Assessment of the Eastern Scotian Shelf (4VW) Haddock Stock with Projections to 1982

## by

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#### Abstract

Research surveys and commercial catch rates both indicate that this stock has increased dramatically in recent years and is now at the highest level since the surveys began in 1970. Three strong year classes $(1975,1976,1977)$ have been providing substantial recruitment which will continue into 1982. Projections indicate that the 1981 TAC of 23,000 mt should be retained for 1982.

Résumé Les relevés par navire de recherche et les taux de capture commerciaux indiquent que ce stock a augmenté de façon dramatique ces dernières années. II est présentement à son plus haut niveau depuis le début des relevés en 1970. Trois abondantes classes d'âges $(1975,1976,1977)$ ont donné un recrutement substantiel, qui se poursuivra en 1982. D'après les prévisions, le TPA pour 1982 devrait être maintenu au niveau de 1981, soit $23,000 \mathrm{tm}$.


## Introduction

The haddock (Melanograrmus aeglefinus (L.)) population on the Scotian Shelf is separated into two major stocks, one occupying ICNAF Division 4X and the other centered in Division 4 VW with some seasonal migration into 4T (Needler 1930 ; McCracken 1956, 1963; Halliday and McCracken 1970). Nominal catches of haddock from Divisions 4TVW for the period 1954 through 1964 ranged from $22,000 \mathrm{mt}$ to $30,000 \mathrm{mt}$ except for a catch of $38,000 \mathrm{mt}$ in 1959 (Table 1). Following a catch of $56,000 \mathrm{mt}$ in 1965 this fishery declined to a minimum catch of $1,360 \mathrm{mt}$ in 1976 (Figure 1). Recent allocations and reported catches for ICNAF Division 4VW are presented in Table 2. Catches in Division 4T have been well below $1,000 \mathrm{mt}$ during this period, and are not subject to quota restrictions.

The major fluctuations in nominal catches are largely determined by the otter trawl fleet. Nominal catches by longliners increased during the 1970's when stock size was lowest, with the highest and third highest catches of the past two decades reported in 1978 and 1979 (Table 3).

Catches by foreign vessels dominated the fishery in the mid-1960's and have continued to play a significant role. The majority of these foreign catches were bycatches to fisheries directed at other species. The bulk ( $43,000 \mathrm{mt}$ ) of the 1965 peak catch was taken by the USSR in connection with their small-mesh silver hake fishery. Bycatches from this fishery have continued to the present, but under increasingly strict control. Although the reported bycatches have been small in recent years, they represent substantial numbers of one and two year old haddock due to the small mesh size.

The distribution of catches in space and time has undergone marked changes. Prior to 1965 approximately $1 / 3$ of the catches were taken in 4T and 4V. This fraction has been much smaller in recent years, paralleling changes in distribution shown by the research surveys (Table 1, Table 12) (Scott 1976, MS 1980). Under quota regulation the Canadian bottom trawl fishery has changed from primarily a spring fishing season to one in which the peak periods are February-June and November-December (Table 4, Figure 2). Seasonal patterns of the Canadian longliner fishery have changed over the past two decades from predominantly a spring and fall fishery to a summer fishery (Table 5, Figure 2b). Thus although the 1981 TAC is within the historical range of 22,000 to $30,000 \mathrm{mt}$, it will not be comparable in terms of its composition by area or month.

This assessment relies heavily on the summer research surveys. Thus it was found necessary to make a special study of the sources of variation in population estimates obtained from these surveys. These results, while preliminary in nature, show that the variance of survey estimates is an increasing function of abundance, and that the predominant source of variation is large sets of young haddock occurring during the periods 0400 to 0800 and 1600 to 2000. While these observations suggest that survey estimates of total population biomass should be quite reliable, they also indicate that very high variances are associated with survey estimates for the strength of incoming year classes.

## Commercial Removals

The breakdown of commercial removals into numbers at age for this stock is complicated by the lack of sampling for haddock removals associated with the USSR directed fishery for silver hake. For this assessment past practices for estimating the removals by the small mesh fishery were continued. However, the discovery of a less than satisfactory fit of an allometric relation between length and weight for the 1980 summer groundfish survey led to the adoption of new parameters for this relationship for the period 1970 - 1979, necessitating a revision of the catch at age matrix for 1970-1979 (the years 1978, 1979 would have been revised in any case to reflect the final NAFO removal statistics). A further change in the calculation for 1980 was the separation of the bottom trawl landings into spring (January - June) and fall (July - December) components to reflect the recent change in seasonal distribution of landings. This change was not warranted for the period 1976-1979 due to the small numbers of samples from the small total landings in this period.

Overall commercial removals at age were determined separately for three major groups: Canadian bottom trawls (OTB), other Canadian gears (mainly longlines (LL)), and small mesh foreign removals. Once total removals (MT) were determined for each category a table of numbers at age (Table 8a) was calculated using standard procedures, vis:

1) a length-weight relation was obtained from the appropriate summer groundfish survey.
2) for Canadian removals the relevant commercial samples were combined and a table of the length-frequency at each age computed. From this table an average weight at age is determined using the relationship with length. Knowing the age composition and mean weight at age it is possible to determine numbers at age in a given removal. These calculations are performed in the AGE-LENGTH system which resides on the central computer at BIO.
3) for small-mesh removals the age composition is assumed to resemble that of the summer research vessel survey for the year in question. The accuracy of this method is questionable as the research surveys employ different gear with different coverage in space and time, but no better method has been proposed.

Age composition by numbers and by weight for removals are presented in Table 10a.
In view of the serious questions regarding the accuracy of these methods it was considered important to attempt to obtain an independent check on the results. While the possibilities are limited, the results of the checks that have been performed are given on the following page.

## Length-Weight Relations

It has been customary, for this stock, to obtain a relationship between length and weight from the summer research vessel surveys. An allometric relationship

$$
\begin{equation*}
W=A L^{B} \tag{1}
\end{equation*}
$$

where $W$ is the (round) weight ( kg ) and $L$ the fork length ( cm ) is fit to the data by linear regression after taking logarithms. No correction for bias is applied when the parameter $A$ is calculated from the regression estimate of $\log \mathrm{A}$. The accuracy of the resulting fit can be checked by comparing average weights for each length group to the predicted value given by (1). For the summer 1980 survey this comparison shows a major systematic trend in the residuals with a tendency to overestimate the weight of fish in the middle size range ( $30-55 \mathrm{~cm}$ ) corresponding roughly to ages 3-6 in the removals (Figure 4).

The most likely explanation for the trend in Figure 4 is that the sampled population is not homogeneous. A variety of mechanisms may be postulated to explain the lack of homogeneity, including higher fishing mortality on rapidly growing members of a year class and density dependence of growth, but it will require a considerable effort to determine which mechanism is responsible.

The actual values of the parameters used in the assessment are presented in Table 8b. Although the individual values vary substantially from year to year, the resulting curves are quite similar for the predominant length classes. This is a common property of fitted allometric curves and is due to a high correlation between the two parameters (Gould 1966). The limited available data do not permit an analysis of seasonal variation in the length-weight relation. It would, in principle, be preferable to apply different relationships to the spring and fall fisheries, but this would required a change in the priorities given the analysis of the spring and fall research survey data so that the data would be available more quickly.

## Sma11 Mesh Removals

Estimates of small mesh removals for the period 1971-1977 were taken from Waldron (MS 1980). These estimates are based on an adjustment to the nominal catches of haddock by the USSR using by-catch rates obtained from the International Observer Program in 1977 and 1979. Estimates for 1970 and 1978 - 79 (using final NAFO statistics) were calculated using the same procedure (Table 6). For 1980 an estimate of haddock bycatches in directed fisheries for silver hake and squid by 8 foreign countries were provided by the International Observer Program (Table 7).

No information about the composition of haddock by-catches was available, so it was necessary to assume an age composition similar to that of the research surveys. As discussed by Waldron (MS 1980), the adjusted removals for 1970 - 1977 are likely to be underestimates. Althought the tonnages are generally small, these removals are believed to include considerable numbers of young fish. Estimates of recruitment from sequential population analysis will consequently be unreliable.

## Comparisons of Average Weight Estimates

It was not possible to resolve all the questions concerning the appropriate length-weight relationship to be used in estimating removals at age. This relationship is used to determine average weights of fish in the removals. The average weight obtained using the length-weight relationship should, in principle, agree closely with the average obtained by dividing the combined sample weight by the number of fish sampled. Table 9 presents the results of this comparison together with the removals at age used in the previous assessment. The revised estimates for removals in Table 8 are somewhat higher than in previous years, due to the use of different parameters in the length-weight relationship. There is good overall agreement between the average weight used to estimate numbers removed for Table 8 and the average weight of fish sampled (Table 9a).

The largest discrepancy in the table occurred for the fall component of the OTB fishery in 1980. Since the summer is a period of intense feeding for haddock it is reasonable to expect changes in condition factor between the spring and fall fishing seasons to produce fish which are heavier at a given length. Thus the parameters of the length-weight relation from a summer research survey may not be suitable. Further investigation of this hypothesis is required when data from the fall research survey series become available. The 1980 estimates are based on provisional statistics and hence are subject to revision in next year's assessment.

## Research Vessel Surveys

Research vessel surveys have been conducted using the A.T. Cameron in NAFO area 4VW each summer since 1970. In 1978 additional fall and spring surveys were instituted using the Lady Hammond. Data are available from 11 summer surveys and the surveys in November, 1978 and March 1979. Catch numbers and biomass remained low for the years 1970-1976 and have increased in recent years (Tables 12, 13). This increase appears to be confined to the area of Western Bank in NAFO Division 4W (Scott 1976, 1980 unpublished).

