## Census of Population Technical Report

## Sampling and Weighting Technical Report

Census of Population, 2016

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. not available for any reference period
.. not available for a specific reference period
... not applicable
0 true zero or a value rounded to zero
$0^{\text {s }}$ value rounded to 0 (zero) where there is a meaningful distinction between true zero and the value that was rounded
${ }^{p}$ preliminary
r revised
x suppressed to meet the confidentiality requirements of the Statistics Act
E use with caution
F too unreliable to be published

* significantly different from reference category ( $p<0.05$ )

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## Introduction

The 2016 Census Program enumerates Canadian households using two main types of questionnaires: the shortform questionnaire and the long-form questionnaire. In 2016, a sample of $25 \%$ of Canadian households received a long-form questionnaire, which included the questions from the short-form questionnaire. The other households received the short-form questionnaire. In 2011, the government decided to eliminate the mandatory character of the long-form census questionnaire. However, it was reintroduced for the 2016 Census.

In addition to the short-form questions, the long-form questionnaire includes a series of questions to paint a full portrait of the Canadian population and households, according to their demographic, social and economic characteristics.

The estimates produced from responses to the questions on both questionnaires are obtained from the entire population via a census. All households that respond to both types of questionnaires contribute to a specific number, such as the population figure for a specific age group.

The estimates produced from responses to at least one question found only on the long-form questionnaire are obtained from the sample. In those cases, only respondent households in the long-form sample contribute to the estimate. The unemployment rate and the highest level of educational attainment are examples of this type of estimate.

The long-form sample is evenly distributed geographically to ensure a high degree of reliability of the estimates for all areas of the country and to grant the same degree of importance to all geographic units of a given size.

This technical report presents the methodology used to produce the estimates based on the 2016 Census of Population long-form sample. Chapter 1 details the collection methods used for the census and for the longform sample. Chapter 2 describes how the sampling was applied for the long-form questionnaire. Chapter 3 explains the data processing procedures. Chapter 4 gives an overview of the procedures used to assign weights to the respondent units in the long-form sample to obtain estimates for the population. Chapter 5 covers different evaluations of the weighting procedures, while Chapter 6 provides an overview of the variance estimation methodology used for the 2016 Census long-form sample. A conclusion is presented in Chapter 7.

## 1. Census data collection

The purpose of data collection for the 2016 Census was to ensure that each of the 15.4 million dwellings in Canada was enumerated and that, for each occupied dwelling, the corresponding household completed a census questionnaire. The census enumerated the entire population of Canada, which consists of Canadian citizens (by birth and by naturalization), landed immigrants and non-permanent residents, as well as the family members living with them. Non-permanent residents are persons living in Canada who have a work or study permit or are claiming refugee status, as well as the family members living with them.

The census also enumerated Canadian citizens and landed immigrants who were temporarily out of the country on Census Day. This included federal, provincial and territorial government employees posted outside Canada; the staff of Canadian embassies abroad; members of the Canadian Forces stationed abroad; and all Canadians aboard merchant vessels.

### 1.1 Collection methodology

Since the 2011 Census, a new collection methodology has been used to collect census data. Referred to as the wave methodology, it involves contacting non-respondent households at key times to remind them to take part in the census and to encourage them to complete the questionnaire. In each wave, households are given the information they need to respond. Since every Canadian household is required by law to answer the census questions, this method is designed to encourage people to respond promptly online, while mitigating the risk of a decline in overall response and reducing the need for costly field follow-up.

This methodology varies with the collection method used to distribute the census materials for a given region. ${ }^{1}$ These collection methods are described in the next section. In 2016, Canadian households had the option of responding online, completing a paper questionnaire (mail-back) or contacting the Census Help Line.

### 1.1.1 Collection methods: Delivery of census questionnaires

The three collection methods used for the Canadian census are mail-out, list/leave and canvasser. To make census collection as efficient as possible, Canada's territory is divided into small geographic units known as collection units (CUs). For the 2016 Census, Canada was divided into approximately 46,000 CUs. There are five types of CUs, and one of three collection methods is assigned to each one. The same collection method is used for all dwellings in a CU.

Table 1.1.1.1
Collection method by type of collection unit

| Type of CU | Collection method |
| :--- | :--- |
| Mail-out | Mail-out |
| List/leave | List/leave |
| Canvasser | Canvasser |
| Collective dwellings | List/leave |
| Reserves | Canvasser |

$\mathrm{CU}=$ Collection unit.

[^0]The collection methods are described below.

### 1.1.1.1 Mail-out

In CUs where the mail-out collection method is used, the postal system is used to deliver census materials. This delivery method ensures effective, coordinated distribution, without needing to recruit and train a large contingent of enumerators. Mail-out CUs are typically in urban areas. While mail-out CUs now include about $82 \%$ of Canadian dwellings, they cover only a tiny fraction of the country's territory. One-quarter of the private dwellings in these CUs are selected to receive the long-form questionnaire; the other dwellings receive the short-form questionnaire.

### 1.1.1.2 List/leave

List/leave CUs are typically in rural areas. In those areas, enumerators prepare a list of dwellings and deliver the census materials. About $17 \%$ of the dwellings in Canada are in list/leave CUs, which cover a large portion of the country's territory. One-quarter of the private dwellings in these CUs are selected to receive the long-form questionnaire. This collection method is also used for collective CUs.

### 1.1.1.3 Canvasser

Canvasser CUs are either Indian reserves or in remote or difficult-to-access places. To limit the number of often costly and logistically complicated trips that enumerators have to make to these places for follow-up, the enumerators do more than prepare dwelling lists. They also complete a questionnaire with each household on the spot. Canvasser CUs cover just over half of Canada's territory, but only about $1 \%$ of dwellings. All private dwellings in these CUs receive the long-form questionnaire.

### 1.1.2 Census wave methodology

The wave methodology was designed to encourage online response and to offer an alternative for households that do not wish to complete their questionnaire online. This wave approach has many advantages for minimizing non-response and increasing the number of questionnaires completed online, which accelerates questionnaire registration, improves question flow and data capture, and ultimately enhances data quality.

The wave methodology is applied differently depending on the collection method. CUs were therefore consolidated by collection method, and a different wave methodology was developed for each of the three groups. Given the nature of canvasser CUs, a single-wave methodology was adopted for these units. The sections that follow and Figure 1.1.2.1 provide an overview of the wave methodology used for the 2016 Census.

Figure 1.1.2.1
Overview of the wave methodology in the 2016 Census


NR = non-response
NRFU = non-response follow-up
Source: Statistics Canada, Census of Population, 2016.

### 1.1.2.1 Mail-out collection units

In mail-out CUs, Wave 1 involved simply sending out a letter asking households to complete the questionnaire online using the secure access code (SAC) provided or to call an automated system on a toll-free line to have a paper questionnaire mailed to them. The Wave 1 letters were delivered by mail one week before Census Day (i.e., on May 2, 2016).

Wave 2 consisted of a reminder letter sent to all Wave 1 households who had not responded by a certain date, by region. The letter reminded the households that they were required by law to complete the census questionnaire. As the Wave 1 letter did, it also provided the SAC and the toll-free telephone number. It was delivered to households from May 11 to 13.

In Wave 3, from May 19 to 26, a paper questionnaire was sent to non-respondent households. The households could still respond online using the SAC printed on the front cover of the questionnaire. The questionnaire was accompanied by a letter indicating that if the questionnaire was not completed by May 31, 2016, an enumerator would contact the household by telephone or in person to complete the questionnaire. The letter also mentioned that if the household refused to answer the census questions, the case could be referred to the Public Prosecution Service of Canada, which would take appropriate action under the Statistics Act.

Wave 4, which began on June 1, 2016, consisted of field non-response follow-up (NRFU) and an automated reminder call at the beginning of the follow-up period. NRFU is described in section 1.1.3 of this document.

### 1.1.2.2 List/leave collection units

A different wave methodology was used for list/leave CUs. In Wave 1, enumerators delivered a paper questionnaire to all the dwellings in these CUs from May 2 to 9 . The questionnaire also provided a SAC, so that the household had the option of responding online.

In Wave 2, from May 11 to 13, all the dwellings in these CUs received a thank you and reminder card in the mail. These were delivered whether they had responded or not, because sending mail to specific civic addresses without the name of the occupant was generally impossible in these areas.

The last wave involved going directly to field non-response follow-up as of May 20.

### 1.1.3 Non-response follow-up

As mentioned in the previous section, the last wave in the wave methodology was NRFU. In that wave, enumerators telephoned and visited households that had not responded to the questionnaire. Each nonrespondent household for which a telephone number was available received an automated reminder call at the beginning of the NRFU period. This message reminded non-respondents of their legal obligation to respond to the census questionnaire.

NRFU was conducted in person by enumerators or over the telephone by call centre staff. Census employees had information at their disposal to manage their work in the Collection Management Portal. Being a computerized system accessible over the Internet, this portal also facilitated the gathering of information on collection progress and costs.

### 1.1.4 Dwelling occupancy verification

Before NRFU, field operations were also carried out for dwelling occupancy verification (DOV) of non-respondent dwellings. The purpose of DOV, which began shortly after Census Day, May 10, 2016, was to identify the greatest possible number of dwellings that were unoccupied on Census Day or cancelled (addresses that were not private or collective dwellings) before NRFU started. Identifying such dwellings close to Census Day should make occupancy classification more accurate and easier to perform. DOV also reduces the NRFU workload, since the unoccupied or cancelled dwellings it identifies do not require follow-up.

Nevertheless, errors in classifying a dwelling as occupied or unoccupied do occur during DOV and NRFU. Some dwellings classified as unoccupied are in fact occupied, and some non-respondent dwellings are unoccupied. As a result, another process, the Dwelling Classification Survey (DCS), is carried out after NRFU. It determines the occupancy status of a sample of dwellings for which no completed questionnaire was received (unoccupied dwellings, non-respondent dwellings or unresolved cases), and it estimates occupancy rates for all dwellings in the same situation across the country. The survey results are used to adjust the Census of Population counts during processing. More information on the DCS is provided in section 3.6.

### 1.2 Questionnaires and forms

Various questionnaires and forms are used to collect data for the Canadian census, and respondents have the choice of completing the census questionnaire online or using a paper questionnaire.

Visitation Record (VR) is used to list every occupied and unoccupied private or collective dwelling, agricultural operation and agricultural operator in the areas of the country where a list of dwellings is created before questionnaires are delivered (i.e., in list/leave collection units; see section 1.1.1.2). The VR serves as an address list for field operation and control purposes during data collection.

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Form 2A (http://www23.statcan.gc.ca/imdb/p3Instr.pl?Function=getInstrumentList\&ltem_Id=295241\&UL=1V\&) is the basic short-form questionnaire used to enumerate private dwellings. Each household is asked to list all household members who belong to the census population and to answer questions for them.

Forms 2A-L (http://www23.statcan.gc.ca/imdb/p3Instr.pl?Function=getInstrumentList\&Item_Id=295122\&UL=1V\&) and 2A-R (http://www23.statcan.gc.ca/imdb/p3Instr.pl?Function=getInstrumentList\&Item_Id=295299\&UL=1V\&) are the long-form questionnaire. The questions on form 2 A are also on these forms, followed by additional questions on various subjects such as education, employment and dwellings. Form 2A-L is distributed to $25 \%$ of Canadian households in mail-out areas and list/leave areas, as described in sections 1.1.1.1 and 1.1.1.2; these are the longform sample households. Form 2A-R has nearly all the same questions as form 2A-L and is used in areas where canvassing is the collection method. All households in these areas are included in the long-form sample.

Form 2C (http://www23.statcan.gc.ca/imdb/p3Instr.pl?Function=getInstrumentList\&Item_Id=295305\&UL=1V\&) is used to enumerate Canadians posted abroad. Form 2C has the same questions as form 2A.

Form 3A (http://www23.statcan.gc.ca/imdb/p3Instr.pl?Function=getInstrumentList\&Item_Id=295255\&UL=1V\&) is an individual census questionnaire used to enumerate persons in collective dwellings. It can also be used to enumerate usual residents in a private household who prefer to complete their own census questionnaire rather than be included in a form 2A.

### 1.3 Collection response rate

The overall collection response rate for the 2016 Census of Population was $98.4 \%$. This rate was calculated directly from the collection results (as of August 4, 2016), i.e., before data processing and quality verification were completed. It represents the number of private dwellings for which a questionnaire was returned, divided by the number of private dwellings that enumerators coded as being occupied. The collection response rate for the long-form sample was $97.8 \%$ (for more information, see the 2016 Census of Population collection response rates (http://www12.statcan.gc.ca/census-recensement/2016/ref/response-rates-eng.cfm)).

## 2. Sampling ${ }^{2}$

When a sample survey is conducted, the sample selection must be planned properly. In sampling, a subset of the survey's target population is selected to receive the questionnaire. The responses of the subset are used to draw inferences for the entire population. Two types of sampling exist: probability sampling and non-probability sampling. Probability sampling is preferable when producing statistical inferences for the entire population is important, since the probability of unit selection can be calculated and the sampling error can be estimated. This chapter discusses the selection of the sample that received the 2016 Census long-form questionnaire.

### 2.1 Long-form sample universe

The census household universe was broken down into three parts: private households, collective households and households outside Canada. The long-form sample universe consists only of private households, including those living in private dwellings attached to collective dwellings in Canada. This universe excludes incompletely enumerated Indian reserves and settlements. Unless otherwise specified, the term "in scope" indicates that a household is part of the long-form sample universe (i.e., private households that are not living in incompletely enumerated Indian reserves and settlements). "Out of scope" refers to households not in the universe (i.e., households living in collective dwellings, outside Canada, or in incompletely enumerated Indian reserves and settlements).

### 2.2 Long-form sampling design

In most cases, the long-form questionnaire was distributed to one-quarter of the households in the long-form universe to gather demographic and socioeconomic data on the Canadian population. The sample was selected from the list of dwellings for the 2016 Census of Population. At the time the sample was selected, the addresses of out-of-scope dwellings were unknown. This meant that some dwellings erroneously received a long-form questionnaire. Once a dwelling was determined to be out of scope, no further collection or processing activities were carried out.

Dwellings were selected to receive the long-form questionnaire according to a stratified systematic sampling design. The sampling design strata were the CUs. For mail-out CUs, the sampling was systematic, with a onequarter sampling fraction. The selection starting point was random. For list/leave CUs, sampling was systematic, and the sampled dwellings were every fourth one on the list, i.e., 4th, 8th, 12th, etc. For example, if a list/leave CU had seven dwellings, only one dwelling was selected. Finally, in canvasser CUs, all households were selected. These CUs were take-all strata.

The sampling design had one exception. Private dwellings attached to collective dwellings were added to the sample with certainty. However, they completed only the short-form questionnaire. Long-form questionnaire responses were later imputed for these households.

Except private households attached to collective dwellings, all households selected for the sample were asked to complete the long-form census questionnaire. Households in private dwellings that were not part of the long-form sample were asked to fill out the short-form questionnaire.

[^1]
## 3. Census data processing

### 3.1 Introduction

This chapter discusses the processing of all the completed questionnaires (all questionnaire types), which encompasses everything from the receipt of the questionnaires through to the creation of an accurate and complete census database. It describes the steps of questionnaire registration, questionnaire imaging and data capture, editing, error correction, failed edit follow-up, coding, dwelling classification and non-response adjustments, linkage of income data, imputation, weighting, and final response rates.

Automated processes, implemented for the 2016 Census, had to be monitored to ensure that all Canadian residences were enumerated once and only once. The Master Control System (MCS) was built to control and monitor the process flow, from collection to data processing. The MCS held a master list of all the dwellings in Canada, where each dwelling was identified with a unique identifier. This system was updated on an ongoing basis with information about each dwelling's status in the census process flow (e.g., delivered, received or processed). Reports were generated daily by the system and made accessible online to managers to ensure that census operations were efficient and effective.

### 3.2 Receipt and registration

Responses received through the Internet or help-line telephone interviews were received directly at the Data Operations Centre (DOC), where the receipt of the responses was registered automatically.

Respondents completing paper questionnaires mailed them back to the DOC. Canada Post registered their receipt automatically in multiple locations in Canada (as part of the normal mail flow process) by scanning the barcode on the front of the questionnaire through the transparent portion of the return envelope. The envelopes were then delivered to the DOC throughout each business day. Canada Post would also send files daily listing all census questionnaires received at each regional processing plant, by date of receipt.

The registration of each returned questionnaire was flagged on the MCS at Statistics Canada. A list of all the dwellings for which a questionnaire had not been received was generated daily by the MCS and transmitted to field operations to prevent follow-up on households that had already completed their questionnaire during NRFU.

### 3.3 Scanning and keying from images

In 2016, all paper census forms (2A, 2C, 2A-L, 2A-R, 3A) were imaged. The following steps were part of the imaging process:

- Document preparation: Mailed-back questionnaires were removed from envelopes and foreign objects (i.e., clips and staples) were detached in preparation for scanning. The questionnaires were batched by form type. Their spine was cut off to separate them into single sheets.
- Scanning: The questionnaires were converted to digital images.
- Automated image quality assessment: An automated system analyzed the images for errors or anomalies. Images failing this process were sent to be reviewed by a document analysis operator.
- Document analysis: At this step, images containing anomalies were presented to an operator for review. The operator could accept the image as is and send it directly to key entry (bypassing automated recognition), or the operator could send the entire questionnaire to be pulled at the check-out step. See below for more details on the key entry and check-out steps.
- Automated recognition: This step attempted to automatically recognize all handwritten responses and marks on the questionnaires.
- Key entry: Operators entered responses that automated recognition could not determine with sufficient accuracy. About $13 \%$ of all responses were sent to keying.
- Check-out: Once the questionnaires were processed through all of the above steps, the paper questionnaires were checked out of the system. Check-out is a quality assurance process that ensures that the images and captured data are of sufficient quality that the paper questionnaires require no subsequent processing. Questionnaires that had been flagged as containing errors were pulled at checkout and reprocessed.


### 3.4 Coverage edits, completion edits and failed edit follow-up

At this stage, a number of automated edits were performed on respondent data. These edits were designed to detect cases where the number of persons counted in the household was incorrect because of an error in collection, a respondent error or a data capture error. Most of these errors occurred on paper questionnaires, including:

- data erroneously entered in the wrong person column
- crossed off data that are captured in error
- data not being provided for every household member listed in the roster at the beginning of the questionnaire.

Errors that can occur both on paper and online include:

- data provided for the same person on more than one questionnaire (e.g., a person completes his or her own 3A questionnaire and is also included on the household 2A questionnaire)
- the receipt of duplicate questionnaires (e.g., a person completes the Internet version and his or her spouse completes the paper version and mails it back).

For about $58 \%$ of edit failures, the system resolved the case automatically. This was done when the error was such that the solution was obvious. The solutions included deleting false person data that were created because of respondent or capture error and deleting duplicate responses. The remainder of the edit failure cases were forwarded to processing clerks for resolution. An interactive system enabled the clerks to compare data across questionnaires and examine the images of paper questionnaires to detect data capture or respondent errors. Edit failures were resolved by deleting invalid or duplicate persons or by adding missing persons (i.e., creating blank person records), as necessary and appropriate.

Following the coverage edits, another set of automated edits was run. These edits detected cases where too many questions had missing responses or where data had not been provided for all the usual residents in the household, including cases where missing persons were added by coverage edit clerks. Households that failed these edits were followed up with. An interviewer called the respondent to resolve coverage issues and obtain missing responses, using a computer-assisted telephone interviewing application. For households that responded to the long-form questionnaire, only data missing for the short-form questions were followed up on. The data obtained through this follow-up activity were introduced into the system for subsequent processing steps. If the follow-up was unsuccessful, the data were imputed in the edit and imputation step (see section 3.9).

### 3.5 Coding

The census questionnaires contained questions for which answers could be checked off a list, as well as questions requiring a written response. Each written response was automatically assigned a numerical code according to Statistics Canada reference files, code sets and standard classifications. Reference files for the automated match process were built using actual responses from past censuses, as well as administrative files. Specially trained coders and subject-matter specialists resolved cases where a code could not be automatically assigned. The following questions required coding on both the long- and short-form questionnaires:

- relationship to Person 1
- home language
- mother tongue.

The following questions required coding for the long-form sample only:

- place of birth
- citizenship
- non-official languages
- ethnic origin
- population group
- First Nation/Indian band
- place of residence one year ago
- place of residence five years ago
- place of birth of parents
- major field of study
- location of study
- industry
- occupation
- place of work
- language of work.

A total of about 69 million write-ins were coded from the 2016 Census questionnaires. Overall, about $85 \%$ were coded automatically, although the autocoding rate varied considerably from one question to the next.

### 3.6 Classification and non-response adjustments for unoccupied and non-response dwellings

The Dwelling Classification Survey (DCS) was used to estimate the rate of enumerator error in classifying private dwellings in mail-out and list/leave census CUs as occupied or unoccupied. This information was used to make adjustments to the census database. The DCS selected a random sample of 1,730 mail-out and list/leave CUs. Enumerators revisited these CUs in June and July 2016 to reassess the occupancy status as of Census Day of each private dwelling for which no response was received. The DCS estimated that $15.0 \%$ of the $1,187,392$ private dwellings classified as unoccupied were actually occupied and that $36.9 \%$ of the 284,966 private dwellings with no response that were classified as occupied or that had an unknown occupancy status were actually unoccupied. Estimates based on the DCS sample were used to adjust the occupancy status for individual dwellings. This resulted in an increase of $2.6 \%$ in the number of occupied private dwellings and a decrease of $6.2 \%$ in the number of unoccupied dwellings at the Canada level.

After this adjustment of the occupancy status by the DCS, occupied private dwellings with total non-response had the number of usual residents (if not known) and all the responses to the census questions imputed. The
responses were borrowed from another responding household within the same CU. This process, called whole household imputation (WHI), imputed $99.9 \%$ of the total non-response households. Using a single donor under WHI was more efficient computationally and was less likely to produce implausible results than using several donors as part of the main edit and imputation process. Nevertheless, the other $0.1 \%$ of the total non-response households where no donor household was found under the WHI process was imputed as part of the main edit and imputation process.

The WHI process has another component that is separate from the use of the DCS estimates to adjust the census database. The non-DCS areas-CUs that have interviewer-administered census questionnaires (i.e., Indian reserve, canvasser and collective CUs)—require a different imputation strategy. In these areas only, all unoccupied dwellings are assumed to be truly unoccupied and all non-responding dwellings are assumed to be truly occupied. This implies that unoccupied dwellings are assumed to be classified correctly and no imputations are done. Private dwellings with an occupancy status classified as unknown are also assumed to be unoccupied. On the other hand, private dwellings with no response that were classified by enumerators as being occupied are all assumed to be occupied, and the geographically nearest neighbour is used as the donor household for these dwellings. No restrictions were placed on the household size for these imputations, as was done in the DCS areas. At the Canada level (for DCS and non-DCS areas), $2.6 \%$ of occupied private dwellings were imputed through the WHI process.

More details on the DCS and the WHI process will be available in the Coverage Technical Report, Census of Population, 2016 (http://www12.statcan.gc.ca/census-recensement/2016/ref/98-303/index-eng.cfm), Statistics Canada Catalogue no. 98-303-X, which will be released in 2019.

### 3.7 Obtaining income data

For the first time, in 2016, administrative data were the only source of information on income for the Census Program. This not only reduced response burden, but also increased the quality and quantity of the income data available. The information on individuals' income was compiled from administrative data for the entire population aged 15 and older, rather than from a sample, as was done in 2011 and 2006. Regular, recurring taxable and non-taxable income received during the 2015 calendar year was included. One-time receipts, such as lump-sum withdrawals from registered retirement savings plans and other savings plans, lump-sum insurance settlements, lump-sum pension benefits, capital gains or losses, inheritances, and lottery winnings, were excluded.

