest attainment levels has declined. would reduce education-related excess mortality, assuming that the relative risks remained unchanged. However, while this study provides baseline data on the nature and extent of educationrelated inequalities in mortality, it cannot determine if those inequalities have persisted, increased or decreased over time. Only future linkages of mortality data to more recent censuses (or to the National Household Survey) can provide the data needed to assess such changes.

Conclusion

This study demonstrates important differences in mortality rates by level of education for most causes of death. Causes more closely associated with health risk behaviours tended to have a steeper gradient in mortality by education than did causes not as strongly associated with those behaviours. These

results build on previous research by providing evidence by cause-specific groups, and confirm the existence of a consistent gradient in mortality by education across most causes of death. With the extension of the 1991-to-2006 Canadian census mortality follow-up study to include linkage to cancer data, future work could examine the nature and extent of educational inequalities in cancer incidence and survival.

Acknowledgement

Funding for this analysis was provided by the Public Health Agency of Canada. Funding for the creation of the Canadian census mortality follow-up study was provided by the Canadian Population Health Initiative of the Canadian Institute of Health Information (original study), the Healthy Environment and Consumer Safety Branch of Health Canada (study extensions), and the Health Analysis Division of Statistics Canada. Finally, the authors acknowledge Canada's provincial and territorial registrars of vital statistics, who furnish the death data for the Canadian Mortality Database.

References

- Commission on Social Determinants of Health. Closing the Gap in a Generation: Health Equity through Action on the Social Determinants of Health. Final Report of the Commission on Social Determinants of Health. Geneva: World Health Organization, 2008.
- Ferguson HB, Bovaird S, Mueller MP. The impact of poverty on educational outcomes for children. *Paediatrics and Child Health* 2007; 1, 2(8): 701–6.
- Phipps S, Lethbridge L. Income and the Outcomes of Children. Analytical Studies Branch Research Paper Series (Catalogue 11F0019MIE, No. 28) Ottawa: Statistics Canada, 2006.
- Brooks-Gunn J, Duncan GJ, Britto PR. Are socioeconomic gradients for children similar to those for adults? In: Keating DP, Hertzman C, eds. *Developmental Health and the Wealth of Nations*. New York: The Guildford Press, 1999: 94-124.
- Hoddinott J, Lethbridge L, Phipps S. Is History Destiny? Resources, Transitions and Child Education Attainments in Canada (Catalogue RH63-1/555-12-02E) Hull, Quebec: Human Resources Development Canada, 2002.
- Link BG, Phelan J. Social conditions as fundamental causes of disease. *Journal of Health and Social Behavior* 1995; 35(extra issue): 80-94.

Cause-specific mortality by education in Canada: A 16-year follow-up study • Research article

- Lantz PM, House JS, Lepkowski JM, et al. Socioeconomic factors, health behaviors, and mortality: results from a nationally representative prospective study of US adults. *Journal of the American Medical Association* 1998; 279(21): 1703-8.
- Lantz PM, Golberstein E, House JS, Morenoff J. Socioeconomic and behavioral risk factors for mortality in a national 19-year prospective study of U.S. adults. Social Science and Medicine 2010; 70(10): 1558-66.
- 9. Rogers RG, Everett BG, Zajacova A, Hummer RA. Educational degrees and adult mortality risk in the United States. *Biodemography and Social Biology* 2010; 56(1): 80-99.
- Canadian Council on Learning. Health Literacy in Canada: Initial Results from the International Adult Literacy and Skills Survey 2007. Ottawa: Canadian Council on Learning, 2007.
- Menvielle G, Leclerc A, Chastang JF, Luce D. Socioeconomic inequalities in cause specific mortality among older people in France. BMC Public Health 2010; 10: 260.
- 12. Adler NE, Newman K. Socioeconomic disparities in health: pathways and policies. *Health Affairs* 2002; 21(2): 60-76.
- Canadian Population Health Initiative. Reducing Gaps in Health: A Focus on Socio-Economic Status in Urban Canada. Ottawa: Canadian Institute for Health Information, 2008.
- Chief Public Health Officer's Report on the State of Public Health in Canada 2008 (Catalogue HP2-10/2008) Ottawa: Minister of Health, 2008.
- McIntosh CN, Finès P, Wilkins R, Wolfson MC. Income disparities in health-adjusted life expectancy for Canadian adults, 1991 to 2001. Health Reports 2009; 20(4): 55-64.
- Mackenbach JP, Stirbu I, Roskam AJ, et al. Socioeconomic inequalities in health in 22 European countries. New England Journal of Medicine 2008; 358(23): 2468-81.
- Skalická V, van Lenthe F, Bambra C, et al. Material, psychosocial, behavioural and biomedical factors in the explanation of relative socio-economic inequalities in mortality: evidence from the HUNT study. *International Journal of Epidemiology* 2009; 38(5): 1272-84.
- Roos LL, Magoon J, Gupta S, et al. Socioeconomic determinants of mortality in two Canadian provinces: multilevel modelling and neighborhood context. Social Science and Medicine 2004; 59(7): 1435-47.

- Aronson K, Howe G, Fair M, Carpenter M. Occupational Surveillance in Canada: Cause-specific Mortality among Workers, 1965 to 1991 (Catalogue 84-546-XCB) Ottawa: Statistics Canada, 2000.
- Martel L, Bélanger A. An analysis of the change in dependence-free life expectancy in Canada between 1986 and 1996. In: Bélanger A, Gilbert S, eds. Report on the Demographic Situation in Canada 1998-1999 (Current Demographic Analysis, Statistics Canada Catalogue 91-209-XPE) Ottawa: Minister of Industry, 1999:164-186 (and Erratum).
- Chen J, Beavon D, Wilkins R. Mortality of retired public servants in Canada. *Proceedings* of the Social Statistics Section, Annual Meeting of the American Statistical Association. Chicago, 1996: 86-91.
- Wolfson M, Rowe G, Gentleman JF, Tomiak M. Career earnings and death: a longitudinal analysis of older Canadian men. *Journal of Gerontology* 1993; 48(4): S167-79.
- Wigle DT, Semenciw RM, Wilkins K, et al. Mortality study of Canadian male farm operators: non-Hodgkin's lymphoma mortality and agricultural practices in Saskatchewan. *Journal of the National Cancer Institute* 1990; 82(7): 575-82.
- Wilkins R, Tjepkema M, Mustard C, Choinière R. The Canadian census mortality follow-up study, 1991 through 2001. *Health Reports* 2008; 19(3): 25-43.
- Peters PA, Tjepkema M. 1991-2011 Canadian census mortality and cancer follow-up study. Proceedings of Statistics Canada Symposium 2010. Social Statistics: The Interplay among Censuses, Surveys and Administrative Data (Catalogue 11-522-XCB) Ottawa: Statistics Canada, 2011: 150-6.
- Tjepkema M, Wilkins R. Remaining life expectancy at age 25 and probability of survival to age 75, by socio-economic status and Aboriginal ancestry. *Health Reports* 2011; 22(4): 31-6.
- Fair M. Generalized Record Linkage System— Statistics Canada's Record Linkage Software. Austrian Journal of Statistics 2004; 33(1 and 2): 37-53.
- World Health Organization. Manual of the International Statistical Classification of Diseases, Injuries and Causes of Death. Ninth Revision. Geneva: World Health Organization, 1977
- World Health Organization. International Statistical Classification of Diseases and Related Health Problems, Tenth Revision. Geneva: World Health Organization, 1992.

- World Health Organization. Global Burden of Disease: 2004 Update. Geneva: World Health Organization, 2008.
- Office of National Statistics. Report: deaths related to drug poisoning: England and Wales, 1999-2003. Health Statistics Quarterly 2005; Spring(25):52-9.
- Stirbu I, Kunst AE, Bopp M, et al. Educational inequalities in avoidable mortality in Europe. *Journal of Epidemiology and Community Health* 2010; 64(10): 913-20.
- Spiegelman M. Introduction to Demography. Revised Edition. Cambridge, Massachusetts: Harvard University Press, 1968.
- Elo IT, Preston SH. Educational differentials in mortality: United States, 1979-85. Social Science and Medicine 1996; 42(1): 47-57.
- Canadian Heart Health Strategy and Action Plan. Building a Heart Healthy Canada. Ottawa: Canadian Heart Health Strategy and Action Plan, 2009. Available at: http://www. chhs.ca/en/reports.
- Health Canada. Economic Burden of Illness in Canada, 1998 (Catalogue H21-136/1998E) Ottawa: Health Canada, 2002.
- Phelan JC, Link BG, Diez-Roux A, et al. "Fundamental causes" of social inequalities in mortality: a test of the theory. *Journal of Health and Social Behavior* 2004; 45(3): 265-85.
- 38. Phelan JC, Link BG, Tehranifar P. Social conditions as fundamental causes of health inequalities: theory, evidence, and policy implications. *Journal of Health and Social Behavior* 2010; 51(Suppl.): S28-40.
- Syme SL. Reducing racial and social-class inequalities in health: the need for a new approach. *Health Affairs* 2008; 27(2): 456-9.
- Lewington S, Clarke R. Qizilbash N, et al. Age-specific relevance of usual blood pressure to vascular mortality: a meta-analysis of individual data for one million adults in 61 prospective studies. *Lancet* 2002; 360(9349): 1903-13.
- McGrail KM, van Doorslaer E, Ross NA, Sanmartin C. Income-related health inequalities in Canada and the United States: a decomposition analysis. *American Journal* of *Public Health* 2009; 99(1). doi:10.2105/ AJPH.2007.129361
- Kim HJ, Ruger JP. Socioeconomic disparities in behavioral risk factors and health outcomes by gender in the Republic of Korea. BMC Public Health 2010; 10(1): 195.

Appendix

Table A Cohort members aged 25 or older, person-years at risk, and deaths ascertained, by age group and educational attainment at baseline, by sex, Canada, 1991 to 2006

		Men		Women			
Age group and educational attainment	Cohort members	Person- years at risk	Number of deaths	Cohort members	Person- years at risk	Number of deaths	
Total 25 or older							
Less than secondary graduation	474,900	6,249,140	138,071	478,500	6,563,790	110,472	
Secondary graduation	510,400	7,318,980	69,084	484,000	7,096,100	47,128	
Postsecondary diploma	168,300	2,457,660	15,493	253,000	3,734,390	19,960	
University degree	204,600	2,942,760	18,339	161,100	2,379,250	8,432	
25 to 44							
Less than secondary graduation	175,000	2,616,330	8,066	176,200	2,651,730	5,243	
Secondary graduation	309,800	4,649,830	9,981	307,200	4,649,350	6,010	
Postsecondary diploma	113,700	1,708,250	2,690	164,600	2,487,060	2,654	
University degree	127,000	1,868,330	2,330	117,100	1,749,410	1,516	
45 to 64							
Less than secondary graduation	179,600	2,484,030	43,532	165,700	2,402,990	24,071	
Secondary graduation	149,900	2,144,350	25,829	122,400	1,810,250	12,369	
Postsecondary diploma	43,100	625,240	5,949	65,500	975,020	5,482	
University degree	60,800	888,540	6,268	34,700	517,610	2,342	
65 to 74							
Less than secondary graduation	78,700	851,130	48,956	80,800	1,015,830	35,470	
Secondary graduation	36,500	415,050	20,772	35,300	459,080	13,749	
Postsecondary diploma	8,400	99,940	4,213	14,300	189,510	5,051	
University degree	12,100	147,970	5,649	5,900	79,360	1,986	
75 or older							
Less than secondary graduation	41,500	297,650	37,517	55,700	493,230	45,688	
Secondary graduation	14,200	109,760	12,502	19,000	177,410	15,000	
Postsecondary diploma	3,100	24,220	2,641	8,700	82,800	6,773	
University degree	4,800	37,930	4,092	3,400	32,880	2,588	

Note: Cohort member counts rounded to nearest 100; person-years at risk to nearest 10. Source: 1991 to 2006 Canadian census mortality and cancer follow-up study.

ELECTRONIC PUBLICATIONS AVAILABLE AT

www.statcan.gc.ca

Informal caregiving for seniors

by Annie Turner and Leanne Findlay

Abstract

Based on data from the 2008/2009 Canadian Community Health Survey-Healthy Aging, this study examines the characteristics of people aged 45 or older who reported caring for a senior. It also describes the nature of the care provided and the positive and negative aspects of caregiving. More than one-third (35%) of Canadians aged 45 or older reported caring for a senior with a short- or long-term health condition or limitation. Compared with non-caregivers, those providing care to a senior were more likely to be women. They tended to be younger and more likely to live in higher-income households and to be postsecondary graduates. More than half the people receiving care were parents or parents-in-law, and they usually did not live with the caregiver. The most common form of care provided was transportation. A third of caregivers had been providing assistance for at least five years. Virtually all (95%) of them reported positive aspects of caregiving, but more than half (56%) experienced challenges and difficulties.

Keywords

Aged, caregiving, elderly, geriatrics, social support

Authors

Annie Turner (1-613-951-4365; annie.turner@ statcan.gc.ca) is with the Social and Aboriginal Statistics Division and Leanne Findlay is with the Health Analysis Division at Statistics Canada, Ottawa, Ontario, K1A 0T6.

As Canadians age, informal caregiving becomes increasingly important to the well-being of seniors. According to the 2008/2009 Canadian Community Health Survey (CCHS)—Healthy Aging, an estimated 3.8 million Canadians who were aged 45 or older (35%) were providing informal care to a senior with a short- or long-term health condition. Informal caregivers—family and friends who provide unpaid assistance with tasks such as transportation and personal care—help seniors remain in their homes, thereby reducing demands on the health care system.¹ Moreover, remaining in one's home is usually the preference of seniors themselves.^{2,3}

Caring for someone with a health condition or limitation, particularly cognitive impairments such as Alzheimer's disease and dementia, can cause physical and emotional problems and create financial and social burdens for the caregiver. This may be especially true for caregivers who, themselves, are seniors. On the other hand, providing care can give individuals pleasure and pride, enhance their self-worth, and help them to build relationships with the care recipient.

Based on data from the 2008/2009 CCHS-Healthy Aging, this study

compares the characteristics of caregivers with those of their contemporaries who are not caregivers (see *The data*). In addition, the characteristics of the care that caregivers provide are outlined, as are the positive and negative aspects of caregiving.

Caregivers

In 2008/2009, women made up just over half (57%) of people aged 45 or older who were providing care to a senior (Table 1). Almost three-quarters (73%) of these caregivers were aged 45 to 64, although a quarter were seniors themselves; in

The data

The data are from the 2008/2009 Canadian Community Health Survey (CCHS)—Healthy Aging, a cross-sectional survey about factors that contribute to healthy aging. Information was collected from 30,865 people aged 45 or older living in private occupied dwellings in the ten provinces. The survey excluded full-time members of the Canadian Forces and residents of the three territories, Indian reserves or Crown lands, institutions and some remote regions.

This study deals with people who reported that they provided care to a senior with a short- or long-term condition or limitation. Respondents who primarily provided care to someone younger than age 65 were excluded from this analysis. Some respondents reported caring for more than one person. The questions pertained to the person to whom, in the past 12 months, the caregiver had dedicated the most time and resources. It is possible that the caregiver who responded to the survey was not the only one providing care to that person.