The variability associated with research survey estimates of abundance must be taken into account when attempting to use survey results, particularily as the variation increases more rapidly than the estimated
abundance (Table 12). Figure 5 presents the relationship between the weighted mean catch per standard tow $W_{h} \bar{y}_{h}$, and the corresponding weighted variance, $W_{h}^{2} S_{h}^{2} / n_{h}$, in the last four years (this notation is explained in Appendix
1). In 1980, for example, strata 63 and 55 contributed respectively 48 and $17 \%$ of the overall mean, and 74 and $19 \%$ of the overall variance. Within a year there are differences among strata (Figure 6). Within each strata the catches are seldom homogeneous. In 4VW extraordinary sets occur between 4:00 and 8:00 and between 16:00 and 20:00 (Figure 7).

Our overall impression of abundance may be largely influenced by whether the survey happens to hit one, two, or three large sets, which may in turn depend on the time of day that certain strata are sampled. Preliminary examination of these extraordinary sets indicates that they are predominantly small fish. Colton (1965) has suggested that daily vertical migrations of young haddock may cause differences between day and night sets. The data presented here suggest a different pattern. In May, 1981 a special cruise by the Lady Hammond was conducted on Emerald Bank. A total of 15 tows were taken at 3 hour intervals in an area of high stock density. For haddock of less than 23 cm fork length a pronounced bimodal pattern with peaks in the morning and evening was observed (Figure 8). There was no clear pattern in catches of larger fish. These data confirm the patterns found in the research survey data and suggest that the effect will be primarily in the younger age groups. Thus while estimates of fishable biomass from the surveys should be influenced only slightly by diurnal patterns, these may be responsible for much of the variance in recruitment indices obtained from the surveys.

There are references in the literature to differences between night time and day time catches in haddock and other groundfish species (Woodhead 1961; Parish et al 1963). These references indicate a pattern consistent with that found in $4 X$ haddock (Figure 7b), but the pattern observed in 4VW haddock (Figure 7a, Figure 8) has not been reported for other haddock stocks. Thus the inability of the research surveys to predict recruitment in 4 VW haddock may not carry over to other areas and stocks.

## Recruitment

Recruitment in haddock is notoriously variable and generally year class strength has a bimodal distribution (Hennemuth et a1. 1981). Catch projections in haddock are sensitive to assumptions about year class strength, except when there are a number of strong year classes entering together. The research surveys are commonly used to provide an indication of recruitment as an aid to projections. However, even when spring and fall data are considered, it is difficult to quantify the size of incoming year classes. Examination of Figure 9 shows that while the 1978 year class is clearly very poor, there is no discernible difference between the 1975-77 year classes. Only with many years of data is it possible to recognize that the 1974 year class was weak. This apparent weakness may have been due to high fishing mortalities from small mesh gear.

The high catch rates for small ( $<23 \mathrm{~cm}$ ) haddock obtained in the 1981 special cruise of the Lady Hammond indicate that the 1980 year class is strong. This observation is supported by high by-catch rates for small haddock reported by the International Observer Program in June, 1981.

## Weights at Age

Mean weights at age for the population are required for projections and in order to make comparisons of population biomass obtained from sequential population analysis and from research surveys. Because commercial gear will tend to select the larger fish at a given age, mean weight at age for the population will not, in general, be the same as that in the removals. However, since projections require mean weights in the removals, estimates of mean weight at age should generally follow commercial patterns. For this assessment the mean lengths at age in the commercial removals were used as a basis for fitting a von Bertalanfly growth curve

$$
\begin{equation*}
L=L_{\infty}\left(1-e^{k\left(t_{0}-t\right)}\right) \tag{2}
\end{equation*}
$$

using (weighted by $\frac{1}{\text { sample size) nonlinear regression. The weighting is }}$ required to reduce the influence of high variances associated with larger fish due to the small sample sizes. Separate growth curves were obtained for each category of removal (e.g., OTB, LL, and small mesh) and converted to weight at age using the appropriate allometric relationship. These relationships were then evaluated at various ages and combined to obtain the table of weights at age (Table 15). To determine how representative these weights were for the population a comparison was made between the observed total biomass from the research surveys and an estimate of biomass obtained by applying the mean weights at age to the research survey population estimates. While there is evidently some bias present (Figure 10), the overall accuracy of this procedure was adequate for comparisons between survey biomass estimates and estimates obtained by sequential population analysis.

## Sequential Population Analysis

Traditional sequential population analysis depends heavily on the stabilization of numbers towards the early years of a cohort as cumulative fishing mortality increases to overcome the potential pitfalls of:

- subjectivity, particularly with regard to partial recruitment patterns sought when adjusting terminal f's,
- the types of adjustments made, and the
- excess of parameters over the number of data points when calibrating against catch rates for age aggregated data.

Calibration of sequential population analysis against catch rate data (such as that provided by the research vessel surveys) presents major technical difficulties. Research vessel survey estimates for the 4VW stock indicate that abundance has increased substantially in the past few years. The variance of these estimates increases roughly as the square of abundance, so that a regression line need not pass near points where abundance is large. Since most of the recent increase in population is concentrated in the younger ages, and since small fish are expected to be more available to the research gear, it is desirable to calibrate younger ages separately, using a regression model which allows for errors in both dependent and independent variables.

The SURVIVOR method (Doubleday MS 1980) implemented by Rivard (1980) (Appendix 2) overcomes some of the central difficulties to calibrating SPA against catch rate data. It seeks calibration constants of the form

$$
\begin{equation*}
N(I, J)=K(I) A(I, J) \tag{3}
\end{equation*}
$$

where $N$ is SPA abundance for age $I$ in year $J$, $A$ is a catch rate, and $K$ the slope which is determined by regression of $\log N$ against Log A. A separate value of $K$ may be obtained for each age in a "calibration block". Typically one finds that $K$ levels off above a certain age. This allows several ages to be combined to compute a fully recruited calibration constant. In 4VW haddock the age at which $K$ levels off was determined from trial runs to be 4. This reflects the application of research survey age composition for small mesh fisheries and is lower than the value obtained in 4 X haddock. As the importance of small mesh removals is decreasing, this age may need to be revised upwards in the future.

Any new technique must be thoroughly tested before it can form the basis of an assessment. Thus we will present the results both of a conventional sequential population analysis and an analysis based on the "SURVIVOR" method. Fortunately the two methods agree closely. The "SURVIVOR" method will be discussed first as it may be unfamiliar to the reader.

## Calibration

The results obtained from the SURVIVOR program will depend on the calibration block that is chosen. Because the program uses a logarithmic transformation it is necessary to exclude ages in which commercial or research catches are zero. This limited the range of ages which could be used to 2 - 8, so conventional procedures were used to extend the analysis to older ages.

Preliminary trials showed that the best results were obtained using a calibration block restricted to ages 2-5 and that the magnitude and variance of the calibration constants increased as the block size was decreased by removing the most recent year from the block (Table 16). With a block including 1970 - 1977 the largest coefficient of variation for estimated survivors in 1980 was associated with an outlier at age 7 . When this point is ignored the greatest variability is associated with age 2.

High variability is expected at age 2 in both commercial removals (due to inability to adequately account for small mesh removals) and in the research survey abundance index (due to diurnal variation in catchability of young haddock in 4 VW ). The coefficients of variation became excessive when block size was restricted to years prior to 1977. Thus a calibration block based on ages $2-5$ and the years 1970-1977 was chosen as the best compromise between the coefficients of variation in K for 1980 and loss of the stabilizing effect of sequential population analysis. The results are presented in Table 17.

## Extension to 01der Ages

Initial guesses for $F$ values in the last year were obtained from the estimated survivors by solving the cohort formula for fishing mortality as a function of catches and survivors in the last year. It was assumed that partial recruitment would be sigmoidal. This was consistent with the results of the SURVIVOR program.

F values for the oldest ages were calculated by the usual iterative procedure based on cohort analysis in which new values were obtained as the average over three fully recruited ages. This algorithm converged in two or three iterations and produced similar patterns for various choices of the fully recruited ages. The results are presented in Table 18.

This extension allows comparison with other indices of abundance, but does not alter the conditions in the final year. The relationship (3) between cohort and research numbers is also maintained without a change in the value of $K$ (Figure 11).

## Conventional Sequential Population Analysis

A conventional sequential population analysis was performed starting with the same partial recruitment as the previous year's assessment. Fine tuning was based on a comparison of research $1+2$ numbers to sequential population analysis age 1 and research to sequential 3+ numbers. It was found necessary to adjust the partial recruitment at age 3 from 0.18 to 0.2 to obtain a good relationship between the generated population and research indices. The best overall relationships were obtained using a terminal F of 0.25 (Figure 12). In comparison, the "SURVIVOR" based method yielded a terminal F of 0.55 .

The ability of the relationships used to discriminate between different choices of input parameters was poor despite the fact that the points used in the regression are distributed fairly evenly over the observed range of abundances. This was due to the presence of three strong year classes which make up the recent increases in abundance. By adjusting the relative strengths of these year classes it was possible to obtain good relationships over a wide range of terminal $F$ values.

## Reconciliation

The two approaches used to determine a 1980 fishing mortality produce different results. This is seen clearly by considering 2+ biomass (Figure 13). The traditional SPA produces the best agreement for the most recent year, while the "SURVIVOR" based estimates are consistently low, but better reflect the overall pattern. In essence, the traditional method relies heavily on the most recent year's survey. Examination of the population matrix for the traditional SPA (Table 19) shows that the 1977 year class is considerably larger than any other in the series. This may be an overestimate. However, it is also possible that underestimates of small mesh removals during the early part of the decade have produced an unrealistic view of year class strength in the early 1970's. Because research survey data are highly variable at these ages it is, however, unwarranted to conclude that the 1977 year class is exceptional on the basis of the large stratified mean catch per tow at age 1 in 1978.

