

Expansion Économique Régionale

> Working Paper No.5 WAGES & SALARIES IN CANDIDE-R July 1975



# ECONOMIC DEVELOPMENT ANALYSIS DIVISION

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Working Paper No.5

WAGES & SALARIES
IN CANDIDE-R

July 1975



This working document represents a partial regionalization of the CANDIDE 1.1 model. The acronym CANDIDE refers to the Canadian Disaggregated Interdepartmental Econometric model.

The CANDIDE-R version of the model outlined in this document is designed to help build an appreciation of the regional diversity of Canada. The authors draw attention to the tentative nature of the econometric work reported upon. So as to avoid attributing official status to the views expressed, prior consultation respecting quotation would be appreciated.

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#### WAGES AND SALARIES IN CANDIDE-R

# 1.0 Introduction 1

The objectives of the regional model's salaries and wages block are the same as those of its counterpart in CANDIDE 1.1. That is, the block estimates the most important component of personal income, and calculates unit labour costs, which in turn are used in determining industry prices.

The regionalization of salaries and wages in CANDIDE-R should allow achievement of these two objectives. The regional salaries and wages model was thus incorporated in CANDIDE-R following the widespread theory that disparity is the appropriate variable for taking account of regional differences in the framework of a national study.

Salaries and wages were therefore regionalized by a limited feedback technique<sup>2</sup>. In summary, the original equations in CANDIDE 1.1 are used to estimate national

<sup>1.</sup> We thank Michel Bergeron, professor of Economics, University of Sherbrooke, for his judicious comments.

See An Overview to CANDIDE-R, Section 3, "Regionalization".

averages at the first stage; at the second stage the regional block determines regional wages, and finally in the last stage regional salaries and wages are adjusted so that their sum corresponds to that obtained in the first stage.

The conceptual framework of regional wage and salary determination has inspired a so-called model of selected determinants. 3 Such a model relies on a choice of selected variables including factors which are likely to shift the labour supply curve, others the labour demand curve and elements which indicate disequilibrium in the market in question. These three types of explanatory variables are combined in a multiplicative rather than linear format in order to quantify the net effect of their Taking into account the important role of productivity in the long run determination of salaries4, the model retains relative productivity as the first determinant which influences regional wages through the displacement of labour demand curves. The second type of determinant comes from variations in real wages paid by industry at the national level, which in turn shift regional labour supply curves. Finally the model has a third variable, the unemployment rate, whose role is to introduce the concept of disequilibrium.

# A Salary Model of Selected Determinants

In most macroeconomic models, including

See Kuh, E., "A Productivity Theory of Wages - An Alternative to the Phillips Curve", Review of Economic Studies,

vol. XXXIV, October 1967, pp. 333-360.

Cowling, K. and D. Metcalf, "Wage-unemployment Relationships: A Regional Analysis for the U.K. 1960-1965", Oxford Bulletin of Economics and Statistics, February 1967, pp. 31-39. Smith, V.K. and R.A. Patton, "Sub-Market Labor Adjustment and Economic Impulses: A Note on the Ohio experience", Regional Studies, vol. 5, 1971, pp. 91-93.

CANDIDE 1.1, wages and salaries are determined within the analytical framework of Phillips or Kuh<sup>5</sup>. The national block determines the wage rates of three industries using a dynamic Phillips curve relationship. The rate of change of wages in one industry is related to the rate of change of the Consumer Price Index and to the inverse of the unemployment rate in the current and lagged periods. well the wage rate in the corresponding American industry is introduced with a distributed lag, under the assumption that the demand for wage parity through collective negotiations and the very close links between the two economies are translated into analogous movements in salary increases. For the remaining nine industries, the national block determines the wage bill as a function of variables more likely to reflect medium term trends and thus better suited to this aspect of the model. tors are productivity, production and/or employment and also a variable to account for price escalation. gionalization of CANDIDE-R does not sacrifice any of the national relationships and takes more direct account of factors associated with medium term growth of salaries.

## 2.1 Introductory Comments

In order to adequately specify the mechanism of wage and salary determination in industry i in region j, we must consider a large range of factors, each as

<sup>5.</sup> Kuh, E., op. cit.

Lipsey, R.G., "The Relation Between Unemployment and the Rate of Change of Money Wage Rates in the United Kingdom, 1862-1957: A Further Analysis", Economica, Vol. XXVII, February 1960.

important as the next. First there is the presence of national equations whose purpose is to determine national averages; second there are the regional interdependencies of wage and salary markets and thus of their structure; third, data limitations exist on a certain number of variables; and finally, there is the explicitly mediumterm nature of CANDIDE.