The information on census respondents could be linked to two types of Canada Revenue Agency (CRA) files, depending on whether respondents were (1) taxfilers, for whom all income information could be extracted from income tax files, including T1 general returns, tax slips and government programs administered by the CRA, or (2) non-filers, for whom the only information available came from tax slips and government programs administered by the CRA. In 2016, the information for $94.8 \%$ of the population aged 15 and older in private households was linked to a CRA administrative file. Specifically, the information for $85.2 \%$ of the population was linked to a taxfiler's file, and the information for $9.6 \%$ of the population was linked to a non-filer's file.

For more information on how income data were obtained, see the Income Reference Guide, Census of Population, 2016 (http://www12.statcan.gc.ca/census-recensement/2016/ref/guides/004/98-500-x2016004-eng.cfm), catalogue no. 98-500-X2016004.

### 3.8 Non-response

A non-response status may differ during the collection and processing phases. The main differences arise because the occupancy status can change between collection and processing, and because the household must answer a minimum number of questions to be considered a respondent in the processing phase. Unless otherwise specified, the term "non-response" refers to non-response in the data processing phase. The same applies when response is referred to rather than non-response.

For the 2016 Census long-form questionnaire, two types of households were considered non-respondent:

- households from the sample that answered only the questions common to both types of questionnaires, i.e., only the short-form questions
- households that did not answer any questions.

This refers to total non-response, which is processed differently depending on the collection method and the type of household.

Finally, partial non-response is when the long-form questionnaire is partially completed. This type of response is processed by imputation. An overview of this method is presented in the next section.

### 3.9 Edit and imputation

The data collected in any survey or census contain some omissions or inconsistencies. For example, a respondent may be unwilling to answer a question, fail to remember the right answer or misunderstand the question. Other errors, such as incorrect coding, can also occur.

The final clean-up of data, done in the edit and imputation process, was fully automated using the Canadian Census Edit and Imputation System (CANCEIS) (Statistics Canada 2014) for all census topics. Two imputation methods were applied. The first method, called "deterministic imputation," involved assigning specific values under certain conditions when problems were clear and unambiguous to resolve. Detailed edit rules were applied to identify these conditions, and the variables involved in the rules were assigned predetermined values. The second method, called "minimum-change nearest-neighbour donor imputation," applied a series of detailed edit rules that identified any missing or inconsistent responses. When a record with missing or inconsistent responses was identified, another record that met the edit rules and had most characteristics in common with the record with an error was selected. Data from this donor record were borrowed and used to make the minimum number of changes to the variables to resolve all cases of missing or inconsistent responses.

### 3.10 Weighting

The 2016 Canadian Census Program consisted of a census of population and a sample survey for which onequarter of Canadian private households were selected. Households not sampled for the survey received a shortform questionnaire, while sampled households received a long-form questionnaire. In addition to the short-form questions, the long-form questionnaire gathered sociocultural information, as well as information on daily activities, mobility, place of birth, education, labour market activity, etc. Weighting was used to represent the entire population based on the information gathered from the sample.

The first step in the weighting process was to assign a design weight to each household that reflected its probability of being sampled. These weights then underwent an initial adjustment for coverage and total nonresponse. This adjustment was applied to the weights of respondent households. Finally, a second adjustment, referred to as final calibration, was made to establish closer agreement between the estimates obtained from respondent households in the sample and the census counts for a number of characteristics from the short-form questionnaire or from administrative data sources. The weighting methodology is described in detail in Chapter 4. All private households attached to collective dwellings and all private households in a canvasser CU were selected for the long-form sample and received a design weight of 1 . They were then excluded from the coverage and nonresponse adjustment processes, as well as from the final calibration process.

Long-form sample households with a non-zero weight at the end of the weighting process were the respondent households, along with the households who were assigned a design weight of 1, i.e., private households attached to collective dwellings and all private households in canvasser CUs. These households made up the set of households that contributed to the long-form estimates.

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### 3.11 Final response rates

Table 3.11.1 presents the final response rates for private households in the 2016 Census of Population, for Canada and for each province and territory, followed by non-weighted and weighted response rates for the longform sample based on the definition of non-response given in section 3.8.

The final response rate is the ratio of the numerator to the denominator, where

- the numerator is the number of private dwellings for which a questionnaire was completed ${ }^{3}$
- the denominator is the number of private dwellings classified as occupied, according to the census database.

The final classification of a dwelling's occupancy status is based on an analysis of the data gathered by field staff, data provided by respondents and the results of a study into the quality of occupancy status in the DCS (see section 3.6). The response rates indicated in Table 3.11.1 differ from the collection response rates, which were previously published and were mentioned in section 1.3 , in that they take data processing and dwelling occupancy verification into account in identifying non-respondent households. These response rates are therefore considered final.

Weighted response rates were produced for the long-form sample. They are based on the following weights as the numerator and denominator:

- the numerator is the design weight
- the denominator is the weights calibrated to the census total for the geographic area (which is the number of private households in occupied dwellings according to the census).

[^2]Table 3.11.1
Final response rates for private households from the 2016 Census of Population and the long-form sample
$\begin{array}{lrrr}\hline & \begin{array}{r}\text { Response rate - } \\ \text { short-form } \\ \text { questionnaire }\end{array} & \begin{array}{r}\text { Non-weighted } \\ \text { response rate - } \\ \text { long-form } \\ \text { questionnaire only }\end{array} & \begin{array}{r}\text { response rate }- \\ \text { long-form }\end{array} \\$\cline { 2 - 4 } questionnaire only\end{array}$]$

[^3]Source: Statistics Canada, 2016 Census of Population and 2016 Census long-form sample.

## 4. Estimation from the census long form sample

Any sampling process requires an associated estimation procedure for scaling sample data up to the population level and for ensuring that survey estimates are representative of the population. The choice of an estimation procedure is generally governed by both operational and theoretical constraints. From the operational viewpoint, the procedure must be feasible within the processing system of which it is a part, and from the theoretical viewpoint, the procedure should minimize the statistical error of the estimates it produces.

The estimation procedure produces a set of weights, and the weight for each sample unit corresponds to the number of units in the population that the sample unit represents. These weights are applied to the sample data to produce millions of estimates from the census long-form sample. Estimates are summary measures such as totals, averages, proportions and medians calculated from the sample for various characteristics of interest.

### 4.1 Considerations in the choice of an estimation procedure

### 4.1.1 Operational considerations

Mathematically, an estimation procedure can be described by an algebraic formula, or estimator, that shows how the estimate for the population is calculated as a function of the observed sample values and other information from the sample design or external data sources. Most of the time, this estimator is a simple function of weights and of the variable of interest for the responding units. Using a unique set of weights to produce all estimates guarantees a certain level of consistency among the different estimates of the survey.

Therefore, the approach taken for the census long-form sample (and in most sample surveys) was to split the estimation procedure into two steps: (a) the calculation of weights (known as the weighting procedure) and (b) the use of weights to produce estimates, such as the estimation of a particular population count by summing the weights of those persons or households with the characteristic of interest. Most of the mathematical complexity is contained in step (a), which is performed just once. Meanwhile, step (b) is reduced to a simple process, such as summing weights whenever tabulation is required. Since the weight attached to each sample unit is the same for any tabulation involving that unit, consistency between different estimates based on sample data is assured.

### 4.1.2 Theoretical considerations

For a given sample design and a given estimation procedure, one can, from sampling theory, make a statement about the chances that a certain interval will contain the unknown population value being estimated. A primary criterion in the choice of an estimation procedure is the minimization of the width of such intervals for a given level of confidence so that these statements about the unknown population values are as precise as possible. A common measure of precision for comparing estimation procedures is known as the standard error. Provided that certain conditions are met, intervals of plus or minus two standard errors from the estimate will contain the true population value for approximately $95 \%$ of all possible samples.

As well as minimizing standard error, a second objective in the choice of an estimation procedure for the longform sample is to ensure, as far as possible, that sample estimates for census characteristics are consistent with the corresponding known census values. Fortunately, these two objectives are usually complementary in the sense that sampling error tends to be reduced by ensuring that sample estimates for certain basic characteristics are consistent with the corresponding population figures. However, while this is true in general, forcing longform sample estimates for census characteristics to be consistent with corresponding census figures for very small subgroups can have a detrimental effect on the standard error of estimates for the sample characteristics themselves. For example, if in several dissemination areas only a few subjects have a given characteristic, such as birth in a certain country, ensuring consistency between the sample estimates and the census counts for that place of birth would unduly increase the standard error for the rest of the characteristics.

In cases where no information about the population being sampled is available other than that collected for sample units and unit non-response has not occurred, the estimation procedure would be restricted to weighting the sample units inversely to their probability of selection. For example, if a unit had a one-in-four chance of selection, then that selected unit would receive a weight of 4 . When unit non-response is observed, the weight must be further adjusted according to the estimated probability of response of the unit, for example. In practice, some supplementary knowledge about the population (e.g., its total size and possibly its breakdown by a certain variable-perhaps by province and territory) is often available. Such information can be used to improve the estimation formula so as to produce estimates with a greater chance of being close to the unknown population value. In the case of the census long-form sample, a large amount of very detailed information about the population being sampled is available from the census short-form data at every geographic level. This wealth of population information is used in the coverage, non-response and calibration adjustments to improve the estimates made from the long-form sample.

Nevertheless, the long-form sample estimates for census characteristics cannot be made consistent with all the census counts at every geographic level. Differences between sample estimates and census counts become visible when a cross-tabulation of a sample variable and the corresponding census variable is produced. The tabulation of sample-based estimates of totals for particular characteristics will not necessarily agree with the equivalent census count tabulations for those characteristics.

Adjusting the weights, by the minimal amounts possible, to achieve perfect agreement between long-form estimates and census counts for certain characteristics and subgroups is known as "calibration."

### 4.2 Weighting areas

The various adjustments to design weights were made independently by weighting area. The geographic areas used for this purpose were aggregate dissemination areas (ADAs) and super aggregate dissemination areas (SADAs). ADAs are a new dissemination geography created for the 2016 Census. SADAs were created specifically for the weighting procedures by ADA aggregation. The geographic subdivisions used in 2011 and earlier were constructed differently.

### 4.2.1 Aggregate dissemination areas

ADAs are a new dissemination geography created for the 2016 Census. Their purpose is to enable the dissemination of a greater quantity of data at a detailed geographic level across the country. In total, Canada was divided into 5,386 ADAs, and households were selected for the long-form sample in 5,143 ADAs. Of the 243 ADAs without sampled households, 235 consisted solely of out-of-scope households. The other eight ADAs had only a handful of in-scope households, and none of them were selected.

ADAs satisfy the following delineation criteria:

1. ADAs cover the entire country and, where possible, have a population count of 5,000 to 15,000 (based on the population counts from the previous census).
2. ADAs respect provincial and territorial borders, as well as the boundaries of census divisions (CDs), census metropolitan areas (CMAs) and census agglomerations (CAs) subdivided into census tracts (CTs) in effect for the 2016 Census.
3. ADAs are based on one of three 2016 Census dissemination geographic areas: dissemination areas (DAs), census subdivisions (CSDs) or census tracts (CTs):

- Within CMAs and CAs with CTs, adjacent CTs are combined to meet the ADA population criterion.
- In areas without CTs (areas outside CMAs and the largest CAs) where CSDs have a population of fewer than 15,000, adjacent CSDs are combined to meet the ADA population criterion.
- In areas without CTs where CSDs have a population of over 15,000 , adjacent DAs are combined within these CSDs to meet the ADA population criterion.


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4. Every CSD that consists of an Indian reserve and a small number of other areas where the canvasser method is required constitute distinct ADAs.
"For more information about Aggregate Dissemination Areas, refer to the Dictionary, Census of Population, 2016, (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/index-eng.cfm) Catalogue no. 98-301-X."

Table 4.2.1.1 shows the degree to which ADAs with households in the long-form sample were properly adjusted to CSDs. The first scenario occurred in most cases, since ADAs were designed above all to respect the boundaries of CTs and CSDs. Scenario 4 is the only one where CSD boundaries were not respected. CTs were not included in the table because they were all in the first scenario except one, which was in scenario 3.

Table 4.2.1.1
Number of census subdivisions within the boundaries of ADAs with households in the long-form sample, 2016 Census

| Scenario | Description | Census subdivision |  |
| :--- | :--- | ---: | ---: | ---: |
|  | number <br> The CSD was small enough to be fully contained in an ADA, and <br> this ADA only had complete CSDs. No CSDs in the ADA were part <br> of another ADA. | 4,512 | 92.40 |
| 2 | The CSD was small enough to be fully contained in an ADA, but another <br> CSD in the same ADA was part of a different ADA. | 81 | 1.66 |
| 3 | The CSD was large enough to contain full ADAs. No ADAs were also part <br> of another CSD. | 261 | 5.35 |
| 4 | The CSD was part of two or more ADAs. | $\mathbf{2 9}$ | $\mathbf{0 . 5 9}$ |
| Total |  | $\mathbf{4 , 8 8 3}$ | $\mathbf{1 0 0 . 0 0}$ |

CSD = Census subdivision.
ADA = Aggregate dissemination area.
Source: Statistics Canada, 2016 Census long-form sample.

Table 4.2.1.2 shows the distribution of ADAs with households in the long-form sample by province or territory.

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Table 4.2.1.2
Number of ADAs with households in the long-form sample, by province or territory

| Region | Number of ADAs |
| :--- | ---: |
| Newfoundland and Labrador | 81 |
| Prince Edward Island | 21 |
| Nova Scotia | 146 |
| New Brunswick | 124 |
| Quebec | 1,118 |
| Ontario | 1,655 |
| Manitoba | 216 |
| Saskatchewan | 256 |
| Alberta | 516 |
| British Columbia | 916 |
| Yukon | 28 |
| Northwest Territories | 40 |
| Nunavut | 26 |
| Canada | 5,143 |

ADA = Aggregate dissemination area.
Source: Statistics Canada, 2016 Census long-form sample.

Table 4.2.1.3 shows the number of ADAs by the number of in-scope households in the census. The majority of ADAs with households in the long-form sample had from 2,000 to 4,999 households. A considerable number of ADAs had small populations.

Table 4.2.1.3
Distribution of ADAs with households in the long-form sample, by number of in-scope households

| In-scope households | Number <br> of ADAs | Percent |
| :--- | ---: | ---: | ---: |
| 0 to 499 | 976 | 18.98 |
| 500 to 999 | 117 | 2.27 |
| 1,000 to 1,999 | 366 | 7.12 |
| 2,000 to 2,999 | 1,339 | 26.04 |
| 3,000 to 3,999 | 1,229 | 23.90 |
| 4,000 to 4,999 | 664 | 12.91 |
| 5,000 to 5,999 | 300 | 5.83 |
| 6,000 to 6,999 | 98 | 1.91 |
| 7,000 to 7,999 | 32 | 0.62 |
| 8,000 to 8,999 | 14 | 0.27 |
| 9,000 to 9,999 | 3 | 0.06 |
| $10,000+$ | 5,143 | 0.10 |
| Total | $\mathbf{5}$ | $\mathbf{1 0 0 . 0 0}$ |

ADA = Aggregate dissemination area.
Source: Statistics Canada, 2016 Census of Population.

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Table 4.2.1.4 presents the number of ADAs by range of numbers of households that responded to the 2016 Census long-form questionnaire. For the ADAs with the fewest respondents, a specific type of processing was applied to have enough households for weighting purposes (see section 4.5). Overall, ADAs had more respondents than the weighting areas used in 2011.

Table 4.2.1.4
Distribution of ADAs with households in the long-form sample, by number of respondent households for the long-form questionnaire

| Respondent households | Number of ADAs | Percent |
| :---: | :---: | :---: |
| 0 to 99 | 605 | 11.76 |
| 100 to 199 | 265 | 5.15 |
| 200 to 299 | 158 | 3.07 |
| 300 to 399 | 142 | 2.76 |
| 400 to 499 | 311 | 6.05 |
| 500 to 599 | 537 | 10.44 |
| 600 to 699 | 628 | 12.21 |
| 700 to 799 | 604 | 11.74 |
| 800 to 899 | 518 | 10.07 |
| 900 to 999 | 400 | 7.78 |
| 1,000 to 1,099 | 304 | 5.91 |
| 1,100 to 1,199 | 229 | 4.45 |
| 1,200 to 1,299 | 162 | 3.15 |
| 1,300 to 1,399 | 101 | 1.96 |
| 1,400 to 1,499 | 73 | 1.42 |
| 1,500+ | 106 | 2.06 |
| Total | 5,143 | 100.00 |

ADA = Aggregate dissemination area.
Source: Statistics Canada, 2016 Census long-form sample.

### 4.2.2 Super aggregate dissemination areas

SADAs were created specifically for weighting 2016 Census data, so that certain weighting procedures for which a large number of observations is desirable could be conducted.

SADAs were created according to the following rules (in order of priority):

1. SADAs are created by combining ADAs (mandatory).
2. SADAs respect provincial and territorial borders (mandatory).
3. SADAs have a population of 50,000 to 150,000 persons (except for CDs with a population of 40,000 to 50,000 persons that constitute their own SADA.
4. SADA population excluding persons living in canvasser CUs).
5. SADAs respect the boundaries of CDs.
6. SADAs respect the boundaries of CMAs and CAs.
7. SADAs respect the boundaries of CSDs.
8. SADAs are single contiguous entities.
9. SADA are as compact as possible.

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The first two rules were mandatory, and rules 3 to 9 were followed where possible. A total of 409 SADAs were created.

Table 4.2.2.1 shows the distribution of SADAs by province or territory.
Table 4.2.2.1
Number of SADAs, by province or territory

| Region | Number of SADAs |
| :--- | ---: |
| Newfoundland and Labrador | 8 |
| Prince Edward Island | 2 |
| Nova Scotia | 13 |
| New Brunswick | 8 |
| Quebec | 97 |
| Ontario | 150 |
| Manitoba | 15 |
| Saskatchewan | 14 |
| Alberta | 44 |
| British Columbia | 55 |
| Yukon | 1 |
| Northwest Territories | 1 |
| Nunavut | 1 |
| Canada | 409 |

SADA = Super aggregate dissemination area.
Note: In the case of the three territories, the SADA corresponds to the territory.
Source: Statistics Canada, 2016 Census long-form sample.
Table 4.2.2.2 shows the degree to which SADAs were properly adjusted to CDs and CMAs. SADAs respected the boundaries of the majority of CDs (scenarios 1 and 3 ) and the boundaries of three-quarters of CMAs. The other CMAs were part of at least two SADAs (scenario 4).

Table 4.2.2.2
Number of census divisions and census metropolitan areas within SADA boundaries, 2016 Census

| Scenario | Description | Census divisions |  | Census <br> metropolitan areas |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | number | percent | number | percent |
| 1 | The CD or CMA was small enough to be fully contained within a SADA, and the SADA included only complete CDs or CMAs. No CDs or CMAs in the SADA were part of another SADA. | 249 | 84.98 | 2 | 5.71 |
| 2 | The CD or CMA was small enough to be fully contained within a SADA, but another CD or CMA in the same SADA was also part of another SADA. | 2 | 0.68 | 0 | 0.00 |
| 3 | The CD or CMA was large enough to contain complete SADAs. No SADAs were also part of another CD or CMA. | 40 | 13.65 | 25 | 71.43 |
| 4 | The CD or CMA was part of two or more SADAs. | 2 | 0.68 | 8 | 22.86 |
| Total |  | 293 | 100.00 | 35 | 100.00 |

$C D=$ Census division.
CMA = Census metropolitan area.
SADA = Super aggregate dissemination area.
Source: Statistics Canada, 2016 Census of Population.

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Table 4.2.2.3 shows the number of SADAs by the number of in-scope persons.

Table 4.2.2.3
Distribution of SADAs with households in the long-form sample, by number of in-scope individuals

| In-scope individuals | Number <br> of SADAs | Percent |
| :--- | ---: | ---: | ---: |
| 30,000 to 39,999 | 2 | 0.49 |
| 40,000 to 49,999 | 26 | 6.36 |
| 50,000 to 59,999 | 23 | 5.62 |
| 60,000 to 69,999 | 45 | 11.00 |
| 70,000 to 79,999 | 106 | 25.92 |
| 80,000 to 89,999 | 67 | 16.38 |
| 90,000 to 99,999 | 46 | 11.25 |
| 100,000 to 149,999 | 94 | $\mathbf{2 2 . 9 8}$ |
| Total | $\mathbf{4 0 9}$ | $\mathbf{1 0 0 . 0 0}$ |

SADA = Super aggregate dissemination area.
Source: Statistics Canada, 2016 Census of Population.

### 4.3 Design weights

The design weight for each household in the long-form sample was calculated differently, depending on the collection method of the CU where the corresponding dwelling was located.

If the method of collection was:

- mail-out, the design weight was equal to the inverse of the survey fraction, giving a weight of 4
- list/leave, the design weight was equal to the ratio of the number of private dwellings enumerated to the number of private dwellings sampled in the CU, and this gave a weight of approximately 4 for $98 \%$ of the selected households, although the weights varied from 1 to 7
- canvasser, the design weight was 1.

Households living in private dwellings attached to collective dwellings were an exception to the rule. As mentioned in section 2.2, all of these households were included in the sample. They were considered take-all, so their design weight was 1.

### 4.3.1 Weights for households counted in the sample

Sampled households with a design weight of 1 did not have their weight adjusted. These households kept their weight of 1 after the weighting procedures were completed (coverage and non-response, as well as calibration to census counts). They either were located in canvasser CUs or were private households that were attached to a collective dwelling.

Total non-response and partial non-response for these households were addressed by imputation. Once the missing data were imputed, these households were considered to be respondents for estimation purposes (although they were considered to be non-respondents for the calculation of response rates in section 3.11).

### 4.4 Coverage and total non-response adjustment

The several ways of treating non-response in surveys can be divided into two main categories: imputation and reweighting. The former is usually applied for the treatment of item missing values and the latter for the treatment of total non-response. A household was considered to be a respondent to the long form when it answered at least one of the long-form questions. With the high response rate to the long form, any non-response adjustment method would have had, for the most part, only a modest impact on the final survey weights and estimates. Coverage and total non-response for households in Indian reserve and canvasser enumeration CUs were compensated for with imputation procedures and, for the most part, with whole household imputation (WHI) as described in section 3.6. In the rest of the country, reweighting procedures were used. The rest of this chapter describes those weighting procedures.

The main purpose of coverage and non-response adjustments is to minimize the impact of any potential biases from lack of complete coverage (or from duplicates) and from unit non-response. For the adjustment to actually reduce the potential bias, a rich set of information about the non-respondents is very useful. Otherwise, the non-response adjustment that can be applied is limited, and the potential bias will not be greatly lessened. Geographical information was known for every non-responding household and long-form sample non-respondent (i.e., respondents in the long-form sample who answered the short-form questions but not the long-form questions). The information on non-respondents was thus somewhat limited. Fortunately, before the coverage and non-response adjustments, the process of WHI occurred. An important part of WHI is to impute the short-form characteristics for all non-respondents to the short form. This included long-form sample non-respondents. This additional information served as the basis for the long-form sample non-response adjustment.

The method used to adjust for coverage and total non-response in the long-form sample was a reweighting calibration-based procedure applied to the design weights. The procedure can be divided into the following main steps:

1. selection of calibration constraints for steps 2 and 3
2. non-linear calibration coverage adjustment
3. estimation of a non-response propensity based on non-linear calibration for non-response
4. application of a score method based on the propensity of step 3.

Steps 1 to 4 were applied independently in each SADA. In other words, the non-response adjustment was applied by SADA. See section 4.2 for the definition and information about ADAs and SADAs.