Among caregivers, 11.4% were also receiving care. These respondents were included in the sample for this study, although subsequent analyses suggested that excluding them would not alter the results. The analysis represents caregivers, not care recipients; the information on care recipients is not representative of all Canadians aged 65 or older who receive care.

Descriptive statistics were used to compare caregivers with non-caregivers. Sampling weights were used in all analyses. To account for the complex survey design, a bootstrapping technique was applied for variance estimation.¹⁰

Caregivers were classified by age group: 45 to 54, 55 to 64, 65 to 74, and 75 or older. Their marital status was classified as married/common-law or widowed/separated/divorced/single. Household income deciles were derived by calculating the ratio between the total household income and Statistics Canada's low-income cutoff (LICO) specific to the number of people in the household, the size of the community, and the survey year. Their highest educational attainment was categorized as: less than secondary graduation, secondary graduation/some postsecondary, and postsecondary graduation. Employment data were collected only from respondents younger than age 75. Employment status was based on whether the respondent had worked in the past year.

Caregivers' self-perceived health was based on the question, "In general, would you say your health is:" Response options were: excellent, very good, good, fair or poor. Those with excellent or very good health were defined as having high self-perceived health, and those with good, fair or poor health were defined as having lower self-perceived health. A similar question was used for self-perceived mental health. Both self-perceived physical and mental health were age-standardized to account for the uneven distribution within the age categories.

The characteristics of care recipients examined in this study are age (65 to 74, 75 to 84, and 85 or older), relationship to caregiver (for example, spouse, parent or child), residence (same household, other household or institution), and nature of their health condition (short-term, long-term, other). The CCHS—Healthy Aging did not collect data on care recipients' specific health conditions.

The characteristics of the care provided are the type (transportation, help with housework, personal care, meals and other), duration (less than one year to five or more years), frequency (daily, less than daily, occasionally or rarely), and whether providing care had affected the caregiver's health.

Based on a list of response options, caregivers were asked about the positive and negative aspects of providing care.

fact, 10% of them were aged 75 or older. Even so, the age profile of caregivers was younger, compared with non-caregivers (who included care recipients), 18% of whom were 75 or older.

Perhaps reflecting their younger age profile, caregivers were more likely than non-caregivers to be married or in a common-law relationship (78% versus 71%). They were also more likely to have a higher household income and to be postsecondary graduates. Among those aged 45 to 74, caregivers were less likely than non-caregivers to have been employed in the past year.

Self-perceived health has been shown to be a reliable measure of general health status.¹¹ Higher percentages of caregivers than non-caregivers reported very good or excellent physical and mental health. Nonetheless, it is possible that some degree of self-selection is operating. That is, healthier people may be more capable of being caregivers, and so, more likely to undertake the task.

In multivariate analyses that controlled for sex, age, household income and educational attainment, associations between high self-perceived physical and mental health and being a caregiver were no longer significant (data not shown).

Care recipients

Around three-quarters of caregivers reported that the person whom they assisted was at least 75 years old; one-third were caring for a senior aged 85 or older (Table 2).

Parents and parents-in-law made up more than half (56%) of those receiving informal care. Another 19% of caregivers reported assisting a friend or neighbour, and 11% were caring for a spouse.

Relatively few care recipients (14%) actually lived with the caregiver. A substantial majority (70%) of these care recipients were living in another private household, and 12% were in a health care institution.

Caregiving duties

Transportation was the most common form of care provided, reported by 39% of caregivers. About 20% were assisting with household activities, and around 15%, with personal care.

Although 57% of caregivers described their provision of care as "regular," this was a daily commitment for only 21% of them; 36% provided regular care once a week, once a month, or less than once a month. About a third of caregivers had been providing care for at least five years; almost as many reported that they had been doing so for less than a year.

The literature suggests that those who care for someone with severe cognitive impairment are at elevated risk of experiencing caregiver stress or burden,^{5,7} but because the CCHS did not ask about the care recipients' specific health condition, this issue could not be addressed in the current study.

Table 1 Percentage distribution of selected characteristics of caregivers and noncaregivers, household population aged 45 or older, Canada excluding territories, 2008/2009

	C	aregiver		Non	-caregive	er
		95% confid inter	ence		95% confide interv	ence
	%	from	to	%	from	to
Total	100.0			100.0		
Sex						
Men	43.2	41.3	45.1	50.1	49.0	51.2
Women	56.8	54.9	58.7	49.9	48.8	51.0
Age group						
45 to 54	39.8	37.7		35.9	34.6	37.2
55 to 64	32.8	31.2		26.5	25.6	27.4
65 to 74	17.1	16.1	18.0	19.2	18.6	19.8
75 or older	10.3	9.6	11.1	18.4	17.9	19.0
Marital status						
Married/Common-law	77.9	76.3	79.4	71.3	70.0	72.5
Widowed/Separated/Divorced/Single	22.1	20.6	23.7	28.7	27.5	30.0
Household income decile						
1 (lowest)	5.7	5.0	6.5	12.9	11.9	14.0
2	8.1	7.2	9.2	11.5	10.6	12.4
3	9.4	8.3	10.7	10.7	9.9	11.6
4	10.2	9.0	11.6	10.5	9.6	11.5
5	10.5	9.2	12.1	9.8	8.8	10.9
6	11.1	9.9	12.6	9.2	8.2	10.3
7	10.2	9.0	11.6	8.9	7.9	10.0
8	11.0	9.4		9.5	8.3	10.9
9	12.3	10.6		8.9	7.9	10.1
10 (highest)	11.4	10.0	13.0	8.0	7.0	9.1
Education						
Less than secondary graduation	16.3	15.0	17.8	27.7	26.4	28.9
Secondary graduation or some postsecondary	25.2	23.7		24.9	23.5	26.3
Postsecondary graduation	58.4	56.5	60.3	47.5	45.8	49.1
Employment status [†]						
Not employed in past year	68.8	66.9		65.3	63.9	66.8
Employed in past year	31.2	29.5	33.1	34.7	33.2	36.1
Self-perceived physical health [‡]						
Excellent/Very good	58.0	55.9	60.0	53.2	51.4	54.9
Good/Fair/Poor	42.0	40.0	44.1	46.8	45.1	48.6
Self-perceived mental health [‡]						
Excellent/Very good	76.6	74.7	78.3	72.2	70.7	73.7
Good/Fair/Poor	23.4	21.7	25.3	27.8	26.3	29.3

[†] respondents aged 45 to 74

Source: 2008/2009 Canadian Community Health Survey–Healthy Aging.

Challenges and rewards

Although a relatively small percentage of informal caregivers reported that caregiving had caused or worsened a health problem (8%), more than half (56%) of them encountered difficulties and challenges (Table 3). When they were asked about the most negative

aspect of caregiving, 17% reported that it was emotionally demanding; 12% said that because of caregiving, they did not have enough time for themselves or family; 10% said it created stress; and 7% reported fatigue.

At the same time, almost all (95%) informal caregivers reported positive

Table 2
Percentage distribution of characteristics of care recipient and care provided, household population aged 45 or older, Canada excluding territories, 2008/2009

75 to 84 45.9 43.9 85 or older 32.5 30.6 Relationship to caregiver Parent/Parent-in-law 55.8 54.0 Friend/Neighbour/Other 19.2 17.8 Spouse/Common-law partner 10.5 9.5 Other relative 9.7 8.7 Child (older than age 65) 4.8 3.9 Residence Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased† 4.8 4.1 Health condition Short-term 13.4 11.0	23.3 47.9 34.6 57.7 20.7 11.6
Age group 21.6 20.0 65 to 74 21.6 20.0 75 to 84 45.9 43.9 85 or older 32.5 30.6 Relationship to caregiver Parent/Parent-in-law 55.8 54.0 Friend/Neighbour/Other 19.2 17.8 Spouse/Common-law partner 10.5 9.5 Other relative 9.7 8.7 Child (older than age 65) 4.8 3.9 Residence Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased† 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	23.3 47.9 34.6 57.7 20.7
65 to 74 21.6 20.0 75 to 84 45.9 43.9 85 or older 32.5 30.6 Relationship to caregiver Parent/Parent-in-law 55.8 54.0 Friend/Neighbour/Other 19.2 17.8 Spouse/Common-law partner 10.5 9.5 Other relative 9.7 8.7 Child (older than age 65) 4.8 3.9 Residence Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased† 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	47.9 34.6 57.7 20.7
75 to 84 45.9 43.9 85 or older 32.5 30.6 Relationship to caregiver Parent/Parent-in-law 55.8 54.0 Friend/Neighbour/Other 19.2 17.8 Spouse/Common-law partner 10.5 9.5 Other relative 9.7 8.7 Child (older than age 65) 4.8 3.9 Residence Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased† 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	47.9 34.6 57.7 20.7
Relationship to caregiver Secondaria Parent/Parent-in-law 55.8 54.0 Friend/Neighbour/Other 19.2 17.8 Spouse/Common-law partner 10.5 9.5 Other relative 9.7 8.7 Child (older than age 65) 4.8 3.9 Residence Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased† 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	34.6 57.7 20.7
Relationship to caregiver Parent/Parent-in-law 55.8 54.0 Friend/Neighbour/Other 19.2 17.8 Spouse/Common-law partner 10.5 9.5 Other relative 9.7 8.7 Child (older than age 65) 4.8 3.9 Residence Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased† 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	57.7 20.7
Parent/Parent-in-law 55.8 54.0 Friend/Neighbour/Other 19.2 17.8 Spouse/Common-law partner 10.5 9.5 Other relative 9.7 8.7 Child (older than age 65) 4.8 3.9 Residence Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased† 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	20.7
Friend/Neighbour/Other 19.2 17.8 Spouse/Common-law partner 10.5 9.5 Other relative 9.7 8.7 Child (older than age 65) 4.8 3.9 Residence Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased† 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	20.7
Spouse/Common-law partner 10.5 9.5 Other relative 9.7 8.7 Child (older than age 65) 4.8 3.9 Residence Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased† 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	
Other relative 9.7 8.7 Child (older than age 65) 4.8 3.9 Residence 3.9 Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased [†] 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	ט.וו
Residence 4.8 3.9 Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased¹ 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	10.7
Residence Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased† 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	6.0
Another household 69.6 67.7 Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased [↑] 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	0.0
Same household 13.6 12.3 Health care institution 12.0 10.9 Deceased [↑] 4.8 4.1 Health condition Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	71.4
Health care institution 12.0 10.9 Deceased [↑] 4.8 4.1 Health condition 3.4 11.0 Short-term 83.5 80.7 Other 3.1 2.4	15.0
Deceased† 4.8 4.1 Health condition 3.4 11.0 Short-term 83.5 80.7 Other 3.1 2.4	13.3
Short-term 13.4 11.0 Long-term 83.5 80.7 Other 3.1 2.4	5.6
Long-term 83.5 80.7 Other 3.1 2.4	
Other 3.1 2.4	16.2
	86.0
Care provided 100.0	4.1
,	
Туре	
· · · · · · · · · · · · · · · · · · ·	40.9
Help with activities 20.5 18.9 such as housework	22.3
	17.1
Meal preparation and delivery 11.2 9.9	12.7
	15.5
Frequency	
3 ()	23.0
3, (,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,	37.6
Occasionally/Rarely 43.4 41.2	45.5
Duration	20.5
,	32.5
	23.2 15.2
5 or more years 34.5 32.5	

[†] person cared for in past 12 months was deceased at time of survey

Source: 2008/2009 Canadian Community Health Survey— Healthy Aging.

aspects: 30% said that the most positive aspect of caregiving was personal satisfaction; 26% enjoyed providing assistance; and 19% stated that it made them feel closer to the care recipient.

[‡] age-standardized to account for uneven distribution within age categories

^{...} not applicable

^{..} not applicable

2008/2009

Table 3 Percentage distribution of negative and positive aspects of caregiving, household population aged 45 or older, Canada excluding territories,

		95% confid inter	ence
	%	from	to
Caregiving caused/worsened caregiver's health condition	100.0		
Yes	7.8	6.7	9.0
No	92.2	91.0	93.3
Most negative aspect	100.0		
Emotionally demanding	17.4	15.9	18.9
Not enough time for self/family	12.0	10.5	13.5
Creates stress	9.5	8.2	11.1
Fatigue	6.6	5.5	7.8
Affects family/other relationships	2.4	1.8	3.2
Interferes with work	1.9 ^E	1.4	2.7
Conflicts with social life	0.9 ^E	0.6	1.3
Financial burden	0.5 ^E	0.3	0.8
Other	4.5	3.6	5.5
Did not experience difficulties	44.4	42.2	46.6
Most positive aspect	100.0		
Personal satisfaction	30.4	28.3	32.5
Enjoy providing assistance	25.8	23.9	27.8
Closer to care recipient	18.7	17.2	20.4
Feel needed	16.4	14.8	18.0
Other	3.8	3.0	4.7
No positive aspects	5.0	3.7	6.6

^E use with caution

Source: 2008/2009 Canadian Community Health Survey– Healthy Aging.

Conclusion

In 2008/2009, about one-third of Canadians aged 45 or older were providing care to a senior with a short- or long-term health condition or limitation. Among the negative aspects of caregiving that they reported were that it was emotionally demanding and that it meant they did not have enough time for themselves or their family. On the other hand, substantial numbers reported that they derived personal satisfaction from caregiving and enjoyed providing assistance.

Acknowledgments

The authors thank Heather Gilmour, Guy Gellatly, and members of the focus content team in the Health Statistics Division at Statistics Canada for their comments and feedback on the paper.

References

- Canadian Institute for Health Information. Supporting Informal Caregivers – The Heart of Home Care. Ottawa: Canadian Institute for Health Information, 2010.
- Tang F, Lee, Y. Home and community-based services utilization and aging in place. *Home Health Care Services Quarterly* 2010; 29: 138-54.
- Sabia, J.J. There is no place like home: A hazard model analysis of aging in place among older homeowners in the PSID. Research on Aging 2008; 30(1): 3-35.
- Cranswick K, Dosman D. Eldercare: What we know today. *Canadian Social Trends* (Statistics Canada, Catalogue 11-0080X) 2008; 86: 48-57.
- Baumgarten M, Battista R, Infante-Rivard C, et al. The psychological and physical health of family members caring for an elderly person with dementia. *Journal of Clinical Epidemiology* 1992; 5(1): 61-70.
- De Frias CM, Tuokko H, Rosenberg T. Caregiver physical and mental health predicts reactions to caregiving. Aging and Mental Health 2005; 9(4): 331-6.
- Papastavrou E, Kalokerinou A, Papacosta S, et al. Caring for a relative with dementia; family caregiver burden. *Journal of Advanced Nursing* 2007; 58 (5): 446-57.
- Jull J. Seniors Caring for Seniors: Examining the Literature on Injuries and Contributing Factors Affecting the Health and Well-Being of Older Adult Caregivers. Provided to the Public Health Agency of Canada. Prepared on behalf of the Canadian Association of Occupational Therapists. 2010.
- Kramer B. Gain in the caregiving experience: Where we are? What next?" The Gerontologist 1997; 37(2): 218-32.
- Rust K, Rao JNK. Variance estimation for complex surveys using replication techniques. Statistical Methods in Medical Research 1996; (5): 281-310.
- Shields, M. Determinants of self-perceived health. *Health Reports* 2001; 13(1): 35-52.