It is possible to adjust the "SURVIVOR" based results to correct for the apparent bias by multiplying the population in 1974 - 1980 by a correction factor equal to the geometric mean of the ratio of survey to SPA 2+ biomass, which yields an estimate for the $2+$ biomass in 1980 of 113,000 mt (Table 20). This correction factor can be interpreted as an average factor which would have to be applied to the catch matrix for 1974-1980 to eliminate the bias evident in Figure 13.

## Projections

The 4VW haddock stock is still in a state of flux, making it impossible to determine equilibrium conditions of weight at age, maximum age, and partial recruitment. Two options are available:

1) use historical partial recruitment and make a (subjective) adjustment to compensate for the reduction in small mesh removals, or
2) use a partial recruitment based on more recent data with the understanding that a large uncertainty is associated with the values so obtained due to lack of convergence of the sequential population analysis. Previous assessments have used an age range of 2-11; under equilibrium conditions a range of 2-15 would be more appropriate.

In order to determine the effect that the indicated changes would have, projections were made using both the historical data and a revised partial recruitment with reflects the recent changes in the fishery. The weights at age for both projections were calculated as the average of the 1978 1980 values.

The change in mean weight at age from that used in the previous assessment was minor, and $F_{0.1}$ remained at 0.3 using this historical partial recruitment (Table 21). A change to the age range 2-15 produced only a modest reduction in $\mathrm{F}_{0.1}$ to 0.28 .

Projections were carried out using the input parameters shown in Table 21. In view of the uncertainties associated with the revised partial recruitment it was decided to bracket the answer that would be obtained using a correction factor of 1.3 by applying factors of 1.2 and 1.4 to the 1980 population of Table 19. No attempt was made to estimate the number of fish at ages $12-15$, so the projections are marginally lower than would be obtained using all ages. It was decided to leave the 1980 removals at their reported level as the increased quota in 1980 would be expected to reduce the level of misreporting below the 1974-1980 average.

Considering the major differences in the input parameters it is encouraging that the projections are all consistent with a TAC of $23,000 \mathrm{mt}$. While it is too soon to say precisely what the equilibrium conditions for this stock will be, these results suggest that the transition to equilibrium conditions will not require major changes in quota advice. It should be noted that the main source of differences between these projections is the sizes of the 1975 - 1977 year classes. While it is difficult to fix the size of any one year class with precision, the fact that recruitment occurs over a period of several years allows an overestimate of one year class to be compensated by an underestimate in an adjacent year.

## Discussion

The 4VW haddock stock has responded dramatically to the imposition of quota restrictions and by-catch regulation of small mesh fisheries. While there is qualitative agreement between commercial catch rates (Table 11) and the research vessel surveys for the fishahle stock, (Figure 3), the surveys provide the only means of estimating recruitment. Unfortunately it was shown that the surveys do not provide quantitative measures of recruitment. Not only are they influenced by diurnal changes in availability of young fish, but the uncertainties over small mesh removals in the early 1970's make calibration of the surveys chancy.

The 1982 projected catches are determined essentially by the strengths of the 1975 - 1977 year classes, and thus are subject to errors resulting from our inability to accurately estimate year class strength. For this assessment the presence of three strong year classes provides some protection from errors in estimating the strength of a single year class. This happy situation is not likely to persist, and it is therefore essential to develop improved methods for determining year class size. In this stock the utility of historical data is seriously reduced by the uncertainties over commercial removals. Thus it may be some time before a recruitment index can be calibrated against sequential population analysis.

In the face of these uncertainties it is very encouraging to find that radically different approaches to the analysis lead to essentially identical results. There is every reason to expect the quality of the data, particularly that related to commercial removals, to be better for the 1980's than the 1970's when removals are almost certainly underestimated.

There is no doubt that the 4 VW haddock stock is presently at the highest level since 1970. It remains to be seen, however, whether the stock will return to its historical distribution as biomass approaches the level prevalent in the 1950 's. Growth rates, as indicated by mean weight at age (Table 15), appear to be higher now than in the early 1970's. If these are not an artifact of higher fishing mortality on the larger fish at a given age then increased stock biomass should be accompanied by a decline in growth rates. Thus this stock may remain in a state of flux for some years to come.

## Acknow Tedgments

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Appendix 1. Standard Notation for use in Groundfish Research Trawl Surveys
$\mathrm{L} \quad=$ number of strata sampled
$N_{h} \quad=$ total number of sample units in the hth stratum
$\begin{aligned} n_{h} \quad= & \text { total number of units sampled in the hth stratum } \\ & \left(h=1,2, \ldots, L_{h}\right)\end{aligned}$
$N=\sum_{h=1}^{L} \quad N_{h}=$ total number of sample units in survey
$n=\sum_{h=1}^{L} \quad n_{h}=$ total number of observations in survey
$W_{h}=\frac{N_{h}}{N} \quad=$ stratum weight
$f_{h}=\frac{n_{h}}{N_{h}} \quad=$ sampling fraction in the hth stratum
$y_{\text {hi }} \quad=$ ith observation in the hth stratum $\left(1=1,2, \ldots, n_{h}\right)$
$\bar{y}_{h}=\sum_{i=1}^{n_{h}} y_{h_{i} / n_{h}} \quad=$ sample mean in the hth stratum
$s_{h}^{2}=\sum_{i=1}^{n_{n}},\left(y_{h i}-\bar{y}_{h}\right)^{2} / n_{h} \quad=$ sample variance in the $h$ th stratum
L
$\bar{y}_{s t}=\sum_{h=1} W_{h} \bar{y}_{h} \quad=\begin{aligned} & \text { estimate of the population mean per unit } \\ & \text { (i.e. stratified mean catch per tow). }\end{aligned}$

A parametric value of the variance of an estimator may be denoted either by Greek letters, i.e.,

$$
\text { st, or by the following convention } \operatorname{VAR}\left(\bar{y}_{s t}\right)
$$

## Appendix 2. SURVI VOR

The "SURVIVOR" program used in this document differs from that originally described by Rivard (1980) in two ways:

1) the original program contained an error in the loop which computes integrated catch. The result of this error was that the final term in the summation was omitted in the original program.
2) the calculation of residuals was changed to agree with its formula in Rivard (1980). The original program introduced a small bias in mean of the residuals. This had no effect on the calculation of estimated survivors, but could influence the interpretation of the results, particularly the effect of changing the calibration block.

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Table 1. Nominal catches (mt) of eastern Scotian Shelf haddock (4TVW) by NAFO Division and country, as reported by NAFO.

(1) Provisional $4 / 81$
(2) Quotas
(3) Catches for 1953-58 are for 4 Vn and 4 Vs combined.

Table 2. Recent fishery allocations and the respective reported catch for 4VW haddock. (Metric Tons)

| Year | Vessel Size | Allocation | Reported Catch | \% of Allocation |
| :--- | :---: | :---: | :---: | :---: |
|  |  |  |  |  |
| 1975 | all vessels | - | - | - |
| 1976 | all vessels | 1,250 | $1,263^{1}$ | 101 |
| 1977 | all vessels | 1,700 | $3,086^{1}$ | 182 |
| 1978 | all vessels | $1,700^{*}$ | $4,459^{1}$ | 262 |
| 1979 | all vessels | $1,700^{*}$ | $2,915^{1}$ | 171 |
| 1980 | $>125 \mathrm{ft}$. | 11,500 | $10,522^{2}$ | 91 |
|  | $<125 \mathrm{ft}$. | 3,400 | $2,105^{2}$ | 62 |
| 1981 | $>125 \mathrm{ft}$. | 16,000 | $11,915^{3}$ | 74 |
|  | $<125 \mathrm{ft}$. | 4,500 | $542^{3}$ | 12 |

* Closed Fisheries
${ }^{1}$ Maritimes and Quebec
${ }^{2}$ Maritimes
${ }^{3}$ At Tantic Quota Report - 29 Apri 1981

Table 3. Nominal catches ( mt ) of eastern Scotian Shelf haddock in $4 V$ and $4 W$ (4TVW) by gear type for Canada (M, Q \& Nfld.) as reported by NAFO.

| Year | Otter <br> Trawler | Longliner | Danish Seiner | Miscellaneous | Tota 1 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1960 | 20835 | 1077 | 23 | 696 | 22631 |
| 1961 | 22060 | 448 | 52 | 1377 | 23937 |
| 1962 | 16453 | 665 | 76 | 705 | 17899 |
| 1963 | 11943 | 511 | 147 | 526 | 13127 |
| 1964 | 10679 | 70 | 62 | 874 | 11685 |
| 1965 | 8033 | 352 | 66 | 160 | 8611 |
| 1966 | 10222 | 233 | 19 | 130 | 10604 |
| 1967 | 7855 | 126 | 25 | 573 | 8579 |
| 1968 | 8819 | 296 | 16 | 364 | 9495 |
| 1969 | 8603 | 289 | 30 | 341 | 9263 |
| 1970 | 5056 | 479 | 20 | 262 | 5817 |
| 1971 | 8709 | 538 | 77 | 179 | 9503 |
| 1972 | 2141 | 528 | 76 | 138 | 2883 |
| 1973 | 2459 | 628 | 28 | 232 | 3347 |
| 1974 | 543 | 493 | 17 | 162 | 1215 |
| 1975 | 593 | 873 | 10 | 82 | 1558 |
| 1976 | 383 | 657 | 10 | 75 | 1125 |
| 1977 | 2198 | 729 | 26 | 170 | 3123 |
| 1978 | 4009 | 1069 | 67 | 364 | 5509 |
| 1979 | 1745 | 1232 | 64 | 142 | 3183 |
| $1980{ }^{1}$ | 13073 | 899 | 169 | 320 | 14535 |