The first step consisted of a forward selection of calibration constraints in the SADA. It was performed as follows:

- The set of potential constraints was derived from the variables common to both the short form and the long form, as well as from some administrative data obtained with record linkage strategies (where all units of the long-form population undergo the linkage procedures). The requirements of the non-linear calibration method used in the second and third steps meant that only constraints at the SADA level, and the number of households and persons in each ADA of the SADA, were considered.
- In each SADA, two mandatory constraints were selected first: the number of households in the SADA (TOTHHLD) and the number of persons in the SADA (TOTPERS).
- The ADA-specific constraints—number of households (HHADA) and number of persons (PPADA)—were evaluated for selection.
- All other potential SADA constraints were evaluated; priority was given to the ones that split the SADA population as closely as possible into halves.

The selection process excluded constraints that occurred in fewer than 250 households in the SADA and constraints that were redundant or almost redundant in terms of collinearity with those constraints or with constraints already selected. Constraints that were redundant with constraints already selected were excluded since they did not add any new information. Given those filters, the order of priority used in the evaluation of
constraints ensured that the constraints selected complemented each other and corrected for any potential coverage differential between the long form and the short form, as well as for census total non-response.

The second step applied a coverage non-linear calibration adjustment to the whole sample in the SADA (i.e., respondents and non-respondents). The long-form sample weighted counts, for the constraints selected in the first step, were made to coincide with the corresponding population counts. The purpose of this step was to correct for any potential coverage differential between the long-form sample and its complement (i.e., the set of households receiving only the short form). One way in which overcoverage can occur is if some individuals are counted in two different households. The coverage for the two populations could also be different if, for example, occupied dwellings were more likely to be incorrectly treated as unoccupied dwellings for the long form than for the short form. Another objective of this step was to isolate as much as possible the sampling error. Without this step, the non-response calibration carried out in the next step would confound the non-response error with the sampling error. This step makes the sample estimates coincide with the population estimates. In addition, the same control totals are used in both calibration procedures. As a result, the non-response propensity estimation done next does not have to correct (directly or indirectly) for the sampling error. Combining a correction for the sampling error and for the non-response error in the next step would have been inappropriate. The calibration procedure would have failed if the weight of any respondent was required to decrease to match the census counts, because the estimated propensity would have been greater than 1 . Moreover, the score method applied in the last step required an estimate of the response propensity alone. To the extent that the variable of interest was related to the selected constraints, the sampling variance was also reduced by this step.

After these two steps, the main non-response adjustment took place. The weights, adjusted in the previous step, of non-respondents were set to 0 and the weights of respondents were increased so that the weighted sums in the SADA coincided with the corresponding population counts for the selected constraints. A logistic link function between the response propensity and the characteristics used in calibration enabled the implicit estimation of the response propensity. Folsom and Singh (2000) proposed this non-linear calibration method as a way of adjusting for non-response while ensuring that both estimates coincided with selected population counts and that the estimated response probabilities were between 0 and 1 . This last condition does not necessarily hold when linear calibration is used for non-response adjustment. To the extent that the response propensity was related to the selected constraints, this step reduced the potential non-response bias without increasing the variance.

The inverse of the estimated response probabilities obtained in the previous step could be directly used to adjust the weights for non-response. However, the score method was used for the last step of the non-response adjustment to smooth the estimated probabilities from the previous step. This further ensured the quality of the non-response adjustment and avoided too extreme adjustments. For each ADA, homogeneous weighting classes were formed according to the estimated response probabilities. In each class, the weighted harmonic mean of the response probabilities was calculated. The harmonic mean was used because it is less affected by outliers in the estimated response probabilities. The inverse of this mean was applied to the weights of respondents in the class as the non-response adjustment.

In summary, the coverage and total non-response adjustment was a product of two quantities: the coverage adjustment and the inverse of the score-method harmonic mean.

### 4.5 Final calibration

Final calibration is a linear calibration that was done to minimize the sampling variability of estimates derived from long-form questionnaire responses, while ensuring consistency between estimated totals and Census of Population totals. This weighting step was necessary, since ensuring consistency between estimated totals and Census of Population totals was important for a large number of variables and geographic areas, i.e., satisfying calibration constraints.

Only the weights for households in mail-out or list/leave CUs were calibrated, since these households were sampled. Exceptions to this rule were households in these CUs that lived in a private dwelling attached to a collective dwelling. Since all these households were included in the long-form sample and all the long-form questionnaire responses for these households were imputed, no calibration was done. The final weights for these
households were therefore equal to 1 . The weights produced by the calibration process were the final weights used to calculate the long-form estimates, and these weights applied to households as well as families and persons. In other words, all families and persons from the same household received the household weight. For this final adjustment, the variability of the calibrated weights needed to be limited to avoid having an excessive portion of the weight applied to a single household or person. Therefore, weights were constrained to range from 1 to 20 .

Calibration constraints were defined at the person, household and census family levels. In 2016, the notion of constraints was also expanded by the addition of two levels to the hierarchy of geographic units, i.e., ADAs and SADAs. These two levels were added to maximize the overall consistency between estimated totals and Census of Population totals, while minimizing the number of calibration constraints. This should help to reduce the variability of estimates. Appendix C lists all of the ADA and SADA constraints that were taken into consideration during the calibration process. Characteristics available from the census, administrative sources and the long-form questionnaire and for which consistency was attempted included age, sex, marital status, common-law status, household size, dwelling type and official language spoken.

Constraints were chosen in both types of geographic areas simultaneously, independent of each other. Calibration was then performed using all of the selected constraints. In 2016, the addition of calibration constraints for two geographic levels (ADAs and SADAs), the removal of mother tongue constraints, and the addition of constraints based on linked administrative data meant that the number of possible constraints was different than for the 2011 Census. A total of 270 constraints were defined for ADAs and 200 for SADAs. Various factors drove the choice of geographic level for calibration constraints. This choice was made in collaboration with subject-matter experts. For example, some constraints were defined only for SADAs, since they would not have been populated enough at the ADA level. Other constraints, such as age groups, were chosen in a way that ensured they were not only populated enough but also not too similar when assessed by the selection process.

To facilitate their calibration, small ADAs were combined before the selection of calibration constraints to ensure a minimum of 60 long-form respondent households per ADA. Small ADAs that fell entirely within a CSD were initially combined with other ADAs in the same SADA. Next, small ADAs in CDs were combined with other ADAs in the same SADA. Finally, the remaining small ADAs were combined with an ADA from an adjacent SADA. The ADA grouping procedure produced 4,180 groups of ADAs with 60 or more respondent households.

The first step in the process to select calibration constraints was to categorize each of the constraints into one of three groups:

Mandatory constraints: These constraints had to be used in the calibration because the census counts had to agree with the long-form estimates at the geographic levels that are usual aggregates of ADAs and SADAs (e.g., Canada, provinces and territories). The number of persons and the number of households in the ADAs and SADAs were the two mandatory constraints.

Low-response constraints: Constraints evaluated for a population of 200 or fewer households were not used in the calibration because they can make survey estimates unstable.

All other constraints: These constraints were examined further to see whether they should be used in the calibration.

The second step was to determine which constraints from the third group should be used in the calibration process, in addition to the mandatory constraints. The constraints from the third group were added one by one, by repeatedly choosing the constraint that divided the population of the SADA or ADA in two as evenly as possible.
Constraints that were too linearly dependent were excluded. To avoid introducing a bias in the point estimates and to avoid increasing their variance, the number of selected constraints was limited. Evaluations determined that this number had to be smaller than the square root of the number of respondent households involved in the constraint.

After the calibration constraints to be used were selected, a final edit was done to check whether the set of constraints chosen at the ADA and SADA levels was free of collinearity.

The calibration itself was then carried out for the final set of constraints from the second step. The weights adjusted for coverage and non-response were modified as little as possible, so that the weighted estimates would be equal to census totals for these constraints. Statistics Canada's Generalized Estimation System (GES) was used to carry out the calibration.

Sample estimates can differ from census counts for a few reasons, particularly for small areas, even after the calibration step. A few of these reasons are given below.

- Constraints excluded during the constraint selection process: As described above, possible constraints could be excluded for having low counts, for being linearly dependent (or overly dependent) on other chosen constraints or for being linearly dependent (or overly dependent) on low-response constraints. This led to some differences between census counts and long-form estimates for these variables when a perfect linear dependency with the chosen constraints was not present.
- Sub-weighting area: In 2016, the ADA was the smallest weighting area for which agreement was attempted between the census counts and the long-form estimates. Any entity smaller than an ADA, such as the majority of DAs, is referred to as a sub-weighting area. These sub-weighting areas could have discrepancies between the census counts and the long-form estimates.


### 4.6 Details on the selection of constraints ${ }^{4}$

Constraints were selected twice during the weighting process: during the coverage and non-response adjustment, which requires the use of non-linear calibration techniques, and during the final calibration. The variables making up the constraints were essentially the same, but the inclusion or exclusion of constraints varied between the two weighting steps, since their respective objectives were different. Basically, constraints were not selected using the exact same criteria, and the weighting areas varied depending on the weighting step.

This section explains how the constraints were selected during these weighting steps and indicates how often certain constraints were excluded during the coverage and non-response adjustment and during the final calibration. Constraint selection was carried out independently in the 408 SADAs that had sampled households with an adjusted weight, as well as in the 4,238 corresponding ADAs. One SADA did not undergo a weight adjustment since it had only counted households (the SADA corresponding to Nunavut). This explains why the following tables present results for 408 SADAs, while the total number of SADAs indicated in section 4.2 was 409. During the coverage and non-response adjustment, 4,646 weighting areas were defined, i.e., 408 SADAs and 4,238 ADAs, while 4,588 weighting areas were defined at the final calibration step. The number of weighting areas differed depending on the weighting step because, during the final calibration, small ADAs were combined into "calibration ADAs" to ensure a minimum of 60 households per ADA, for a total of 4,180 ADAs at the calibration step.

### 4.6.1 Process for the coverage and non-response adjustment

The coverage and non-response adjustment procedure uses calibration to adjust the survey weights. The rationale is that, if estimates based on respondents match as much as possible the census counts for auxiliary variables, then the non-response bias of estimates associated to those variables will be reduced. See section 4.4 for more information on the coverage and non-response adjustment.

As mentioned in section 4.4, the coverage and non-response adjustment was carried out independently in the SADA, where several variables were used to define the constraints, and in the ADA for two variables. A sequential procedure was used to select the constraints. In each SADA, two mandatory constraints were selected first: the number of households in the SADA (TOTHHLD) and the number of persons in the SADA (TOTPERS). After that,

[^4]the ADA-specific constraints—number of households (HHADA) and number of persons (PPADA)—in each of the ADAs were considered. Then, all the other potential SADA constraints were considered, with priority given to the ones that split the SADA (approximately) in half. The non-mandatory constraints were entered sequentially into the adjustment as long as they did not fall into the criteria listed in Table 4.6.3.1. The final list of calibration constraints that were considered for the coverage and non-response adjustment is presented in Appendix C.

A constraint that is excluded frequently usually has a larger difference between the census count and the nonresponse adjusted sample estimate than a constraint that is used more often. This can be seen by comparing Appendix C with Table 5.2.1. Appendix C lists all the potential variables or constraints, the number of times each constraint was used for calibration and the number of times that constraints were excluded for one of the reasons enumerated in Table 4.6.3.1. A constraint was excluded from the coverage and non-response adjustment under five criteria: "No population," "Small population," "Linearly dependent," "High collinearity" and "Explanatory redundancy." A constraint may have been excluded from calibration for one of the reasons above and yet been calibrated at the end of the process. This occurs for example when the constraint is collinear with the selected constraints. In this case, Appendix C shows that constraint as "Calibrated."

### 4.6.2 Procedures for the final calibration adjustment

The purpose of the final calibration is to adjust the household weights so that the long-form estimates are as close as possible to the census counts for many common characteristics. In addition to producing agreement between the estimates, an appropriate choice of constraints reduces variance. Appendix C provides the complete list of possible constraints, and section 4.5 describes the calibration performed on the long-form estimates. The criteria applied to the selection of constraints are similar to those applied to the selection of constraints for the coverage and non-response adjustment, with a few differences as presented in Table 4.6.3.1.

Calibration was performed simultaneously for SADAs and calibration ADAs. In each SADA, calibration constraints were defined at the SADA and ADA levels. All constraints were evaluated in each SADA and excluded only if necessary. The total number of persons (TOTPERS) and the total number of households (TOTHHLD) were the only two mandatory constraints. This meant that they could not be excluded for any of the ADAs.

Statistics Canada's GES was sent a total of 132,777 preselected constraints at the national level, so that it could perform the final linear calibration. This represents an average of 27 constraints per ADA and an average of 31 constraints per SADA. The mandatory constraints were selected in all the weighting areas. The process selected income constraints the most often, particularly the household income constraint and the low-income household constraint. The constraints selected the least often were primarily the year of immigration and the country of origin.

### 4.6.3 Comparison of the procedures for the two adjustment and selection steps

Criteria were applied in the selection of constraints at each weighting step. These criteria are indicated in Table 4.6.3.1 by weighting step.

Table 4.6.3.1
Criteria applied in selecting coverage, non-response and final calibration adjustment constraints

| Criteria | Adjustment for coverage <br> and non-response | Final calibration |
| :--- | :--- | :--- |
| No population according to <br> the census counts: If the constraint had <br> no population in the weighting area, then <br> the estimate after adjustment must also <br> be 0 for that constraint. These constraints | Applied at the SADA/ADA level. | Applied at the SADA/ADA level. |
| are not classified as excluded but rather |  |  |
| as ineligible to the adjustment process. | Applied at the SADA/ADA level. | Applied at the SADA/ADA level. |
| Small population according to the | The number of households in | The number of households in the |
| census counts: If a constraint involves | Then |  |
| less than a certain number of households | the population is larger than 0 | population is more than 0 but less <br> in the population of the weighting area, <br> then it is considered small and the |
| but less than 250 in the weighted <br> than 200 in the weighted area. <br> constraint is excluded. Including such |  |  |
| a constraint would unduly increase the |  |  |
| variance. However, constraints with small |  |  |
| population can be implicitly calibrated |  |  |
| and in this case are included in the total |  |  |
| number of calibrated constraints. |  |  |

Linearly dependent: If the value of a constraint can be calculated by combining the values of other constraints, one of these constraints is not necessary and must be deleted during the adjustment process because of its linear dependency. However, constraints that are excluded because of their linear dependency are implicitly calibrated. They are therefore included in the total number of calibrated constraints.
High collinearity: If a constraint value can be almost calculated by the combination of other constraint values, then at least one of those constraints must be avoided in the adjustment process. Such a constraint is not perfectly calibrated.

Applied at the SADA level. The selection of constraints can be compared to the selection of explanatory variables in a linear regression. The variance inflation factor ${ }^{1}$ (VIF) and the condition number ${ }^{2}$ are thus used to detect high collinearity.

Applied at the SADA level. The selection of constraints can be compared to the selection of explanatory variables in a linear regression. The variance inflation factor ${ }^{1}$ (VIF) and the condition number ${ }^{2}$ are thus used to detect high collinearity.

Applied at the SADA/ADA level.
Two dependency checks are
conducted to identify linearly
dependent constraints. The
first check is done when the
constraints at the SADA/ADA
level are selected, and the second
check includes all the constraints
chosen at both levels of the
geographic hierarchy (SADAs
and ADAs).
Applied at the SADA/ADA
level. Two linear dependency
checks are conducted to identify
constraints that are close to being
linearly dependent. The first check
is done when the constraints at
the SADA level and the ADA level
are selected, and the second
check includes all the constraints
chosen at both levels of the
hierarchy simultaneously (SADAs
and ADAs).

Table 4.6.3.1
Criteria applied in selecting coverage, non-response and final calibration adjustment constraints

| Criteria | Adjustment for coverage <br> and non-response | Final calibration |
| :--- | :--- | :--- |
| Explanatory redundancy: If a | Applied at the SADA. <br> constraint explains the non-response <br> (almost) as well as other constraint(s) <br> is applied (a form of logistic <br> already selected, then the non-response <br> calibration procedure would fail. This is <br> equession) to test the <br> convergence of the logistic <br> regression. | $\mathrm{N} / \mathrm{A}$ |
| does not add any information about the |  <br> non-response mechanism, beyond what <br> is explained by the already-selected <br> constraints, then it should not be included. |  |

1. The VIF quantifies the increase in variance of regression coefficients due to collinearity.
2. The condition number quantifies the degree to which a matrix is close to singularity.

SADA = Super aggregate dissemination area.
ADA = Aggregate dissemination area .

Appendix C indicates the status of each constraint selected in at least one of the geographic areas once the selection of constraints was carried out at the weighting step. The geography column indicates the geographic level to which the constraint was applied. When a constraint was applied to both geographic levels, the totals reflect both levels with no distinction. A constraint could have been excluded from the calibration process and still have been calibrated. In that case, the constraint was considered to be calibrated.

### 4.6.4 Analysis of calibration during the coverage and non-response adjustment

This section summarizes the number of constraints selected and excluded. The reasons for not selecting constraints are also summarized. Additionally, the section sheds some light on the reasons why some constraints are frequently excluded.

Persons born in certain places tend to be more concentrated in certain parts of the country, to the point that many SADAs have little or no population with a given place of birth. As a result, the constraints for place of birth were often not selected because they had small census counts. Similarly, the constraint involving the French official language (OLN_FR) has little or no population in certain regions of the country. Consequently, it was often excluded.

The constraints most often excluded because of high collinearity were:

- "Persons in the ADA" (PPADA)
- "Children in a census family" (CHILD)
- "Census families with children" (CHILDFAM)
- "One-person households" (HHSIZE1)
- "Female" (FEMALE)
- "Male" (MALE)
- "Census families without children" (NOCLDFAM)
- "Females aged 14 years and younger" (FEMALELT15)
- "Persons in an economic family" (INEFAM)
- "Persons in a household that are not part of an economic family" (NOINEFAM)
- "Males aged 14 years and younger" (MALELT15)
- "Persons aged 10 to 14 years" (AGE14)
- "Five-or-more-person households" (HHSIZEGE5)
- "Persons not in a census family" (NOTINFAM)
- "Persons aged 5 to 9 years" (AGE9)
- "Persons aged 0 to 4 years" (AGE4)
- "Six-or-more-person households" (HHSIZEGE6)
- "Adults in a census family" (ADULTCF)
- "Persons in a couple (spouse, partner)" (COUPLE)
- "Two-person households" (HHSIZE2).

The procedure excluded these constraints automatically, since they could be determined very accurately from a combination of the mandatory constraints, "Households" (TOTHHLD) and "Persons" (TOTPERS), together with other constraints that were selected often, such as some constraints for age, marital status, household size, sex by age and persons in census family. These constraints might also have been explained too well by a combination of the variables selected and of the small constraints.

The actual differences between the census counts and the non-response adjusted estimates are examined in section 5.2.

Table 4.6.4.1 shows the number of times that each reason for dropping or removing a SADA-level constraint occurred. The total number of constraints excluded is the sum of the "Small population," "High collinearity" and "Explanatory redundancy" categories. The "No population" category is not included in the total because it does not actually represent excluded constraints. The average number of constraints excluded per SADA is the total for each category divided by 408, the number of SADAs where coverage and non-response adjustments were done.

Table 4.6.4.1
Summary statistics on SADA-level constraints in 2016 coverage and non-response adjustment

| Constraint | Calibrated | $\begin{array}{r} \text { No } \\ \text { population } \end{array}$ | Excluded |  |  | Total excluded |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Small population | High collinearity | Explanatory redundancy |  |
| Total constraints | 30,328 | 2,865 | 29,742 | 11,550 | 49 | 41,341 |
| Average number of constraints per SADA | 74.3 | 7.0 | 72.9 | 28.3 | 0.1 | 101.3 |

SADA = Super aggregate dissemination area.
Note: Two SADAs were excluded: the Wood Buffalo SADA and a second one that consists of on-reserve households only.
Source: Statistics Canada, 2016 Census long-form sample.

On average, 74.3 SADA-level constraints were calibrated per SADA. An average of 72.9 constraints were discarded per SADA because of small population and 28.3 because of high collinearity.

### 4.6.5 Analysis of the final calibration

Other than cases where the population was nil, each time that a constraint was excluded, the calibration process did not attempt to make census counts and long-form estimates agree for that constraint in that weighting area. The gap between the census count and the long-form estimate was usually larger for a constraint that was excluded frequently than for a constraint that was excluded less often.

Table 4.6.5.1 shows how often a constraint was removed or excluded at the weighting area level according to each criterion. The total number of constraints excluded is equal to the sum of the values for the various removal criteria. The average number of constraints excluded per weighting area is simply equal to the total for the category divided by the number of weighting areas.

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Table 4.6.5.1
Summary statistics on constraint selection status at the weighted area level in the final weight adjustment, 2016

| Constraints | Weighted area | Calibrated | No population | Excluded |  | Total excluded | Number of weighted areas |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Small population | High collinearity |  |  |
|  | ADA | 210,307 | 112,442 | 283,947 | 99,319 | 383,266 | 4,180 |
| Number of constraints | SADA | 39,262 | 2,698 | 27,292 | 10,588 | 37,880 | 408 |
| Average number | ADA | 50.3 | 27.1 | 67.9 | 23.8 | 91.7 | 4,180 |
| of constraints <br> by SADA/ADA | SADA | 96.2 | 6.6 | 66.9 | 26.0 | 92.0 | 408 |

SADA = Super aggregate dissemination area.
ADA = Aggregate dissemination area.
Note: Two SADAs were excluded: the Wood Buffalo SADA and a second one that consists of on-reserve households only.
Source: Statistics Canada, 2016 Census long-form sample.

## 5. Evaluation of the weighting procedures

As described in Chapter 4, the first step in weighting the long-form sample was to assign design weights to households. Weights were assigned differently depending on the collection method of the CU where the household was living. Adjustments were applied only to private households with a design weight greater than 1. These private households were not attached to a collective dwelling in a mail-out or list/leave CU. All of the results presented in this chapter were calculated for this subset of households. The final weight for the other private households therefore corresponds to the design weight and remains at the initial value of 1 .

In short, each household was assigned a design weight that was determined by the long-form sample design. Some adjustments were then necessary to address coverage and total non-response. Non-linear calibration was performed during this adjustment to estimate the parameters for non-response models. After being adjusted for coverage and total non-response, the weights were adjusted further in the final calibration process to produce the final weights. The final weights enabled generally better agreement between the census counts and the long-form estimates.

The next few sections examine the distribution of the weights and, for various characteristics, the discrepancies between the census counts and the sample estimates at various aggregate geographic levels.

### 5.1 Distribution of the weights

Chart 5.1.1 and tables 5.1.1 and 5.1.2 illustrate the distribution of the design weights, the weights adjusted for coverage and non-response, and the weights adjusted in the final calibration. The weights are grouped by 0.5 length intervals, with the exception of the first and last intervals. The chart shows the percentage of times the weights appear in each interval. Nearly all the design weights ranged from 3.75 to 4.25 . This is because of the long-form sample design, in which approximately one in four households received a long-form questionnaire in most areas. The impact on the weights of the coverage and non-response adjustments and the final calibration adjustments can also be seen. A very noticeable difference is that the percentage of households that had a design weight from 3.75 to 4.25 was significantly reduced from $100 \%$ to $67 \%$, ending at close to $30 \%$. The final weights were more evenly distributed within categories 1 to 20 .

Logically, the non-response adjustment process should tend to increase the weights to compensate for the nonresponding units. This did occur for most cases. The changes between the design weights and the coverage and non-response adjusted weights can be observed in Table 5.1.1. This table shows that most of the units that left the $[3.75,4.25)$ range moved to the $[4.25,4.75)$ or $[4.75,5.25)$ ranges. However, the coverage and non-response adjustment process moved the weights from the $3.75+$ range to the $[1,2.75)$, $[2.75,3.25$ ) or $[3.25,3.75$ ) ranges for a few units. The main reason is that the procedure included an adjustment for overcoverage and undercoverage. To the extent that a few population groups may have experienced overcoverage, the weights would have been reduced in those areas.