It is important to consider the choice of unit of measurement of salaries and wages. Should we use the wage bill, the wage rate or the rate of change of the wage rate? As is well known, wage equations are commonly estimated using the rate of change of the wage rate as the dependent variable. The rationalization for this usage, as given by Lipsey<sup>6</sup>, rests on the hypothesis that in situations of competition, the rate of change of the price of a good (in this case labour) depends on the excess demand for that good. As Agarwala, Drinkwater, Khosla and McMenomy have observed, this hypothesis offers only a theory of prices in disequilibrium about a stationary equilibrium value. In practice, we should have more success in attempting to explain wage time series if we examine the divergences from equilibrium rather than the magnitude of the disequilibrium.

#### 2.2 The Model

In order to incorporate the aspects discussed

<sup>6.</sup> Lipsey, R.G., op. cit.

Agarwala et al., "A Neo-classical Approach to the Determination of Prices and Wages", <u>Econometrica</u>, August 1972.

above, first, a peripheral model was constructed of labour supply and demand in industry (i,j), which is structurally different for all (i,j) and which crystallizes the state of regional interdependence in order to determine wages in industry (i,j).

In order to better understand the approach taken for the basic specification of our model, it is necessary to spell out some of the details. Assume that labour demand in industry (i,j) is a positive function of production (D) and is independent of real wages, that the supply of labour (S) is a positive function of real wages. The positive unemployment rate can be used to analyse the dynamics of the market in disequilibrium.

These assumptions together give us Figure 1 which allows us to visualize for certain industries the possibility of a disequilibrium real wage rate, the size of which is measured by the prevailing unemployment rate in that industry. (Note that unemployment in industry i is given by  $\mathbf{E}_0$ - $\mathbf{E}_1$  in Figure 1).

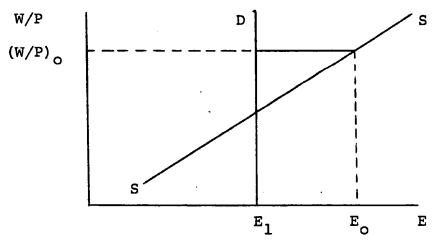


Figure 1.

The estimation of such a model requires the specification of a technique for disequilibrium estimation. The Phillips curve approach is one way to get around the problem, but this procedure was rejected as it is not well suited to a medium term outlook. While it would be possible to specify a disequilibrium estimation technique, equally interesting results can be obtained by selecting supply and demand variables which, through affecting the position of supply and demand curves, influence the level of wages.

### 2.2.1 Model Specification

Assume the following relationship for industry (i,j):

$$W_{ijt}/E_{ijt} = e^{\delta} U_{jt}^{\alpha} [(X_{ijt}/E_{ijt})/(X_{it}/E_{it})]^{\beta}$$

$$[(W_{it}/E_{it})/CPI_{t}]^{\gamma}CPI^{\eta} e^{\mu t}$$
(1)

where  $W_{ijt}$  = Wage bill in industry i, region j, in period t

 $\delta$  = constant

U<sub>jt</sub> = Unemployment rate in region j, period t

 $X_{i+}$  = Industry output in industry i, region j

E<sub>iit</sub> = Employment in industry i, region j

 $W_{it}$  = Wage bill in national industry i

CPI = Consumer Price Index

 $\mu_t$  = Error term.

The hypothesis implicit in equation (1) rests on the following assumptions: that the variable U

represents a state of disequilibrium in industry (i,j); that the "productivity" variable will cause the demand curve to shift, and that the real wage variable  $(W_i/E_i)/CPI$  will cause the labour supply curve to shift. We hypothesize as well that the relative improvement in the productivity of workers in industry (i,j) with respect to those in the Canadian industry i is translated onto Figure 1 by an increase in labour demand in industry (i,j) in anticipation of an increase in productivity in this same industry (i,j) to take into consideration the level of wages and salaries paid in industry (i,j) and the real wage in the Canadian industry, when considering their participation in the labour force.

Considering the productivity of the national industry i as a result of the regional performances, all growth in Canadian output which is accompanied by an increase in the average productivity of industry i should affect the relative regional productivities of industry i. If the growth in national output in industry i is concentrated in region j, it is clear that this will lead to a relative improvement in the productivity of industry (i,j) and will cause an increase in labour demand in the same industry. Under these circumstances, if the labour supply curve does not move, the wages in industry (i,j) should be forced up, taking into consideration that since labour demand is somewhat specialized it will not lower

<sup>8.</sup> This assumption is discussed below. In the meantime, it is noted that this "hypothesis" does not necessarily imply the use of distributed lagged variables since annual data is used.

proportionately the pool of unemployed wishing to work in industry (i,j) at the going wage. As a result, according to the assumption that the increase in labour demand does not reduce to zero the number of unemployed, one can anticipate an increase in wages in industry (i,j).