Table 4. Nominal catches (mt) of eastern Scotian Shelf haddock in 4 V and 4 W by month by Canadian (MQ) otter trawlers, as reported to NAFO

| Year | Jan | Feb | Mar | Apri1 | May | June | July | Aug | Sept | Oct | Nov | Dec | Total |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1960 | 578 | 3372 | 4827 | 1328 | 1177 | 597 | 1427 | 1678 | 1543 | 1199 | 1665 | 1442 | 20833 |  |
| 1961 | 1387 | 2761 | 5029 | 6605 | 1389 | 324 | 508 | 489 | 859 | 927 | 1022 | 488 | 21788 |  |
| 1962 | 626 | 1863 | 4749 | 2401 | 1164 | 615 | 954 | 1079 | 1015 | 739 | 654 | 449 | 16308 |  |
| 1963 | 664 | 236 | 388 | 4444 | 1357 | 645 | 844 | 1079 | 1004 | 434 | 659 | 237 | 11991 |  |
| 1964 | 406 | 1531 | 1473 | 1557 | 1155 | 378 | 688 | 1082 | 804 | 359 | 342 | 638 | 10413 |  |
| 1965 | 347 | 819 | 1005 | 1114 | 986 | 350 | 1563 | 644 | 109 | 206 | 338 | 363 | 7844 |  |
| 1966 | 369 | 463 | 3301 | 1821 | 2151 | 264 | 247 | 138 | 136 | 63 | 262 | 101 | 9316 |  |
| 1967 | 198 | 294 | 4038 | 800 | 258 | 85 | 263 | 237 | 100 | 526 | 661 | 187 | 7647 |  |
| 1968 | 254 | 546 | 3302 | 782 | 730 | 901 | 602 | 114 | 317 | 391 | 650 | 408 | 8997 |  |
| 1969 | 888 | 1183 | 3108 | 1472 | 852 | 183 | 132 | 106 | 61 | 117 | 81 | 349 | 8532 |  |
| 1970 | 425 | 480 | 1436 | 1459 | 141 | 86 | 398 | 110 | 74 | 78 | 115 | 227 | 5029 |  |
| 1971 | 408 | 772 | 4740 | 1946 | 147 | 225 | 47 | 39 | 16 | 20 | 32 | 200 | 8592 |  |
| 1972 | 103 | 90 | 1022 | 280 | 105 | 221 | 19 | 56 | 26 | 18 | 49 | 128 | 2117 |  |
| 1973 | 93 | 155 | 1218 | 313 | 150 | 282 | 4 | 2 | 23 | 16 | 32 | 107 | 2395 |  |
| 1974 | 45 | 78 | 58 | 20 | 24 | 103 | 18 | 43 | 35 | 28 | 30 | 40 | 522 |  |
| 1975 | 25 | 71 | 68 | 124 | 65 | 20 | 85 | 9 | 40 | 34 | 20 | 27 | 588 |  |
| 1976 | 15 | 1 | 18 | 39 | 76 | 102 | 4 | 32 | 17 | 22 | 13 | 42 | 381 |  |
| 1977 | 44 | 90 | 79 | 57 | 217 | 37 | 49 | 114 | 184 | 180 | 297 | 796 | 2144 |  |
| 1978 | 118 | 151 | 669 | 1121 | 193 | 25 | 124 | 113 | 58 | 62 | 226 | 66 | 2926 |  |
| 1979 | 26 | 76 | 157 | 43 | 357 | 136 | 120 | 112 | 45 | 110 | 193 | 268 | 1643 |  |
| 19807 | 107 | 1165 | 2391 | 1099 | 316 | 486 | 744 | 676 | 411 | 1108 | 1444 | 1466 | 11413 |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

${ }^{1}$ Preliminary

Table 5. Nominal catches (mt) of eastern Scotian Shelf haddock in 4 V and 4 W by month by Canadian (MQ) longliners, as reported to NAFO

| Year | Jan | Feb | Mar | April | May | June | July | Aug | Sept | Oct | Nov | Dec | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1960 | 50 | 122 | 76 | 66 | 18 | - | 133 | 80 | 172 | 25 | 190 | 145 | 1077 |
| 1961 | 3 | 36 | 23 | 35 | 6 | 1 | 8 | 13 | 63 | 64 | 159 | 81 | 492 |
| 1962 | 13 | 1 | 74 | 2 | 7 | 6 | 25 | 33 | 67 | 145 | 206 | 86 | 665 |
| 1963 | 25 | 3 | 4 | 49 | 9 | 17 | 26 | 30 | 49 | 85 | 68 | 52 | 417 |
| 1964 | 3 | - | 3 | 5 | 3 | - | - | - | - | 23 | 12 | 22 | 71 |
| 1965 | 17 | 41 | 27 | 65 | 23 | 10 | 5 | 23 | 28 | 39 | 53 | 21 | 352 |
| 1966 | - | 24 | 71 | 11 | - | - | 7 | 12 | 9 | 52 | 30 | 17 | 233 |
| 1967 | 3 | 1 | 19 | 10 | - | 3 | 7 | 5 | 15 | 29 | 25 | 9 | 126 |
| 1968 | 10 | 19 | 17 | 42 | 10 | 10 | 12 | 42 | 42 | 49 | 38 | 6 | 297 |
| 1969 | 1 | 1 | 8 | 8 | 4 | 9 | 25 | 56 | 39 | 68 | 53 | 17 | 289 |
| 1970 | 19 | 4 | 43 | 22 | 12 | 12 | 25 | 57 | 120 | 110 | 40 | 15 | 479 |
| 1971 | - | 14 | 12 | 33 | 18 | 26 | 94 | 61 | 106 | 107 | 38 | 29 | 538 |
| 1972 | - | $\stackrel{-}{-}$ | 3 | 9 | 17 | 26 | 102 | 88 | 73 | 111 | 81 | 18 | 528 |
| 1973 | 1 | 6 | 115 | 149 | 47 | 40 | 39 | 62 | 56 | 78 | 59 | 17 | 669 |
| 1974 | 10 | 4 | 16 | 20 | 27 | 44 | 74 | 78 | 59 | 71 | 63 | 27 | 493 |
| 1975 | 31 | 37 | 69 | 78 | 93 | 81 | 74 | 138 | 88 | 105 | 57 | 24 | 875 |
| 1976 | 20 | 36 | 93 | 113 | 71 | 56 | 106 | 85 | 72 | 70 | 57 | 17 | 796 |
| 1977 | 15 | 33 | 55 | 36 | 42 | 86 | 65 | 92 | 72 | 116 | 100 | 34 | 746 |
| 1978 | 31 | 63 | 78 | 104 | 121 | 116 | 175 | 166 | 105 | 53 | 49 | 8 | 1069 |
| 1979 | 5 | 18 | 123 | 109 | 129 | 110 | 148 | 215 | 142 | 94 | 109 | 30 | 1232 |
| 1980 | 4 | 2 | 36 | 98 | 103 | 111 | 125 | 189 | 98 | 89 | 32 | 12 | 899 |

${ }^{1}$ Preliminary

Table 6. Derivation of adjusted nominal catches (mt) of eastern Scotian Shelf haddock (4TVW) for 1971-79 by USSR, as determined by International Observer Program.

\begin{tabular}{|c|c|c|c|c|c|}
\hline Year

1 \& | Nominal Catch of 4 VW S. Hake in directed $S$. Hake fishery ${ }^{1}$ |
| :--- |
| 2 | \& Nominal Catch of 4 VW Haddock in directed $S$. Hake fishery ${ }^{1}$ 3 \& Nominal Catch of 4VW Haddock in all USSR fisheries ${ }^{2}$ 4 \& Estimated Catch of 4 VW Haddock in S. Hake directed fishery ${ }^{2}$ 5 \& Adjusted Catch of 4VW Haddock in all USSR fisheries (Colums $(4-3)+5)$ <br>

\hline 1970 \& 164013 \& 670 \& 670 \& 1230 \& 1230 <br>
\hline 1971 \& 122413 \& 279 \& 475 \& 918 \& 1114 <br>
\hline 1972 \& 107969 \& 102 \& 106 \& 810 \& 814 <br>
\hline 1973 \& 268511 \& 76 \& 76 \& 2014 \& 2014 <br>
\hline 1974 \& 81181 \& 91 \& 132 \& 609 \& 650 <br>
\hline 1975 \& 95298 \& 52 \& 52 \& 715 \& 715 <br>
\hline 1976 \& 70054 \& 18 \& 24 \& 525 \& 531 <br>
\hline 1977 \& 20145 \& 10 \& 14 \& 121 \& 125 <br>
\hline 1978 \& 41243 \& 132 \& 142 \& 371 \& 381 <br>
\hline $1979^{3}$ \& 43936 \& ? \& 104 \& 395 \& 395 <br>
\hline
\end{tabular}

1 As reported to NAFO
2 As estimated from Observer Program
${ }^{3}$ Reported in NAFO STATLANT Output, March 1981.

Table 7. Derivation of adjusted nominal catch (mt) of eastern Scotian Shelf haddock (4TVW) for 1980 by foreign countries, as determined by International Observer Program.