An important element to notice in Table 5.1.1 is that the non-responding units originally had positive weights, since they were selected for the sample. The non-response adjusting process assigned them a weight of 0 and redistributed their original weights among responding units. The non-responding units correspond to the line labelled "0 (non-respondents)" in Table 5.1.1 and were removed from Table 5.1.2, since they were not used in the calibration process. Table 5.1.2 presents the changes between the coverage and non-response adjusted weights and the calibrated weights.

According to Table 5.1.2, most weights experienced only a small modification during the calibration process. In fact, $76.6 \%$ of cases either stayed in the same range or moved only one range up or down. The most stable range was 1 to 2.75 , where $64 \%$ of the households with a non-response adjusted weight between 1 and 2.75 stayed in that category after calibration. The second most stable category was $5.75+$, where $61.8 \%$ of households with a non-response adjusted weight between 5.75 and 9.5 stayed in that category (although the calibration range goes up to 20 rather than 9.5).

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Finally, whereas the non-response adjusted weights varied from 1 to 9.5 , the range of the final weights was from 1 to 20. This variability in the weights is far smaller than the variability in the final weights in the 2011 National Household Survey (see Statistics Canada, 2015), where the maximum weight was capped at 100.

Chart 5.1.1
Distribution of design weights, coverage and non-response adjusted weights, and final weights


Notes: All households with a design weight of 1 were excluded from the weighting process. These households either were located in an Indian reserve or a canvasser enumeration collection unit or were private households attached to a collective dwelling.
The "[" symbol means the number is included in the interval and the ")" symbol means it is not included in the interval.
Source: Statistics Canada, 2016 Census long-form sample.

Table 5.1.1
Distribution of design weights and coverage and non-response adjusted weights

| Coverage and non-response <br> adjusted weights | Design weights |  |  |
| :--- | ---: | ---: | ---: |
| $\left[\begin{array}{ll}{[3.75,4.25)} & {[4.25,7.00]}\end{array}\right.$ | Total |  |  |
| 0 (non-respondents) | 122,807 | 19 | $\mathbf{1 2 2 , 8 2 6}$ |
| $1.00,2.75)$ | 1,548 | 0 | $\mathbf{1 , 5 4 8}$ |
| $[2.75,3.25)$ | 12,363 | 0 | $\mathbf{1 2 , 3 6 3}$ |
| $[3.25,3.75)$ | 201,534 | 2 | $\mathbf{2 0 1 , 5 3 6}$ |
| $[3.75,4.25)$ | $1,921,456$ | 60 | $\mathbf{1 , 9 2 1 , 5 1 6}$ |
| $4.25,4.75)$ | $1,082,895$ | 239 | $\mathbf{1 , 0 8 3 , 1 3 4}$ |
| $[4.75,5.25)$ | 103,892 | 85 | $\mathbf{1 0 3 , 9 7 7}$ |
| $[5.25,5.75)$ | 12,783 | 14 | $\mathbf{1 2 , 7 9 7}$ |
| $[5.75,20.00]$ | 3,730 | 5 | $\mathbf{3 , 7 3 5}$ |
| Total | $\mathbf{3 , 4 6 3 , 0 0 8}$ | $\mathbf{4 2 4}$ | $\mathbf{3 , 4 6 3 , \mathbf { 4 3 2 }}$ |

Notes: All households with a design weight of one are excluded from the weighting process. These households either come from Indian reserve and canvasser enumeration CUs or are private households attached to collective.
The "[" symbol means the number is included in the interval and the ")" symbol means it is not included in the interval.
Source: Statistics Canada, 2016 Census long-form sample.

Table 5.1.2
Distribution of coverage and non-response adjusted weights and final weights

| Coverage and non-response adjusted weights |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Final weights | $\begin{array}{r} {[1.00,} \\ 2.75) \end{array}$ | $\begin{array}{r} {[2.75,} \\ 3.25) \end{array}$ | $\begin{array}{r} {[3.25,} \\ 3.75) \end{array}$ | $\begin{array}{r} {[3.75,} \\ 4.25) \end{array}$ | $\begin{array}{r} {[4.25,} \\ 4.75) \end{array}$ | $\begin{array}{r} {[4.75,} \\ 5.25) \end{array}$ | $\begin{array}{r} {[5.25,} \\ 5.75) \end{array}$ | $\begin{array}{r} {[5.75,} \\ 9.50) \end{array}$ | Total |
| [1.00, 2.75) | 990 | 3,750 | 18,082 | 42,843 | 12,597 | 1,402 | 196 | 71 | 79,931 |
| [2.75, 3.25) | 340 | 3,805 | 38,693 | 138,180 | 33,513 | 2,196 | 205 | 37 | 216,969 |
| [3.25, 3.75) | 142 | 2,910 | 63,351 | 414,026 | 115,038 | 5,456 | 454 | 86 | 601,463 |
| [3.75, 4.25) | 42 | 1,266 | 50,385 | 649,869 | 275,758 | 12,551 | 878 | 135 | 990,884 |
| [4.25, 4.75) | 27 | 413 | 21,470 | 444,950 | 334,973 | 22,686 | 1,522 | 219 | 826,260 |
| [4.75, 5.25) | 7 | 151 | 6,648 | 165,518 | 196,585 | 25,532 | 2,354 | 364 | 397,159 |
| [5.25, 5.75) | 0 | 48 | 1,986 | 47,211 | 75,313 | 17,880 | 2,601 | 514 | 145,553 |
| [5.75, 20.00] | 0 | 20 | 921 | 18,919 | 39,357 | 16,274 | 4,587 | 2,309 | 82,387 |
| Total | 1,548 | 12,363 | 201,536 | 1,921,516 | 1,083,134 | 103,977 | 12,797 | 3,735 | 3,340,606 |

Notes: All households with a design weight of one are excluded from the weighting process. These households either come from Indian reserve and canvasser enumeration CUs or are private households attached to collective.
The "[" symbol means the number is included in the interval and the ")" symbol means it is not included in the interval.
Source: Statistics Canada, 2016 Census long-form sample.

### 5.2 Discrepancies between census counts and long-form estimates, Canada

Chapter 4 describes the methods used to calculate the final household weights, and section 5.1 shows some of the relationships between design weights, non-response adjusted weights and final weights. The coverage and non-response adjustment reduced the discrepancies between the census counts and the corresponding longform estimates for the constraints considered (see Appendix C). Following those adjustments, calibration further reduced or eliminated those discrepancies for certain variables (constraints). However, some discrepancies remain, since constraints are sometimes excluded discarded (see section 4.6). The relative difference between census counts and long-form estimates also called discrepancy is defined as:

$$
\text { Discrepancy }=\frac{\text { (long-form estimate }- \text { census count) }}{\text { census count }} \times 100 \%
$$

The numerator in the above expression (long-form estimate - census count) is referred to as the difference between census counts and long-form estimates. By dividing this value by the census count, the difference between census counts and long-form estimates relative to the size of the census count can be seen. In other words, the ratio represents the percentage that the characteristic was overestimated (a positive value) or underestimated (a negative value).
Table 5.2 .1 shows the 2016 Canada-level differences between census counts and long-form estimates for the constraints considered for the design weights, the coverage and non-response adjusted weights and the final weights. The census count for the total persons characteristic (TOTPERS) is lower than the published figure of the 2016 Census $(35,151,728)$. This difference is the result of the weighting process, which used only private households (not those living in dwellings attached to collective dwellings) and excluded Indian reserve and canvasser enumeration CUs to use a common population in the comparison.

Table 5.2.1 also shows percentage relative differences (called discrepancies) for estimates based on final weights. Looking at these discrepancies sheds more light on the differences. Over $88 \%$ of the cases in Table 5.2.1 had a discrepancy from $-1 \%$ to $1 \%$, and over $97 \%$ of them from $-5 \%$ to $5 \%$.

Chart 5.2.2 shows, for all the constraints, the difference between the census counts and each of the three estimates: design-weighted (red), non-response adjusted (purple) and final (blue). The x-axis represents the population size of the constraint, and the three lines show for each constraint:

- the difference between the sum of the design weights and the census count
- the sum of the non-response adjusted weights and the census count
- the sum of the final weight and the census count.

The constraints are sorted, from left to right, by increasing population size.
Chart 5.2.3 shows the percentage discrepancies between census counts and final estimates for all the constraints by population size. For the medium-sized and large constraints, the discrepancies are all small. Only certain small constraints have relatively large discrepancies. For example, the difference between the census count and the design weight estimate for the "Single person, never married" (COMMONLAWNO_SINGLE) constraint, with a population size of $13,624,243$, is about $-171,000$ (red). In contrast, the non-response adjustment reduces that difference to only about $-6,000$ (purple), and the final calibration further reduces the difference to about $-2,000$ (blue).

The most important observation from Chart 5.2.2 is that the coverage and non-response adjustment carries a big improvement over the design-weighted estimates. Although it is not apparent in the chart, the coverage adjustment does most of the job. The difference between census counts and long-form estimates for design weights tended to be (much) greater than the difference between census counts and long-form estimates for the non-response adjusted weights. This, in turn, tended to be greater than the corresponding difference using the final weights. This shows the importance of the non-response adjustment and calibration processes. As mentioned in section 4.6, a difference between the census count and long-form estimate could occur in a SADA or ADA for a characteristic if its constraint is excluded during calibration. In other words, the process did not control on the excluded constraint for a given area. If the constraint is excluded in many areas, these differences could partially cancel each other out, or they could cumulate to create a large difference at the Canada level. Total persons (TOTPERS) and total households (TOTHHLD) were the only mandatory constraints for which agreement between census counts and long-form estimates had to be guaranteed for all ADAs. As a result, the final weight difference and discrepancy for these characteristics were 0 . However, all other constraints had to be excluded in some areas.

Section 4.6 pointed out the constraints that were excluded frequently and where high differences or discrepancies might lie. The effect of dropping those constraints can be seen in Table 5.2.1. The SADA-only constraints that were most often excluded ${ }^{5}$ were:

- "Persons in an economic family" (INEFAM)
- "Persons in a household that are not part of an economic family" (NOINEFAM)
- "Persons in a couple (spouse, partner)" (COUPLE)
- "Persons in a two-or-more-person household that are not part of an economic family" (NOINEFAMHHSIZEGT1)
- "Adults in a census family" (ADULTCF)
- "Persons not in a census family" (NOTINFAM)
- "Children in a census family" (CHILD)
- "Persons in a two-or-more-person household that are not part of a census family" (NOTINFAMHHSIZEGT1)
- "Lone parents" (LONEPAR).

[^5]Not surprisingly, the largest differences (more than 2,000 in absolute value) were obtained precisely for these constraints. The largest difference was obtained for "Persons in an economic family" (INEFAM) $(14,493)$ and for its complement, "Persons in a household that are not part of an economic family" (NOINEFAM). However, because the census counts were so high, the percentage discrepancies for all these constraints were not too high (ranging from $-0.93 \%$ to $0.06 \%$ ). In the 2011 National Household Survey, the maximum difference was 166,801 (see Statistics Canada, 2015).

The largest percentage discrepancies were obtained for some of the places of birth. Since many places of birth are not found in high numbers in many parts of the country, these variables were often excluded during calibration. This led to some differences in the census counts and long-form estimates. Furthermore, because many places of birth have relatively low census counts, the relative differences are magnified. This resulted in greater percentage discrepancies. The largest percentage discrepancy, $12.87 \%$, was for the POBG2_49 constraint, which corresponds to "Mongolia." This constraint was rarely selected because it had either no population or a small population in the area. Other places of birth with high discrepancies, all of which were also rarely selected, were "Botswana, Lesotho and Swaziland" (POBG2_9) with 10.53\%; "Canada" (POBG2_13) with $8.56 \%$; "Malta" (POBG2_44) with 6.22\%; "Guadeloupe, Martinique and Saint Martin (French part)" (POBG2_15) with 6.1\%; "Andorra, Gibraltar and Spain" (POBG2_2) with $4.75 \%$; and "Burma (Myanmar)" (POBG2_12) with 4.53\%.

All the household and family characteristics had small percent discrepancies, with the largest being an underestimation of $0.34 \%$, for households of size 6 or more (HHSIZEGE6). The largest difference for this type of characteristics was observed for households of size 1, with an underestimation of 2,025 households. Other household characteristics with relatively large differences ${ }^{6}$ (but still small discrepancies) are "Two-person households" (HHSIZE2), "Census families without children" (NOCLDFAM), "Census families" (TOTCFAM), "Six-or-more-person households" (HHSIZEGE6) and "Five-or-more-person households" (HHSIZEGE5). The closest estimate for a non-mandatory constraint of this type belonged to "Households living in a single-detached house" (SNGLDET), with a difference of 56 and a discrepancy of $0.001 \%$.

Other than places of birth, the person-level characteristics with the largest discrepancies were the constraints "Persons in a two-or-more-person household that are not part of an economic family" (NOINEFAMHHSIZEGT1), with a discrepancy of $-0.93 \%$, and "Persons in a two-or-more-person household that are not part of a census family" (NOTINFAMHHSIZEGT1), with a discrepancy of $-0.32 \%$. Other than the ones mentioned above, the person-level characteristics with an overestimation or underestimation of at least 1,000 persons were:

- "Single persons (never married) not in a common law relationship" (COMLAWNO_SINGLE)
- "Single persons (never married) under 15 years and not in a common law relationship" (COMLAWNO_ SINGLE_LT15)
- "Persons in a common law relationship" (COMLAW_YE)
- "Married persons" (MARRIED)
- "Place of birth - Belarus, Moldova, Russian Federation and Ukraine" (POBG2_28).

However, the discrepancies for all of these were small, the largest being $0.74 \%$, for the last.

[^6]
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The closest estimates for non-mandatory person-level constraints belonged to the Age 35 to 44 (AGE35_44) and Age 45 to 54 (AGE45_54) groups and to persons whose employment income is above the 75 th percentile in the SADA (EMPIN_SADA_P100). These characteristics were estimated without sampling error at the Canada level. Other person-level constraints with very small differences and discrepancies ${ }^{7}$ included:

- "Persons aged 55 to 59 years" (AGE59)
- "Persons with an annual employment income above the 50th percentile for the SADA" (EMPIN_SADA_ GT50)
- "Persons with an annual employment income above the 50th percentile and equal to or below the 75th percentile for the SADA" (EMPIN_SADA_P75)
- "Persons aged 20 to 24 years" (AGE24)
- "Persons with an annual employment income equal to or below the 50th percentile for the SADA" (EMPIN_SADA_LE50)
- "Persons aged 15 to 29 years" (AGE15_29)
- "Persons aged 50 to 64 years" (AGE50_64).

Finally, among the SADA-only constraints, the largest difference (-640) and percentage discrepancy ( $-0.02 \%$ ) was obtained for the constraint "Households with an annual income at or below the 25th percentile for the SADA" (HHINC_SADA_P25). All these constraints were selected relatively often. The constraint that was selected the least often, "Persons with an annual employment income above the 25 th percentile and equal to or below the 50th percentile for the SADA" (EMPIN_SADA_P50), was selected in $62.7 \%$ of the ADAs (as EMPIN_P50).

Table 5.2.1
Census counts and long-form estimate differences and discrepancies - Canada

| Characteristic | Census counts | Design weights |  | Coverage and non-response adjusted weights |  | Final weights |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | estimates | difference | estimates | difference | estimates | difference | discrepancy <br> (\%) |
| ADULTCF | 17,873,064 | 17,796,322 | -76,742 | 17,880,659 | 7,595 | 17,879,104 | 6,040 | 0.03 |
| AGE00_14 | 5,679,559 | 5,629,183 | -50,376 | 5,684,239 | 4,680 | 5,679,522 | -37 | 0.00 |
| AGE4 | 1,845,178 | 1,828,386 | -16,792 | 1,848,647 | 3,469 | 1,845,517 | 339 | 0.02 |
| AGE9 | 1,962,324 | 1,945,126 | -17,198 | 1,962,857 | 533 | 1,962,244 | -80 | 0.00 |
| AGE14 | 1,872,057 | 1,855,671 | -16,386 | 1,872,734 | 677 | 1,871,761 | -296 | -0.02 |
| AGE15_24 | 4,155,227 | 4,098,212 | -57,015 | 4,152,363 | -2,864 | 4,155,209 | -18 | 0.00 |
| AGE15_29 | 6,387,808 | 6,297,713 | -90,095 | 6,383,708 | -4,100 | 6,387,805 | -3 | 0.00 |
| AGE19 | 1,970,121 | 1,949,762 | -20,359 | 1,970,676 | 555 | 1,970,110 | -11 | 0.00 |
| AGE24 | 2,185,106 | 2,148,450 | -36,656 | 2,181,687 | -3,419 | 2,185,099 | -7 | 0.00 |
| AGE25_34 | 4,513,475 | 4,453,702 | -59,773 | 4,512,230 | -1,245 | 4,513,547 | 72 | 0.00 |
| AGE29 | 2,232,581 | 2,199,501 | -33,080 | 2,231,345 | -1,236 | 2,232,597 | 16 | 0.00 |
| AGE30_49 | 9,047,387 | 8,960,912 | -86,475 | 9,045,676 | -1,711 | 9,047,405 | 18 | 0.00 |
| AGE34 | 2,280,894 | 2,254,201 | -26,693 | 2,280,885 | -9 | 2,280,950 | 56 | 0.00 |
| AGE35_44 | 4,454,400 | 4,412,465 | -41,935 | 4,452,867 | -1,533 | 4,454,400 | 0 | 0.00 |
| AGE39 | 2,243,631 | 2,223,020 | -20,611 | 2,242,432 | -1,199 | 2,243,735 | 104 | 0.00 |
| AGE44 | 2,210,769 | 2,189,445 | -21,324 | 2,210,436 | -333 | 2,210,665 | -104 | 0.00 |

[^7]
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Table 5.2.1
Census counts and long-form estimate differences and discrepancies - Canada

| Characteristic | Census counts | Design weights |  | Coverage and non-response adjusted weights |  | Final weights |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | estimates | difference | estimates | difference | estimates | difference | discrepancy (\%) |
| AGE45_54 | 4,936,819 | 4,905,005 | -31,814 | 4,936,661 | -158 | 4,936,819 | 0 | 0.00 |
| AGE49 | 2,312,093 | 2,294,247 | -17,846 | 2,311,923 | -170 | 2,312,055 | -38 | 0.00 |
| AGE50_64 | 7,436,076 | 7,406,957 | -29,120 | 7,435,705 | -371 | 7,436,084 | 8 | 0.00 |
| AGE54 | 2,624,726 | 2,610,758 | -13,968 | 2,624,738 | 12 | 2,624,764 | 38 | 0.00 |
| AGE55_64 | 4,811,350 | 4,796,199 | -15,151 | 4,810,967 | -383 | 4,811,321 | -29 | 0.00 |
| AGE59 | 2,568,136 | 2,559,441 | -8,695 | 2,569,510 | 1,374 | 2,568,127 | -9 | 0.00 |
| AGE64 | 2,243,214 | 2,236,757 | -6,457 | 2,241,457 | -1,757 | 2,243,194 | -20 | 0.00 |
| AGE65PL | 5,443,624 | 5,445,343 | 1,719 | 5,444,038 | 414 | 5,443,637 | 13 | 0.00 |
| AGE74 | 3,302,938 | 3,302,147 | -791 | 3,303,259 | 321 | 3,302,967 | 29 | 0.00 |
| AGE75PL | 2,140,686 | 2,143,195 | 2,509 | 2,140,779 | 93 | 2,140,670 | -16 | 0.00 |
| APTLT5 | 2,533,749 | 2,515,260 | -18,489 | 2,532,535 | -1,214 | 2,533,675 | -74 | 0.00 |
| CHILD | 10,082,408 | 9,998,775 | -83,633 | 10,102,364 | 19,956 | 10,085,159 | 2,751 | 0.03 |
| CHILDFAM | 5,722,912 | 5,681,084 | -41,828 | 5,728,040 | 5,128 | 5,722,653 | -259 | 0.00 |
| COMLAW_YE | 3,440,562 | 3,416,138 | -24,424 | 3,447,069 | 6,507 | 3,442,087 | 1,525 | 0.04 |
| COMLAWNO_DIV | 1,749,631 | 1,741,004 | -8,627 | 1,749,867 | 236 | 1,749,963 | 332 | 0.02 |
| COMLAWNO_ OTHERS | 3,793,163 | 3,778,466 | -14,697 | 3,792,762 | -401 | 3,793,576 | 413 | 0.01 |
| COMLAWNO_SEP | 688,126 | 685,664 | -2,462 | 687,481 | -645 | 688,056 | -70 | -0.01 |
| COMLAWNO_ SINGLE | 13,624,243 | 13,453,334 | -170,909 | 13,618,465 | -5,778 | 13,622,342 | -1,901 | -0.01 |
| COMLAWNO SINGLE_GE15 | 7,944,684 | 7,824,151 | -120,533 | 7,934,226 | -10,458 | 7,942,819 | -1,865 | -0.02 |
| COMLAWNO_WID | 1,355,406 | 1,351,797 | -3,609 | 1,355,414 | 8 | 1,355,557 | 151 | 0.01 |
| COMLAWYE_ MARRIED | 16,577,048 | 16,508,308 | -68,740 | 16,582,139 | 5,091 | 16,578,536 | 1,488 | 0.01 |
| COUPLE | 16,304,196 | 16,244,300 | -59,896 | 16,317,065 | 12,869 | 16,313,204 | 9,008 | 0.06 |
| $\begin{aligned} & \text { EMPIN_SADA_ } \\ & \text { GT50 } \end{aligned}$ | 10,115,337 | 10,071,665 | -43,672 | 10,115,480 | 143 | 10,115,329 | -8 | 0.00 |
| $\begin{aligned} & \text { EMPIN_SADA_ } \\ & \text { LE50 } \end{aligned}$ | 10,116,607 | 10,006,750 | -109,857 | 10,115,995 | -612 | 10,116,600 | -7 | 0.00 |
| EMPIN_SADA_P0 | 13,762,510 | 13,661,692 | -100,818 | 13,761,890 | -620 | 13,762,525 | 15 | 0.00 |
| $\begin{aligned} & \text { EMPIN_SADA_PO_ } \\ & \text { GE15 } \end{aligned}$ | 8,082,951 | 8,032,509 | -50,442 | 8,077,651 | -5,300 | 8,083,003 | 52 | 0.00 |
| EMPIN_SADA_P25 | 5,058,711 | 5,008,873 | -49,838 | 5,058,197 | -514 | 5,059,199 | 488 | 0.01 |
| EMPIN_SADA_P50 | 5,057,896 | 4,997,877 | -60,019 | 5,057,797 | -99 | 5,057,401 | -495 | -0.01 |
| EMPIN_SADA_P75 | 5,057,676 | 5,030,940 | -26,736 | 5,058,134 | 458 | 5,057,668 | -8 | 0.00 |
| $\begin{aligned} & \text { EMPIN_SADA_ } \\ & \text { P100 } \end{aligned}$ | 5,057,661 | 5,040,726 | -16,935 | 5,057,347 | -314 | 5,057,661 | 0 | 0.00 |
| FEMALE | 17,258,035 | 17,153,083 | -104,952 | 17,262,580 | 4,545 | 17,258,007 | -28 | 0.00 |
| FEMALEGE15 | 14,489,305 | 14,407,839 | -81,466 | 14,489,472 | 167 | 14,489,532 | 227 | 0.00 |
| FEMALELT15 | 2,768,730 | 2,745,244 | -23,486 | 2,773,108 | 4,378 | 2,768,475 | -255 | -0.01 |
| $\begin{aligned} & \text { HHADACSD } \\ & \text { (CSDxADA) } \end{aligned}$ | 13,932,820 | 13,372,132 | -560,688 | 13,932,820 | 0 | 13,932,820 | 0 | 0.00 |
| $\begin{aligned} & \text { HHINC_SADA_ } \\ & \text { GT50 } \end{aligned}$ | 6,966,238 | 6,936,358 | -29,880 | 6,966,557 | 319 | 6,966,487 | 249 | 0.00 |