^{...} not applicable

Overweight and obesity in children and adolescents: Results from the 2009 to 2011 Canadian Health Measures Survey

by Karen C. Roberts, Margot Shields, Margaret de Groh, Alfred Aziz and Jo-Anne Gilbert

Abstract

Background

The 2009 to 2011 Canadian Health Measures Survey provides the most recent measured body mass index (BMI) data for children and adolescents. However, different methodologies exist for classifying BMI among children and youth. Based on the most recent World Health Organization classification, nearly a third of 5- to 17-year-olds were overweight or obese. The prevalence of obesity differed between boys and girls (15.1% versus 8.0%), most notably those aged 5 to 11, among whom the percentage of obese boys (19.5%) was more than three times that of obese girls (6.3%). These estimates indicate a higher prevalence of overweight/ obesity among children than do estimates based on International Obesity Task Force cut-offs. Although the prevalence of overweight and obesity among children in Canada has not increased over the last decade, it remains a public health concern, given the tendency for excess weight to persist through to adulthood and lead to negative health outcomes.

Keywords

Body mass index, child, adolescent, population surveillance

Authors

Karen C. Roberts (1-613-946-5436; karen.c.roberts@phac-aspc.gc.ca) and Margaret de Groh are with the Health Promotion and Chronic Disease Prevention Branch at the Public Health Agency of Canada, Ottawa, Ontario, K1A 0K9. Margot Shields was formerly with the Health Analysis Division at Statistics Canada. Alfred Aziz and Jo-Anne Gilbert are with the Health Products and Food Branch at Health Canada.

Since the late 1970s, the prevalence of overweight and obesity has risen among children and adolescents in Canada.¹ Excess weight in childhood has been linked to insulin resistance, type 2 diabetes, hypertension, poor emotional health, and diminished social well-being.^{2,3} As well, obese children tend to become obese adults, making childhood obesity a public health concern.⁴⁻⁶

Routine surveillance of overweight and obesity is important for the development and assessment of efforts aimed at reducing excess weight in children and adolescents. The most common approach to classifying weight is the body mass index (BMI), which estimates adiposity based on weight relative to height.⁷⁻¹⁰ The use of measured, rather than reported, height and weight to derive BMI is strongly recommended, especially for children and adolescents.¹¹

Since the Canada Health Survey (age 0 and up) in 1978/1979, only a few national population-level surveys have directly measured the height and weight of children and adolescents: the 1981 Canada Fitness Survey (age 7 or older), the 1988 Campbell's Survey on the Well-being of Canadians (age 7 or older), the 2004 Canadian Community Health Survey (CCHS), Cycle 2.2 Nutrition (age 2 or older), and the 2007 to 2009 Canadian Health Measures

Survey (CHMS) (age 6 or older). The most recent CHMS cycle (2009 to 2011) included children aged 3 or older.

BMI classification guidelines for adults have been in place for decades,9 with cut-offs for specific categories based on scientific evidence of increasing health risks with increased BMI. Establishing a standard BMI classification system for children has been more challenging, because of variations in growth rates and the difficulty of linking estimated adiposity levels in childhood to weightrelated health outcomes that tend to manifest later in life. A number of classification systems for use at the population level have been developed to estimate overweight and obesity in children.12 Since 2004, Canada has used the age-/sex-specific classification cut-offs established by the International Obesity Task Force (IOTF).^{13,14} In 2007, the World Health Organization (WHO) released a new set of age-/sex-specific

The data

Estimates are based on data from the second cycle (2009 to 2011) of the Canadian Health Measures Survey (CHMS). The CHMS is an ongoing survey designed to provide comprehensive direct health measures data at the national level.¹⁵ Ethics approval was obtained from Health Canada's Research Ethics Board.¹⁶ The 2009 to 2011 CHMS covered the population aged 3 to 79 in private households. It excluded residents of Indian Reserves, institutions and some remote regions, and full-time members of the regular Canadian Forces. More than 96% of the population aged 3 to 79 is represented.¹⁷

Data were collected at 18 sites across Canada from August 2009 to December 2011. In addition to a questionnaire administered in the respondent's home, the survey involved physical measures (including height and weight) in a mobile examination centre. Participation was voluntary. Written informed consent was obtained from respondents aged 14 or older. For younger children, a parent or legal guardian provided written consent, in addition to written assent from the child (where possible). The CHMS Cycle 2 Data User Guide¹⁷ contains details about the 2009 to 2011 survey content and sample design.

Of the households selected for the survey, 75.9% agreed to participate. In each responding household, one or two members were selected: 90.5% of selected household members completed the household questionnaire, and 81.7% of the responding household members participated in the subsequent physical measures component. The final response rate, after adjusting for the sampling strategy, was 55.5%.¹⁷ This article is based on 2,123 respondents aged 5 to 17, for whom measured values of height and weight were collected.

Height was measured to the nearest 0.1 centimetre using a ProScale M150 digital stadiometer (Accurate Technology Inc., Fletcher, USA), and weight, to the nearest 0.1 kilogram with a Mettler Toledo VLC with Panther Plus terminal scale (Mettler Toledo Canada, Mississauga, Canada).

Body mass index was derived as weight in kilograms divided by height in metres squared. Based on BMI, children and adolescents were classified according to thinness, normal weight, overweight or obesity using two sets of age- and sex-specific cut-offs, one set specified by the WHO, ¹⁸ and the other, by the IOTF. ^{13,14}

The WHO cut-off criteria used to classify children younger than age 5 as overweight or obese¹⁹ differ slightly from those used for children aged 5 or older,¹⁸ and the WHO does not recommend combining across age groups. Because the sample size for 3- and 4-year-old children in the 2009 to 2011 CHMS was too small to provide reliable estimates using these cut-offs, this age group was not included in this report.

All estimates were based on weighted data. Statistical analyses were performed using SAS and SUDAAN software. Standard errors, coefficients of variation and 95% confidence intervals were calculated with the *bootstrap* technique.^{20,21} The number of degrees of freedom was specified as 13 to account for the 2009 to 2011 CHMS sample design.¹⁷ Significance levels were set at p <0.05.

Table 1 Percentage distribution of children and adolescents, by body mass index (BMI) category (based on World Health Organization cut-offs), age group and sex, household population aged 5 to 17, 2009 to 2011

	TI	ninnes	8	Norr	nal wei	ght	Ove	erweig	ht	0	besity		
		95% confidence interval		confidence confidence			ence	95% confidence interval				95% confidence interval	
	%	from	to	%	from	to	%	from	to	%	from	to	
Total Age group (years)	2.2 ^E	1.1	4.1	66.4	62.8	69.8	19.8	16.6	23.4	11.7	9.9	13.7	
5 to 11	F			65.5	61.7	69.2	19.7	16.4	23.4	13.1	10.5	16.3	
12 to 17	F			67.2	60.2	73.6	19.9	15.0	25.8	10.2	7.3	14.1	
Boys Age group (years)	F			62.3	56.3	68.0	19.4	15.1	24.4	15.1	12.6	17.9	
5 to 11	F			59.0	51.9	65.7	19.8	14.8	26.0	19.5	15.5	24.1	
12 to 17	F			65.6	55.3	74.6	18.9 ^E	12.6	27.5	10.7*	7.5	15.0	
Girls Age group (years)	1.0 ^E	0.6	1.6	70.8	64.6	76.3	20.2	15.8	25.6	8.0†	5.7	11.1	
5 to 11	1.5 ^E	0.7	3.1	72.6 [†]	69.8	75.2	19.6	16.1	23.6	6.3 ^{†E}	4.1	9.8	
12 to 17	F			69.0	58.5	77.9	20.9	14.9	28.6	9.6 ^E	6.0	15.1	

^{*} significantly different from ages 5 to 11 (p<0.05)

Source: 2009 to 2011 Canadian Health Measures Survey.

[†] significantly different from boys (p<0.05)

E use with caution

F too unreliable to be published

^{...} not applicable

Table 2
Percentage distribution of children and adolescents, by body mass index category based on World Health Organization (WHO) and International Obesity Task Force (IOTF) cut-offs, age group and sex, household population aged 5 to 17, 2009 to 2011

			Thinr	ness				Noi	mal weigl	nt				0	verw	eight					Obe	sity		
		NHO			OTF			WHO		IOTF			١	WHO			IOTF			WHO			IOTF	
		95% confid inter	ence		95% confidenter	ence		95% confider interva		95 confid	lence			95% confider interva			95 confid inte	dence		95° confid inter	ence		95 confic	dence
Characteristics	%	from	to	%	from	to	%	from	to %	from	to		%	from	to	%	from	to to	%	from	to	%	from	ı to
Total	2.2 ^E	1.1	4.1	1.6 ^E	0.8	3.2	66.4	62.8 6	9.8 73.6	69.7	77.3	1	19.8	16.6 2	3.4	16.4	13.4	19.9	11.7	9.9	13.7	8.4	6.8	3 10.2
Age group (years) 5 to 11 12 to 17	F F			1.0 ^E F	0.5	1.9	65.5 67.2	61.7 6 60.2 7			79.9 77.0		19.7 19.9	16.4 2 15.0 2		14.7 18.0	12.1 13.8	17.9 3 23.1	13.1 10.2	10.5 7.3	16.3 14.1	7.9 8.9		3 10.5 3 12.3
Sex Boys Girls	F 1.0 ^E	0.6	 1.6	F 1.2 ^E	0.7	 2.2	62.3 70.8	56.3 6 64.6 7			78.6 79.9		19.4 20.2	15.1 2 15.8 2		15.8 17.0		21.1 21.8	15.1 8.0	12.6 * 5.7	17.9 11.1	9.5 7.1		1 12.2) 10.0

^{*} significantly different from boys (p<0.05)

Source: 2009 to 2011 Canadian Health Measures Survey.

classification cut-offs for children and adolescents aged 5 to 19.18

Although the IOTF classification has been used extensively, a systematic review has found that it underestimates obesity.22 Furthermore, the IOTF classification is only appropriate for use at the population level and cannot be used to assess excess weight at the individual level.13 The WHO growth charts¹⁸ have gained acceptance for use at the individual level, and in 2010, key professional associations recommended that health care professionals employ them to monitor the growth of Canadian children.²³ Adaptation and implementation of the WHO growth charts is underway in several jurisdictions (for example, British Columbia, Alberta, Saskatchewan, Yukon, New Brunswick and Nova Scotia).24 Estimating overweight and obesity based on the WHO growth references ensures that the methods used to determine excess weight in children and adolescents are consistent at the individual and population levels.

With measured height and weight data from the 2009 to 2011 CHMS, this report presents population estimates of overweight and obesity among Canadian children and adolescents based on the WHO cut-off values and compares them with the IOTF thresholds (see *The data*).

Obesity prevelance

According to the WHO approach, close to one third (31.5%) of 5- to 17-year-olds, an estimated 1.6 million, were classified as overweight (19.8%) or obese (11.7%) in 2009 to 2011 (Table 1). The percentage who were overweight was similar across age groups. However, the prevalence of obesity differed between boys and girls (15.1% versus 8.0%), most notably at ages 5 to 11, among whom the percentage of boys who were obese (19.5%) was more than three times

the percentage of girls who were obese (6.3%) (Table 1).

WHO versus IOTF approaches

The WHO cut-offs identified a greater percentage of children as overweight or obese than did the IOTF cut-offs: 31.5% versus 24.8% (Table 2). At ages 5 to 11, the difference was more pronounced than at ages 12 to 17. According to the WHO cut-offs, an estimated 32.8% of 5- to 11-year-olds were overweight or obese, compared with an estimated 22.6% based on the IOTF cut-offs.

Table 3
Mean body mass index (BMI) and percentage distribution by BMI category (based on World Health Organization cut-offs) of children and adolescents, household population aged 6 to 17, 2004, 2007 to 2009, and 2009 to 2011

		2004		2007	7 to 200)9	2009	9 to 20°	11
		95% confid inter	ence		95% confidenter	ence		95° confid inter	ence
Characteristics		from	to		from	to		from	to
Mean BMI	20.19	20.03	20.35	20.09	19.55	20.63	20.03	19.67	20.40
BMI category (%) Thinness Normal weight Overweight	1.4 ^E 63.8 21.4	1.0 61.9 19.9	2.0 65.7 23.1	1.6 ^E 66.4 17.7	0.8 60.4 13.9	3.2 71.9 22.2	2.3 ^E 66.6 19.5		4.5 70.3 23.6
Obesity	13.3	12.1	14.7	14.3	11.5	17.5	11.6	9.8	13.7

E use with caution

Note: There are no significant differences over time.

Source: 2004 Canadian Community Health Survey – Nutrition; 2007 to 2009 Canadian Health Measures Survey; 2009 to 2011 Canadian Health Measures Survey.

E use with caution

F too unreliable to be published

^{...} not applicable

A comparison of the classification systems showed that 72% of the children classified as obese based on the WHO approach would also be classified as obese based on the IOTF approach; the remaining 28% would be classified as overweight. Likewise, 66% of the children classified as overweight based on the WHO approach would also be classified as overweight based on the IOTF approach; the remaining 34% would be classified as normal weight.

The higher prevalence of obesity observed using the WHO approach is consistent with previous reports. 12,25 In a summary of the results of a number of studies, Reilly et al. noted that many of them demonstrated that the IOTF classification underestimates the prevalence of excess weight, particularly obesity, in children and adolescents. 22

No significant differences were observed in the estimates of overweight and obesity among children and adolescents aged 6 to 17 when data from the 2004 CCHS, the 2007 to 2009

CHMS, and the 2009 to 2011 CHMS were compared using the WHO cutoffs (Table 3) or IOTF cut-offs (data not shown).

This analysis concerns only one measure of adiposity—BMI. A recent Canadian study²⁶ showed that over time, waist circumference among Canadians of all ages has increased more than BMI. Evidence for adults indicates that changes in the distribution of body fat, such as increases in waist circumference, are associated with elevated health risk,²⁷ and suggests that even if the population prevalence of BMI does not change, changes in the distribution of body fat may increase health risk.²⁶

Conclusion

The factors associated with overweight and obesity are complex,⁷ and include health behaviours, such as eating habits and daily physical activity, and broader social, environmental and biological determinants that influence these health

behaviours.^{28,29} However, the sample size did not permit examination of trends in rates by these characteristics.

The 2009 to 2011 CHMS provides the most recent BMI data, based on measured height and weight, for children and adolescents in Canada. According to the WHO approach, close to a third of 5- to 17-year-olds were identified as overweight or obese, compared with about a quarter according to the IOTF Classification differences cut-offs. between approaches were greatest at ages 5 to 11. Although these estimates have not changed significantly in recent years. more data points are needed to determine if the pace of increase in prevalence is slowing, as has been observed in some countries.³⁰ Regardless, the estimates remain high and are a public health concern, given the tendency for excess weight in childhood to persist through to adulthood.