| Country | Directed Species $2$ | Nominal Catch of 4VW Haddock in all fisheries ${ }^{1}$ $3$ | Nominal Catch of 4VW Haddock in all fisheries ${ }^{2}$ | Nominal Catch of 4VW Haddock in directed fisheries ${ }^{3}$ | Nominal Catch of Directed species ${ }^{3}$ $6$ | Nominal Catch of Directed species ${ }^{2}$ $7$ | $\begin{gathered} \text { Ratio } \\ (5 \div 6) \end{gathered}$ | Estimated bycatch of 4VW Haddock in directed fishery $(8 \times 7)$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Bulgaria | S. H. | - | 3.22 | 2.07 | 516.4 | 687.4 | 0.004 | 2.75 |
| Cuba | S. H. | 20 | 19.27 | 14.60 | 1105.7 | 2110.6 | 0.013 | 27.44 |
| France | Sq. | - | 0.41 | 1.13 | 366.9 | 364.0 | 0.003 | 1.09 |
| Japan | Sq. | 10 | 5.84 | 29.16 | 8115.8 | 15835.7 | 0.004 | 63.34 |
| Poland | Sq. | - | - | - | 488.5 | 488.8 | - | ~ |
| Portuga 1 | Sq. | - | 1.88 | 3.60 | 931.3 | 1646.7 | 0.004 | 6.59 |
| Spain | Sq. | - | 1.44 | 1.07 | 869.3 | 4817.0 | 0.001 | 4.82 |
| USSR | S. H. | 247 | 161.6 | 202.90 | 22990.1 | 38450.2 | 0.009 | 346.05 |
| Total |  | 277 | 193.7 | 254.53 | - | - | - | 452.08 |

[^0]2. As reported to FLASH
${ }^{3}$ As reported to International Observer Program


Table 8b. Parameters used to calculate removals at age

| PERIOD | A | B | RESEARCH SURVEY |  |
| :--- | :---: | :---: | :---: | :---: |
| $1970-1975$ | 0.00885 | 3.039 | Summer | 1975 |
| 1976 | 0.00553 | 3.160 | $"$ | 1976 |
| 1977 | 0.00909 | 3.041 | $"$ | 1977 |
| 1978 | 0.00580 | 3.152 | $"$ | 1978 |
| 1979 | 0.00640 | 3.126 | $"$ | 1979 |
| 1980 | 0.00369 | 3.268 | $"$ | 1980 |

Table 9a. Check on average weights used to calculate removals at age for Table 8.

| Period |  | Gear | Sample Wt. (kg round) ${ }^{1}$ | $\begin{aligned} & \text { Number in } \\ & \text { Sample } \end{aligned}$ | Average <br> (A) of Sample | (kg round) <br> (B) Determined Using Length-Weight Relation | Ratio A/B |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Jan-June | '80 | OTB | 9,152 | 6,668 | 1.37 | 1.406 | 0.98 |
| July-Dec | '80 | ОТВ | 9,062 | 7,104 | 1.28 | 1.098 | 1.16 |
| Jan-Dec | '80 | LL | 2,497 | 1,538 | 1.62 | 1.638 | 0.99 |
| Jan-Dec | '79 | 0TB | 3,084 | 2,169 | 1.42 | 1.380 | 1.03 |
|  | '79 | LL | 4,105 | 2,766 | 1.48 | 1.448 | 1.02 |
| " | '78 | OTB | 9,547 | 7,688 | 1.24 | 1.282 | 0.97 |
| " | '78 | LL | 2,127 | 1,289 | 1.65 | 1.638 | 1.01 |
| " | '77 | отв | 5,857 | 4,506 | 1.30 | 1.236 | 1.05 |
| " | '76 | отв | 1,429 | 1,727 | 0.83 | 0.799 | 1.04 |
| " | ${ }^{7} 75$ | OTB | 1,480 | 1,273 | 1.16 | 1.254 | 0.93 |
| " | '74 | OTB | 435 | 364 | 1.20 | 1.183 | 1.01 |
| " | '73 | отв | 2,613 | 1,898 | 1.38 | 1.364 | 1.01 |
| " | '72 | OTB | 2,686 | 1,868 | 1.44 | 1.482 | 0.97 |
| " | '71 | отв | 10,040 | 7,245 | 1.39 | 1.477 | 0.94 |
| " | '70 | отв | 3,119 | 2,172 | 1.44 | 1.378 | 1.04 |

[^1]Table 9b. PREVIOUS REMOVALS AT AGE ('000) FOR THE 4VW HADDOCK STOCK 2/7/81

|  | 1970 | 1971 | 1972 | 1973 | 1974 | 1975 | 1976 | 1977 | 1978 | 1979 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 231 | 179 | 453 | 290 | 26 | 318 | 348 | 36 | 105 | 1 |
| 2 | 85 | 430 | 100 | 1048 | 216 | 36 | 484 | 127 | 182 | 198 |
| 3 | 443 | 586 | 447 | 439 | 624 | 319 | 124 | 832 | 565 | 567 |
| 4 | 1069 | 1490 | 505 | 1050 | 158 | 572 | 195 | 363 | 1432 | 1322 |
| 5 | 886 | 1890 | 619 | 638 | 269 | 218 | 253 | 430 | 299 | 650 |
| 6 | 888 | 835 | 429 | 560 | 200 | 140 | 140 | 201 | 499 | 125 |
| 7 | 768 | 647 | 225 | 240 | 71 | 49 | 122 | 66 | 143 | 98 |
| 8 | 238 | 818 | 124 | 127 | 27 | 23 | 27 | 38 | 42 | 7 |
| 9 | 53 | 194 | 101 | 48 | 7 | 6 | 6 | 12 | 19 | 4 |
| 10 | 22 | 26 | 11 | 75 | 8 | 3 | 8 | 7 | 6 | 1 |
| 11 | 10 | 14 | 11 | 1 | 8 | 1 | 7 | 6 | 5 | 1 |

${ }^{1}$ CAFSAC Res. Doc. 80/61

Table 10a. Percent age composition (numbers) for commercial removals.

| 1 | 1 | 5 | 3 | 7 | $6:$ | ? | 18 | 36 | i | 3 | $\bigcirc$ | $\bigcirc$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | 1 | 2. | 7 | 5 | 20 | 9 | 3 | 3.4 | 6 | 3 | 5 | 0 |
| 3 | 1 | 9 | 8 | 1.5 | 10 | 40 | 19 | 8 | 39 | 13 | 17 | 20 |
| 4 | 1 | 23 | 21 | 17 | 25 | 10 | 3.1 | 3 l | i6 | 个6 | 43 | 36 |
| 5 | 1 | 19 | 26 | 22 | i5 | 19 | 13 | 13 | 21 | 9 | 24 | 26 |
| 6 | 1 | 19 | 17. | 16 | 1.3 | 12 | 9 | 8 | 1.0 | 1.5 | 6 | 14 |
| 7 | 1 | 16 | 9 | 8 | 6 | 5 | 3 | 6 | 4 | 4 | 5 | 2 |
| 8 | 1 | J | 11 | 5 | 3 | 2 | ; | 2 | 2 | i | i | 1 |
| 9 | 1 | 1. | 3 | 4 | 1. | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 10 | 1 | 1 | 0 | 0 | 2 | i | 0 | 0 | 0 | 0 | 0 | 0 |
| 11 | 1 | 0 | 0 | 0 | 0 | i | 0 | 0 | 0 | 0 | 0 | 0 |
| 12 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 13 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 2 | 0 | 0 | 0 | 0 |
| 14 | 1 | 0 | 0 | 0 | $\bigcirc$ | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 15 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1.6 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

Table 10b. Percent age composition (round weight) of removals.

| 1 | 1 | 0 | 0 | 0 | $\bigcirc$ | 0 | $\bigcirc$ | 0 | 0 | 0 | 0 | 0 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 9 | , | 0 | 0 | 0 |
| 3 | 1 | 3 | 3 | : | 4 | 1.9 | 11. | 6 | 27 | 8 | 7 | 11 |
| 4 |  | 15 | 14 | 13 | 25 | 10 | 42 | it | 16 | 01 | 38 | 29 |
| 5 | 1 | 18 | 2.6 | 24 | 22. | 28 | 21 | 23 | 26 | i2 | 31 | 30 |
| 6 | 1 | 23 | 1.4 | 23 | 20 | 2.4 | 1.3 | 2.3 | 1.5 | 25 | 10 | 22 |
| 7 | 1 | 25 | 1.4 | 14 | 14 | 10 | 7 | 18 | 7 | 9 | 11 | 4 |
| 8 | 1 | 10 | 20 | 9 | 6 | 5 | 4 | 6 | 5 | 3 | j | 2 |
| 9 | 1 | 2 | 6 | 9 | 3 | 1. | 1. | 1. | 1 | 1 | 1. | 1 |
| 10 | 1 | 1 | 1 | 0 | 6 | 2 | 0 | 2 | i. | i | 0 | 0 |
| 11 | 1 | 1 | 1 | i | 0 | i. | 0 | i | : | 0 | 0 | 0 |
| 12 | 1 | 1. | 1. | 0 | 0 | 0 | 0 | 1. | 1. | 0 | 0 | 0 |
| 13 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 14 |  | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 15 | 1 | 0 | 0 | 0 | ) | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 16 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

Table 11. Catch rates (mt/hr fishing) exhibited by Canadian otter trawl fishery (side and stern) during the February-June period from 1970 to the present.

| Year | Tonnage Class |  |  |
| :---: | :---: | :---: | :---: |
|  | 3 | 4 | 5 |
| 1970 | 0.349 | 0.366 | 0.475 |
| 1971 | 0.158 | 0.332 | 0.497 |
| 1972 | 0.143 | 0.245 | 0.311 |
| 1973 | - | 0.212 | 0.364 |
| 1974 | - | 0.201 | 0.240 |
| 1975 | 0.124 | 0.275 | 0.298 |
| 1976 | 0.080 | 0.318 | 0.191 |
| 1977 | 0.188 | 0.228 | 0.413 |
| 1978 | 0.413 | 0.665 | 0.737 |
| 1979 | 0.189 | 0.492 | 0.484 |
| 1980 | 0.863 | 1.690 |  |