Table 5.2.1
Census counts and long-form estimate differences and discrepancies - Canada

| Characteristic | Census counts | Design weights |  | Coverage and non-response adjusted weights |  | Final weights |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | estimates | difference | estimates | difference | estimates | difference | discrepancy (\%) |
| HHINC_SADA_ |  |  |  |  |  |  |  |  |
| LE50 | 6,966,582 | 6,927,519 | -39,064 | 6,966,263 | -319 | 6,966,333 | -249 | 0.00 |
| HHINC_SADA_P25 | 3,483,439 | 3,458,080 | -25,359 | 3,482,595 | -844 | 3,482,799 | -640 | -0.02 |
| HHINC_SADA_P50 | 3,483,143 | 3,469,439 | -13,704 | 3,483,667 | 524 | 3,483,534 | 391 | 0.01 |
| HHINC_SADA_P75 | 3,483,232 | 3,472,828 | -10,404 | 3,483,857 | 625 | 3,483,764 | 532 | 0.02 |
| HHINC_SADA_- |  |  |  |  |  |  |  |  |
| P100 | 3,483,006 | 3,463,530 | -19,476 | 3,482,700 | -306 | 3,482,723 | -283 | -0.01 |
| HHSIZE1 | 3,939,604 | 3,932,270 | -7,334 | 3,938,772 | -832 | 3,937,579 | -2,025 | -0.05 |
| HHSIZE2 | 4,801,532 | 4,788,730 | -12,802 | 4,802,550 | 1,018 | 4,803,338 | 1,806 | 0.04 |
| HHSIZE3 | 2,119,446 | 2,103,306 | -16,140 | 2,119,554 | 108 | 2,120,058 | 612 | 0.03 |
| HHSIZE4 | 1,927,188 | 1,915,750 | -11,438 | 1,928,438 | 1,250 | 1,928,027 | 839 | 0.04 |
| HHSIZE5 | 739,505 | 732,595 | -6,910 | 742,931 | 3,426 | 739,636 | 131 | 0.02 |
| HHSIZEGE5 | 1,145,050 | 1,123,820 | -21,230 | 1,143,506 | -1,544 | 1,143,819 | -1,231 | -0.11 |
| HHSIZEGE6 | 405,545 | 391,225 | -14,320 | 400,575 | -4,970 | 404,183 | -1,362 | -0.34 |
| INEFAM | 28,720,122 | 28,545,611 | -174,511 | 28,750,644 | 30,522 | 28,734,615 | 14,493 | 0.05 |
| IR_LINK_NO | 33,637,976 | 33,390,492 | -247,484 | 33,637,564 | -412 | 33,638,859 | 883 | 0.00 |
| IR_LINK_YE | 356,478 | 349,615 | -6,863 | 355,801 | -677 | 355,595 | -883 | -0.25 |
| LIM_NO | 29,187,451 | 28,997,962 | -189,489 | 29,188,145 | 694 | 29,188,235 | 784 | 0.00 |
| LIM_YE | 4,807,003 | 4,742,145 | -64,858 | 4,805,220 | -1,783 | 4,806,219 | -784 | -0.02 |
| LONEPAR | 1,568,868 | 1,552,022 | -16,846 | 1,563,594 | -5,274 | 1,565,900 | -2,968 | -0.19 |
| MALE | 16,736,419 | 16,587,024 | -149,395 | 16,730,785 | -5,634 | 16,736,447 | 28 | 0.00 |
| MALEGE15 | 13,825,590 | 13,703,085 | -122,505 | 13,819,654 | -5,936 | 13,825,399 | -191 | 0.00 |
| MALELT15 | 2,910,829 | 2,883,939 | -26,890 | 2,911,130 | 301 | 2,911,048 | 219 | 0.01 |
| MARRIED | 13,136,486 | 13,092,169 | -44,317 | 13,135,070 | -1,416 | 13,136,449 | -37 | 0.00 |
| NOCLDFAM | 3,998,054 | 3,993,087 | -4,967 | 3,994,087 | -3,967 | 3,999,849 | 1,795 | 0.04 |
| NOINEFAM | 5,274,332 | 5,194,497 | -79,835 | 5,242,721 | -31,611 | 5,259,839 | -14,493 | -0.27 |
| NOINEFAMHHSIZEGT1 | 1,334,728 | 1,262,226 | -72,502 | 1,303,949 | -30,779 | 1,322,260 | -12,468 | -0.93 |
| NOTINFAM | 6,038,982 | 5,945,010 | -93,972 | 6,010,343 | -28,639 | 6,030,190 | -8,792 | -0.15 |
| NOTINFAMHHSIZEGT1 | 2,099,378 | 2,012,740 | -86,638 | 2,071,570 | -27,808 | 2,092,611 | -6,767 | -0.32 |
| OLN_BI | 6,187,606 | 6,160,319 | -27,287 | 6,187,142 | -464 | 6,187,821 | 215 | 0.00 |
| OLN_EN | 23,157,389 | 22,959,992 | -197,397 | 23,156,234 | -1,155 | 23,157,275 | -114 | 0.00 |
| OLN_FR | 4,015,036 | 3,989,439 | -25,597 | 4,015,372 | 336 | 4,014,760 | -276 | -0.01 |
| OLN_NO | 634,423 | 630,357 | -4,066 | 634,618 | 195 | 634,598 | 175 | 0.03 |
| POBG2_1 | 14,784 | 14,945 | 161 | 14,798 | 14 | 14,841 | 57 | 0.38 |
| POBG2_2 | 7,790 | 7,381 | -409 | 7,390 | -400 | 7,420 | -370 | -4.75 |
| POBG2_3 | 27,825 | 27,490 | -335 | 27,315 | -510 | 27,564 | -261 | -0.94 |
| POBG2_4 | 61,352 | 61,578 | 226 | 61,928 | 576 | 61,993 | 641 | 1.04 |
| POBG2_5 | 6,443 | 6,217 | -226 | 6,259 | -184 | 6,238 | -205 | -3.18 |

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Table 5.2.1
Census counts and long-form estimate differences and discrepancies - Canada

| Characteristic | Census counts | Design weights |  | Coverage and non-response adjusted weights |  | Final weights |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | estimates | difference | estimates | difference | estimates | difference | discrepancy (\%) |
| POBG2_6 | 34,071 | 33,889 | -182 | 33,847 | -224 | 33,734 | -337 | -0.99 |
| POBG2_7 | 68,871 | 68,511 | -360 | 69,240 | 369 | 68,917 | 46 | 0.07 |
| POBG2_8 | 4,069 | 4,012 | -57 | 4,142 | 73 | 4,114 | 45 | 1.10 |
| POBG2_9 | 1,360 | 1,452 | 92 | 1,516 | 156 | 1,503 | 143 | 10.53 |
| POBG2_10 | 32,812 | 32,871 | 59 | 32,625 | -187 | 32,781 | -31 | -0.10 |
| POBG2_11 | 100,655 | 100,325 | -330 | 100,502 | -153 | 100,519 | -136 | -0.14 |
| POBG2_12 | 7,474 | 7,156 | -318 | 7,298 | -176 | 7,135 | -339 | -4.53 |
| POBG2_13 | 2,230 | 2,034 | -196 | 2,032 | -198 | 2,039 | -191 | -8.56 |
| POBG2_14 | 1,694 | 1,785 | 91 | 1,779 | 85 | 1,754 | 60 | 3.57 |
| POBG2_15 | 1,405 | 1,508 | 103 | 1,505 | 100 | 1,491 | 86 | 6.10 |
| POBG2_16 | 27,097 | 26,029 | -1,068 | 26,732 | -365 | 27,016 | -81 | -0.30 |
| POBG2_17 | 22,097 | 22,521 | 424 | 22,907 | 810 | 22,708 | 611 | 2.77 |
| POBG2_18 | 2,909 | 2,876 | -33 | 2,896 | -13 | 2,892 | -17 | -0.58 |
| POBG2_19 | 22,252 | 22,054 | -198 | 22,189 | -63 | 22,073 | -179 | -0.80 |
| POBG2_20 | 24,130 | 24,675 | 545 | 24,381 | 251 | 24,165 | 35 | 0.15 |
| POBG2_21 | 873,285 | 868,958 | -4,327 | 873,080 | -205 | 873,254 | -31 | 0.00 |
| POBG2_22 | 109,015 | 108,270 | -745 | 108,163 | -852 | 108,570 | -445 | -0.41 |
| POBG2_23 | 15,319 | 14,917 | -402 | 15,094 | -225 | 15,262 | -57 | -0.37 |
| POBG2_24 | 42,218 | 42,067 | -151 | 42,520 | 302 | 42,632 | 414 | 0.98 |
| POBG2_25 | 14,794 | 13,853 | -941 | 14,093 | -701 | 14,366 | -428 | -2.90 |
| POBG2_26 | 65,533 | 66,428 | 895 | 65,646 | 113 | 65,666 | 133 | 0.20 |
| POBG2_27 | 56,296 | 56,892 | 596 | 56,841 | 545 | 56,950 | 654 | 1.16 |
| POBG2_28 | 139,455 | 140,786 | 1,331 | 140,543 | 1,088 | 140,480 | 1,025 | 0.74 |
| POBG2_29 | 70,138 | 69,530 | -608 | 70,139 | 1 | 69,877 | -261 | -0.37 |
| POBG2_30 | 33,952 | 34,553 | 601 | 33,707 | -245 | 33,709 | -243 | -0.72 |
| POBG2_31 | 103,283 | 102,569 | -714 | 103,462 | 179 | 102,855 | -428 | -0.41 |
| POBG2_32 | 186,698 | 186,575 | -123 | 187,555 | 857 | 187,251 | 553 | 0.30 |
| POBG2_33 | 112,041 | 110,677 | -1,364 | 111,862 | -179 | 111,726 | -315 | -0.28 |
| POBG2_34 | 19,235 | 19,630 | 395 | 19,665 | 430 | 19,480 | 245 | 1.27 |
| POBG2_35 | 72,931 | 72,295 | -636 | 73,132 | 201 | 72,782 | -149 | -0.20 |
| POBG2_36 | 74,350 | 73,931 | -419 | 74,129 | -221 | 74,570 | 220 | 0.30 |
| POBG2_37 | 112,543 | 111,512 | -1,031 | 112,225 | -318 | 112,106 | -437 | -0.39 |
| POBG2_38 | 86,917 | 87,053 | 136 | 87,116 | 199 | 87,155 | 238 | 0.27 |
| POBG2_39 | 28,117 | 27,361 | -756 | 27,542 | -575 | 27,852 | -265 | -0.94 |
| POBG2_40 | 124,170 | 123,731 | -439 | 124,693 | 523 | 124,738 | 568 | 0.46 |
| POBG2_41 | 2,234 | 2,204 | -30 | 2,240 | 6 | 2,258 | 24 | 1.08 |
| POBG2_42 | 154,606 | 153,966 | -640 | 155,098 | 492 | 154,564 | -42 | -0.03 |
| POBG2_43 | 609,741 | 607,631 | -2,110 | 609,653 | -88 | 609,822 | 81 | 0.01 |
| POBG2_44 | 2,761 | 3,017 | 256 | 2,965 | 204 | 2,933 | 172 | 6.22 |
| POBG2_45 | 60,546 | 60,919 | 373 | 60,721 | 175 | 60,639 | 93 | 0.15 |
| POBG2_46 | 51,880 | 51,081 | -799 | 51,305 | -575 | 51,651 | -229 | -0.44 |

Table 5.2.1
Census counts and long-form estimate differences and discrepancies - Canada

| Characteristic | Census counts | Design weights |  | Coverage and non-response adjusted weights |  | Final weights |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | estimates | difference | estimates | difference | estimates | difference | discrepancy <br> (\%) |
| POBG2_47 | 125,745 | 124,733 | -1,012 | 125,786 | 41 | 125,600 | -145 | -0.12 |
| POBG2_48 | 348,474 | 348,730 | 256 | 348,747 | 273 | 348,490 | 16 | 0.00 |
| POBG2_49 | 1,342 | 1,500 | 158 | 1,466 | 124 | 1,515 | 173 | 12.87 |
| POBG2_50 | 1,044 | 1,064 | 20 | 1,068 | 24 | 1,078 | 34 | 3.23 |
| POBG2_51 | 14,222 | 13,744 | -478 | 13,803 | -419 | 13,902 | -320 | -2.25 |
| POBG2_53 | 266,119 | 265,374 | -745 | 264,686 | -1,433 | 265,339 | -780 | -0.29 |
| POBG2_54 | 115,745 | 114,997 | -748 | 115,501 | -244 | 115,690 | -55 | -0.05 |
| POBG2_55 | 20,529 | 20,181 | -348 | 20,361 | -168 | 20,472 | -57 | -0.28 |
| POBG2_56 | 74,226 | 74,585 | 359 | 74,351 | 125 | 74,304 | 78 | 0.10 |
| POBG2_57 | 27,373 | 28,272 | 899 | 28,058 | 685 | 27,813 | 440 | 1.61 |
| POBG2_58 | 13,459 | 13,286 | -173 | 13,275 | -184 | 13,306 | -153 | -1.13 |
| POBG2_59 | 40,929 | 40,823 | -106 | 40,995 | 66 | 40,986 | 57 | 0.14 |
| POBG2_60 | 123,382 | 125,358 | 1,976 | 123,900 | 518 | 123,828 | 446 | 0.36 |
| POBG2_61 | 13,814 | 13,787 | -27 | 13,814 | 0 | 13,816 | 2 | 0.02 |
| POBG2_62 | 14,021 | 14,015 | -6 | 14,103 | 82 | 13,920 | -101 | -0.72 |
| POBG2_63 | 9,250 | 9,276 | 26 | 9,190 | -60 | 8,962 | -288 | -3.11 |
| POBG2_64 | 880,611 | 880,776 | 165 | 880,196 | -415 | 880,351 | -260 | -0.03 |
| POBG2_65 | 31,092 | 30,739 | -353 | 30,820 | -272 | 30,852 | -240 | -0.77 |
| POBG2_66 | 292,617 | 294,336 | 1,719 | 292,849 | 232 | 292,496 | -121 | -0.04 |
| POBG2_67 | 172,316 | 170,704 | -1,612 | 172,312 | -4 | 172,232 | -84 | -0.05 |
| POBG2_68 | 21,518 | 21,146 | -372 | 21,201 | -317 | 21,317 | -201 | -0.94 |
| POBG2_69 | 70,746 | 69,813 | -933 | 70,611 | -135 | 70,649 | -97 | -0.14 |
| POBG2_70 | 38,949 | 37,477 | -1,472 | 37,941 | -1,008 | 38,369 | -580 | -1.49 |
| POBG2_71 | 105,555 | 105,197 | -358 | 105,258 | -297 | 105,289 | -266 | -0.25 |
| POBG2_M3 | 27,477,998 | 27,240,476 | -237,522 | 27,478,602 | 604 | 27,479,831 | 1,833 | 0.01 |
| POBG3_2 | 315,727 | 313,684 | -2,043 | 315,604 | -123 | 315,346 | -381 | -0.12 |
| POBG3_3 | 52,168 | 51,502 | -666 | 52,616 | 448 | 52,702 | 534 | 1.02 |
| POBG3_4 | 144,752 | 144,355 | -397 | 145,063 | 311 | 144,824 | 72 | 0.05 |
| POBG3_5 | 172,246 | 173,386 | 1,140 | 171,963 | -283 | 172,386 | 140 | 0.08 |
| POBG3_6 | 1,026,914 | 1,021,550 | -5,364 | 1,026,782 | -132 | 1,027,359 | 445 | 0.04 |
| POBG3_7 | 258,618 | 257,390 | -1,228 | 258,524 | -94 | 258,841 | 223 | 0.09 |
| POBG3_8 | 526,099 | 524,544 | -1,555 | 525,838 | -261 | 525,741 | -358 | -0.07 |
| POBG3_9 | 224,749 | 223,501 | -1,248 | 225,241 | 492 | 224,445 | -304 | -0.14 |
| POBG3_10 | 306,076 | 307,622 | 1,546 | 306,124 | 48 | 305,802 | -274 | -0.09 |
| POBG3_12 | 48,365 | 47,675 | -690 | 47,680 | -685 | 48,041 | -324 | -0.67 |
| POBG3_13 | 263,337 | 262,581 | -756 | 261,909 | -1,428 | 262,571 | -766 | -0.29 |
| POBG3_14 | 288,406 | 288,209 | -197 | 288,243 | -163 | 288,093 | -313 | -0.11 |
| POBG3_15 | 817,934 | 815,377 | -2,557 | 818,608 | 674 | 818,128 | 194 | 0.02 |
| POBG3_16 | 42,289 | 42,275 | -14 | 42,511 | 222 | 42,489 | 200 | 0.47 |
| POBG3_17 | 1,022,321 | 1,023,931 | 1,610 | 1,022,082 | -239 | 1,022,232 | -89 | -0.01 |
| POBG3_18 | 298,701 | 298,685 | -16 | 298,557 | -144 | 298,837 | 136 | 0.05 |
| POBG3_19 | 208,492 | 209,073 | 581 | 209,002 | 510 | 208,606 | 114 | 0.05 |

Table 5.2.1
Census counts and long-form estimate differences and discrepancies - Canada

| Characteristic | Census counts | Design weights |  | Coverage and non-response adjusted weights |  | Final weights |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | estimates | difference | estimates | difference | estimates | difference | discrepancy (\%) |
| POBG3_21 | 112,196 | 109,730 | -2,466 | 111,021 | -1,175 | 111,510 | -686 | -0.61 |
| POBG3_22 | 212,520 | 211,823 | -697 | 213,051 | 531 | 212,398 | -122 | -0.06 |
| PPADACSD (CSDxADA) | 33,994,454 | 32,767,748 | -1,226,706 | 33,993,365 | -1,089 | 33,994,454 | 0 | 0.00 |
| SNGLDET | 7,434,272 | 7,406,572 | -27,700 | 7,435,570 | 1,298 | 7,434,328 | 56 | 0.00 |
| TOTCFAM | 9,720,966 | 9,674,172 | -46,794 | 9,722,126 | 1,160 | 9,722,502 | 1,536 | 0.02 |
| TOTHHLD | 13,932,820 | 13,863,876 | -68,944 | 13,932,820 | 0 | 13,932,820 | 0 | 0.00 |
| TOTPERS | 33,994,454 | 33,740,107 | -254,347 | 33,993,365 | -1,089 | 33,994,454 | 0 | 0.00 |
| TPERGE15 | 28,314,895 | 28,110,924 | -203,971 | 28,309,127 | -5,768 | 28,314,932 | 37 | 0.00 |
| YRIMD_1900 | 1,038,231 | 1,038,998 | 767 | 1,037,883 | -348 | 1,038,201 | -30 | 0.00 |
| YRIMD_1981 | 273,907 | 275,157 | 1,250 | 274,631 | 724 | 274,167 | 260 | 0.09 |
| YRIMD_1986 | 491,794 | 491,464 | -330 | 491,411 | -383 | 491,648 | -146 | -0.03 |
| YRIMD_1991 | 726,588 | 725,927 | -661 | 726,496 | -92 | 726,876 | 288 | 0.04 |
| YRIMD_1996 | 680,866 | 683,744 | 2,878 | 681,300 | 434 | 680,314 | -552 | -0.08 |
| YRIMD_2001 | 849,532 | 846,726 | -2,806 | 849,728 | 196 | 849,356 | -176 | -0.02 |
| YRIMD_2006 | 961,099 | 956,822 | -4,277 | 960,315 | -784 | 960,960 | -139 | -0.01 |
| YRIMD_2011 | 1,150,540 | 1,141,489 | -9,051 | 1,150,015 | -525 | 1,150,117 | -423 | -0.04 |
| YRIMD_M5 | 343,899 | 339,304 | -4,595 | 342,984 | -915 | 342,983 | -916 | -0.27 |
| YRIMG1_1981 | 765,701 | 766,620 | 919 | 766,042 | 341 | 765,815 | 114 | 0.01 |
| YRIMG1_1991 | 1,407,454 | 1,409,672 | 2,218 | 1,407,796 | 342 | 1,407,190 | -264 | -0.02 |
| YRIMG1_2001 | 1,810,631 | 1,803,548 | -7,083 | 1,810,043 | -588 | 1,810,317 | -314 | -0.02 |

[^8]Chart 5.2.2
Differences between census counts and design-weighted, coverage and non-response adjusted, and final estimates

Difference (in thousands)


Note: All households with a design weight of 1 were excluded from the weighting process. These households either were located in an Indian reserve or a canvasser enumeration collection unit or were private households attached to a collective dwelling.
Source: Statistics Canada, 2016 Census long-form sample.

Chart 5.2.3
Discrepancy between census counts and final estimates, as a percentage of census counts
Discrepancy (\%)


Note: All households with a design weight of 1 were excluded from the weighting process. These households either were located in an Indian reserve or a canvasser enumeration collection unit or were private households attached to a collective dwelling.
Source: Statistics Canada, 2016 Census long-form sample.

## 6. Variance estimation

The error in an estimate is the difference between the estimate and the actual value of what is being estimated. There are several sources of error in the long-form sample survey, including sampling error and total non-response error. Sampling error stems from the fact that the estimates are based on observations from a sample and not from the Census of Population. Total non-response error occurs when households selected in the sample do not respond to the survey.

Error has a random component, measured by variance, and a systematic component, measured by bias. Variance measures how much the estimate varies from the average that would result from hypothetical repetitions of the survey process. Variance can be estimated using data from the sample. Bias is the difference between the average estimate that would result from hypothetical repetitions of the survey process and the actual value of the characteristic being estimated. The sampling and estimation methods used in the long-form sample survey produce a negligible bias.

Some estimation methods are more precise than others in estimating a particular characteristic of the population, so they can affect error. The estimated variance can be used to produce several quality indicators that are often used to measure the accuracy of an estimate. For example, it can be used to calculate standard errors, confidence intervals and coefficients of variation. Standard error was the quality indicator produced for the 2016 Census longform estimates. Standard error corresponds to the square root of the variance.

These measures of variability must be carefully distinguished from other measures of quality that are not, strictly speaking, measures of variability. Examples of such measures are the final response rates presented in section 3.11 and the global non-response rate to the 2016 Census long-form questionnaire. The response rate is an indicator of the risk associated with household non-response error. The global non-response rate is an indicator of the risk of error attributable to household non-response and item non-response. For more information, see Chapter 11 (http://www12.statcan.gc.ca/census-recensement/2016/ref/98-304/chap11-eng.cfm) of the Guide to the Census of Population, 2016 (Statistics Canada 2017).

Since the long-form sample is geographically stratified into take-some strata (mail-out and list/leave CUs) and take-all strata (canvasser CUs), two variance estimators are used. The first variance estimator is used to estimate the variance in take-some geographic areas (see section 6.3.1), and the second estimator is used to estimate the variance attributable to total non-response in take-all areas (see section 6.3.2). For the remainder of this chapter, the term variance is used to designate the sampling and total non-response variance in take-some geographic areas or the total non-response variance in take-all areas.