References

- Shields M. Overweight and obesity among children and youth. *Health Reports* 2006; 17(3): 27-42.
- Singh AS, Mulder C, Twisk JW, et al. Tracking of childhood overweight into adulthood: a systematic review of the literature. *Obesity Review* 2008; 9(5): 474-88.
- Reilly JJ, Methven E, McDowell ZC, et al. Health consequences of obesity. Archives of Disease in Childhood 2003; 88(9): 748-52.
- Whitaker RC, Wright JA, Pepe MS, et al. Predicting obesity in young adulthood from childhood and parental obesity. New England Journal of Medicine 1997; 337(13): 869-73.
- Nader PR, O'Brien M, Houts R, et al. Identifying risk for obesity in early childhood. Pediatrics 2006; 118(3): e594-601.
- Daniels SR, Greer FR, Committee on Nutrition. Lipid screening and cardiovascular health in childhood. *Pediatrics* 2008; 122(1): 198-208.
- Hall KD, Sacks G, Chandramohan D, et al. Quantification of the effect of energy imbalance on bodyweight. *Lancet* 2011; 378(9793): 826-37.
- Barlow SE, Expert Committee. Expert committee recommendations regarding the prevention, assessment, and treatment of child and adolescent overweight and obesity: summary report. *Pediatrics* 2007; 120(Suppl. 4): S164-92.
- Health Canada. Canadian Guidelines for Body Weight Classification in Adults (Catalogue H49-179). Ottawa: Health Canada, 2003.
- National Collaborating Centre for Primary Care, Centre for Public Health Excellence at NICE. Obesity—Guidance on the Prevention, Identification, Assessment and Management of Overweight and Obesity in Adults and Children. London: National Institute for Health and Clinical Excellence, 2006.
- Shields M, Gorber SC, Janssen I, Tremblay MS. Obesity estimates for children based on parent-reported versus direct measures. *Health Reports* 2011; 22(3): 47-58.
- Shields M, Tremblay MS. Canadian childhood obesity estimates based on WHO, IOTF and CDC cut-points. *International Journal of Pediatric Obesity* 2010; 5(3): 265-73.

- Cole TJ, Bellizzi MC, Flegal KM, Dietz WH. Establishing a standard definition for child overweight and obesity worldwide: international survey. *British Medical Journal* 2000; 320(7244): 1240-3.
- Cole TJ, Flegal KM, Nicholls D, Jackson AA. Body mass index cut offs to define thinness in children and adolescents: international survey. British Medical Journal 2007; 335(7612): 194.
- Giroux S. Canadian Health Measures Survey: Sampling strategy overview. *Health Reports* 2007; 18(Suppl): 31-6.
- Day B, Langlois R, Tremblay M, Knoppers BM. Canadian Health Measures Survey: Ethical, legal and social issues. *Health Reports* 2007; 18(Suppl.): 37-51.
- Statistics Canada. Canadian Health Measures Survey (CHMS) Data User Guide: Cycle 2, September 2012. Available at: www.statcan.gc.ca.
- de Onis M, Onyango AW, Borghi E, et al. Development of a WHO growth reference for school-aged children and adolescents. Bulletin of the World Health Organization 2007; 85(9): 660-7.
- World Health Organization. Training Course on Child Growth Assessment. Geneva: World Health Organization, 2008.
- Rao JNK, Wu CFJ, Yue K. Some recent work on resampling methods for complex surveys. Survey Methodology 1992; 18(2): 209-17.
- Rust KF, Rao JNK. Variance estimation for complex surveys using replication techniques. Statistical Methods in Medical Research 1996; 5: 281-310.
- Reilly JJ, Kelly J, Wilson DC. Accuracy of simple clinical and epidemiological definitions of childhood obesity: systematic review and evidence appraisal. *Obesity Reviews* 2010; 11(9): 645-55.
- Dietitians of Canada, Canadian Paediatric Society, College of Family Physicians of Canada, et al. Promoting optimal monitoring of child growth in Canada: using the new WHO growth charts. Canadian Journal of Dietetic Practice and Research 2010; 71(1): e1-3.

- 24. Public Health Agency of Canada. Actions Taken and Future Directions 2011 - Curbing Childhood Obesity: A Federal, Provincial and Territorial Framework for Action to Promote Healthy Weights. Ottawa: Public Health Agency of Canada, 2011. 25.
- Twells LK, Newhook LA. Obesity prevalence estimates in a Canadian regional population of preschool children using variant growth references. BMC Pediatrics 2011; 11: 21.
- Janssen I, Shields M, Craig CL, Tremblay MS. Changes in the obesity phenotype within Canadian children and adults, 1981 to 2007-2009. Obesity (Silver Spring) 2012; 20(4): 916-9.
- Janssen I, Heymsfield SB, Ross R. Application of simple anthropometry in the assessment of health risk: implications for the Canadian Physical Activity, Fitness and Lifestyle Appraisal. *Canadian Journal of Applied Physiology* 2002; 27(4): 396-414.
- Mozaffarian D, Hao T, Rimm EB, et al. Changes in diet and lifestyle and long-term weight gain in women and men. New England Journal of Medicine 2011; 364(25): 2392-2404.
- Prince SA, Kristjansson EA, Russell K, et al. Relationships between neighborhoods, physical activity, and obesity: A multilevel analysis of a large Canadian city. *Obesity* (Silver Spring) 2012; January 19 (epub ahead of print).
- Olds T, Maher C, Zumin S, et al. Evidence that the prevalence of childhood overweight is plateauing: data from nine countries. *International Journal of Pediatric Obesity* 2011; 6(5-6): 342-60.

ELECTRONIC PUBLICATIONS AVAILABLE AT

www.statcan.gc.ca

Area-based methods to calculate hospitalization rates for the foreign-born population in Canada, 2005/2006

by Gisèle Carrière, Paul A. Peters and Claudia Sanmartin

Abstract

Background

Hospital records lack information about country of birth. This study describes a method for calculating hospitalization rates by the percentage of foreign-born in Census Dissemination Areas (DAs).

Data and methods

Data from the 2006 Census were used to classify DAs by the percentage of the foreign-born population who lived in them. Quintile and tercile thresholds were created to classify DAs as having low to high percentages of foreign-born residents. This information was appended to the 2005/2006 Hospital Morbidity Database via postal codes. Age-sex standardized hospitalization rates were calculated for low to high foreign-born concentration DAs, nationally and subnationally.

Results

Nationally, quintile thresholds had better discriminatory power to detect variations in hospitalization rates by foreign-born concentration, but tercile thresholds produced reliable results at subnational levels. All-cause hospitalization rates were lowest among residents of the high foreign-born concentration terciles. Similar gradients emerged in hospitalization rates for heart disease, diseases of the circulatory system, and mental health conditions. The pattern varied more at the subnational level.

Interpretation

With this approach, administrative data can be used to calculate hospitalization rates by foreign-born concentration.

Keywords

Administrative data, ecological studies, hospital records, immigration, public health surveillance

Authors

Gisèle Carrière (1-604-666-5907; gisele.carriere@ statcan.gc.ca) is with the Health Analysis Division at Statistics Canada, Vancouver, British Columbia, V6B 6C7. Paul A. Peters and Claudia Sanmartin are with the Health Analysis Division at Statistics Canada, Ottawa, Ontario, K1A 0T6.

By 2031, it is projected that 28% of Canada's population could be foreign-born, up from about 20% in 2006.¹ Understanding patterns of health care use among this growing segment of the population is important for the planning and delivery of services. While evidence suggests better health among the foreign-born compared with people born in Canada,²-9 much of that research is based on survey data.³-6,¹0,¹¹¹ Those studies are typically constrained by small sample sizes that limit area-level comparisons. Furthermore, analysis based on survey data may be subject to recall bias or affected by linguistic and cultural barriers.

Hospital administrative records cover all acute-care hospitalizations and allow analysis at detailed levels of geography. However, these records do not contain information about patients' country of birth and immigration status. In the absence of individual-level information, area-based methods can be applied to study hospital use patterns for areas with high concentrations of foreign-born individuals. Such approaches have been used in Canada to analyze health outcomes by neighbourhood socioeconomic status and concentration of the Aboriginal population. 12-14

Because immigrants have tended to settle in large urban areas, ¹⁵ and because the concentration of the foreign-born has increased over time, ¹⁶⁻¹⁸ area-based methods can be applied to study hospitalization patterns in areas with high concentrations of foreign-born individuals. While hospital data for specific regions have been linked to immigration data, ² such work has not been undertaken at the national level.

This study describes an area-based method of calculating standardized, comparable hospitalization rates for areas with varying concentrations of foreign-born, at national and subnational levels.¹⁹ Based on previous research,²⁻⁷ the hypothesis is that hospitalization rates are likely to be lower in areas with high percentages of foreign-born residents.

Methods

Data sources

Counts of the foreign-born population are from the 2006 Census long-form questionnaire, which was administered to 20% of households (non-institutionalized population). The information collected in the questionnaire included country of birth, immigrant status, and socio-economic and demographic characteristics.

Hospitalization data are from the Hospital Morbidity Database, which covers all inpatient acute-care hospital discharges in Canada. The data are compiled by the Canadian Institute of Health Information. Hospital records from fiscal year 2005/2006 (April 1, 2005 through March 31, 2006) were used because they are closest in time to the census year. These records contain medical information such as diagnoses and procedures, and patient information such as date of birth, sex, and importantly for this analysis, postal codes. character postal codes are available for all provinces/territories except Quebec, for which only the first three characters are provided.

Definition of foreign-born

In this study, "foreign-born" refers to those who either (1) ever held the legal status of immigrant to Canada, or (2) were non-permanent residents (NPRs). NPRs are people from another country who, at the time of the census, held a work or study permit or were refugee claimants, or who had applied for landed immigrant status but had not yet been accepted, as well as family members living with them in Canada.²⁰ From a health perspective, NPRs more closely resemble immigrants than people born in Canada, so they are combined with immigrants to represent the foreign-born. In 2006, NPRs made up 4% of the foreign-born population,

and less than 1% of the total Canadian population.

Level of geography

This analysis requires that areas with high percentages of foreign-born residents be distinguished from areas with low percentages of foreign-born residents. A small geographic unit is needed because population homogeneity across an area tends to increase with geographic size,²¹ thereby diluting associations between foreign-born concentration and hospitalization rates.

To some extent, the foreign-born population in Canada is spatially concentrated. Immigrants tend to settle in Census Metropolitan Areas (CMAs), such as Toronto and Vancouver, 15 and many remain in these "gateway" cities. 22

This analysis is based on Dissemination Areas (DAs), the smallest level of geography (400 to 700 residents) for which aggregate census information is available.²³ In 2006, DAs totalled 54,626; aggregate information about

population characteristics was available for 92% of them (50,214).

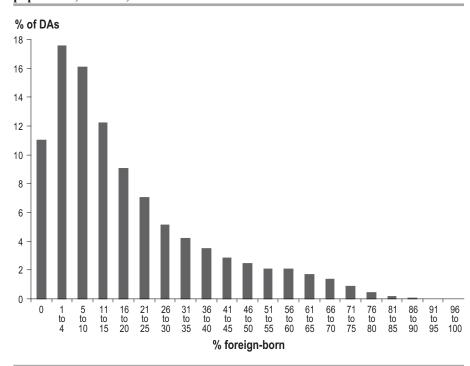
Development of area-based measure of foreign-born

Development of an area-based measure of foreign-born concentration involves three steps: 1) calculation of foreign-born population concentrations; 2) selection of the base population to measure the distribution of concentration values; and 3) selection of a quantile to delineate thresholds that will be used to classify areas according to their foreign-born concentration.

Calculation of foreign-born population concentration

Foreign-born concentration is the percentage of individuals in each DA in 2006 who were foreign-born. In 45% of DAs, no more than 10% of the population were foreign-born, on the other hand, in about 9% of DAs, at least 50% of the population were foreign-born (Figure 1).

Figure 1
Distribution of Dissemination Areas (DAs), by percentage foreign-born in DA population, Canada, 2006



Note: Foreign-born include those with landed immigrant status (ever) and non-permanent residents. **Source:** 2006 Census of Population.

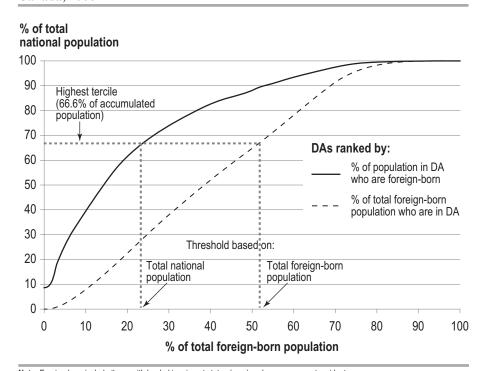
Selection of base population

To establish concentration thresholds, all DAs across Canada were ordered from lowest to highest DA percentage of foreign-born. Next, this percentage foreign-born (the DA measure) was distributed across each of two possible base populations: the total national population and the total foreignborn population. By distributing the percentage foreign-born measure this way, a specific foreign-born percentage value could be discerned to divide the base population into a given quantile, for example, thirds (terciles) or fifths (quintiles). The foreign-born population is preferred as the base population because the resulting thresholds that define any given quantile contain greater foreignborn concentrations than do thresholds derived if the total national population had been used as the base. The advantage of using the foreign-born population as the base population is illustrated in Figure 2. If the total national population is used as the base population, thresholds that delineate areas that contain terciles of the total population show that the highest tercile (where 66.6% of the national population has accumulated) cut-off occurs when the percentage foreign-born DA measure is equal to 24%. That is, DAs in which at least 24% of the population are foreign-born would be considered "high concentration of foreign-born." However, thresholds that delineate areas containing thirds of the foreign-born population yield a "highconcentration" cut-off of 52% foreignborn. Given these results, the distribution of the percentage foreign-born across the foreign-born population was used to establish thresholds to classify each DA.

Selection of quantile

The choice of quantiles (thirds or fifths of the foreign-born) to set foreign-born concentration thresholds for classifying DAs was based on two criteria: 1) the range of percentage foreign-born within a given quantile; and 2) the feasibility of calculating hospitalization rates at national and sub-national (provincial/regional/territories or CMA) levels.

Figure 2 Cumulative distribution of percentage foreign-born in Dissemination Areas (DAs), by percentage of total national or total foreign-born population, Canada, 2006



Note: Foreign-born include those with landed immigrant status (ever) and non-permanent residents. **Source:** 2006 Census of Population.

Because admission to hospital is a relatively rare event—on average, fewer than one person in ten is hospitalized in any year—population counts must be large enough to produce stable estimates.

Thresholds for terciles and quintiles are presented in Table 1. By definition, each level of either quantile contains approximately equal numbers of the foreign-born: about 2.1 million in each tercile, and about 1.3 million in each quintile. All DAs across Canada were ordered from lowest to highest percentage of the foreign-born population in each, and the foreign-born population was then classified into terciles and quintiles. While the absolute number of foreign-born individuals is the same in each level of a quantile, the percentage of the total population in each level who are foreign-born (concentration) varies. For example, while approximately 2.1 million foreign-born are in each tercile, they make up 9.8% of the total

population in the "low-concentration" tercile, but 63.7% of the total population in the "high-concentration" tercile.