Table 12a. Catch (number of fish) per standard tow by groundfish surveys in 4VW

A - mean for all 27 strata
B - excluding strata with only one set

| MONTH | YEAR | STRATIFIED MEAN CATCH PER STANDARD TOW |  | STRATIFIED ESTIMATE OF THE VARIANCE OF B | ( $n$ ) |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | A | B |  |  |
| June | 1970 | 8.87 | 8.91 | 6.12 | 26 |
| June | 1971 | 8.00 | 8.00 | 6.18 | 27 |
| June | 1972 | 4.41 | 4.41 | 1.39 | 27 |
| June | 1973 | 3.71 | 3.71 | 1.13 | 27 |
| June | 1974 | 6.16 | 6.36 | 7.18 | 26 |
| June | 1975 | 8.90 | 8.90 | 8.59 | 27 |
| June | 1976 | 8.80 | 9.50 | 7.74 | 25 |
| June | 1977 | 25.87 | 25.99 | 56.08 | 26 |
| June | 1978 | 37.75 | 37.91 | 209.48 | 26 |
| Nov. | 1978 | $47.3{ }^{1}$ | - | - | - |
| June | 1979 | 28.91 | 28.97 | 78.82 | 27 |
| March | 1979 | $18.56{ }^{1}$ | - | - | - |
| June | 1980 | 37.74 | 37.71 | 128.69 | 27 |

1 Corrected for differences in fishing power, for area 4VsW (strata 43-66)

Table 12b. Catch (kg) per standard tow by summer groundfish survey in 4VW

AREA
(Strata)

| YEAR | $\frac{4 V n}{40-42}$ | $43-46$ |  | $\frac{4 V s}{27-52}$ | $\frac{4 V^{1}}{50-52}$ | $\frac{4 W}{53-66}$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| 1970 | 2.38 | - | 15.13 | 5.11 | 10.59 | 8.03 |
| 1971 | 0.00 | 0.19 | 1.50 | 0.52 | 9.01 | 5.18 |
| 1972 | 0.00 | 1.92 | 2.63 | 1.67 | 4.16 | 3.04 |
| 1973 | 0.15 | 0.13 | 0.90 | 0.36 | 5.07 | 2.94 |
| 1974 | 0.00 | - | 1.54 | 0.44 | 9.09 | 5.19 |
| 1975 | 0.37 | 0.00 | 2.57 | 0.83 | 12.33 | 7.14 |
| 1976 | 1.64 | - | 2.57 | 1.13 | 8.07 | 4.94 |
| 1977 | 3.08 | 0.17 | 7.28 | 2.90 | 28.86 | 17.15 |
| 1978 | 0.87 | - | 0.03 | 0.21 | 40.38 | 22.26 |
| 1979 | 0.11 | 0.16 | 0.97 | 0.38 | 40.18 | 22.22 |
| 1980 | 0.50 | 0.60 | 2.39 | 1.09 | 60.07 | 33.43 |

$1_{4 V}=0.2367 \times 4 \mathrm{Vn}+0.4758(43-46)+0.2875(47-52)$
$2_{4 V W}=0.4512 \times 4 V+0.5488 \times 4 W$

GROUNDFISH SURVEY STRATIFIED MEAN CATCH PER TOW
4VW HADDOCK SUMMER SURVEY
FOR YEARS 1970 - 1980, AGES 0-11, 12+,UK
VESSEL: A.T. CAMERON
Table 13. 4VW HADDOCK R.V. SURVEY STRATIPIED MEAN CATCH PER TOW 19/ 4/81

|  | 1970 | 1971 | 1972 | 1973 | 1974 | 1975 | 1976 | 1977 | 1978 | 1979 | 1980 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | 0.09 | 0.05 | 0.01 | 0.00 | 0.19 | 0.06 | 0.25 | 0.21 | 0.00 | 1.24 | 1.20 |
| 1 | 2.27 | 1.44 | 1.10 | 0.48 | 0.30 | 3.85 | 2.74 | 5.01 | 8.25 | 0.07 | 2.93 |
| 2 | 0.84 | 3.02 | 0.74 | 1.59 | 1.79 | 0.55 | 3.01 | 9.49 | 9.23 | 7.61 | 0.23 |
| 3 | 1.53 | 1.00 | 1.08 | 0.48 | 2.42 | 1.60 | 0.42 | 7.49 | 12.34 | 8.28 | 12.40 |
| 4 | 1.70 | 1.32 | 0.49 | 0.45 | 0.44 | 1.43 | 0.82 | 1.02 | 6.93 | 8.61 | 11.60 |
| 5 | 0.82 | 0.52 | 0.41 | 0.16 | 0.45 | 0.37 | 0.79 | 1.62 | 0.43 | 2.41 | 7.21 |
| 6 | 0.52 | 0.30 | 0.31 | 0.33 | 0.23 | 0.68 | 0.18 | 0.60 | 0.41 | 0.31 | 1.74 |
| 7 | 0.58 | 0.14 | 0.12 | 0.07 | 0.17 | 0.17 | 0.19 | 0.17 | 0.10 | 0.25 | 0.28 |
| 8 | 0.30 | 0.21 | 0.06 | 0.08 | 0.07 | 0.07 | 0.05 | 0.10 | 0.01 | 0.08 | 0.10 |
| 9 | 0.13 | 0.01 | 0.03 | 0.03 | 0.04 | 0.04 | 0.01 | 0.00 | 0.00 | 0.00 | 0.02 |
| 10 | 0.03 | 0.00 | 0.02 | 0.04 | 0.03 | 0.04 | 0.01 | 0.06 | 0.01 | 0.03 | 0.00 |
| 11 | 0.03 | 0.00 | 0.00 | 0.00 | 0.03 | 0.00 | 0.01 | 0.01 | 0.01 | 0.02 | 0.00 |
| $12+1$ | 0.04 | 0.00 | 0.00 | 0.00 | 0.00 | 0.02 | 0.05 | 0.02 | 0.00 | 0.01 | 0.00 |
| $U K$ \| | 0.00 | 0.00 | 0.04 | 0.00 | 0.01 | 0.01 | 0.27 | 0.08 | 0.03 | 0.00 | 0.04 |

Table $14^{\cdot}$ PERCENT AGE COMPOSITION OF RESEARCH SURVEY CATCHES (NOS.) 8/ 5/81

|  | 1970 | 1971 | 1972 | 1973 | 1974 | 1975 | 1976 | 1977 | 1978 | 1979 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 25.94 | 18.06 | 25.25 | 12.95 | 5.09 | 43.71 | 33.27 | 19.62 | 21.87 | 0.27 |
| 2 | 9.66 | 38.04 | 17.05 | 42.93 | 29.96 | 6.20 | 36.54 | 37.13 | 24.45 | 27.50 |
| 3 | 17.51 | 12.53 | 24.79 | 12.89 | 40.53 | 18.15 | 5.14 | 29.30 | 32.72 | 29.93 |
| 4 | 19.48 | 16.56 | 11.22 | 12.12 | 7.39 | 16.24 | 10.00 | 3.98 | 18.38 | 31.09 |
| 5 | 9.32 | 6.58 | 9.35 | 4.44 | 7.61 | 4.25 | 9.57 | 6.33 | 1.13 | 8.72 |
| 6 | 5.98 | 3.71 | 7.02 | 8.88 | 3.78 | 7.76 | 2.13 | 2.34 | 1.08 | 1.13 |
| 7 | 6.64 | 1.73 | 2.81 | 1.85 | 2.82 | 1.97 | 2.33 | 0.67 | 0.28 | 0.89 |
| 8 | 3.38 | 2.64 | 1.39 | 2.22 | 1.13 | 0.81 | 0.60 | 0.39 | 0.03 | 0.30 |
| 9 | 1.43 | 0.14 | 0.76 | 0.74 | 0.64 | 0.44 | 0.13 | 0.00 | 0.00 | 0.00 |
| 10 | 0.31 | 0.00 | 0.38 | 0.98 | 0.50 | 0.48 | 0.13 | 0.22 | 0.03 | 0.12 |
| 11 | 0.34 | 0.00 | 0.00 | 0.00 | 0.54 | 0.00 | 0.13 | 0.02 | 0.04 | 0.06 |