Standard errors for various estimates and geographic areas can be viewed or downloaded from the Census Profile Standard Error Supplement, Canada, provinces, territories, census divisions (CDs) and aggregate dissemination areas (ADAs), 2016 Census (http://www12.statcan.gc.ca/census-recensement/2016/dp-pd/adaprof/index. cfm?Lang=E\&TABID=0). The sampling and estimation methods used for the 2016 long-form survey are different from those used for the 2011 National Household Survey (NHS). Since the size and impact of non-response were very different in 2011, they affected estimate variability. For more information, refer to the methodological note comparing standard errors for 2016 estimates with those for 2011 and 2006 estimates (http://www12.statcan.gc.ca/ census-recensement/2016/ref/se-et-eng.cfm) for some variables.

### 6.1 Elements to consider in choosing a variance estimation method

A very high number of diverse estimates were produced, and quality indicators for these estimates needed to be established within a reasonable timeframe. As a result, a resampling variance estimator was used, derived from the modified partially balanced repeated replication method (Judkins 1990). The method consists of drawing samples (or replicates) from the original sample. Weights are calculated for each replicate, and the weights undergo the same coverage, non-response and calibration adjustments as the original sample. The resulting weights are called replicate weights. Estimates are then produced for each replicate, and the variance is estimated using replicate estimates and the average of these estimates.

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Figure 6.1.1 gives an overview of replication variance estimation when $R$ samples are used.
Figure 6.1.1
Overview of replication variance estimation


Figure 6.1.1 gives an overview of the replication variance estimation methodology used in the 2016 Census. The replication variance estimation method simulates the selection of several samples to estimate sampling variance.

More specifically, the figure shows the long-form questionnaire universe representing the population of interest and the long-form questionnaire sample. The sample is situated within the universe to indicate whether it corresponds to a subset of the population of interest. This sample is used to estimate a characteristic of the population of interest, such as the number of people who are members of a visible minority. The theta symbol is used to represent the true value of this characteristic. A circumflex on the theta indicates that the value is an estimate of this characteristic. This value is known as theta hat.

The R other samples placed outside the universe are linked to the long-form questionnaire sample with arrows. The arrows indicate that these samples are taken from the long-form questionnaire sample. The characteristic of interest is re-estimated based on these $R$ sub-samples. The theta hat $R$ values, referred to as theta hat one, theta hat two, up to theta hat $R$, are used to calculate the estimated theta hat variance.

In this figure, we define:

- theta, $\boldsymbol{\theta}$, the true value of the characteristic in the population, which can be a total, an average, a quantile, etc.;
- theta hat, $\hat{\boldsymbol{\theta}}$, the value of $\boldsymbol{\theta}$ estimated using the main weights;
- theta hat $r, \hat{\boldsymbol{\theta}}^{(r)}$, the value of $\boldsymbol{\theta}$ estimated using the replication weights $r, r=1, \ldots, \mathrm{R}$;
- theta hat bar, $\overline{\hat{\theta}}$, the average value of the $R$ replication estimates $\hat{\theta}^{(r)}$ and
- $\operatorname{vâr}(\hat{\theta})$, the estimated value of the variance of $\hat{\theta}$.


### 6.2 Variance estimator

The replicate estimator chosen for the long-form sample survey was derived from Fay's balanced half-sample method (Judkins 1990). This method determines the creation of replicates, the calculation of replicate weights and the multiplication factor used to estimate variance.

To produce variance estimates for the long-form sample estimates, two sets of replicate weights were created: the first had 32 replicate weights and the second had 100 replicate weights. The set of 32 replicate weights was produced to estimate the standard errors of standard products that are calculated under operational constraints (i.e., the need to publish a large number of standard errors within a reasonable timeframe). The set of 100 replicate weights was made available to Statistics Canada analysts, research data centre analysts who have access to microdata and users who request custom products, to provide more precise variance estimators.

The replication variance estimator can be calculated in two ways, one of which is more conservative than the other. The first method consists of adding the squared differences between the replication estimates, $\hat{\theta}^{(r)}$, and either the average of the replication estimates, $\overline{\hat{\theta}}$, or the estimate from the primary sample, $\hat{\boldsymbol{\theta}}$, and multiplying this sum by a certain factor. The second method, which uses the estimate from the primary sample, is more conservative. In the Computer-assisted Product Specification System used to publish statistics, the variance estimator is calculated using the average of the replication estimates.

For example, two variance estimators to estimate a total from a set R of replicates are given in the equations below:

$$
\begin{aligned}
& \text { Vâr }_{1}(\hat{T})=\frac{1}{R / 2} \sum_{r=1}^{\mathrm{R}}\left(\hat{T}^{(r)}-\overline{\hat{T}}\right)^{2}, \\
& \text { Vâr }_{2}(\hat{T})=\frac{1}{R / 2} \sum_{r=1}^{\mathrm{R}}\left(\hat{T}^{(r)}-\hat{T}\right)^{2}
\end{aligned}
$$

where

$$
\begin{gathered}
\hat{T}=\sum_{k \in s} w_{k} y_{k}, \\
\hat{T}^{(r)}=\sum_{k \in s} w_{k}^{(r)} y_{k}, \text { and } \\
\overline{\hat{T}}=\sum_{r=1}^{R} \hat{T}^{(r)} / \mathrm{R} .
\end{gathered}
$$

The final weight of the sample is represented by $\boldsymbol{w}_{k}, \boldsymbol{w}_{k}^{(r)}$ is the final weight of replicate $r, y_{k}$ is the value of characteristic y for unit $k$, and s is the long-form sample.

The number of degrees of freedom of the variance estimator is approximated by the number of squared differences $\left(\hat{T}^{(r)}-\hat{T}\right)^{2}$ for the variance estimator, i.e., 32 or 100 . The number of degrees of freedom gives an idea of the precision of the variance estimator and is used in calculating confidence intervals for long-form estimates. Since the method used to estimate the variance of the estimate can use few replicates, Student's law of probability was used to determine the quantile of the confidence level to be used in calculating the interval. In cases where the number of households participating in the estimate is larger than $R$, the number of degrees of freedom should be $R$. In cases where the number of households is less than $R$ but equal to or greater than 20 , the number of degrees of freedom corresponds to the number of households. Producing confidence intervals when there are fewer than 20 households is not recommended.

### 6.3 Replicate weight adjustment

### 6.3.1 Mail-out and list/leave collection units

As mentioned in section 6.2 , replicate weights were calculated for all long-form sample households. The replicates were partially balanced. They were balanced by resampling strata, which were created by combining CUs to obtain 600 to 1,800 households per resampling stratum.

Fay's modified balanced half-sample method, as described by Rao and Shao (1999), requires an epsilon value in the calculation of replicate weights to control the perturbation of the replicate weights. This perturbation results in all sampled households participating in every replicate, unlike other more popular replication methods. This facilitates the calibration of the replicate weights and, occasionally, the calculation of point estimates for each replicate (e.g., the denominator of a ratio estimator for a given replicate will not have a nil value if the corresponding denominator was not nil with the final weight). Adding an epsilon factor to the calculation of replicate weights meant the large survey fraction used to select the long-form sample could be taken into account. The technical details of the variance estimation process were provided by Devin and Verret (2016).

The replicate weights underwent the same adjustments as the primary sample design weight. They were adjusted for coverage and total non-response following the same methodology that was used for the primary sample weight (see section 4.4). The resulting replicate weights were then calibrated to census counts, once again following the same methodology that was used for the primary weight (see section 4.5).

### 6.3.2 Indian reserve ${ }^{8}$ and canvasser enumeration collection units

As described in Chapter 2, all the households in Indian reserve and canvasser enumeration CUs were selected with certainty. As such, they originally had a design weight equal to 1 . A coverage adjustment was not needed. All these households were selected for the long form, and therefore differential coverage between the short and long form could not occur. Total non-response in these areas was treated with the process of whole household imputation (WHI), described in Chapter 3. In other words, the data of a non-responding household were replaced by the data of another responding household from the same CU (except for geography variables, which were known for non-respondents). As a consequence, reweighting for households in Indian reserve and canvasser enumeration CUs was not needed.

Calibration was not needed in these areas, because the long form was a census. Consequently, all households in Indian reserve and canvasser enumeration CUs maintained their original weight equal to 1 in the final weighting scheme. For more information on incompletely enumerated Indian reserves and Indian settlements, please refer to Appendix 1.2 (http://www12.statcan.gc.ca/census-recensement/2016/ref/98-304/app-ann1-2-eng.cfm) of the Guide to the Census of Population, 2016, Catalogue no. 98-304-X.

Although sampling variability did not occur for households in Indian reserve and canvasser enumeration CUs, WHI variability did occur. Variance estimation in these areas was computed using a similar method to that of the rest of the country, with the following exceptions. First, the response probability by household size combination in each census division was estimated as the number of responding households divided by the number of in-scope households. Then the base replicate weights were created as in the rest of the country, except that all respondents for which the response probability was equal to 1 were placed in every replicate. Respondents with estimated response probabilities less than 1 were not considered certainties and were treated as sampled elements (i.e., they were randomly divided among the replicates). Non-respondent households imputed by WHI were also divided among replicates, and each one was assigned the replicate inclusion indicator corresponding to its donor in a manner similar to that of Shao and Tang (2011). This caused the weights to vary from one replicate to the other

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instead of the values. Finally, the replicate weights were calibrated to the number of households and number of persons in the SADA. As a result, the estimated variance of those two quantities was equal to 0 at the SADA level and at more aggregate levels, such as Canada (since those two constraints are mandatory in the rest of the country).

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## 7. Conclusion

The reintroduction of the mandatory long form allowed for several desirable characteristics in terms of sampling and weighting. Mainly, the sampling design was greatly simplified compared with that of the 2011 National Household Survey (NHS) (see Statistics Canada, 2015 for more details on the 2011 NHS sample design). As a direct result, the weighting procedure was also simplified, and the final weights were less variable than those of the 2011 NHS. Additionally, the mandatory nature of the 2016 long form presumably increased the response rate. Reducing the amount of non-response is always important, as this may have a direct effect on the bias of estimates. Therefore, minimizing the non-response rate is important to keep the non-response bias as small as possible.

The combination of the one-in-four sampling fraction and the overall unweighted final response rate of $96.1 \%$ meant that the 2016 long-form sample obtained more responses overall than the 2011 NHS (which had a one-inthree sampling fraction and a response rate of 68.6\%). The response rates across geographic areas were also more homogeneous. The much higher and more uniform response rate enabled the simplification of the weighting procedures.

All these features, coupled with an improved calibration procedure, meant that the maximum difference between census counts and final-weight estimates $(14,493)$ was about $1 / 10$ of the maximum difference for the 2011 NHS $(166,801)$. The improvement in the calibration procedure was a mixture of choosing fewer calibration variables, larger weighting areas (i.e., SADAs) and a better selection of constraints (such as the prioritization of constraints that split the population approximately in half).

As a consequence of these improvements, the range of the final weights was reduced (compared with the 2011 NHS). This resulted in the reduction of the standard error for most characteristics. For the variables related to the calibration constraints that were selected often, the variance was reduced by virtue of the calibration method. For variables not highly related to the calibration constraints, the variance was still moderated by the limitation of the range of the weights.

Finally, the release of replicate weights implies that users can easily estimate the standard error for any particular variable of interest. For the 2011 NHS, the chosen method of variance estimation-Taylor series linearization combined with a corrective upward adjustment factor (from Monte Carlo simulation)—made computing standard errors for non-standard products difficult.

## Appendix A - Glossary

The definitions of the main census terms, variables and concepts mentioned in this document are presented here. Users can also refer to the Dictionary, Census of Population, 2016 (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/index-eng.cfm), Catalogue no. 98-301-X for additional information.

Aggregate dissemination area (ADA) (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/geo053-eng. cfm): The aggregate dissemination area (ADA) is a new dissemination geography created for the 2016 Census. ADAs cover the entire country and, where possible, have a population between 5,000 and 15,000 based on the previous census population counts. ADAs are created from existing dissemination geographic areas and are formed from census tracts (CTs), census subdivisions (CSDs) or dissemination areas (DAs). ADAs respect provincial, territorial, census division (CD), census metropolitan area (CMA) and census agglomeration (CA) with census tract (CT) boundaries.

The intent of the ADA geography is to ensure the availability of census data, where possible, across all regions of Canada.

Census division (CD) (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/geo008-eng.cfm): Group of neighbouring municipalities joined together for the purposes of regional planning and managing common services (such as police or ambulance services). These groupings are established under laws in effect in certain provinces of Canada. Census division (CD) is the general term for provincially legislated areas (such as county, municipalité régionale de comté and regional district) or their equivalents. In other provinces and the territories where laws do not provide for such areas, Statistics Canada defines equivalent areas for statistical reporting purposes in cooperation with these provinces and territories. Census divisions are intermediate geographic areas between the province or territory level and the municipality (census subdivision).

Census family (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/fam004-eng.cfm): "Census family" is defined as a married couple and the children, if any, of either and/or both spouses; a couple living common law and the children, if any, of either and/or both partners; or a lone parent of any marital status with at least one child living in the same dwelling and that child or those children. All members of a particular census family live in the same dwelling. A couple may be of opposite or same sex. Children may be children by birth, marriage, commonlaw union or adoption regardless of their age or marital status as long as they live in the dwelling and do not have their own married spouse, common-law partner or child living in the dwelling. Grandchildren living with their grandparent(s) but with no parents present also constitute a census family.

Census subdivision (CSD) (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/geo012-eng.cfm): Census subdivision (CSD) is the general term for municipalities (as determined by provincial or territorial legislation) or areas treated as municipal equivalents for statistical purposes (e.g., Indian reserves, Indian settlements and unorganized territories). Municipal status is defined by laws in effect in each province and territory in Canada.

Census tract (CT) (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/geo013-eng.cfm): Census tracts (CTs) are small, relatively stable geographic areas that usually have a population of less than 10,000 persons, based on data from the previous Census of Population Program. They are located in census metropolitan areas and in census agglomerations that had a core population of 50,000 or more in the previous census.

A committee of local specialists (for example, planners, health and social workers, and educators) initially delineates census tracts in conjunction with Statistics Canada. Once a census metropolitan area (CMA) or census agglomeration (CA) has been subdivided into census tracts, the census tracts are maintained even if the core population subsequently declines below 50,000 .

Collection unit (CU):9 Collection units are small geographic units used for the collection of census data. Collection units cover all the territory of Canada.

Collective dwelling (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/dwelling-logements002-eng. cfm): Refers to a dwelling of a commercial, institutional or communal nature. It may be identified by a sign on the premises or by an enumerator speaking with the person in charge, a resident, a neighbour, etc. Included are lodging or rooming houses, hotels, motels, tourist establishments, nursing homes, hospitals, staff residences, military bases, work camps, jails, group homes, and so on.

Dissemination area (DA) (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/geo021-eng.cfm): A dissemination area (DA) is a small, relatively stable geographic unit composed of one or more adjacent dissemination blocks with an average population of 400 to 700 persons based on data from the previous Census of Population Program. It is the smallest standard geographic area for which all census data are disseminated. DAs cover all the territory of Canada.

Dwelling (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/dwelling-logements004-eng.cfm): A dwelling is defined as a set of living quarters. Two types of dwellings are identified in the census, collective dwellings and private dwellings. The former pertains to dwellings which are institutional, communal or commercial in nature. The latter, private dwellings, refers to a separate set of living quarters with a private entrance either from outside the building or from a common hall, lobby, vestibule or stairway inside the building. The entrance to the dwelling must be one that can be used without passing through the living quarters of some other person or group of persons.

Economic family (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/fam011-eng.cfm): "Economic family" refers to a group of two or more persons who live in the same dwelling and are related to each other by blood, marriage, common-law union, adoption or a foster relationship. A couple may be of opposite or same sex. By definition, all persons who are members of a census family are also members of an economic family. Examples of the broader concept of economic family include the following: two co-resident census families who are related to one another are considered one economic family; co-resident siblings who are not members of a census family are considered as one economic family; and, nieces or nephews living with aunts or uncles are considered one economic family.

Household (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/households-menage007-eng.cfm): Household refers to a person or group of persons who occupy the same dwelling and do not have a usual place of residence elsewhere in Canada or abroad. The dwelling may be either a collective dwelling or a private dwelling. The household may consist of a family group such as a census family, of two or more families sharing a dwelling, of a group of unrelated persons or of a person living alone. Household members who are temporarily absent on reference day are considered part of their usual household.

Private dwelling (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/dwelling-logements005-eng.cfm): "Private dwelling" refers to a separate set of living quarters with a private entrance either from outside the building or from a common hall, lobby, vestibule or stairway inside the building. The entrance to the dwelling must be one that can be used without passing through the living quarters of some other person or group of persons.

The dwelling must meet the two conditions necessary for year-round occupancy:

1. a source of heat or power (as evidenced by chimneys, power lines, oil or gas pipes or meters, generators, woodpiles, electric lights, heating pumps or solar panels)
2. an enclosed space that provides shelter from the elements, as evidenced by complete and enclosed walls and roof, and by doors and windows that provide protection from wind, rain and snow.
[^10]Dwellings that do not meet the conditions necessary for year-round occupancy are marginal dwellings. Private dwellings are classified into regular private dwellings and occupied marginal dwellings. Regular private dwellings are further classified into three major groups: occupied dwellings (occupied by usual residents), dwellings occupied solely by foreign residents and/or by temporarily present persons, and unoccupied dwellings. Marginal dwellings are classified as occupied by usual residents or occupied solely by foreign residents and/or by temporarily present persons. Marginal dwellings that were unoccupied on May 10, 2016, are not counted in the housing stock.

Private dwelling occupied by usual residents (http://www12.statcan.gc.ca/census-recensement/2016/ref/ dict/dwelling-logements006-eng.cfm): Refers to a private dwelling in which a person or a group of persons is permanently residing. Also included are private dwellings whose usual residents are temporarily absent on May 10, 2016. Unless otherwise specified, all data in housing products are for private dwellings occupied by usual residents, rather than for unoccupied private dwellings or dwellings occupied solely by foreign and/or by temporarily present persons.

Private household (http://www12.statcan.gc.ca/census-recensement/2016/ref/dict/households-menage014-eng. cfm): "Private household" refers to a person or group of persons who occupy the same dwelling and do not have a usual place of residence elsewhere in Canada or abroad. The household universe is divided into two subuniverses on the basis of whether the household is occupying a collective dwelling or a private dwelling. The latter is a private household.

For census purposes, households are classified into three groups: private households, collective households and households outside Canada.

Unless otherwise specified, all data in census products are for private households only.
Super aggregate dissemination area (SADA): ${ }^{10}$ Super aggregate dissemination areas are created specifically for weighting 2016 Census data. They respect pre-established rules, some of which are mandatory and others optional. SADAs are created by combining aggregate dissemination areas (ADAs) and are contained within provincial and territorial boundaries. All individuals living in census collection units (CUs) are excluded from the SADA population. SADAs are created, inasmuch as possible, with a target population of between 50,000 and 150,000 in mind. Census divisions (CD) with a population of 40,000 to 50,000 comprise their own SADA. In addition, where possible, SADAs respect the boundaries of-in order of priority-census division (CDs), census metropolitan areas (CMAs),census agglomerations (CAs) and census subdivisions (CSD). Lastly, SADAs should be created by combining adjacent ADAs (where possible) and must be as compact as possible.

[^11]
## Appendix B - The history of sampling in the Canadian census

Sampling was first used in the Canadian census in 1941. A housing schedule was completed for every 10th dwelling. The information from 27 questions on the separate housing schedule was integrated with the data in the personal and household section of the population schedule for the same dwelling. This enabled cross-tabulation of sample and basic characteristics. Also in the 1941 Census, sampling was used at the processing stage to obtain early estimates of earnings of wage-earners, of the distribution of the population of working age and of the composition of families in Canada. In this case, a sample of every 10th enumeration area across Canada was selected and all population schedules in these areas were processed in advance.

The census of housing was again conducted on a sample basis in 1951. This time, every fifth dwelling (those whose identification numbers ended in a 2 or 7) was selected to complete a housing document containing 24 questions. In the 1961 Census, persons aged 15 years and older in a $20 \%$ sample of private households were required to complete a population sample questionnaire containing questions on internal migration, fertility and income. Sampling was not used in the smaller censuses of 1956 and 1966.

The 1971 Census saw several major innovations in the method of census-taking. The primary change was from the traditional canvasser method of enumeration to the use of self-enumeration for the majority of the population. This change was prompted by the results of several studies in Canada and elsewhere (Fellegi 1964; Hansen et al. 1959), which indicated that the effect of the enumerator was a major contribution to the variance of census figures in a canvasser census. Consequently, the use of self-enumeration was expected to reduce the variance of census figures by reducing the effect of the enumerator and by giving the respondent more time and privacy in which to answer the census questions-factors that might be expected to yield more accurate responses.

The second aspect of the 1971 Census that differentiated it from any earlier census was its content. The number of topics covered and the number of questions asked were greater than in any previous census. Considerations of cost, respondent burden and timeliness versus the level of data quality to be expected using self-enumeration and sampling led to a decision to collect all but certain basic characteristics on a one-third sample basis in the 1971 Census. In all but the more remote areas of Canada, every third private household received the "long questionnaire," which contained all the census questions. The remaining private households received the "short questionnaire" containing only the basic questions covering name, relationship to head of household, sex, date of birth, marital status, mother tongue, type of dwelling, tenure, number of rooms, water supply, toilet facilities and certain census coverage items. All households in pre-identified remote enumeration areas and all collective dwellings received the long questionnaire. A more detailed description of the consideration of the use of sampling in the 1971 Census is given in Sampling in the Census (Dominion Bureau of Statistics 1968).

The 1976 Census had considerably less content than the 1971 Census. Furthermore, the 1976 questionnaire did not include the questions that cause the most difficulty in collection (e.g., income) or that are costly to code (e.g., occupation, industry and place of work). Therefore, the benefits of sampling in terms of cost savings and reduced respondent burden were less clear than for the 1971 Census. Nevertheless, after estimating the potential cost savings to be expected with various sampling fractions and considering the public relations issues related to a reversion to $100 \%$ enumeration after a successful application of sampling in 1971, Statistics Canada decided to use the same sampling procedure in 1976 as in 1971.

Most of the methodology used in the 1971 and 1976 censuses was kept for the 1981 Census, except that the sampling rate was reduced from every third occupied private household to every fifth. Studies done at the time showed that the resulting reduction in data quality (measured in terms of variance) would be tolerable and would not be significant enough to offset the benefits of reduced cost and respondent burden and improved timeliness (Royce 1983). The one-in-five sampling rate was maintained for every census from 1981 to 2006.

In 2011, information previously collected by the mandatory long-form census questionnaire was collected on a voluntary basis, via the National Household Survey (NHS). With this change, every household was required to answer the 10 questions that were contained in the 2011 Census questionnaire, while $30 \%$ of households were selected to respond to the NHS. As well, NHS non-responding households were subsampled for follow-up at a rate of one in three. The increased sampling fraction to $30 \%$ was implemented in anticipation of a lower response rate

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to the NHS. For the 2016 Census, the government reinstated the census long-form questionnaire as mandatory, replacing the NHS. The sampling fraction was changed in 2016 to one in four, compared with one in five for the previous census long-form questionnaire in 2006, to mitigate the risk of the response rate not recovering to its previously high levels.

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## Appendix C - Constraints used in or excluded from the weighting process ${ }^{11}$

The following is a list of the possible constraints defined at the ADA and SADA levels. A total of 270 possible constraints were defined at the ADA level and 200 at the SADA level. Any constraints that were not defined at either the ADA level or the SADA level are not taken into account in the table below. This is why this table does not contain 4,283 or 4,180 ADAs and 408 SADAs, like previous tables.

Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| ADULTCF | Adults in a census family | SADA | 105 | 302 | SADA | 138 | 270 |
| AGE00_14 | Persons aged 0 to 14 years | SADA | 288 | 120 | BOTH | 4,473 | 115 |
| AGE4 | Persons aged 0 to 4 years | SADA | 70 | 338 | BOTH | 868 | 3,720 |
| AGE9 | Persons aged 5 to 9 years | SADA | 64 | 344 | BOTH | 836 | 3,752 |
| AGE14 | Persons aged 10 to 14 years | SADA | 52 | 356 | BOTH | 820 | 3,768 |
| AGE15_24 | Persons aged 15 to 24 years | SADA | 310 | 98 | SADA | 390 | 18 |
| AGE15_29 | Persons aged 15 to 29 years | $\ldots$ | ... | ... | ADA | 4,084 | 96 |
| AGE19 | Persons aged 15 to 19 years | SADA | 268 | 140 | BOTH | 822 | 3,766 |
| AGE24 | Persons aged 20 to 24 years | SADA | 252 | 156 | BOTH | 885 | 3,703 |
| AGE25_34 | Persons aged 25 to 34 years | SADA | 373 | 35 | SADA | 401 | 7 |
| AGE29 | Persons aged 25 to 29 years | SADA | 324 | 84 | BOTH | 1,406 | 3,182 |
| AGE30_49 | Persons aged 30 to 49 years | $\ldots$ | ... | ... | ADA | 4,148 | 32 |
| AGE34 | Persons aged 30 to 34 years | SADA | 348 | 60 | BOTH | 1,575 | 3,013 |
| AGE35_44 | Persons aged 35 to 44 years | SADA | 360 | 48 | SADA | 408 | 0 |
| AGE39 | Persons aged 35 to 39 years | SADA | 369 | 39 | BOTH | 1,455 | 3,133 |
| AGE44 | Persons aged 40 to 44 years | SADA | 389 | 19 | BOTH | 1,384 | 3,204 |

[^12]
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Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| AGE45_54 | Persons aged 45 to 54 years | SADA | 406 | 2 | SADA | 408 | 0 |
| AGE49 | Persons aged 45 to 49 years | SADA | 403 | 5 | BOTH | 1,656 | 2,932 |
| AGE50_64 | Persons aged 50 to 64 years | ... | ... | ... | ADA | 4,026 | 154 |
| AGE54 | Persons aged 50 to 54 years | SADA | 405 | 3 | BOTH | 1,918 | 2,670 |
| AGE55_64 | Persons aged 55 to 64 years | SADA | 390 | 18 | SADA | 402 | 6 |
| AGE59 | Persons aged 55 to 59 years | SADA | 385 | 23 | BOTH | 1,683 | 2,905 |
| AGE64 | Persons aged 60 to 64 years | SADA | 371 | 37 | BOTH | 1,491 | 3,097 |
| AGE65PL | Persons aged 65 years and older | SADA | 390 | 18 | BOTH | 4,361 | 227 |
| AGE74 | Persons aged 65 to 74 years | SADA | 400 | 8 | BOTH | 2,128 | 2,460 |
| AGE75PL | Persons aged 75 years and older | SADA | 391 | 17 | BOTH | 2,109 | 2,479 |
| APTLT5 | Households living in an apartment in a building with less than five storeys | SADA | 406 | 2 | BOTH | 2,669 | 1,919 |
| CHILD | Children in a census family | SADA | 7 | 400 | SADA | 39 | 369 |
| CHILDFAM | Census families with children | SADA | 16 | 391 | BOTH | 3,145 | 1,443 |
| $\begin{aligned} & \text { COMLAW_ } \\ & \text { YE } \end{aligned}$ | Persons in a common law relationship | SADA | 252 | 155 | BOTH | 2,708 | 1,880 |
| COMLAWNO_ <br> DIV | Divorced persons not in a common law relationship | SADA | 405 | 2 | BOTH | 1,181 | 3,407 |
| COMLAWNO_ OTHERS | Divorced, separated or widowed persons not in a common law relationship | SADA | 399 | 8 | BOTH | 2,954 | 1,634 |
| COMLAWNO_ SEP | Separated persons not in a common law relationship | SADA | 399 | 8 | BOTH | 609 | 3,979 |

Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| COMLAWNO SINGLE | Single persons (never married) not in a common law relationship | SADA | 204 | 203 | BOTH | 3,496 | 1,092 |
| COMLAWNO SINGLE_LT15 | Single persons (never married) under 15 years and not in a common law relationship | ... | ... | ... | BOTH | 4,473 | 115 |
| COMLAWNO WID | Widowed persons not in a common law relationship | SADA | 398 | 9 | BOTH | 763 | 3,825 |
| COMLAWYE MARRIED | Persons married or in a common law relationship | SADA | 200 | 207 | BOTH | 3,241 | 1,347 |
| COUPLE | Persons in a couple (spouse, partner) | SADA | 111 | 296 | SADA | 139 | 269 |
| EMPIN_GT50 | Persons with an annual employment income above the 50th percentile for the ADA | ... | ... | ... | ADA | 4,110 | 70 |
| EMPIN_LE50 | Persons with an annual employment income equal to or below the 50th percentile for the ADA | ... | ... | ... | ADA | 4,094 | 86 |
| EMPIN_P0 | Persons with no annual employment income, at the ADA level | ... | ... | ... | ADA | 4,129 | 51 |
| $\begin{aligned} & \text { EMPIN_P0_ } \\ & \text { GE15 } \end{aligned}$ | Persons aged 15 years and older with no annual employment income, at the ADA level | ... | ... | ... | ADA | 4,069 | 111 |

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Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| $\begin{aligned} & \text { EMPIN_PO_ } \\ & \text { LT15 } \end{aligned}$ | Persons aged 14 years and younger with no annual employment income, at the ADA level | ... | ... | ... | ADA | 4,083 | 97 |
| EMPIN_P25 | Persons with an annual employment income equal to or below the 25th percentile for the ADA | ... | ... | ... | ADA | 2,636 | 1,544 |
| EMPIN_P50 | Persons with an annual employment income above the 25th percentile and equal to or below the 50th percentile for the ADA | ... | ... | ... | ADA | 2,619 | 1,561 |
| EMPIN_P75 | Persons with an annual employment income above the 50th percentile and equal to or below the 75th percentile for the ADA | ... | ... | ... | ADA | 3,713 | 467 |
| EMPIN_P100 | Persons with an annual employment income above the 75th percentile for the ADA | ... | ... | ... | ADA | 3,718 | 462 |
| EMPIN_SADA <br> GT50 | Persons with an annual employment income above the 50th percentile for the SADA | SADA | 405 | 3 | SADA | 407 | 1 |

Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| EMPIN_SADA_ LE50 | Persons with an annual employment income equal to or below the 50th percentile for the SADA | SADA | 406 | 2 | SADA | 407 | 1 |
| EMPIN_SADA_ P0 | Persons with no annual employment income, at the SADA level | SADA | 407 | 1 | SADA | 406 | 2 |
| EMPIN_SADA <br> P0_GE15 | Persons aged 15 years and older with no annual employment income, at the SADA level | . | ... | ... | SADA | 392 | 16 |
| EMPIN_SADA <br> P0_LT15 | Persons aged 14 years and younger with no annual employment income, at the SADA level | ... | ... | ... | SADA | 390 | 18 |
| EMPIN_SADA_ P25 | Persons with an annual employment income equal to or below the 25th percentile for the SADA | SADA | 341 | 67 | SADA | 192 | 216 |
| $\begin{aligned} & \text { EMPIN_SADA_ } \\ & \text { P50 } \end{aligned}$ | Persons with an annual employment income above the 25th percentile and equal to or below the 50th percentile for the SADA | SADA | 339 | 69 | SADA | 192 | 216 |

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Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| EMPIN_SADA _ P75 | Persons with an annual employment income above the 50th percentile and equal to or below the 75th percentile for the SADA | SADA | 386 | 22 | SADA | 407 | 1 |
| $\begin{aligned} & \text { EMPIN_SADA_ } \\ & \text { P100 } \end{aligned}$ | Persons with an annual employment income above the 75th percentile for the SADA | SADA | 387 | 21 | SADA | 408 | 0 |
| FEMALE | Females | SADA | 39 | 369 | BOTH | 4,400 | 188 |
| FEMALEGE15 | Females aged 15 years and older | SADA | 386 | 22 | BOTH | 3,054 | 1,534 |
| FEMALELT15 | Females aged 14 years and younger | SADA | 35 | 373 | BOTH | 2,942 | 1,646 |
| HHADA (ADA) | Households in the ADA | ADA | 4,043 | 195 | ... | ... |  |
| HHADACSD | Households that fall within the CSD and the ADA | $\ldots$ | ... | ... | ADA | 4,938 | 1,994 |
| HHINC_GT50 | Households with an annual income above the 50th percentile for the ADA | ... | ... | ... | ADA | 4,138 | 42 |
| HHINC_LE50 | Households with an annual income at or below the 50th percentile for the ADA | ... | ... | ... | ADA | 4,138 | 42 |
| HHINC_P25 | Households with an annual income at or below the 25th percentile for the ADA | ... | $\ldots$ | ... | ADA | 4,022 | 158 |

Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| HHINC_P50 | Households with an annual income above the 25th percentile and at or below the 50th percentile for the ADA | ... | ... | ... | ADA | 4,024 | 156 |
| HHINC_P75 | Households with an annual income above the 50th percentile and at or below the 75th percentile for the ADA | ... | ... | ... | ADA | 4,014 | 166 |
| HHINC_P100 | Households with an annual income above the 75th percentile for the ADA | ... | ... | ... | ADA | 4,014 | 166 |
| $\begin{aligned} & \text { HHINC_SADA_ } \\ & \text { GT50 } \end{aligned}$ | Households with an annual income above the 50th percentile for the SADA | SADA | 407 | 0 | SADA | 407 | 1 |
| $\begin{aligned} & \text { HHINC_SADA_ } \\ & \text { LE50 } \end{aligned}$ | Households with an annual income at or below the 50th percentile for the SADA | SADA | 407 | 0 | SADA | 407 | 1 |
| HHINC_SADA_ P25 | Households with an annual income at or below the 25th percentile for the SADA | SADA | 407 | 0 | SADA | 407 | 1 |
| HHINC_SADA_ P50 | Households with an annual income above the 25th percentile and at or below the 50th percentile for the SADA | SADA | 407 | 0 | SADA | 407 | 1 |

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Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| $\begin{aligned} & \text { HHINC_SADA_ } \\ & \text { P75 } \end{aligned}$ | Households with an annual income above the 50th percentile and at or below the 75th percentile for the SADA | SADA | 402 | 5 | SADA | 407 | 1 |
| $\begin{aligned} & \text { HHINC_SADA_ } \\ & \text { P100 } \end{aligned}$ | Households with an annual income above the 75th percentile for the SADA | SADA | 402 | 5 | SADA | 407 | 1 |
| HHSIZE1 | One-person households | SADA | 28 | 379 | BOTH | 766 | 3,822 |
| HHSIZE2 | Two-person households | SADA | 114 | 293 | BOTH | 3,985 | 603 |
| HHSIZE3 | Three-person households | SADA | 394 | 13 | BOTH | 3,634 | 954 |
| HHSIZE4 | Four-person households | SADA | 376 | 31 | BOTH | 3,639 | 949 |
| HHSIZE5 | Five-person households | SADA | 131 | 276 | BOTH | 991 | 3,597 |
| HHSIZEGE5 | Five-or-more-person households | SADA | 15 | 392 | BOTH | 386 | 4,202 |
| HHSIZEGE6 | Six-or-moreperson households | SADA | 13 | 394 | BOTH | 277 | 4,311 |
| INEFAM | Persons in an economic family | SADA | 51 | 356 | SADA | 111 | 297 |
| IR_LINK_NO | Persons who could not be linked to the Indian Register | SADA | 207 | 200 | BOTH | 515 | 4,073 |
| IR_LINK_YE | Persons who could be linked to the Indian Register | SADA | 207 | 200 | BOTH | 451 | 4,137 |
| LIM_NO | Persons not in a low income economic family | SADA | 399 | 8 | BOTH | 3,267 | 1,321 |
| LIM_YE | Persons in a low income economic family | SADA | 399 | 8 | BOTH | 3,267 | 1,321 |
| LONEPAR | Lone parents | SADA | 118 | 289 | SADA | 149 | 259 |

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Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| MALE | Males | SADA | 39 | 369 | BOTH | 4,400 | 188 |
| MALEGE15 | Males aged 15 years and older | SADA | 295 | 113 | BOTH | 3,054 | 1,534 |
| MALELT15 | Males aged 14 years and younger | SADA | 53 | 355 | BOTH | 2,938 | 1,650 |
| MARRIED | Married persons | SADA | 346 | 62 | BOTH | 2,761 | 1,827 |
| NOCLDFAM | Census families without children | SADA | 40 | 367 | BOTH | 1,774 | 2,814 |
| NOINEFAM | Persons in a household that are not part of an economic family | SADA | 51 | 356 | SADA | 111 | 297 |
| NOINEFAMHHSIZEEQ1 | Person in a oneperson household that is not part of an economic family | ... | ... | ... | SADA | 103 | 305 |
| NOINEFAMHHSIZEGT1 | Persons in a two-or-more-person household that are not part of an economic family | ... | ... | ... | SADA | 160 | 248 |
| NOTINFAM | Persons not in a census family | SADA | 75 | 332 | SADA | 143 | 265 |
| NOTINFAMHHSIZEGT1 | Persons in a two-or-more-person household that are not part of a census family | ... | ... | ... | SADA | 98 | 310 |
| OLN_BI | Official languages English and French | SADA | 374 | 33 | BOTH | 1,874 | 2,714 |
| OLN_EN | Official language English | SADA | 78 | 329 | BOTH | 559 | 4,029 |
| OLN_FR | Official language French | SADA | 74 | 333 | BOTH | 782 | 3,806 |
| OLN_NO | Official language neither | SADA | 224 | 183 | BOTH | 455 | 4,133 |
| POBG2_1 | Place of birth Albania | SADA | 10 | 299 | BOTH | 64 | 4,524 |

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Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| POBG2_2 | Place of birth Andorra, Gibraltar and Spain | SADA | 22 | 366 | BOTH | 101 | 4,487 |
| POBG2_3 | Place of birth Australia and New Zealand | SADA | 13 | 383 | BOTH | 191 | 4,397 |
| POBG2_4 | Place of birth Austria, Germany and Liechtenstein | SADA | 25 | 382 | BOTH | 262 | 4,326 |
| POBG2_5 | Place of birth Estonia, Latvia and Lithuania | SADA | 17 | 332 | BOTH | 62 | 4,526 |
| POBG2_6 | Place of birth Belgium and Netherlands | SADA | 14 | 393 | BOTH | 241 | 4,347 |
| POBG2_7 | Place of birth Belize, El Salvador, Guatemala and Honduras | SADA | 26 | 381 | BOTH | 173 | 4,415 |
| POBG2_8 | Place of birth Bhutan | SADA | 3 | 122 | BOTH | 10 | 4,578 |
| POBG2_9 | Place of birth Botswana, Lesotho and Swaziland | SADA | 13 | 254 | BOTH | 11 | 4,577 |
| POBG2_10 | Place of birth Brazil | SADA | 15 | 389 | BOTH | 138 | 4,450 |
| POBG2_11 | Place of birth Bulgaria and Romania | SADA | 45 | 361 | BOTH | 179 | 4,409 |
| POBG2_12 | Place of birth Burma (Myanmar) | SADA | 6 | 285 | BOTH | 39 | 4,549 |
| POBG2_13 | Place of birth Canada | SADA | 34 | 335 | BOTH | 40 | 4,548 |
| POBG2_14 | Place of birth Anguilla, Bermuda, British Virgin Islands, Cayman Islands, Montserrat and Turks and Caicos Islands | SADA | 9 | 303 | BOTH | 26 | 4,562 |

Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| POBG2_15 | Place of birth Guadeloupe, Martinique and Saint Martin (French part) | SADA | 4 | 181 | BOTH | 17 | 4,571 |
| POBG2_16 | Place of birth Democratic Republic of the Congo and Republic of the Congo | SADA | 16 | 366 | BOTH | 85 | 4,503 |
| POBG2_17 | Place of birth Cameroon, Central African Republic, Chad and Gabon | SADA | 16 | 350 | BOTH | 71 | 4,517 |
| POBG2_18 | Place of birth - <br> Angola and <br> Sao Tome and <br> Principe | SADA | 14 | 289 | BOTH | 44 | 4,544 |
| POBG2_19 | Place of birth Kazakhstan, Kyrgyzstan, Tajikistan, Turkmenistan and Uzbekistan | SADA | 17 | 356 | BOTH | 89 | 4,499 |
| POBG2_20 | Place of birth Chile | SADA | 19 | 384 | BOTH | 162 | 4,426 |
| POBG2_21 | Place of birth China, Hong Kong, Macao and Taiwan | SADA | 65 | 342 | BOTH | 269 | 4,319 |
| POBG2_22 | Place of birth Colombia, Ecuador and Peru | SADA | 69 | 338 | BOTH | 256 | 4,332 |
| POBG2_23 | Place of birth Costa Rica, Nicaragua and Panama | SADA | 16 | 379 | BOTH | 127 | 4,461 |
| POBG2_24 | Place of birth Czech Republic, Hungary and Slovakia | SADA | 20 | 380 | BOTH | 161 | 4,427 |
| POBG2_25 | Place of birth - <br> Burundi and Rwanda | SADA | 14 | 340 | BOTH | 68 | 4,520 |

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Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| POBG2_26 | Place of birth Eritrea, Kenya, Tanzania, Uganda and Zambia | SADA | 33 | 358 | BOTH | 159 | 4,429 |
| POBG2_27 | Place of birth - <br> Comoros, <br> Djibouti, <br> Madagascar, <br> Malawi, Mauritius, <br> Seychelles, <br> Somalia and <br> Zimbabwe | SADA | 28 | 378 | BOTH | 150 | 4,438 |
| POBG2_28 | Place of birth - <br> Belarus, Moldova, <br> Russian <br> Federation and Ukraine | SADA | 43 | 364 | BOTH | 176 | 4,412 |
| POBG2_29 | Place of birth Egypt, South Sudan and Sudan | SADA | 22 | 377 | BOTH | 131 | 4,457 |
| POBG2_30 | Place of birth Ethiopia | SADA | 28 | 325 | BOTH | 111 | 4,477 |
| POBG2_31 | Place of birth France, Luxembourg and Monaco | SADA | 27 | 380 | BOTH | 197 | 4,391 |
| POBG2_32 | Place of birth Cambodia, Laos and Viet Nam | SADA | 59 | 348 | BOTH | 149 | 4,439 |
| POBG2_33 | Place of birth Cuba, Dominican Republic and Haiti | SADA | 34 | 372 | BOTH | 185 | 4,403 |
| POBG2_34 | Place of birth Greece | SADA | 16 | 374 | BOTH | 131 | 4,457 |
| POBG2_35 | Place of birth Guyana and Suriname | SADA | 45 | 327 | BOTH | 140 | 4,448 |
| POBG2_36 | Place of birth Holy See (Vatican City State), Italy and San Marino | SADA | 54 | 349 | BOTH | 241 | 4,347 |

Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| POBG2_37 | Place of birth Bahamas, Jamaica and Puerto Rico | SADA | 76 | 314 | BOTH | 221 | 4,367 |
| POBG2_38 | Place of birth - <br> Antigua and Barbuda, Aruba, Barbados, Bonaire, Sint Eustatius and Saba, Curaçao, Dominica, Grenada, Saint Kitts and Nevis, Saint Lucia, Sint Maarten (Dutch part), Saint Vincent and the Grenadines, Trinidad and Tobago and United States Virgin Islands | SADA | 39 | 357 | BOTH | 186 | 4,402 |
| POBG2_39 | Place of birth Japan | SADA | 27 | 369 | BOTH | 175 | 4,413 |
| POBG2_40 | Place of birth North Korea and South Korea | SADA | 24 | 376 | BOTH | 120 | 4,468 |
| POBG2_41 | Place of birth Liberia | SADA | 12 | 274 | BOTH | 27 | 4,561 |
| POBG2_42 | Place of birth Algeria, Libya, Morocco and Tunisia | SADA | 23 | 382 | BOTH | 134 | 4,454 |
| POBG2_43 | Place of birth - <br> Brunei <br> Darussalam, Indonesia, Malaysia, <br> Philippines, <br> Singapore and <br> Timor-Leste | SADA | 87 | 320 | BOTH | 266 | 4,322 |
| POBG2_44 | Place of birth Malta | SADA | 10 | 270 | BOTH | 36 | 4,552 |
| POBG2_45 | Place of birth Mexico | SADA | 17 | 390 | BOTH | 183 | 4,405 |

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Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| POBG2_46 | Place of birth - |  |  |  |  |  |  |
|  | Bahrain, Qatar, |  |  |  |  |  |  |
|  | Saudi Arabia, |  |  |  |  |  |  |
|  | United Arab |  |  |  |  |  |  |
|  | Emirates and |  |  |  |  |  |  |
|  | Yemen | SADA | 24 | 356 | BOTH | 156 | 4,432 |
| POBG2_47 | Place of birth - |  |  |  |  |  |  |
|  | Lebanon, Syria | SADA | 33 | 373 | BOTH | 164 | 4,424 |
| POBG2_48 | Place of birth - |  |  |  |  |  |  |
|  | Afghanistan, |  |  |  |  |  |  |
|  | Cyprus, Iran, Iraq, |  |  |  |  |  |  |
|  | Israel, Jordan, |  |  |  |  |  |  |
|  | Kuwait, Oman, |  |  |  |  |  |  |
|  | Turkey and West |  |  |  |  |  |  |
|  | Bank and Gaza |  |  |  |  |  |  |
|  | Strip (Palestine) | SADA | 94 | 309 | BOTH | 218 | 4,370 |
| POBG2_49 | Place of birth - |  |  |  |  |  |  |
|  | Mongolia | SADA | 3 | 161 | BOTH | 12 | 4,576 |
| POBG2_50 | Place of birth - |  |  |  |  |  |  |
|  | Mozambique | SADA | 18 | 243 | BOTH | 20 | 4,568 |
| POBG2_51 | Place of birth - |  |  |  |  |  |  |
|  | Nepal | SADA | 11 | 315 | BOTH | 33 | 4,555 |
| POBG2_54 | Place of birth - |  |  |  |  |  |  |
|  | Poland | SADA | 64 | 339 | BOTH | 222 | 4,366 |
| POBG2_55 | Place of birth - |  |  |  |  |  |  |
|  | Oceania Region |  |  |  |  |  |  |
|  | (excluding |  |  |  |  |  |  |
|  | Australia and |  |  |  |  |  |  |
|  | New Zealand) | SADA | 13 | 348 | BOTH | 57 | 4,531 |
| POBG2_56 | Place of birth - |  |  |  |  |  |  |
|  | Portugal | SADA | 52 | 343 | BOTH | 222 | 4,366 |
| POBG2_57 | Place of birth - |  |  |  |  |  |  |
|  | Argentina, Bolivia, |  |  |  |  |  |  |
|  | Paraguay and |  |  |  |  |  |  |
|  | Uruguay | SADA | 19 | 383 | BOTH | 154 | 4,434 |
| POBG2_58 | Place of birth - |  |  |  |  |  |  |
|  | Åland Islands, |  |  |  |  |  |  |
|  | Denmark, Faroe |  |  |  |  |  |  |
|  | Islands, Finland, |  |  |  |  |  |  |
|  | Iceland, Norway |  |  |  |  |  |  |
|  | and Sweden | SADA | 14 | 383 | BOTH | 158 | 4,430 |
| POBG2_59 | Place of birth - |  |  |  |  |  |  |
|  | Namibia and |  |  |  |  |  |  |
|  | Republic of |  |  |  |  |  |  |
|  | South Africa | SADA | 13 | 373 | BOTH | 146 | 4,442 |

Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| POBG2_60 | Place of birth Sri Lanka | SADA | 27 | 342 | BOTH | 93 | 4,495 |
| POBG2_61 | Place of birth Switzerland | SADA | 18 | 387 | BOTH | 165 | 4,423 |
| POBG2_62 | Place of birth - <br> Thailand | SADA | 20 | 379 | BOTH | 138 | 4,450 |
| POBG2_63 | Place of birth Armenia, Azerbaijan and Georgia | SADA | 9 | 324 | BOTH | 48 | 4,540 |
| POBG2_64 | Place of birth Bangladesh, India and Pakistan | SADA | 38 | 364 | BOTH | 161 | 4,427 |
| POBG2_65 | Place of birth - <br> Union of Soviet Socialist Republics, Former | SADA | 20 | 365 | BOTH | 138 | 4,450 |
| POBG2_66 | Place of birth Guernsey, Ireland, Isle of Man, Jersey, Sark and United Kingdom | SADA | 10 | 397 | BOTH | 127 | 4,461 |
| POBG2_67 | Place of birth United States | SADA | 220 | 187 | BOTH | 443 | 4,145 |
| POBG2_68 | Place of birth Venezuela | SADA | 16 | 364 | BOTH | 111 | 4,477 |
| POBG2_69 | Place of birth Gambia, Ghana, Nigeria and Sierra Leone | SADA | 13 | 371 | BOTH | 106 | 4,482 |
| POBG2_70 | Place of birth Benin, Burkina Faso, Côte d'Ivoire, Guinea, Mali, Mauritania, Niger, Senegal and Togo | SADA | 9 | 372 | BOTH | 73 | 4,515 |

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Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| POBG2_71 | Place of birth - |  |  |  |  |  |  |
|  | Bosnia and |  |  |  |  |  |  |
|  | Herzegovina, |  |  |  |  |  |  |
|  | Croatia, Kosovo, |  |  |  |  |  |  |
|  | Republic of |  |  |  |  |  |  |
|  | Macedonia, |  |  |  |  |  |  |
|  | Montenegro, |  |  |  |  |  |  |
|  | Serbia and |  |  |  |  |  |  |
|  | Slovenia | SADA | 55 | 346 | BOTH | 221 | 4,367 |
| POBG3_1 | Place of birth - |  |  |  |  |  |  |
|  | Canada | SADA | 34 | 335 | BOTH | 40 | 4,548 |
| POBG3_2 | Place of birth - |  |  |  |  |  |  |
|  | Caribbean and |  |  |  |  |  |  |
|  | Bermuda | SADA | 154 | 253 | BOTH | 374 | 4,214 |
| POBG3_3 | Place of birth - |  |  |  |  |  |  |
|  | Central Africa | SADA | 27 | 375 | BOTH | 142 | 4,446 |
| POBG3_4 | Place of birth - |  |  |  |  |  |  |
|  | Central America | SADA | 128 | 279 | BOTH | 305 | 4,283 |
| POBG3_5 | Place of birth - |  |  |  |  |  |  |
|  | Eastern Africa | SADA | 120 | 287 | BOTH | 300 | 4,288 |
| POBG3_6 | Place of birth - |  |  |  |  |  |  |
|  | Eastern Asia | SADA | 205 | 202 | BOTH | 739 | 3,849 |
| POBG3_7 | Place of birth - |  |  |  |  |  |  |
|  | Eastern Europe |  |  |  |  |  |  |
|  | (excluding Union |  |  |  |  |  |  |
|  | of Soviet Socialist |  |  |  |  |  |  |
|  | Republics, |  |  |  |  |  |  |
|  | Former) | SADA | 180 | 227 | BOTH | 381 | 4,207 |
| POBG3_8 | Place of birth - |  |  |  |  |  |  |
|  | West Central Asia |  |  |  |  |  |  |
|  | and the Middle |  |  |  |  |  |  |
|  | East | SADA | 191 | 215 | BOTH | 470 | 4,118 |
| POBG3_9 | Place of birth - |  |  |  |  |  |  |
|  | Northern Africa | SADA | 89 | 318 | BOTH | 262 | 4,326 |
| POBG3_10 | Place of birth - |  |  |  |  |  |  |
|  | Northern Europe | SADA | 247 | 160 | BOTH | 461 | 4,127 |
| POBG3_12 | Place of birth - |  |  |  |  |  |  |
|  | Oceania | SADA | 47 | 356 | BOTH | 215 | 4,373 |
| POBG3_14 | Place of birth - |  |  |  |  |  |  |
|  | South America | SADA | 176 | 231 | BOTH | 364 | 4,224 |
| POBG3_15 | Place of birth - |  |  |  |  |  |  |
|  | Southeast Asia | SADA | 249 | 158 | BOTH | 652 | 3,936 |
| POBG3_16 | Place of birth - |  |  |  |  |  |  |
|  | Southern Africa | SADA | 17 | 370 | BOTH | 158 | 4,430 |

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Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| POBG3_17 | Place of birth Southern Asia | SADA | 218 | 185 | BOTH | 647 | 3,941 |
| POBG3_18 | Place of birth Southern Europe | SADA | 182 | 225 | BOTH | 407 | 4,181 |
| POBG3_19 | Place of birth - <br> Union of <br> Soviet Socialist <br> Republics, <br> Former | SADA | 112 | 295 | BOTH | 307 | 4,281 |
| POBG3_20 | Place of birth United States | SADA | 220 | 187 | BOTH | 443 | 4,145 |
| POBG3_21 | Place of birth - <br> Western Africa | SADA | 68 | 339 | BOTH | 211 | 4,377 |
| POBG3_22 | Place of birth Western Europe | SADA | 195 | 212 | BOTH | 449 | 4,139 |
| PPADA (ADA) | Persons in the ADA | ADA | 42 | 4,196 | ... | ... |  |
| PPADACSD | Persons with geographic overlap between census subdivision and ADA | ... | ... | - ... | ADA | 4,881 | 2,051 |
| SNGLDET | Households living in a singledetached house | SADA | 392 | 16 | BOTH | 3,704 | 884 |
| TOTCFAM | Census families | SADA | 333 | 74 | BOTH | 1,672 | 2,916 |
| TOTHHLD | Households | SADA | 408 | 0 | BOTH | 4,588 | 0 |
| TOTPERS | Persons | SADA | 408 | 0 | BOTH | 4,588 | 0 |
| TPERGE15 | Persons aged 15 years and older | SADA | 288 | 120 | BOTH | 4,474 | 114 |
| YRIMD_1900 | Immigrants who landed prior to 1981 | SADA | 360 | 47 | SADA | 376 | 32 |
| YRIMD_1981 | Immigrants who landed from 1981 to 1985 | SADA | 229 | 178 | SADA | 277 | 131 |
| YRIMD_1986 | Immigrants who landed from 1986 to 1990 | SADA | 224 | 183 | SADA | 276 | 132 |
| YRIMD_1991 | Immigrants who landed from 1991 to 1995 | SADA | 238 | 169 | SADA | 293 | 115 |

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Table C. 1
Statistics on the use of calibration constraints, by constraint

| Variable/ constraint | Description | Coverage and non-response adjustment |  |  | Final calibration |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Area | Number of calibrated constraints | Number of excluded constraints | Area | Number of calibrated constraints | Number of excluded constraints |
| YRIMD_1996 | Immigrants who landed from 1996 to 2000 | SADA | 231 | 176 | SADA | 290 | 118 |
| YRIMD_2001 | Immigrants who landed from 2001 to 2005 | SADA | 259 | 148 | SADA | 309 | 99 |
| YRIMD_2006 | Immigrants who landed from 2006 to 2010 | SADA | 258 | 149 | SADA | 310 | 98 |
| YRIMD_2011 | Immigrants who landed from 2011 to 2016 | SADA | 300 | 107 | SADA | 333 | 75 |
| YRIMD_M5 | Persons with no year of immigration | ... | ... | ... | SADA | 239 | 169 |
| YRIMG1_1900 | Immigrants who landed prior to 1981 | SADA | 360 | 47 | SADA | 376 | 32 |
| YRIMG1_1981 | Immigrants who landed from 1981 to 1990 | SADA | 310 | 97 | SADA | 338 | 70 |
| YRIMG1_1991 | Immigrants who landed from 1991 to 2000 | SADA | 320 | 87 | SADA | 350 | 58 |
| YRIMG1_2001 | Immigrants who landed from 2001 to 2010 | SADA | 346 | 61 | SADA | 369 | 39 |
| YRIMG1_2011 | Immigrants who landed from 2011 to 2016 | SADA | 300 | 107 | SADA | 333 | 75 |
| ... not applicable <br> SADA = Super aggregate dissemination area. <br> ADA = Aggregate dissemination area. <br> BOTH = SADA/ADA. <br> Source: Statistics Canada, 2016 Census long-form samp |  |  |  |  |  |  |  |

## Appendix D - Wood Buffalo

## Introduction

On May 1, 2016, a wildfire began southwest of Fort McMurray, Alberta. On May 3, the wildfire swept through the community, destroying many homes and buildings and forcing the largest wildfire evacuation in Alberta's history. Statistics Canada suspended census data collection (referred to as "field data collection") in the evacuated areas. At that time, many responses had already been received from the evacuated areas.

Statistics Canada used a set of measures to ensure that the residents of the Wood Buffalo census subdivision (CSD) (referred to as the specialized municipality of Wood Buffalo or Wood Buffalo) were included in the 2016 Census of Population. Data for the evacuated areas were derived from a combination of sources: direct field data collection and administrative data sources. Statistics Canada worked closely with provincial and local authorities in Alberta to obtain access to administrative records to help validate the data derived from administrative data sources available at Statistics Canada.

This appendix provides additional details about the methodology used in the Wood Buffalo CSD as the result of the wildfire. In particular, it covers weighting and estimation procedures. More details on the collection procedures can be found in Appendix 1.4 (http://www12.statcan.gc.ca/census-recensement/2016/ref/98-304/app-ann1-4-eng. cfm ) of the Guide to the Census of Population, 2016, Catalogue no. 98-304-X.

## Reference date

For the 2016 Census, the reference date for data reporting was May 10, 2016. For residents of the areas evacuated during the wildfire, the reference date was May 1, 2016. This new date was deemed to better represent the situation in the Wood Buffalo CSD. Consequently, a slight discrepancy may have arisen in responses received after the fire began but before field data collection resumed in August. This is considered to be a minor issue.

## Administrative data

Wherever possible and when no direct response had been received for a dwelling, data from various administrative data sources were used with a reference date as close as possible to May 2016, for variables related to a person's name, date of birth, sex and marital status.

Since the administrative data files did not contain the same language content that was collected on the census questionnaire, record linkages between the administrative sources and the 2011 Census database were performed. For successful linkages, the 2011 responses to the language questions were used as proxies for the 2016 language questions. Responses to census questions for which no comparable information could be obtained from administrative data files, such as relationship to Person 1 and common-law status, were derived during data processing.

## Data collection

In August 2016, data collection was reinstated in Wood Buffalo, and census representatives went door to door to complete census questionnaires. Efforts were focused on collecting data for the one in four dwellings included in the long-form questionnaire sample. To further improve data quality, field data collection was performed for dwellings in the areas for which no administrative data were available and for collective dwellings. In areas where enumerators prepared a list of dwellings and delivered census materials (list/leave areas), field data collection was done for all dwellings.

## Processing

If a census response was obtained for the residents of a dwelling, the data collected from that respondent took precedence over any available administrative data. For cases without a census response, data from administrative sources were considered a response to the same extent as a direct response obtained through traditional collection methods during data processing and for the calculation of response rates. The same procedures as described in Chapter 3 were applied to both data direct from respondents and data from administrative sources.

## Data quality

For population and dwelling counts, the Wood Buffalo CSD data went through the same quality assessments as the overall census data. A supplementary pre-validation activity was performed by Statistics Canada once data from field collection and administrative sources were combined. This additional step was done to certify that the alternative methods developed for this exceptional situation provided satisfactory results.

## Estimation

So that a calibration adjustment did not cause an estimate bias, the sample and census data needed to be conceptually comparable. In Wood Buffalo, the collection response rate for the short-form questionnaire (2A) was lower than for the long-form questionnaire (2A-L). This occurred because the objective was, when possible, to complete 2A questionnaires using administrative data sources and to complete 2A-L questionnaires with collection data.

For the 2 A form, $27 \%$ of the data came from direct responses, $67 \%$ came from administrative data sources and $6 \%$ were imputed through the Dwelling Classification Survey (see Chapter 2). For the 2A-L form, $79 \%$ of the data came from direct responses, $8 \%$ came from 2A questionnaire administrative sources, $6 \%$ came from a mix of the 2 A form administrative data sources and the imputed 2A-L forms, $6 \%$ came from 2 A form imputations through the DCS, and $1 \%$ came from 2A form responses only.

Some census concepts are more difficult to derive from administrative data sources, namely families, marital status, common law status and household income characteristics. Therefore, the difference in contribution from administrative data sources to the 2A form and to the 2A-L form resulted in sizeable differences between census counts and sample estimates (before weight adjustments), as expected.

When weighting diagnostics were performed, the weighting distribution was observed to be atypical or bimodal for Wood Buffalo. This issue was found to be the result of more households being of size 1 in the administrative database (a database built by Statistics Canada from administrative sources used to populate the 2 A form) than expected when compared with the census long-form data.

Sizeable discrepancies were also found for characteristics derived from the Indian Register and immigration administrative data linked to census data. In these cases, the differences stemmed in part from the fact that these data can be linked to the administrative database much more easily than to returned questionnaires. This is because the Indian Register and immigration administrative data were used to form the administrative database.

To resolve these issues, certain constraints were deactivated for Wood Buffalo to prevent the introduction of biases in the estimates. The deactivated constraints relate to characteristics involving families, marital status, common law status, household size and household income, and to characteristics derived from the Indian Register and immigration administrative data linked to census records.

The following table shows the differences between the census counts and long-form estimates for those characteristics with large discrepancies. Note that the census counts were largely gathered from administrative sources. This is particularly the case for variables related to the 2A-L form, such as year of immigration, where the so-called census counts were strictly administrative data counts.

Table D. 1
Census counts and long-form estimate differences and discrepancies - Wood Buffalo
$\left.\begin{array}{llrlrr}\hline & \text { Description } & \begin{array}{r}\text { Census counts } \\ \text { (mostly from } \\ \text { administrative } \\ \text { sources) }\end{array} & \begin{array}{r}\text { Estimates } \\ \text { (2A-L) }\end{array} & \begin{array}{r}\text { Relative } \\ \text { difference } \\ \text { (\%) }\end{array} & \text { Difference }\end{array}\right\}$

ADA = Aggregate dissemination area.
Source: Statistics Canada, 2016 Census long-form sample.

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This table lists the constraints both used for and excluded from the calibration process in Wood Buffalo. A value of "Yes" indicates that the variable was used in the process, while a value of "No" indicates that the variable was excluded.

Table D. 2
Calibration constraints excluded in Wood Buffalo

| Variable/Constraint | Excluded in Wood Buffalo |
| :---: | :---: |
| ADULTCF | No |
| AGE00_14 | Yes |
| AGE14 | Yes |
| AGE15_24 | Yes |
| AGE19 | Yes |
| AGE24 | Yes |
| AGE25_34 | Yes |
| AGE29 | Yes |
| AGE34 | Yes |
| AGE35_44 | Yes |
| AGE39 | Yes |
| AGE4 | Yes |
| AGE44 | Yes |
| AGE45_54 | Yes |
| AGE49 | Yes |
| AGE54 | Yes |
| AGE55_64 | Yes |
| AGE59 | Yes |
| AGE64 | Yes |
| AGE65PL | Yes |
| AGE74 | Yes |
| AGE75PL | Yes |
| AGE9 | Yes |
| APTLT5 | Yes |
| CHILD | No |
| CHILDFAM | No |
| COMLAW_YE | No |
| COMLAWNO_DIV | No |
| COMLAWNO_OTHERS | No |
| COMLAWNO_SEP | No |
| COMLAWNO_SINGLE | No |
| COMLAWNO_WID | No |
| COMLAWYE_MARRIED | No |
| COUPLE | No |
| EMPIN_GT50 | Yes |
| EMPIN_LE50 | Yes |
| EMPIN_P0 | Yes |
| EMPIN_P100 | Yes |
| EMPIN_P25 | Yes |
| EMPIN_P50 | Yes |

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Table D. 2
Calibration constraints excluded in Wood Buffalo

| Variable/Constraint | Excluded in Wood Buffalo |
| :---: | :---: |
| EMPIN_P75 | Yes |
| FEMALE | Yes |
| FEMALEGE15 | Yes |
| FEMALELT15 | Yes |
| HHADACSD (ADA) | Yes |
| HHINC_GT50 | No |
| HHINC_LE50 | No |
| HHINC_P100 | No |
| HHINC_P25 | No |
| HHINC_P50 | No |
| HHINC_P75 | No |
| HHSIZE1 | No |
| HHSIZE2 | No |
| HHSIZE3 | No |
| HHSIZE4 | No |
| HHSIZE5 | No |
| HHSIZEGE5 | No |
| HHSIZEGE6 | No |
| INEFAM | No |
| IR_LINK_NO | No |
| IR_LINK_YE | No |
| LIM_NO | No |
| LIM_YE | No |
| LONEPAR | No |
| MALE | Yes |
| MALEGE15 | Yes |
| MALELT15 | Yes |
| MARRIED | Yes |
| NOCLDFAM | No |
| NOINEFAM | No |
| NOTINFAM | No |
| OLN_BI | Yes |
| OLN_EN | Yes |
| OLN_FR | Yes |
| OLN_NO | Yes |
| POBG2_1 | No |
| POBG2_10 | No |
| POBG2_11 | No |
| POBG2_12 | No |
| POBG2_13 | No |
| POBG2_14 | No |
| POBG2_15 | No |
| POBG2_16 | No |
| POBG2_17 | No |

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Table D. 2
Calibration constraints excluded in Wood Buffalo

| Variable/Constraint | Excluded in Wood Buffalo |
| :--- | :--- |
| POBG2_18 | No |
| POBG2_19 | No |
| POBG2_2 | No |
| POBG2_20 | No |
| POBG2_21 | No |
| POBG2_22 | No |
| POBG2_23 | No |
| POBG2_24 | No |
| POBG2_25 | No |
| POBG2_26 | No |
| POBG2_27 | No |
| POBG2_28 | No |
| POBG2_29 | No |
| POBG2_3 | No |
| POBG2_30 | No |
| POBG2_31 | No |
| POBG2_32 | No |
| POBG2_33 | No |
| POBG2_34 | No |
| POBG2_35 | No |
| POBG2_36 | No |
| POBG2_37 | No |
| POBG2_38 | No |
| POBG2_39 | No |
| POBG2_4 | No |
| POBG2_40 | No |
| POBG2_41 | No |
| POBG2_42 | No |
| POBG2_43 | No |
| POBG2_44 | No |
| POBG2_45 | No |
| POBG2_46 | No |
| POBG2_47 | No |
| POBG2_48 | No |
| POBG2_49 | No |
| POBG2_5 | No |
| POBG2_50 | No |
| POBG2_51 | No |
| POBG2_54 | No |
| POBG2_55 | No |
| POBG2_56 | POBG2_57 |

Table D. 2
Calibration constraints excluded in Wood Buffalo

| Variable/Constraint | Excluded in Wood Buffalo |
| :---: | :---: |
| POBG2_6 | No |
| POBG2_60 | No |
| POBG2_61 | No |
| POBG2_62 | No |
| POBG2_63 | No |
| POBG2_64 | No |
| POBG2_65 | No |
| POBG2_66 | No |
| POBG2_67 | No |
| POBG2_68 | No |
| POBG2_69 | No |
| POBG2_7 | No |
| POBG2_70 | No |
| POBG2_71 | No |
| POBG2_8 | No |
| POBG2_9 | No |
| POBG3_1 | No |
| POBG3_10 | No |
| POBG3_12 | No |
| POBG3_14 | No |
| POBG3_15 | No |
| POBG3_16 | No |
| POBG3_17 | No |
| POBG3_18 | No |
| POBG3_19 | No |
| POBG3_2 | No |
| POBG3_20 | No |
| POBG3_21 | No |
| POBG3_22 | No |
| POBG3_3 | No |
| POBG3_4 | No |
| POBG3_5 | No |
| POBG3_6 | No |
| POBG3_7 | No |
| POBG3_8 | No |
| POBG3_9 | No |
| PPADACSD (ADA) | Yes |
| SNGLDET | Yes |
| TOTCFAM | No |
| TOTHHLD | Yes |
| TOTPERS | Yes |
| TPERGE15 | Yes |
| TPERLT15 | Yes |
| YRIMD_1900 | No |

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Table D. 2
Calibration constraints excluded in Wood Buffalo

| Variable/Constraint | Excluded in Wood Buffalo |
| :--- | :--- |
| YRIMD_1981 | No |
| YRIMD_1986 | No |
| YRIMD_1991 | No |
| YRIMD_1996 | No |
| YRIMD_2001 | No |
| YRIMD_2006 | No |
| YRIMD_2011 | No |
| YRIMG1_1900 | No |
| YRIMG1_1981 | No |
| YRIMG1_1991 | No |
| YRIMG1_2001 | No |
| YRIMG1_2011 | No |

## Notes:

Yes indicates the variable was used in the process.
No indicates it was excluded.

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[^0]:    1. Data collection for the Wood Buffalo census subdivision was exempt from this methodology. On May 1, 2016, a forest fire broke out southwest of Fort McMurray, Alberta. Within two days (by May 3), it had ravaged the area, destroying many homes and buildings. Statistics Canada introduced an alternative strategy for the Wood Buffalo census subdivision, which encompasses Fort McMurray, because of the evacuation of residents during the collection period. For more information on the collection method used for the Wood Buffalo census subdivision, see Appendix D.
[^1]:    2. For more information on the history of sampling in Canadian censuses, see Appendix B.
[^2]:    3. Non-respondent sampled households in Wood Buffalo for which administrative data were used (see Appendix D), as well as private dwellings attached to collective dwellings, which by default received only a short-form questionnaire, were considered as having completed their questionnaire for the purposes of short-form response rates. Since they did not complete the long-form questionnaire, they were considered non-respondents for the purposes of calculating long-form questionnaire response rates.
[^3]:    Note: All private households and occupied dwellings are included in the calculation of these response rates, without exception.

[^4]:    4. The selection of constraints specific to the Wood Buffalo area was an exception, since this area had its own process. That is why the selection of constraints for this census subdivision was excluded from this chapter and is presented in Appendix D.
[^5]:    5. Excluded in more than $60 \%$ of the SADAs.
[^6]:    6. More than 1,200 in absolute value.
[^7]:    7. Difference less than 10 in absolute value and discrepancy less than $0.005 \%$.
[^8]:    Notes: Estimates are based on variables related to knowledge of official languages-OLN_BI, OLN_EN and OLN_FR—and their weights were calculated using the responses from before the language data update. A study showed that the use of these unadjusted variables as calibration constraints has a negligible impact on the quality of the weighted estimates. For further information on this impact, see Appendix 1.8 of the Guide to the Census of Population, 2016 (http://www12. statcan.gc.ca/census-recensement/2016/ref/98-304/app-ann1-8-eng.cfm).
    All households with a design weight of one are excluded from the weighting process. These households either come from Indian reserve and canvasser enumeration CUs or are private households attached to collective.
    Source: Statistics Canada, 2016 Census long-form sample.

[^9]:    8. The exception to this characteristic was the units in incompletely enumerated Indian reserves and Indian settlements, which were excluded from the target population and whose weight was set to 0 , without any further modification to the dataset or weights.
[^10]:    9. This definition is not found in the Dictionary, Census of Population, 2016, because the dictionary consists mainly of dissemination terms and this is a collection term.
[^11]:    10. This definition is not found in the Dictionary, Census of Population, 2016, because the dictionary consists mainly of dissemination terms and this is a collection term.
[^12]:    11. See Appendix $D$ for the Wood Buffalo constraints.