For the *quintile* thresholds, quintile 1 (lowest foreign-born concentration) consists of DAs in which the percentage of the population who were foreign-born was 19.0% or less; quintile 5 (highest foreign-born concentration) consists of DAs in which the percentage of the population who were foreign-born was more than 62.0%. The percentage of the population who were foreign-born in the lowest *tercile* is less than 27.0%, and in the highest, more than 51.8%.

Hospitalization rates were calculated based on quintiles and terciles (Table 1). Quintile-based thresholds result in better discrimination of national all-cause hospitalization rate differences between the highest and lowest quantiles (rate ratio=0.64) than do tercile-based thresholds (rate ratio=0.69). For both quantiles, 95% confidence intervals

Area-based methods to calculate hospitalization rates for the foreign-born population in Canada, 2005/2006 • Methodological insights

Table 1 Age-sex standardized all-cause hospitalization rates (per 10,000), by quantile of percentage of foreign-born population, Canada, 2005/2006

	Forei	gn-born				i%			5%
Foreign-born		% of quantile	% foreign-born	Hospitalization	confidence inte	rval	Rate	confic	rval
concentration quantile	Number populat		range	rate	from	to	ratio	from	to
Canada	6,377,250	20.8		753.0	752.0	753.9			
Quintile									
1 (lowest)	1,282,165	7.1	19.0 or less	840.7	839.4	842.0	1.00		
2	1,270,545	24.4	More than 19.0 to 32.0	663.8*	661.6	666.0	0.79	0.79	0.79
3	1,277,825	38.3	More than 32.0 to 46.0	621.8*	619.1	624.5	0.74	0.74	0.74
4	1,278,465	53.8	More than 46.0 to 62.0	586.7*	583.6	589.9	0.70	0.69	0.70
5 (highest)	1,268,250	69.8	More than 62.0	539.0*	536.4	543.5	0.64	0.64	0.65
Tercile									
1 (lowest)	2,124,740	9.8	27.0 or less	814.0	812.9	815.2	1.00		
2	2,137,055	37.6	More than 27.0 to 51.8	621.9*	619.8	623.9	0.76	0.76	0.77
3 (highest)	2,115,455	63.7	More than 51.8	558.5*	555.9	561.2	0.69	0.68	0.69
Selected Census Metropolitan Areas [†]									
Total	4,600,320	34.1		616.2	614.9	617.6			
1 (lowest)	814,030	14.0	27.0 or less	655.4	653.3	657.5	1.00		
2	1,708,150	38.7	More than 27.0 to 51.8	607.9*	605.6	610.3	0.93	0.92	0.93
3 (highest)	2,078,140	63.9	More than 51.8	555.4*	552.8	558.1	0.85	0.84	0.85

[†] Halifax, Montreal, Toronto, Hamilton, Winnipeg, Calgary, Vancouver

Notes: Excludes pregnancy-related conditions. Rates are age-sex standardized to 2006 national population. Foreign-born include those with landed immigrant status (ever) and non-permanent residents. Source: 2006 Census of Population; 2005/2006 Hospital Morbidity Database.

indicate that the national hospitalization rates were stable. However, at the subnational level (province/region/territory and CMA), rates were less stable (wider confidence intervals) based on quintiles than on terciles, particularly for areas with smaller populations (total and foreign-born). For example, in Alberta, hospitalization rates produced using tercile thresholds had a smaller standard error for areas with the highest foreign-born concentration than did rates produced using quintile thresholds (data not shown). As a result, the tercile measure was selected for this analysis.

Criteria for immigrant concentration within DAs could also be defined at the sub-national level (provincial/regional or CMA) to produce jurisdiction-specific thresholds. These would be more sensitive to the spatial distribution of immigrants within each jurisdiction. Jurisdiction-specific terciles were created for provinces/regions and selected CMAs. A comparison of threshold levels revealed significant variation in the definition of high, medium and

low concentration of immigrants. For example, thresholds for including DAs in the high foreign-born concentration tercile range from a low of 10.7% in the Atlantic Region to a high of 57.8% in Ontario (data not shown), whereas the nationally derived threshold is 51%. This variation limits the ability to compare hospitalization rates across jurisdictions within similar terciles owing to cross-classification of DAs. For example,

DAs with 25% foreign-born would be considered medium-concentration in some areas, but high-concentration in others. Because the objective of this study is to produce nationally comparable results for each level of geography, the analyses are based on nationally defined terciles.

Table 2 Age-sex standardized all-cause hospitalization rates (per 10,000), by selected causes and foreign-born concentration tercile, Canada, 2005/2006

		culator		co	Heart Indition	S	Mental health behavioura disorders		al
Foreign-born concentration		confi	5% dence rval		95 confic inte	lence		95 confic inte	dence
tercile	Rate	from	to	Rate	from	to	Rate	from	to
Canada	120.1	119.7	120.5	66.9	66.6	67.2	50.7	50.4	50.9
Low foreign-born	128.4	128.0	128.9	71.5	71.2	71.9	54.7	54.3	55.0
Medium foreign-born	100.0*	99.2	100.8	54.9*	54.3	55.5	43.8*	43.3	44.4
High foreign-born	93.4*	92.3	94.5	53.2*	52.4	54.1	37.1*	36.4	37.7

^{*} significantly different from low foreign-born concentration tercile (p<0.05)

Note: Foreign-born include those with landed immigrant status (ever) and non-permanent residents.

Source: 2006 Census of Population; 2005/2006 Hospital Morbidity Database.

^{*} significantly different from lowest foreign-born concentration quantile (p<0.05)

^{..} not applicable

Applying the area-based measure to hospital data

Using the PCCF+ application developed at Statistics Canada, a 2006 Census DA code was assigned to each hospital separation record based on the patient's residential postal code.^{24,25} The DA code was used to classify each hospital record into a foreign-born concentration tercile.

Causes of hospitalization were determined from the "most responsible diagnosis" (excluding pregnancyrelated) coded either to the International Statistical Classification of Diseases and Related Health Problems, Tenth Canada $(ICD-10)^{26}$ Revision, to the International Classification of Diseases, 9th Revision, Clinical Modification (ICD-9)²⁷ for Quebec. The following cause-specific hospitalizations were defined: circulatory system diseases; selected heart conditions (heart failure, pulmonary edema; ischemic heart disease including acute myocardial infarction); and mental and behavioural disorders. These causes were chosen for analysis because earlier research reported differences in the prevalence between ethnic minority populations and other Canadians.¹¹ and because of the need for information about the mental health of immigrants.28

Hospitalization rates were calculated as the number of hospitalizations per 10,000 population in DAs with high, medium or low percentages of foreignborn residents. The number of hospital separations and the corresponding population denominators are shown in Appendix Table A for Canada and by province/region/territory and selected CMAs, based on nationally defined foreign-born concentration terciles. No DAs in the Atlantic Region had a "high" concentration of foreign-born residents. In Saskatchewan, fewer than 700 people were in the high foreign-born concentration tercile.

Rates were standardized to the age and sex structure of the 2006 population of Canada using the direct method. Age-sex standardized rates and 95% confidence intervals were calculated for Canada, for each province/region/territory, and for

Table 3
Age-sex standardized all-cause hospitalization rates (per 10,000), by province/region/territory and foreign-born concentration tercile, Canada, 2005/2006

Province/Region/Territory and foreign-born	Hospitalization		nfidence erval	Rate
concentration tercile	rate	from	to	ratio
Canada	753.0	752.0	753.9	
Atlantic	931.2	927.4	935.2	
Low foreign-born	931.0	927.1	934.9	1.00
Medium foreign-born	965.4	903.6	1,031.4	1.04
High foreign-born				
Quebec	717.9	716.0	719.8	
Low foreign-born	742.2	740.1	744.3	1.00
Medium foreign-born	597.6*	592.6	602.6	0.81
High foreign-born	592.2*	583.5	601.0	0.80
Ontario	698.3	696.8	699.8	
Low foreign-born	773.4	771.3	775.4	1.00
Medium foreign-born	620.6*	617.7	623.5	0.80
High foreign-born	560.9*	557.7	564.1	0.73
Manitoba	813.6	808.3	818.9	
Low foreign-born	835.5	829.7	841.3	1.00
Medium foreign-born	679.5*	666.3	692.9	0.81
High foreign-born	703.5*	645.7	766.5	0.84
Saskatchewan	1,059.0	1,052.5	1,065.5	
Low foreign-born	1,059.1	1,052.6	1,065.7	1.00
Medium foreign-born	1,111.2	1,035.5	1,192.4	1.05
High foreign-born [†]	F	F	F	F
Alberta	846.0	842.7	849.4	
Low foreign-born	898.3	894.5	902.2	1.00
Medium foreign-born	632.9*	626.0	639.9	0.70
High foreign-born	617.2*	592.4	643.1	0.69
British Columbia	709.0	706.4	711.6	
Low foreign-born	807.1	803.4	810.7	1.00
Medium foreign-born	616.6*	612.1	621.2	0.76
High foreign-born	518.3*	512.6	524.0	0.64
Territories and Nunavut	990.3	962.7	1,018.7	
Low foreign-born	990.3	962.7	1,018.7	1.00
Medium foreign-born				
High foreign-born				

 $^{^{\}scriptscriptstyle \dagger}$ based on one Dissemination Area (population of 695) and 13 hospitalizations; use with caution

Notes: Excludes pregnancy-related conditions. Foreign-born include those with landed immigrant status (ever) and non-permanent residents

Source: 2006 Census of Population; 2005/2006 Hospital Morbidity Database.

selected CMAs. Confidence intervals for the standardized rates used methods derived from Spiegelman.²⁹

Results

The national all-cause acute-care hospitalization rate was lowest among residents of DAs in the high foreign-born concentration tercile (559 hospitalizations per 10,000 population), and highest among residents of DAs

in the low foreign-born concentration tercile (814 hospitalizations per 10,000 population) (Table 1). This pattern persisted for hospitalizations due to circulatory system diseases, selected heart conditions, and mental and behavioural disorders (Table 2).

Patterns in Ontario, Alberta and British Columbia were similar to the national level, with lower hospitalization rates among residents of areas classified in the high or medium foreign-born

^{*} significantly different from low foreign-born concentration tercile (p<0.05)

F too unreliable to be published

^{...} not applicable

Area-based methods to calculate hospitalization rates for the foreign-born population in Canada, 2005/2006 • Methodological insights

Table 4
Age-sex standardized all-cause hospitalization rates (per 10,000), by selected Census Metropolitan Area (CMA) and foreign-born concentration tercile, Canada, 2005/2006

CMA and foreign-born	Hospitalization	95% con inter		Rate
concentration tercile	rate	from	to	ratio
Selected CMAs	616.2	614.9	617.6	
Vancouver	596.8	593.4	600.1	
Low foreign-born	684.6	677.5	691.9	1.00
Medium foreign-born	602.7*	597.7	607.7	0.88
High foreign-born	516.1*	510.4	521.8	0.75
Calgary	660.5	655.2	665.9	
Low foreign-born	678.2	671.6	685.0	1.00
Medium foreign-born	634.6*	625.1	644.2	0.94
High foreign-born	553.9*	527.2	581.9	0.82
Toronto	587.9	585.7	590.1	
Low foreign-born	639.4	634.0	644.8	1.00
Medium foreign-born	596.7*	593.2	600.3	0.93
High foreign-born	559.0*	555.8	562.3	0.87
Montreal	633.6	631.0	636.2	
Low foreign-born	653.0	649.7	656.3	1.00
Medium foreign-born	595.9*	590.9	600.9	0.91
High foreign-born	594.4*	585.6	603.2	0.91
Halifax	612.8	604.6	621.0	
Low foreign-born	613.2	605.0	621.6	1.00
Medium foreign-born	576.0	509.2	651.6	0.94
High foreign-born				
Hamilton	690.2	684.2	696.4	
Low foreign-born	690.9	683.2	698.7	1.00
Medium foreign-born	691.3	681.0	701.7	1.00
High foreign-born	699.1	656.3	744.8	1.01
Winnipeg	643.6	637.7	649.6	
Low foreign-born	638.1	631.5	644.8	1.00
Medium foreign-born	661.8*	648.3	675.6	1.04
High foreign-born	689.8	632.7	752.1	1.08

^{*} significantly different from low foreign-born concentration tercile (p<0.05)

Notes: Excludes pregnancy-related conditions. Foreign-born include those with landed immigrant status (ever) and non-permanent residents.

Source: 2006 Census of Population; 2005/2006 Hospital Morbidity Database.

concentration terciles, compared with residents of areas in the low foreign-born concentration tercile (Table 3). In Quebec and Manitoba, hospitalization rates for residents of the medium and high foreign-born terciles differed significantly from each other (data not shown), and both differed significantly from the low foreign-born concentration tercile. As well, in Manitoba, the lowest hospitalization rate was among residents of the medium foreign-born concentration tercile.

All-cause hospitalization rates by foreign-born concentration also varied

across CMAs (Table 4). The national pattern prevailed in Vancouver, Calgary, Toronto, and to a lesser extent, Montreal—the lowest hospitalization rates were among residents of areas classified in the high foreign-born concentration tercile. Differences across terciles were not significant in Hamilton or Halifax.

Discussion and limitations

This study demonstrates how an areabased method can be used to examine

Why is this study important?

- A growing percentage of the Canadian population is foreign-born.
- Understanding the health and health services use patterns of this population is increasingly important.
- Administrative health data typically do not contain information about country of birth.

What is already known on this subject?

In the absence of individual-level information, ecological methods have been applied to understand patterns of hospital use among areas having greater or lesser percentages of subpopulations.

What this study adds

- Hospital records that lack information on country of birth can be analysed with aggregate census data to compare hospitalization rates for areas with greater or lesser percentages of foreign-born residents.
- Hospitalization rates tend to be lowest among residents of areas with a high percentage of foreign-born residents.

hospitalization rates in areas with high versus low percentages of foreignborn residents. The approach yields comparable information at national and sub-national levels.

Research generally suggests that it would be reasonable to expect lower hospitalization rates (excluding pregnancy-related) in areas with high percentages of foreign-born residents. Individual-level data show better self-reported health, lower prevalence of chronic conditions, lower age-specific mortality risks, and longer life expectancy among immigrants compared with the

^{...} not applicable

Canadian-born.^{3,7,9,30} The distribution of hospitalization rates in this analysis supports that expectation.

Differences in hospitalization rates between high and low foreignborn concentration terciles should be interpreted in the context of area characteristics, including the composition of the foreign-born population (region or country of birth). Recent studies have reported differences in health status by country of birth and time since immigration.⁶⁻⁸ As well, some segments of the foreign-born may be at higher risk of hospitalization than are others because of poorer health and higher rates of chronic disease.^{4,7,9,11,31,32} Given this evidence, differences in the composition of the foreign-born population are relevant in interpreting differences in hospitalization rates across immigrant terciles. For example, in 2006, more than 70% of immigrants in Vancouver reported that their country of birth was in Asia, compared with 31% of those in Montreal. In Montreal, more than 25% of immigrants reported being of Africa origin, higher than any other jurisdiction in Canada.²³ Differences in the composition of the foreign-born populations may explain variations in the gradient of hospitalization rates across jurisdictions. The steeper gradient between high and medium tercile hospitalization rates in Vancouver than in Montreal, for example, may be due to the higher percentage of Asianborn immigrants in Vancouver. Other

area-level factors that have been found to be associated with higher rates of hospitalization include low-income and a high concentration of Aboriginal peoples.³³

A key limitation of this study is that the hospitalization rates cannot be regarded as rates for foreign-born and non-foreign-born individuals per se, but rather, as rates among people living in areas with varying concentrations of the foreign-born. The analysis suggests an association between hospitalization and percentages of foreign-born in the population, but does not allow for causal inferences.