Table 15a．4VW haddock weight－at－age（ kg ）as derived from the Canadian commercial sampling－both mid－year（a）and beginning of year（b）estimates．

|  | EStzhatien wright at ane |  |  |  |  |  |  |  |  | ／5／81 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1970 | 1971 | 1972 | 1973 | 1974 | 3975 | 1976 | 1977 | 1978 | 1979 | 1980 |
| 1 | 0， 15 | 0.15 | 0.15 | 0， 15 | 0， 15 | 0，15 | 0，09 | 0，13 | 1），08 | 0.11 | 9．08 |
| $\frac{2}{3}$ | $0 \cdot 41$ | 0.11 | 0.41 |  | 0.41 | 0.12 | 0.35 | 0.56 | 0.31 | 0.32 | 0.33 |
| 4 | 0.97 | 0.99 | 1.05 | 1．03 | － | －1．23 |  |  | 0.68 1.19 |  | 0.69 1.02 |
| 51 | 1．23 | 1，30 | 1．70 | 1，38 | 1，60 | 1.58 | 1.58 | 1，5． | 1，53 | 1，60 | 1，42 |
| 61 | 1.56 | 1.65 | i． 77 | 1.84 | 2.05 | 1.99 | 2.07 | i， 88 | 2.00 | 2.04 | 1．88 |
| 71 | 1.98 | 2.01 | 2．18 | 2，17 | $2 \cdot 5$ | 2.45 | 2， 3 年 | 2.31 | 2．45 | 2： 16 | 2.36 |
| 81 | 2．41 | 2،34 | 2.51 | 2.61 | 2.88 | 2．89 | 2.77 | 2.68 | 2.81 | 2．83 | 2.90 |
| 91 | 2，85 | ？ 3.66 | 2，93 | 2，96 | 3，13 | 3.25 | 3，13 | 3，07 | 3，26 | 3.22 | 3， 18 |
| 10 | 3.37 | 3.05 | 3． 25 | 3.19 | 3，39 | 3.27 | 3.36 | 3.45 |  | 呂：99 | ¢ 4.01 |
| 11 | 3,74 +1 | 3，37 | 3.45 | 2．94 | 3.69 |  | $3 \mathrm{3}, 74$ | 3.86 | 4.25 | 3.77 3.70 | 4．83 |
| 12 | 4.09 4.78 | 3，66 | 3.82 | 3.65 | ． 87 | 3.67 4.69 | 3.80 3.70 | 4.17 | 68 | 3.70 4.26 | 3.92 |
| 14 | 5．37 | 4， 1.47 | 3.82 | 5 | 3＋8． |  |  | 4 C | 5 | 4， 48 |  |
| 151 |  | 4，30 |  |  |  |  |  |  |  | 3，6．3 | － |
| 161 | － | － | － | － | － | － | － | － | － | 路 | － |

Table 15b．

| 1 | 1970 |  | AT |  | at | 4 l | ak |  |  |  | 4／5／81 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1971 | 1972 | 1973 | 1974 | 1975 | 1976 | 1977 | 1978 | 1979 | 1980 |  |
| 1 | 0.07 | 0.07 | 0.07 | 0.07 | 0.07 | 0.07 | 0.03 | 0.06 | 0.03 | 0.05 | 0.08 | 0.03 |
| 2 | 0.27 | 0.27 | 0.26 | 0.26 | 0.26 | 0.26 | 0.21 | 0.43 | 0.18 | 0.20 | 0.22 | 0.22 |
| 3 | 0.57 | 0.54 | 0.55 | 0.513 | 0.46 | 0.63 | 0.49 | 0.78 | 0.51 | 0.57 | 0.35 | 0.55 |
| 4 | 0.77 | 0.83 | 0.87 | 0.86 | 0.94 | \}.03 | 0.94 | 1.07 | 0.88 | 0.95 | 0.83 | 0.83 |
| 5 | 1.06 | 1.13 | 1.21 | 1．19 | 1.36 | 1．36 | 1．36 | 1．3A | 1.35 | 1．38 | 1.20 | 1．20 |
| 6 | 1.38 | 1.18 | 1.59 | 1.65 | 1．82 | $\cdots \cdot 77$ | 1．86 | 2.69 | 2．78 | 4.82 | 1.63 | 1.63 |
| 8 | 2．70 | 1．83 | 1．99 | 1：99 | $2 \cdot 32$ | ？：24 | 2，74 | $3 \cdot 10$ | $2 \cdot 23$ | 2.25 | $2 \cdot 11$ | $2 \cdot 11$ |
| 8 | $2 \cdot 2 \cdot 6$ | $2: 49$ | 2， 2,75 | 2：78 | 2：9\％ | －3．68 | 2．ind | 3＋89 | 3 2， 67 | 2.68 | 2.64 | $3+64$ |
| 10 | 3.10 | 2，88 | 3.10 | 3.01 | 3.28 | 3.15 | 3，23 | 3.26 | $3 \times 4$ | 3.35 | 3.74 | 3， 74 |
| 11 | 3.53 | 3＋17 | 3.30 | 2．86 | 3.50 | 4， 33 | 3.60 | 3．68 | 6.05 | 3.63 | 4 9， 5. | 4． 51 |
| 12 | 3， 89 | 3，51 |  | 3：52 |  | 3.58 | 3．69 | 4：00 |  | 3.69 |  | 5.56 |
| 13 | 4.55 | 3.78 | 3.75 | 3，56 | 3.75 | A， 52 | 3．64 | 4， 13 | 3.62 | 4.16 |  |  |
| 14 | 5.10 | 4，3， |  | 4，96 |  |  |  |  | 5.26 | 4.39 | － |  |
| 15 | ＝ | 1．19 |  | － | － |  | － |  |  | A．56 |  |  |
| 17 | － | － | E | － | － | － | － | － | ＝ | － | － | － |

Table 16. Summary of SURVIVOR Runs


* Second run with same parameters using estimated SURVIVORS of first run as inftial data

Table 17. Results of 'SURVIVOR' calibration block 1970 through 1977, ages 2 through 5

EIMAL ITRPATION (2)


ESTIUAMFD SURVIVORS FOR AGE \& (WFIGHTED)

| $Y E A R$ | SUPVIVORS | VARIANCF | STANDARD FIRROR | c.v. (0/0) |
| :---: | :---: | :---: | :---: | :---: |
| 1970 | 727 | 374.380 | 612 | 84.17 |
| 1971 | 84 | 157727 | 397 | 472.95 |
| 1972 | 74 | 11826 | 109 | 147.09 |
| 1973 | 109 | 15359 | 124 | 113.94 |
| 1974 | 80 | 7412 | 85 | 106.98 |
| 1975 | 228 | 14471 | 120 | 52.77 |
| 1976 | 136 | 5474 | 80 | 59.12 |
| 1.977 | 309 | 22279 | 149 | 48.31 |
| 1978 | 10 | 502 | 22 | 223.63 |
| 1979 | 258 | 14142 | 11.9 | 46.19 |
| 1980 | 325 | 23132 | 158 | 51.67 |

FSTIMATFD SURVIVORS FOR 1980 (WRIGHTFD)

| $A F^{F}$ | SURVIVORS | VARIAMC ${ }^{\text {P }}$ | STANDARD EFPOR | C.7. (0/0) |
| :---: | :---: | :---: | :---: | :---: |
| 2 | 958 | 414877 | 544 | 67.26 |
| 3 | 28306 | 214922913 | 14660 | 51.7 ? |
| 4 | 22122 | 104081035 | 10202 | 46.12 |
| 5 | 19605 | 59787873 | 7732 | 29.44 |
| 5 | 4252 | 4421250 | 2103 | 49.45 |
| 7 | 313 | 53786 | 232 | 74.1 ? |
| 8 | 325 | 29132 | 168 | 51.67 |

EIVAL ESTIMATION FOR K

| $A G T$ | $\underline{K}$ | $\underline{L} y(\underline{y})$ | VAR ( $V^{\text {V }}$ (K) | STAMDARD FRTROR | $\underline{D} \cdot \underline{P}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | 4600.67 | 8.3363 | 0.1954 | 0.1563 | 3 |
| 3 | 3617.08 | 8.0818 | 0.2233 | 0.1571 | 3 |
| 4 | 3405.65 | 8.0355 | 0.1954 | 0.1105 | 1.5 |
| 5 | 3405.55 | 8.0355 | 0.1954 | 0.1105 | 15 |
| 6 | 3405.65 | 8.0355 | 0.1954 | 0.1105 | 15 |
| 7 | 3405.65 | 8.0355 | 0.1954 | 0.1105 | 15 |
| 8 | 3405.65 | 8.0355 | 0.1954 | 0.1105 | 15 |

$4 / 5 / 81$

|  | 1970 | 1971 | 1972 | 1973 | 1974 | 1975 | 1976 | 1977 | 1978 | 1979 | 1980 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | 0.3627 | 0.7519 | 0.2357 | 0.2043 | 0.3057 | 0.4610 |  |  | 0.1229 |  | 0.0977 |
| 3 | 0.2882 | -0.1349 | 0.2054 | 0.0695 | 0.8049 | 0.2066 | 0.7378 | 0.4243 | 0.0254 | 0.2767 | 0.2098 |
| 4 | 0.1385 | -0.0108 | 0.4258 | 0.6314 | 0.3007 | 0.7423 | -0.2227 | 0.4106 | 0.6225 | 0.1515 | 0.2945 |
| 5 | 0.2351 | -0.6015 | 0.4713 | -0.8589 | 0.0462 | 0.6674 | 0.6359 | 0.7929 | 0.0001 | 0.0866 | 0.0462 |
| 6 | 0.3311 | -0.4602 | -0.0063 | 0.0962 | 0.3339 | 0.9520 | 0.5937 | 0.7965 | 0.0083 | 0.1624 | 0.0362 |
| 7 | -0.1809 | -0.6441 | 0.3081 | 0.4552 | 0.2568 | 0.7546 | 0.0950 | 1.3427 | -0.3905 | 0.1335 | 0.5771 |
| 8 | -0.0977 | -0.0721 | 0.0364 | 0.2325 | 0.6212 | 0.1945 | -0.0815 | 0.1896 | -0.2496 | 0.1389 | 0.3205 |

MEAN OF RFSIDUALS $=0.03855540169$
STANDARD DFVIATION OF RFSIDUALS $=0.4389019591$
OUTLIPRS OF RESIDUALS $\quad 4 / 5 / 81$

|  | 1970 | 1971 | 1972 | 1973 | 1974 | 1975 | 1976 | 1977 | 1978 | 1979 | 1980 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| -+ | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| 2 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| 4 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| 5 | 0.00 | 0.00 | 0.00 | 0.86 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| 8 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.95 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| 7 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 1.34 | 0.00 | 0.00 | 0.00 |
| 8 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |

Table 17.