Furthermore, the hospital discharge records that Statistics Canada receives from Quebec contain only the first three characters of the postal codes. PCCF+uses population weights to probabilistically assign cases to DAs. This method results in less precise matching than would the full six-character postal code. It may result in greater misclassification of hospital records in urban areas, but has little effect in rural areas which are predominately in the low foreign-born concentration tercile.

Conclusion

The results confirm that this areabased methodology can be employed to compare hospitalization rates among areas having greater versus lesser concentrations of foreign-born than Canadian-born populations. An advantage of this approach is its use of existing data, which makes it a costeffective method for routine surveillance of health care utilization. As well, the measure allows for comparisons between different geographic areas, particularly those with high foreign-born concentrations.

Finally, definitive analysis of health services use by the foreign-born awaits the creation of administrative data with person-level information such as country of birth, year of immigration, income, and educational attainment. The linkage of health care administrative records to other Statistics Canada data holdings such as the Census under the Longitudinal Health and Administrative Data Initiative should address this information gap.

Acknowledgements

The authors thank the acute-care hospitals of each province and territory, which create the discharge abstracts, and the Canadian Institute for Health Information, which provides edited machine-readable hospital morbidity files to Statistics Canada. The Institute's Janet Manuel provided valuable help with the cause groupings. Dr. Elizabeth Muggah provided important comments about analytical content. Russell Wilkins contributed substantially to an earlier version of this paper.

Area-based methods to calculate hospitalization rates for the foreign-born population in Canada, 2005/2006 • Methodological insights

References

- Malenfant ÉC, Lebel A, Martel L. Projections of the Diversity of the Canadian Population, 2006 to 2031 (Catalogue 91-551-X) Ottawa: Statistics Canada, 2010.
- DesMeules M, Gold J, Kazanjian A, et al. New approaches to immigrant health assessment. Canadian Journal of Public Health 2004; 95(3): 122-6.
- Perez CE. Health status and health behaviours among immigrants. *Health Reports* 2002; 13 (Suppl.): 1-13.
- Newbold KB. Self-rated health within the Canadian immigrant population: risk and the healthy immigrant effect. Social Science and Medicine 2005; 60(6): 1359-70.
- Hyman I. *Immigration and Health*. Health Policy Working Paper Series. Working Paper 01-05. Ottawa: Health Canada, 2001.
- Ng E, Wilkins R, Gendron F. Dynamics of immigrant's health in Canada: Evidence from the National Population Health Survey. Healthy Today, Healthy Tomorrow? Findings from the National Population Health Survey (Catalogue 82-618-MWE) Ottawa: Statistics Canada, 2005.
- Ng E. The healthy immigrant effect and mortality rates. *Health Reports* 2011; 22(4): 25-9.
- Creatore MI, Moineddin R, Booth G, et al. Age- and sex-related prevalence of diabetes mellitus among immigrants to Ontario, Canada. *Canadian Medical Association Journal*; 182(8): 781-9. doi:10.1503/ cmaj.091551.
- Wilkins R, Tjepkema M, Mustard C, Choinière R. The Canadian census mortality follow-up study, 1991 through 2001. *Health Reports* 2008; 19(3): 25-43.
- Rotermann M. The impact of considering birthplace in analyses of immigrant health. Health Reports 2011; 22(4): 37-43.
- Chiu M, Austin PC, Manuel DG, Tu J. Comparison of cardiovascular risk profiles among ethnic groups using population health surveys between 1996 and 2007. Canadian Medical Association Journal; 182(8): E301-10. doi:10.1503/cmaj.091676.
- Berthelot JM, Ross NA, Tremblay S. Regional socio-economic context and health. *Health Reports* 2002; 13(Suppl): 1-12.
- Wilkins R, Berthelot JM, Ng E. Trends in mortality by neighbourhood income in urban Canada from 1971 to 1996. *Health Reports* 2002; 13(Suppl): 1-27.
- Hou F, Chen J. Neighbourhood low income, income inequality and health in Toronto. Health Reports 2003; 14(2): 21-34.

- Chui T, Tran K, Maheux H. Immigration in Canada: A Portrait of the Foreign-Born Population, 2006 Census (Catalogue 97-557-XIE) Ottawa: Statistics Canada. 2007.
- Davies WKD, Murdie RA. 1993. Measuring the social ecology of cities. In: Bourne LS, Ley DF, eds. *The Changing Social Geography of Canadian Cities*. Montreal: McGill-Queens University Press.
- Hou F, Myles J. Neighbourhood inequality, relative deprivation and self-perceived health status. *Analytical Studies Branch– Research Paper Series*, No. 228 (Catalogue 11F0019MIE) Ottawa: Statistics Canada, 2004.
- Hou F. Recent immigration and the formation of visible minority neighbourhoods in Canada's largest cities. Neighbourhood inequality, relative deprivation and self-perceived health status. *Analytical Studies Branch–Research Paper Series*, No. 221 (Catalogue 11F0019MIE) Ottawa: Statistics Canada, 2004.
- Peters PA, Oliver LN, Carrière G M. Geozones: An area-based method for analysis of health outcomes. *Health Reports* 2012; 23(1): 55-64.
- Statistics Canada. 2006 Census Dictionary (Catalogue 92-566-X) Ottawa: Statistics Canada, 2010.
- Johnston R, Voas D, Poulsen M. Measuring spatial concentration: the use of threshold profiles. *Environment and Planning B: Planning and Design* 2003; 30: 3-14.
- Hou F. The initial destinations and redistribution of Canada's major foreign-born groups: Changes over the past two decades. *Analytical Studies – Research Paper Series*, No. 254 (Catalogue 11F0019MIE) Ottawa: Statistics Canada, 2005.
- Statistics Canada. Profiles for Canada, Provinces, Territories, Census Divisions, Census Subdivisions and Dissemination Areas, 2006 Census (Catalogue 94-581-XCB2006) Ottawa: Statistics Canada, 2008.
- Wilkins R, Khan S. PCCF + Version 5G User's Guide (Catalogue 82F0086-XDB) Ottawa: Statistics Canada, 2010. Available at: http://abacus.library.ubc.ca/ bitstream/10573/42320/18/msword.pccf5g. pdf. Accessed March 20, 2011.
- Wilkins R, Khan S. PCCF + Version 5H User's Guide (Catalogue 82F0086-XDB) Ottawa: Statistics Canada, 2011. Available at: http://abacus.library.ubc.ca/ bitstream/10573/42442/3/msword.pccf5h. pdf. Accessed March 20, 2011.

- World Health Organization and Canadian Institute for Health Information. *International* Statistical Classification of Diseases and Related Health Problems, Tenth Revision, Canada. Ottawa: Canadian Institute for Health Information, 2009.
- World Health Organization and National Center for Health Statistics. *International Classification of Diseases*, 9th Revision, Clinical Modification, Sixth Edition. Washington, DC: United States Government Printing Office, 2008.
- Hyman I. Setting the stage: Reviewing current knowledge on the health of Canadian immigrants. What is the evidence and where are the gaps? *Canadian Journal of Public Health* 2004; 95(3): I4-8.
- Spiegelman M. Introduction to Demography, Revised Edition. Cambridge Massachusetts: Harvard University Press, 1968.
- Chen J, Wilkins R, Ng E. Life expectancy of Canada's immigrants from 1986 to 1991. Health Reports 1996; 8(3): 29-38.
- Tomlin AM, Tilyard MW, Dovey SM, Dawson AG. Hospital admissions in diabetic and non-diabetic patients: A case-control study. *Diabetes Research and Clinical Practice* 2006; 73(3): 260-7.
- 32. Bo S, Ciccone G, Grassi G, et al. Patients with type 2 diabetes had higher rates of hospitalization than the general population. *Journal of Clinical Epidemiology* 2004; 57(11): 1196-201.
- Carrière G, Garner R, Sanmartin C, LHAD Research Team. Acute-care Hospitalizations and Aboriginal Identity in Canada, 2001/2002. Health Research Working Paper Series (Catalogue 82-622-X, No. 005) Ottawa: Statistics Canada, 2010. Available at: http://www.statcan.gc.ca/pub/82-622-x/82-622-x2010005-eng.pdf. Accessed March 20, 2011.

Area-based methods to calculate hospitalization rates for the foreign-born population in Canada, 2005/2006 • Methodological insights

Appendix

Table A Population and all-cause hospitalizations, by foreign-born concentration tercile, province/region/territory and selected Census Metropolitan Areas, Canada, 2005/2006

		Popula	ation 2006			Hospitaliza	ntions 2005/200	6
	Total	Tercile 1 (lowest % foreign-born)	Tercile 2 (medium % foreign-born)	Tercile 3 (highest % foreign-born)	Total	Tercile 1 (lowest % foreign-born)	Tercile 2 (medium % foreign-born)	Tercile 3 (highest % foreign-born)
Province/Region/Territory								
Atlantic	2,274,315	2,263,740	10,575		220,292	219,352	940	
Quebec	7,480,310	6,269,040	906,290	304,980	546,010	471,809	56,282	17,919
Ontario	12,077,010	6,814,775	2,918,810	2,343,425	835,330	539,592	175,488	120,250
Manitoba	1,090,955	925,255	156,895	8,805	91,991	81,063	10,369	559
Saskatchewan	917,545	909,095	7,755	695	104,148	103,163	972	13
Alberta	3,219,455	2,540,690	634,035	44,730	247,682	211,111	34,152	2,419
British Columbia	4,009,085	2,273,955	1,100,105	635,025	294,239	191,963	69,910	32,366
Territories	100,510	100,510			7,154	7,154		
Census Metropolitan Area								
Vancouver	2,095,580	526,240	944,370	624,970	122,791	35,161	55,869	31,761
Calgary	1,075,680	683,245	356,110	36,325	61,286	40,951	18,642	1,693
Toronto	5,093,485	923,815	1,886,145	2,283,525	279,335	55,128	107,344	116,863
Montreal	3,613,520	2,417,560	890,980	304,980	226,116	152,812	55,320	17,984
Halifax	371,350	365,575	5,775		21,632	21,354	278	
Hamilton	692,120	441,635	235,415	15,070	49,380	30,980	17,412	988
Winnipeg	693,665	542,360	142,500	8,805	45,207	35,326	9,331	550

^{...} not applicable

Notes: Excludes pregnancy-related conditions. Foreign-born include those with landed immigrant status (ever) and non-permanent residents.

Source: 2006 Census of Population; 2005/2006 Hospital Morbidity Database.

ELECTRONIC PUBLICATIONS AVAILABLE AT

www.statcan.gc.ca

Comparison of waist circumference using the World Health Organization and National Institutes of Health protocols

by Jennifer Patry-Parisien, Margot Shields and Shirley Bryan

Abstract

Background

This study compares waist circumference (WC) measured using the World Health Organization (WHO) and National Institutes of Health (NIH) protocols to determine if the results differ significantly, and whether equations can be developed to allow comparison between WC taken at the two different measurement sites.

Data and Methods

Valid WC measurements using the WHO and NIH protocols were obtained for 6,306 respondents aged 3 to 79 from Cycle 2 of the Canadian Health Measures Survey. Linear regression was used to identify factors associated with the difference between the NIH and WHO values. Separate prediction equations by sex were generated using WC_NIH as the outcome and WC_WHO and age as independent variables. Sensitivity and specificity were calculated to examine whether health risk based on the WC_WHO and on WC_NIH predicted measurements agreed with estimates based on WC_NIH actual measured values.

Results

For adults and children, WC_NIH significantly exceeded WC_WHO (1.0 cm for boys, 2.1 cm for girls, 0.8 cm for men and 2.2 cm for women). Predicted NIH values were statistically similar to measured values. Sensitivity (86% to 98%) and specificity (70% to 100%) values for health risk category based on the NIH predicted values were very high, meaning that respondents would be appropriately classified when compared with actual measured values.

Interpretation

The prediction equations proposed in this study can be applied to historical datasets to compare estimates based on WC data measured using the WHO and NIH protocols.

Keywords

Body composition, central obesity, cross-over study, direct measure, sensitivity, specificity

Authors

Jennifer Patry-Parisien (613-951-6010; jennifer. patry-parisien@statcan.gc.ca) and Shirley.Bryan (613-951-4968; shirley.bryan@statcan.gc.ca) are with the Health Statistics Division, and Margot Shields was formerly with the Health Analysis Division at Statistics Canada, Ottawa, Ontario, K1A 0TG.

Taist circumference (WC) is an important independent measure in the assessment of obesity-related health risk.¹ The 2003 Canadian Guidelines for Weight Classification in Adults recommended that WC be measured on all persons with a body mass index (BMI) between 18.5 and 34.9 kg/m², using the World Health Organization (WHO) measurement protocol. In 2006, based on recommendations from an expert panel, the Canadian clinical practice guidelines on the management and prevention of obesity in adults and children suggested that practitioners use the National Institutes of Health (NIH) method to measure WC.² Two years later, the Canadian Society for Exercise Physiology (CSEP) adopted the NIH method as part of the measurement protocols in the Canadian Physical Activity and Fitness Lifestyle Approach.^{3,4}

In March 2007, Statistics Canada launched the Canadian Health Measures Survey (CHMS), which collects directly measured health data on a nationally representative sample of Canadians. Cycle 1 (2007 to 2009) included WC measurements using the WHO protocol.⁵ Cycle 2 (2009 to 2011) used both the WHO and NIH protocols in order to conduct a cross-over study that would compare the measurement techniques

based on a large, nationally representative sample.

This paper presents the results of that cross-over study. It compares the measurements of WC using the WHO and NIH measurement protocols, assesses the effect of measurement site on health risk classification, and evaluates the feasibility of predicting WC based on the NIH protocol from WC based on the WHO protocol.

Comparison of waist circumference using the World Health Organization and National Institutes of Health protocols • Methodological insights

Methods

Data source

The data are from the second cycle of the CHMS. The CHMS is an ongoing survey designed to provide comprehensive, direct health measures at the national level.⁶ Cycle 2 covers the population aged 3 to 79 living in private households. Residents of Indian Reserves, institutions and some remote regions, and full-time members of the Canadian Forces are excluded. More than 96% of the Canadian population is represented. Ethics approval for the CHMS was obtained from Health Canada's Research Ethics Board.

Data for Cycle 2 were collected at 18 locations across Canada from August 2009 through December 2011. In addition to a detailed questionnaire administered in the respondent's home, the survey involved physical measures (including WC, height and weight) several days later at a mobile examination centre. Participation in the survey was voluntary, and written informed consent was obtained from respondents for participation in the physical measures Additional information component. about the content and sample design can be found in the CHMS Cycle 2 Data User Guide.7

Of the households selected for the survey, 75.9% agreed to participate, and 90.5% of selected household members completed the household questionnaire. A total of 6,395 respondents (81.7% of those who completed the household questionnaire) completed the mobile examination centre component. final response rate, after adjusting for the sampling strategy, was 55.5%. This study pertains to 6,306 respondents for whom WC was measured using both the WHO and NIH protocols. Respondents who had a missing value for either or both protocols were excluded (n=89) from the analysis; this included pregnant women, whose WC was not measured.

Waist circumference measurement and classification

CHMS health measures specialists were trained to measure WC using both the

NIH and WHO protocols. WC was measured to the nearest 0.1 cm, directly on the landmarked skin with a flexible, inelastic measuring tape with a tension meter attached.⁷ For the NIH protocol, the measure is taken at the highest point of the iliac crest.8 For the WHO protocol, the measure is taken at the mid-point between the highest point of the iliac crest and the last floating rib⁵ (Figure 1). The two measurements were taken consecutively near the beginning of the visit to the mobile examination centre. The health measures specialists landmarked the location of the tape, marked it with a washable marker, and took the measure at the end of a normal expiration, on the right side of the back, using the reflection of the left side of the body in a mirror to ensure that the tape was horizontal.

Based on their WC measurements, adults aged 20 or older were classified into three health risk categories according to cut-offs recommended by the WHO,⁵ Health Canada,¹ and Obesity Canada.² Those cut-offs were also applied to the NIH protocol. The three categories are: low risk (men, WC 93.9 cm or less; women, WC 79.9 cm or less); increased risk (men, WC 94.0 to 101.9 cm; women, WC 80.0 to 87.9 cm); and high risk (men, WC 102.0 cm or more; women, WC 88.0 cm or more). Adolescents aged 12 to 19 were classified into low-, increased,

and high-risk WC categories according to the age- and sex-specific cut-offs proposed by Jolliffe et al.⁹ These cut-offs were developed using growth curve modeling, and they correspond to the cut-offs at entry into adulthood at age 20.⁹ Comparable WC cut-offs are not available for children younger than 12.

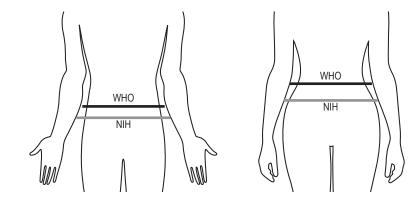
Body mass index measurement and classification

BMI is calculated as weight (kg) divided by height squared (m²). Height was measured to the nearest 0.1 cm using a ProScale M150 digital stadiometer (Accurate Technology Inc., Fletcher, USA), and weight, to the nearest 0.1 kg with a Mettler Toledo VLC with Panther Plus terminal scale (Mettler Toledo Canada, Mississauga, Canada).

Adults aged 18 or older were classified into six BMI categories: underweight (less than 18.5 kg/m²), normal weight (18.5 to 24.9 kg/m²), overweight (25.0 to 29.9 kg/m²), obese class I (30 to 34.9 kg/m²), obese class II (35 to 39.9 kg/m²), and obese class III (40 kg/m² or more). $^{1.2.5}$

Children and teenagers aged 3 to 17 were classified into BMI categories based on growth curves using age- and sex-specific cut-offs of the WHO. The WHO recommends that 5- to 17-year-olds whose BMI is more than two standard deviations (SD) above the mean

Figure 1 Waist circumference measurement sites for men and women based on World Health Organization (WHO) and National Institutes of Health (NIH) protocols



Note: Following the WHO protocol, the measure is taken midway between the highest point of the iliac crest and the bottom of the ribcage. Following the NIH protocol, the measure is taken at the highest point of the iliac crest.

be considered obese, and those whose BMI is between one and two SD above the mean, overweight.¹⁰ Although the WHO recommends a different set of cutoffs for children younger than 5,¹¹ for this analysis the one- and two-SD cutoffs were used to define overweight and obesity for children aged 3 and 4.

Analytical techniques

The feasibility of predicting WC based on the NIH protocol (WC NIH) from WC based on the WHO protocol (WC WHO) was assessed. Scatter plots and linear regression were used to identify factors associated with the difference between the two measurements. difference was significantly associated with three variables: WC WHO, BMI, and age (continuous age for children and adult males, and age group for adult females). Because of the high correlation between WC WHO and BMI, it was not possible to include both in the regression models. WC WHO was retained because R-squared values were higher for the models using WC WHO as a predictor of the difference (data not shown).

The sample was then randomly divided into split-sample A and splitsample B, each containing about 50% of respondents. Split-sample A was used to generate prediction equations using WC NIH as the outcome and WC WHO and age as independent variables. Separate prediction equations were generated for men and women (ages 20 to 79) and for boys and girls (ages 3 to 19). Outliers (respondents for whom the difference between WC NIH and WC WHO was more than three SD from the mean) were excluded when generating the prediction equations; this was the case for 56 (fewer than 2%) of the 3,202 records in splitsample A.

The prediction equations generated from split-sample A were applied to split-sample B. The WC_NIH value calculated from WC_WHO measurement is referred to as "WC_NIH_predicted." To evaluate the success of the prediction equations, the WC_NIH_predicted measurements from split-sample B were

compared with the actual measurements. Outliers in split-sample B were included in this evaluation. The estimates were compared by sex for six age groups: 3 to 5, 6 to 11, 12 to 19, 20 to 39, 40 to 59, and 60 to 79. Comparisons were also made by BMI categories, by sex, for adults and for children.

Using split-sample B, the health risk cut-offs were applied to WC WHO, WC NIH predicted WC NIH and The percentages of measurements. respondents whose WC put them in a high health risk category were compared the three measurements. among Sensitivity and specificity calculated to examine the extent to which health risk estimates based on WC WHO and WC NIH predicted measurements agreed with health risk estimates based on WC NIH measurements (WC NIH was the "gold-standard").

Sensitivity refers to the percentage of true positives—in this case, the percentage of respondents classified at high health risk based on their WC NIH measurements who were also classified at high health risk based on their WC_WHO and WC_NIH_predicted measurements. Specificity refers to the percentage of true negatives—in this case, the percentage of respondents who were not classified at high health risk based on the WC_NIH measurements who were also not classified at high health risk based on their WC WHO and on WC NIH predicted measurements. Estimates were also calculated and compared for the combined increased/ high health risk group.

Prediction equations for detailed age-sex groups were evaluated (data not shown), but the results were similar to those based on the four prediction equations presented in the current study. Crude adjustments were also evaluated, whereby the differences between WHO and NIH measurements were calculated based on split-sample A and then applied to split-sample B. Crude adjustments were made by detailed age-sex groups and by BMI categories (by sex, for adults and for children). The results based on these crude adjustments (data not shown) were not as favourable as those based on the four regression models.

Estimates of percentages, means and regression coefficients were calculated using weighted data. Differences between estimates were tested for statistical significance, which was set at 0.05. Standard errors were estimated with the bootstrap technique; the number of degrees of freedom was specified as 13 to account for the sample design of the data. Weighted estimates were produced to adjust for unequal probabilities of selection and to take advantage of the adjustments made to reduce non-response bias in the CHMS.

Results

Measurements and equations

For men and women aged 20 to 79, the difference between WC_NIH and WC_WHO was negatively associated with WC_WHO; that is, the larger the WC_WHO measurement, the smaller

Table 1
Equations to predict National Institutes of Health waist circumference measures (WC_NIH_predicted) based on waist circumference measured using World Health Organization protocols (WC_WHO), by age group and sex

Age group (years)/ Sex	Equation	R²
3 to 19 Boys Girls	WC_NIH_predicted = -0.89911 + 1.01829*(WC_WHO) + 0.05164*(age) WC_NIH_predicted = -0.70299 + 1.01891*(WC_WHO) + 0.12297*(age)	0.99 0.99
20 to 79 Men Women	WC_NIH_predicted = 3.83072 + 0.98613*(WC_WHO) - 0.03609*(age) WC_NIH_predicted = 3.53771 + 0.98479*(WC_WHO) + 0.21949*(x) (where x is set to 1 if age is 20 to 39; otherwise x=0)	0.99 0.98

Comparison of waist circumference using the World Health Organization and National Institutes of Health protocols • Methodological insights

Table 2 Mean waist circumference based on World Health Organization (WHO) and National Institutes of Health (NIH) protocols, by sex and age group, household population aged 3 to 79, Canada, 2009 to 2011

			Measured			NIH predicted					
Sex/Age group (years)		Difference (NI predicte					95% confidence interval				
	Sample size	NIH	minus WHO WHO)		NIH predicted	minus NIH measured)	from	to			
	Number			Centi	netres						
Boys 3 to 5 6 to 11 12 to 19	702 151 277 274	68.8 52.4 61.0 81.3	67.8 52.0 60.2 79.8	1.0 * 0.4* 0.8* 1.5*	68.7 52.2 60.8 81.1	-0.1 -0.1 -0.2 -0.1	-0.3 -0.4 -0.5 -0.4	0.0 0.1 0.1 0.1			
Girls 3 to 5 6 to 11 12 to 19	670 143 274 253	66.7 51.5 59.5 76.1	64.7 50.8 58.2 73.1	2.1 * 0.7* 1.3* 3.1*	66.6 51.5 59.7 75.6	-0.2 0.1 0.2 -0.5*	-0.4 -0.2 -0.1 -0.9	0.1 0.4 0.5 -0.1			
Men 20 to 39 40 to 59 60 to 79	824 270 303 251	95.3 88.8 97.1 103.2	94.5 87.5 96.4 103.3	0.8 * 1.3* 0.8* -0.1	95.4 89.0 97.1 103.3	0.1 0.2 -0.1 0.1	-0.2 -0.2 -0.6 -0.3	0.4 0.6 0.5 0.4			
Women 20 to 39 40 to 59 60 to 79	908 355 284 269	89.2 85.6 89.4 95.2	87.0 82.8 87.4 93.7	2.2* 2.8* 2.1* 1.5*	89.3 85.3 89.6 95.8	0.1 -0.3 0.2 0.6	-0.3 -0.9 -0.2 -0.02	0.4 0.2 0.5 1.2			

^{*} significantly different from zero (p < 0.05)

Note: Estimates are generated from sub-sample B.

Source: 2009 to 2011 Canadian Health Measures Survey.

Table 3
Difference between waist circumference measured according to National Institutes of Health (NIH) and World Health Organization (WHO) protocols, by sex and age group, household population aged 3 to 79, Canada, 2009 to 2011

	Difference											
Sex/Age	NIH me	asured v	ersus W	HO mea	NIH p	NIH predicted versus NIH measured						
group (years)	≤1 cm	≤2 cm	≤3 cm	≤4 cm	> 4 cm	≤1 cm	≤2 cm	≤3 cm	≤4 cm	>4 cm		
			%					%				
Boys	53	81	92	98	2	69	93	97	99	1		
3 to 5	71	96	100	100	0	74	96	100	100	C		
6 to 11	65	88	96	97	3	72	94	96	98	2		
12 to 19	36	69	87	98	2	64	91	97	98	2		
Girls	34	58	75	85	15	53	80	92	95			
3 to 5	67	87	97	99	1	66	95	99	100	C		
6 to 11	47	75	90	96	4	58	85	98	100	C		
12 to 19	16	38	57	73	27	45	73	85	91	g		
Men	45	75	91	96	4	50	86	94	97	3		
20 to 39	41	73	91	98	2	53	87	97	99	1		
40 to 59	47	76	93	96	4	47	88	95	97	3		
60 to 79	49	78	90	95	5	53	79	90	95	5		
Women	25	48	67	81	19	38	71	89	96	4		
20 to 39	15	36	55	69	31	34	63	87	97	3		
40 to 59	27	52	75	90	10	46	82	93	97	3		
60 to 79	40	59	73	83	17	31	62	82	90	10		

Note: Estimates are generated from sub-sample B. **Source:** 2009 to 2011 Canadian Health Measures Survey.

the difference (Appendix Table A). For men, the association with age was negative. For women, the association with age was not linear, but when age groups were included in the regression model, a positive association emerged for women aged 20 to 39.

For boys and girls aged 3 to 19, a positive relationship was observed between the difference and both WC_WHO and age. That is, higher values of WC_WHO and age were associated with larger differences.

The prediction equations derived from split-sample A to calculate WC_NIH_ predicted based on WC_WHO and age are presented in Table 1.

Measured and predicted waist circumferences

Regardless of age and sex, mean values of WC_NIH significantly exceeded those of WC_WHO: 1.0 cm for boys, 2.1 cm for girls; 0.8 cm for men, and 2.2 cm for women (Table 2). The differences were greatest for girls aged 12 to 19 (3.1 cm) and women aged 20 to 39 (2.8 cm).

Overall, the measured and predicted NIH values were statistically similar (mean differences range from -0.2 cm to 0.1 cm) (Table 2). For the detailed agesex groups, the only significant difference was for girls aged 12 to 19 (-0.5 cm).

At ages 3 to 19, WC_NIH_predicted was within 1 cm of WC_NIH measured for 69% of boys and 53% of girls (Table 3). A difference of more than 2 cm was observed for 7% of boys and 20% of girls overall, and 27% of girls aged 12 to 19.

Half of men had a predicted NIH value within 1 cm of the measured value; a difference of more than 2 cm was observed in 14% of cases. For women, the predicted value was within 1 cm of the measured value in 38% of cases; in 29% of cases, the difference was more than 2 cm.

For all age-sex groups, the predicted value was within 4 cm of the measured value in at least 90% of cases.

Body mass index

Among children and adolescents, differences between measured WC_

WHO and WC_NIH were greater for those classified as obese: a mean difference of 1.6 cm for boys and 2.5 cm for girls (Table 4). Among adults, differences were greater for those in the normal weight range: 1.3 cm for men and 2.7 cm for women.

The only significant difference between the measured and predicted NIH values was for obese boys (-0.4 cm). Although the measured and predicted NIH means were fairly close for women in obese categories II and III, a difference of more than 2 cm was observed in 50% of cases (Table 5).

Health risk

For men and boys, the percentages whose waist circumference put them in a high health risk category were similar whether based on WHO, NIH or NIH-predicted measures (Table 6). For men,

the prevalence of increased/high health risk was slightly elevated when WC was based on NIH rather than on WHO, and the prevalence of increased/high health risk based on the predicted NIH values was similar to the estimate based on the measured NIH values.

For women and girls, the percentages whose waist circumference put them in a high (or increased/high) health risk category were significantly greater based on NIH rather than WHO measures, while estimates based on the predicted NIH values were similar to those based on NIH measured values.