Allalysis of variallce


| SOURCP | S. 5 | $D E$ | MS | P |
| :---: | :---: | :---: | :---: | :---: |
| COMSTAMT | $0 . \overline{1115}$ | 1 |  |  |
| AGE | 0.65135 | 6 | 0.1085 | 0.647 |
| YFAR | 4.23016 | 10 | 0.4230 | 2.522 |
| YRCLASS | 2.04282 | 15 | 0.1277 | 0.761 |
| PESIDUALS | 7.71590 | 44 | 0.1754 |  |
| TOTAL | 14.75174 | 77 |  |  |



Table 19. Results of traditional sequential population analysis

| POPULATION NUMBERS ('000) |  |  |  |  |  |  |  |  |  | 6/ 5/81 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1970 | 1971 | 1972 | 1973 | 1974 | 1975 | 1976 | 1977 | 1978 | 1979 | 1980 |
| 11 | 7782 | 4128 | 8015 | 8213 | 6999 | 31497 | 54131 | 67792 | 80173 | 652 | 4546 |
| 21 | 5469 | 6057 | 3101 | 6295 | 6427 | 5696 | 25438 | 44052 | 55471 | 65512 | 533 |
| 31 | 7016 | 4358 | 4342 | 2350 | 4130 | 5060 | 4614 | 20272 | 35932 | 45265 | 53504 |
| 4 | 7507 | 5133 | 2800 | 3012 | 1431 | 2523 | 3777 | 3647 | 15705 | 28640 | 36632 |
| 5 | 4103 | 4591 | 2278 | 1642 | 1236 | 954 | 1398 | 2919 | 2612 | 10836 | 22366 |
| 61 | 3261 | 2087 | 1360 | 1040 | 616 | 605 | 535 | 923 | 1899 | 1744 | 8266 |
| 71 | 1 3398 | 1420 | 643 | 502 | 215 | 291 | 329 | 307 | 534 | 898 | 1287 |
| 8 | 1821 | 1700 | 332 | 219 | 116 | 77 | 139 | 176 | 166 | 247 | 606 |
| 9 | 207 | 319 | 351 | 102 | 39 | 55 | 35 | 37 | 93 | 79 | 189 |
| 101 | 179 | 87 | 16 | 147 | 21 | 24 | 37 | 23 | 59 | 54 | 60 |
| 111 | 146 | 30 | 37 | 3 | 27 | , | 16 | 24 | 12 | 40 | 40 |
| 1 | 139691 | 29910 | 23274 | 23526 | 21256 | 46733 | 90449 | 140223 | 192654 | 153967 | 128026 |



Table 20. Correction of bias in 'survivor' based population biomass estimates

| YEAR | RESEARCH SURVEY | 2+ BIOMASS (MT) 'SURVIVOR' BASED SPA | RATIO SURVEY/SPA |
| :---: | :---: | :---: | :---: |
| 1974 | 16,151 | 8,102 | 1.99 |
| 1975 | 18,833 | 8,662 | 2.17 |
| 1976 | 14,057 | 14,133 | 0.995 |
| 1977 | 47,950 | 38,165 | 1.26 |
| 1978 | 63,710 | 56,585 | 1.13 |
| 1979 | 65,152 | 79,652 | 0.82 |
| 1980 | 100,416 | 88,036 | 1.14 |
|  |  | geometric | 1.28 |
| $\because$ | corrected SPA estimated of 1980 2+ biomass: 113,000 |  |  |

Table 21. Input conditions and results of projections



Figure 1. Nominal catches of Div. 4VW haddock (mt) value for 1980 is preliminary.


Figure 2a. Seasonal Distribution of 4VW Haddock landings by the otter trawler fleet since 1960. Each year is cumulated with the previously indicated years.


Figure 2b. Seasonal Distribution of 4VW Haddock landings by the longliner fleet since 1960. Each year is cumulated with the previously indicated years.


Figure 3. 1970-1980 trends in standardized (to 1980) catch rates (weight/time) in Canadian Otter Trawler fishery (TC 4 and TC 5) and Canadian Summer Bottom Trawl Survey


Figure 4. RESIDUALS = OBSERVED - PREDICTED WEIGHTS FOR 4VW HADDOCK, JULY 1980, A306-A307

Figure 5a.


|  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| itratua | isar | 550 | HEJCHTEU | WEIGHTED |
|  |  |  |  |  |

Figure 5b.


| FEEULTS OF GROUNDFIEH SURUET FOE 1978 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| stratuk | tbari | 550 | WEIGHTED Ybar: | $\underset{=50}{\text { WEIGHTED }}$ |
|  |  |  |  |  |

STRATIFIED MEAN CATCH PER TOW 37,90553336
STRATIFIED ESTIKATE OF THE VAKIANCE; 209.4848535

Figure 5c.


|  |  |
| :---: | :---: |


| stratux | ybar | 35 | $\underset{\text { YSEAR }}{\text { WEITED }}$ | $\underset{\text { STSO }}{\text { WEIED }}$ |
| :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |

Figure 5d.



| STRATUM | YEAR | 550 | WEIGHTED YGAR | $\begin{gathered} \text { HEIEHTED } \\ \text { ESQ } \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |
|  |  |  | $5 \times$ | 1. |
|  |  |  | $2 \cdot 8 \mathrm{E}$ | 2. |
|  |  |  | \%) $\frac{1}{4}$ |  |
|  |  |  | 8.6E0 | 0,0 |
|  |  |  | C, 0 ES |  |
|  |  |  | - |  |
|  |  |  | 5: |  |
|  |  |  |  |  |
|  |  |  | 4,9E-5 |  |
|  |  |  | 2.0E |  |
|  |  |  |  |  |
|  |  |  | 0.20 C | $0.00^{2}$ |
|  |  |  |  |  |
|  |  |  | $\bigcirc{ }^{\circ} \mathrm{CO}$ | \% |
|  |  |  | 1060 |  |
|  |  |  | $5{ }^{\text {P }}$ |  |
|  |  |  |  |  |
|  |  |  | $5+3$ | $3.2 \mathrm{E}^{2}$ ? |
|  |  |  | 1.EO |  |
|  |  |  | 0, 0 E |  |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  |  |  |
|  |  |  | 0.050 | , |

STRATIFIEI MEAN CATEH PER TOW* 3 3'70856481.
STRATIFIEU ESTIMATE OF THE VARIAREE: $128+6948986$


4VW Haddock r.v. 1980


4VW Haddock r.v. 1980.

Figure 6. Length-frequencies for four dominant strata in summer 1980 survey


Figure 6 (continued). Length-frequencies for four dominant strata in summer 1980 survey


Figure 7a. Diurnal variation in Canadian summer bottom trawl survoy averape catch rates (numbers per standard tow) of 4VW haddock during 1970-80.


Figure 7b. Diurnal variation in Canadian summer bottom trawl survey average catch rates (numbers per standard tow) of 4 X haddock during 1970-80.


Figure 8. Numbers of haddock per 30 min . tow for two length groups: small (<23 cm) and large ( $>23 \mathrm{~cm}$ ) on a special cruise conducted in May, 1981


Figure 9. Year class size as determined from summer, spring and fall research surveys. (A correction factor of
Was applied to spring and fall surveys.)


Figure 10. Comparison of observed and calculated biomass for the 4 VW haddock stock.

Mid-year abundance from cohort. YS. . corrected research catch per standard tow for age 2. Sample correlation 0.9765730531


Mid-year abundance from cohort vs. corrected research catch per standard tow for ages 3. Sample correlation: 0.9526485655


Figure 11. Mid-year abundance from 'SURVIVOR' based cohort versus summer research survey catch (numbers) per standard tow for ages 2-8.

Mid-year abundance from cohort vs. corrected research catch per standard tow for age 4 . Sample correlation: 0.9376787959


Mid-year abundance from cohort vs. corrected research catch per standard tow for age 5. Sample correlation: 0.9840161934


Figure 11. (continued)

Mid-year abundance from cohort vs. corrected research catch per standard tow for age 6.


Mid-year abundance from cohort vs. corrected research catch per standard tow for age 7. Sample correlation: 0.8863754831


Figure 11. (continued)

Mid-year abundance from cohort vs. corrected research catch per standard. tow for age 8. Sample correlation: 0.8182336253


Figure 11. (continued)


| テGan Class | 196 | 1770 | 1771 | 1972 | 1975 | 1974 | 1975 | 1976 | 1977 | 1772 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| FYMATM | 5.27 | 2.18 | 2.69 | 2.27 | 0.85 | 688 | 12.23 | 1,1,24 | 15.86 | 0.30 |
| SFA! | 7792 | 4128 | 9015 | 8213 | 879 | 3147 | 517 | 57792 | 30173 | 65 |
| fremicten | 2188 | 6367 | 8930 | 682 | -308 | 2795 | 3625 | 6916 | 7249 | 306 |

```
EAMFLE CORELLATIOR: 0.98006778
EOFE; 5020,45505
IHTERGEFT+4575.126E50
```

Figure 12a. Comparison of calculated age 1 numbers ('000) with research $1+2$ catch per standard tow for the 4 VM haddock stocks, with least squares regression line (GM regression line was similar).

SAMPL: CORELATTOU: 0.9317229597
5LOFE; 2410,708215
INTEACEFT: 1170.99569

Figure 12b. Comparison of calculated 3+ numbers ('000) with research catch per standard tow for the 4 VW haddock stock, with least squares regression line (GM regression line was similar).



[^0]:    ${ }^{1}$ As reported to NAFO

[^1]:    ${ }^{1}$ excludes samples for which no weight was recorded