Sensitivity and specificity

Sensitivity and specificity were very high when based on NIH predicted values, meaning that in almost all cases, respondents would be classified in the appropriate health risk category—that is,

Table 4
Mean waist circumference based on World Health Organization (WHO) and National Institutes of Health (NIH) protocols, by age group, sex and body mass index (BMI) category, household population aged 3 to 79, Canada, 2009 to 2011

			Measured			NIH predicted				
Age group/			С	Difference (NIH		Difference (NIH predicted	95% confidence interval			
Sex/ BMI category	Sample size	NIH	minus NIH WHO WHO)		NIH predicted	minus NIH measured)	from	to		
	Number			С	entimetres					
Ages 3 to 19										
Boys Normal weight Overweight Obese	438 139 111	63.4 69.7 87.0	62.5 68.7 85.4	0.9* 1.0* 1.6*	63.3 69.6 86.7	-0.1 -0.1 -0.4*	-0.3 -0.5 -0.7	0.1 0.2 -0.01		
Girls Normal weight Overweight Obese	467 137 58	62.5 71.9 90.2	60.5 69.6 87.7	2.0* 2.3* 2.5*	62.3 71.5 90.2	-0.1 -0.4 0.1	-0.4 -1.0 -1.0	0.1 0.2 1.1		
Ages 20 to 79										
Men Normal weight Overweight Obese class I Obese class II/III	219 360 174 67	83.0 95.9 107.6 125.1	81.7 95.1 107.3 125.1	1.3* 0.8* 0.3 0.0	83.0 95.9 107.8 125.3	0.0 0.1 0.2 0.3	-0.5 -0.4 -0.3 -0.4	0.5 0.5 0.7 0.9		
Women Normal weight Overweight Obese class I Obese class II/III	365 270 156 97	78.2 91.6 101.9 117.4	75.5 89.6 100.2 115.9	2.7* 2.0* 1.7* 1.5*	78.0 91.8 102.2 117.8	-0.2 0.2 0.3 0.4	-0.6 -0.3 -0.2 -0.7	0.2 0.7 0.9 1.4		

^{*} significantly different from zero (p < 0.05)

Notes: Estimates are generated from sub-sample B. Estimates for underweight are not included because of small sample sizes Source: 2009 to 2011 Canadian Health Measures Survey.

the same category in which they would be placed based on measured values (Table 7). Sensitivity and specificity were 90% or more, with two exceptions: specificity was somewhat low (70%) for increased/high risk for women aged 60 to 79, and sensitivity for girls aged 12 to 19 was 86% for high risk and 89% for increased/high risk. However, these sensitivity values were an improvement over those based on WHO.

Discussion

In the present study, WC for Canadian adults and children was significantly greater when measured using the NIH protocol than the WHO protocol. The difference was greatest among girls and young women. These findings add to the limited information about WC measurements taken at different sites. 12,13 In a study based on 111 healthy volunteers aged 7 to 83, Wang et al.12 compared measurements at four sitesimmediately below the lowest rib, at the narrowest waist, midway between the lowest rib and iliac crest (WHO), and immediately above the iliac crest (NIH). In that study, males' mean WC at the narrowest waist was significantly lower than at the other three sites. For females, mean WC at each site differed significantly from means at the others, and WC measurements using the NIH protocol significantly exceeded those using the WHO protocol (1.82 cm).¹²

Mason et al.¹³ conducted a more recent study (2009) of 542 healthy volunteers aged 20 to 67 to assess whether WC differed across four commonly used measurement sites. They noted no significant differences between sites for men. For women, the mean for each site differed significantly from the means for the others, except for the means at the sites used for the NIH and WHO protocols, which did not differ.¹³

In the present study, the differences that emerged between the NIH and WHO protocols may be related to the sample size or sample characteristics (the Mason sample consisted of healthy adult volunteers, while the CHMS

Comparison of waist circumference using the World Health Organization and National Institutes of Health protocols • Methodological insights

Table 5
Difference between waist circumference measured according to National Institutes of Health (NIH) and World Health Organization (WHO) protocols, by age group, sex and body mass index (BMI) category, household population aged 3 to 79, Canada, 2009 to 2011

	Difference										
Age group/ Sex/	NIH me	easured	versus V	VHO me	NIH p	NIH predicted versus NIH measured					
BMI category	≤1 cm	≤2 cm	≤3 cm	≤4 cm	> 4 cm	≤1 cm	≤2 cm	≤3 cm	≤4 cm	>4 cm	
			%					%			
Ages 3 to 19											
Boys											
Normal weight	59	85	96	98	2	70	94	98	99	1	
Overweight	50	85	91	98	2	75	94	98	100	0	
Obese	33	61	80	95	5	54	87	92	96	4	
Girls											
Normal weight	35	61	76	87	13	56	82	94	96	4	
Overweight	30	51	74	81	19	51	75	85	94	6	
Obese	38	51	62	77	23	27	79	93	93	7	
Ages 20 to 79											
Men											
Normal weight	39	71	92	98	2	49	88	97	99	1	
Overweight	47	75	91	96	4	51	86	94	95	5	
Obese class I	53	83	93	95	5	56	84	94	99	1	
Obese class II/III	42	81	86	95	5	36	80	87	96	4	
Women											
Normal weight	18	40	61	76	24	37	72	90	97	3	
Overweight	31	55	71	89	11	41	77	92	97	3	
Obese class I	40	55	79	88	12	42	68	87	91	9	
Obese class II/III	25	47	64	72	28	28	50	76	92	8	

Notes: Estimates are generated from sub-sample B. Estimates for underweight are not included because of small sample sizes. **Source:** 2009 to 2011 Canadian Health Measures Survey.

Table 6
Percentage with high and increased/high health risk based on waist circumference according to World Health Organization (WHO) and National Institutes of Health (NIH) protocols, by sex and age group, household population aged 12 to 79, Canada, 2009 to 2011

	H	igh health ris	k	Increased/High health risk			
Sex/Age group (years)	WHO measured	NIH measured	NIH predicted	WHO measured	NIH measured	NIH predicted	
		%		%			
Boys aged 12 to 19	13 ^E	14 ^E	14 ^E	20 ^E	20 ^E	20 ^E	
Girls aged 12 to 19	15* ^E	22 ^E	20 ^E	31*	41	37	
Men	26	29	28	47*	50	48	
20 to 39	12 ^E	16 ^E	15 ^E	24 ^E	27 ^E	26 ^E	
40 to 59	26 ^E	30	27 ^E	54	56	55	
60 to 79	51	50	51	75	76	76	
Women 20 to 39	41* 28*	46 33	46 30	60 * 45*	70 54	67 52	
40 to 59	42*	47	50	65*	77	70	
60 to 79	62*	69	67	79*	85	88	

significantly different from NIH measured (p<0.05)

Note: Estimates are generated from sub-sample B. Source: 2009 to 2011 Canadian Health Measures Survey.

sample is representative of the Canadian population aged 3 to 79).

In a comprehensive review, Ross et al. 14 suggested that the protocol used to measure WC does not substantially influence the association between WC and all-cause mortality, cardiovascular disease mortality, and cardiovascular disease and diabetes morbidity. the present study, the classification of men and boys into the high health risk category was similar regardless of whether WC was measured using the WHO or NIH protocol. However, the prevalence of high health risk for women and girls, and the prevalence of combined increased/high health risk among men were significantly greater when measures were based on the NIH protocols than on the WHO protocols. Similarly, Mason et al.13 reported that the prevalence of abdominal obesity (more than 88 cm for women; more than 102 cm for men) depended on the WC measurement protocol used. When comparing the WHO and NIH protocols, they noted no difference in the prevalence of abdominal obesity for men (32.7% versus 31.8%), but for women, the prevalence was higher based on the NIH protocol (47.0%) than on the WHO protocol (41.1%).¹³ Willis et al.¹⁵ used different WC measurement protocols, but they also noted that classification of health risk depends on which protocol is used. When WC was measured at the umbilicus rather than the minimal waist, 54% more men and 68% more women met the National Cholesterol Education Program criteria for abdominal obesity.¹⁵

To assess the accuracy of the prediction equations proposed in this study, the difference between the measured NIH value and the predicted NIH value was calculated on a portion of the sample. For the majority of cases, the equations yield statistically similar WC values. And although the results show a difference greater than 2 cm for 50% of women in obese class II and III, this would not be a meaningful difference for health-risk assessment at these levels of BMI.

E use with caution

Table 7
Sensitivity and specificity for high and increased/high health risk according to waist circumference, based on World Health Organization (WHO) and predicted National Institutes of Health (NIH) protocols, by sex and age group, household population aged 12 to 79, Canada, 2009 to 2011

		NIH pr	edicted			WHO measured				
	Sensitivity (% true positives)		Specificity (% true negatives)			nsitivity positives)	Specificity (% true negatives)			
Sex/Age group (years)	High risk	Increased /High risk	High risk	Increased /High risk	High risk	Increased /High risk	High risk	Increased /High risk		
Boys aged 12 to 19	98	98	100	100	94	97	100	100		
Girls aged 12 to 19	86	89	99	99	69	77	100	100		
Men	91	95	98	98	86	94	99	98		
20 to 39	92	95	100	99	73	88	100	99		
40 to 59	86	95	98	98	84	94	99	98		
60 to 79	96	96	94	89	96	96	94	90		
Women	95	93	95	93	88	85	100	99		
20 to 39	88	93	99	96	84	83	100	100		
40 to 59	98	90	91	95	89	84	100	99		
60 to 79	95	98	95	70	90	92	100	96		

Note: Estimates are generated from sub-sample B using NIH measured as the standard.

Source: 2009 to 2011 Canadian Health Measures Survey.

and specificity were Sensitivity calculated to examine the extent to which health risk estimates based on WC WHO and WC NIH predicted measurements agreed with those based on WC_NIH measurements. The sensitivity and specificity values were generally very high for WC_NIH_ predicted values, which means that respondents would be correctly classified into the appropriate health risk category based on the predicted NIH values. In a few cases, the absolute differences between the predicted and measured NIH values were large, but from a clinical perspective, the predicted values result in the correct health risk assessment. These findings suggest that the equations generated from the CHMS dataset can be applied to historical WHO data so that WHO and NIH waist circumference data can be compared. These equations can be applied to a broader age range (including 3- to 19-year-olds) than those proposed by Mason et al.¹³

A strength of this analysis is the large sample from the general population, ranging in age from 3 to 79, which made it possible to examine differences in WC by age, sex and BMI.

The differences between the measurement protocols that emerged are not the result of inter- or intra-tester variability. CHMS staff underwent biannual training with a measurement expert, were regularly observed as they measured the respondents, and were monitored by assessments of the technical error of measurement (TEM)16 once a vear. The TEM compares measurements made by CHMS staff with measurements of a gold standard. To ensure high data quality, a low TEM threshold (1.5%) was set, based on the literature.16 On average, the relative TEM result was 1.42% for measurements taken using the WHO protocol, and ranged from 1.79% to 3.06% for measurements taken using the NIH protocol, which are very close to the target.

Another strength of the study is that measurements using the two protocols were taken one after another on the same day. Consequently, factors such as food and beverage consumption, time of day and menstrual cycle did not affect differences between them.

What is already known on this subject?

- Abdominal obesity is associated with increased all-cause and cardiovascular mortality, and cardiovascular disease and diabetes morbidity.
- Waist circumference provides information beyond body mass index in the assessment of obesity-related health risk in clinical settings.
- Waist circumference measurements differ, depending on the measurement protocol used.

What does this study add?

- This study examines the difference between waist circumference measured using the World Health Organization (WHO) and National Institutes of Health (NIH) protocols on a large, representative sample of Canadians aged 3 to 79.
- Waist circumference measures based on the WHO and NIH protocols differ significantly.
- The prediction equations in this study can be used to compare estimates based on the WHO and NIH protocols on a wide range of age groups.

Conclusion

The CSEP and the Canadian clinical practice guidelines have adopted the NIH protocol as the standard method for WC measurement in Canada. The prediction equations proposed in this study can be applied to historical Canadian datasets in which the WHO protocol was used. This will allow researchers to assess WC trends over time.

References

- Health Canada. Canadian Guidelines for Body Weight Classification in Adults (Catalogue H49-179) Ottawa: Health Canada, 2003.
- Lau DC, Douketis JD, Morrison KM, et al. 2006 Canadian clinical practice guidelines on the management and prevention of obesity in adults and children [summary]. Canadian Medical Association Journal 2007; 176: S1-13.
- Canadian Society for Exercise Physiology. The Canadian Physical Activity, Fitness and Lifestyle Approach 3rd edition. Ottawa: Canadian Society for Exercise Physiology, 2003.
- Canadian Society for Exercise Physiology. The Revision of the Measurement of Waist Circumference in the CPAFLA, November 2008. Available at: www.csep.ca [Home -> Publications -> Knowledge Translation].
- World Health Organization. Obesity: Preventing and Managing the Global Epidemic. Report of a WHO consultation on obesity. Geneva: World Health Organization, 2000.
- Giroux, S. Canadian Health Measures Survey: Sampling strategy. *Health Reports*. 2007; 18(supplement): 7-20.

- Statistics Canada. Canadian Health Measures Survey (CHMS) Data User Guide: Cycle 2, September 2012. Available at: www.statcan. gc.ca.
- National Institutes of Health. The Practical Guide to the Identification, Evaluation and Treatment of Overweight and Obesity in Adults. Bethesda, Maryland: National Institutes of Health, 2000.
- Jolliffe CJ, Janssen I. Development of age-specific adolescent metabolic syndrome criteria that are linked to the Adult Treatment Panel III and International Diabetes Federation criteria. *Journal of the American College of Cardiology* 2007; 49: 891-8.
- de Onis M, Onyango AW, Borghi E, et al. Development of a WHO growth reference for school-aged children and adolescents. Bulletin of the World Health Organization 2007; 85(9): 660-7.
- World Health Organization. Training Course on Child Growth Assessment. Geneva: World Health Organization, 2008. Available at: http://www.who.int/childgrowth/training/ module_a_introduction.pdf.

- Wang J, Thornton JC, Bari S, et al. Comparisons of waist circumferences measured at 4 sites. *American Journal of Clinical Nutrition* 2003; 77: 379-84
- Mason C, Katzmaryzyk PT. Variability in waist circumference measurements according to anatomic measurement site. *Obesity* 2009; 17(9): 1789-95.
- Ross R, Berentzen T, Bradshaw AJ, et al. Does the relationship between waist circumference, morbidity and mortality depend on measurement protocol for waist circumference? *Obesity Reviews* 2007; 9: 312-25.
- Willis LH, Slentz CA, Houmard JA, et al. Minimal versus umbilical waist circumference measures as indicators of cardiovascular disease risk. *Obesity* 2007; 15: 753-9.
- Perini TA, de Oliveira GL, dos Santos Ornellas J, Palha de Oliveira F. Technical error of measurement in anthropometry. Revista Brasileira de Medicina do Esporte 2005; 11(1): 81-90.

Appendix

Table A
Regression coefficients for difference between waist circumference based on National Institutes of Health and World Health Organization (WHO) protocols, by age group and sex, household population aged 3 to 79, Canada, 2009 to 2011

	Ages	3 to 19	Ages 2	0 to 79
	Boys	Girls	Men	Women
	- 1	Regression coeffic	ient (B)	
Intercept	-1.20268*	-1.02162*	4.01038*	3.70067*
WHO waist circumference (cm)	0.02574*	0.02349*	-0.01952*	-0.01852*
Age continuous	0.03825*	0.12454*	-0.02881*	
Age group 20 to 39 40 to 79 [†]	 	 	 	0.48826*
Adjusted R squared	0.18	0.25	0.13	0.04

[†] reference group

Source: 2009 to 2011 Canadian Health Measures Survey.

significantly different from zero (p < 0.05)

^{...} not applicable