On the other hand, if the increase in the average productivity of industry (i,j) is not accompanied by an increase in demand, one can easily imagine a situation where workers organize themselves in order to benefit from these improvements in productivity.

The advantage of using such a "productivity" variable is the possibility which it offers to draw from the national model variations in the productivity of the Canadian industry which can be translated regionally through wage increases. The variable  $(W_i/E_i)/CPI$  offers the same Through national wages this variable establishadvantages. es the necessary links between the national and regional mechanisms of wage and salary determination. As well as establishing these links,  $(W_i/E_i)/CPI$  plays a particular role in this model. As mentioned above, it allows workers in industry (i,j), when considering their participation in the labour force, to take into account the wages paid by industry (i,j) with respect to the real national wage of that industry. An increase in W,/CPI will shift the supply curve up to the left and, ceteris paribus, will lead to an increase in W<sub>ii</sub>. This shift in labour supply may be

<sup>9.</sup> A parallel can be drawn between regional and national salaries as these two variables are for all intents and purposes measured in constant dollars. The variable CPIn on the right hand side serves as a deflator for...

explained by an emigration of workers towards regions offering higher real wages, or by retirements from the labour force when the benefits of collective or private negotiations are not forthcoming in the form of salary increases.

Taking into account our analytical framework, if the unemployment variable, U, is significant, it must have a positive coefficient (α) in order to conform with a priori expectations. It is then possible to classify regions according to their wage elasticities with respect to unemployment, as the coefficients from an equation estimated in logarithmic form offer a direct measure of elasticity.

#### 2.2.2 Some Theoretical Aspects of the Model

The general appeal of equation (1) has several dimensions. It was inspired by the theory of Kuh<sup>10</sup>, where the operation of a labour market depends, at the first stage of analysis, on a fundamental variable of demand, that is productivity, and which recognizes, at the next stage, the importance of supply variables in a system of

<sup>... (9.</sup> cont'd) (Wijt/Eijt) to the extent that the estimated value of η is expected to be very close to unity. By not constraining the coefficient η to equal one, the diverse regional industries are allowed to adapt more or less perfectly to take account of an increase in the cost of living.

<sup>10.</sup> Kuh, E., "A Productivity Theory of Wage Levels - An Alternative to the Phillips Curve", Review of Economic Studies, Vol. XXXIV, 1967, pp. 333-360.

simultaneous equations. At the level of region (j), the assumptions of Kuh's theory can essentially be reduced to the following:

$$W_{ijt} = f[U_{ijt}, (P_{ijt} \cdot X_{ijt})/E_{ijt}, CPI_{t}]$$
 (2)

where productivity appears in value terms 11.

Contrary to the Kuh approach, equation (1) considers the underlying structure of the labour market and depends on the functioning of this market in the determination of wages in a regional context. The comparative statics exercise described at the introduction of this model provides a description of the operation of such a labour market. The strength of the theory behind the dynamic forces of the market results from the choice of a framework based on relative prices.

In spite of the problems relating to the theoretical basis for wage and salary determination, the

<sup>11.</sup> This measurement of productivity in value terms facilitates the analysis by allowing use of nominal salaries as dependent variables. Since this model considers the relative productivity of an industry in a region, compared to the national value, the ratio of productivities in value does not differ from the ratio actually used. Moreover, as already indicated, we are working to some extent with the real wage in the assumption than  $\eta \to 1$ . Thus the theory is coherent as we use two real measures for salaries and productivity, in place of two nominal measures.

regional context brings additional constraints. The problem of specification of a regional model reduces to one of capturing the determinants, both autonomous and induced, and the transversal influences which are felt from one region to another. Consequently, recognizing the spatial dimension of labour supply and demand and the partial mobility of factors of production, it is important to allow for changes in relative prices, as well as in the general price level. This is what has been attempted in equation (1) through the choice of variables retained as the central force for shifts in the labour supply and demand curves.

In this perspective, the theory of regional salaries relative to Canadian salaries is related to an adaptation of Friedman's "permanent wage theory".  $^{12}$  Briefly, real wages  $((\text{W}_{\text{it}}/\text{E}_{\text{it}})/\text{CPI}_{\text{t}})$ , can be associated with permanent wages and the Consumer Price Index with the transitory component of wages. Thus  $(\text{W}_{\text{it}}/\text{E}_{\text{it}})/\text{CPI}_{\text{t}}$  takes account of variations in relative wages in real terms and causes modifications in the participation of workers of one region due to their relatively more or less favourable position in the national industry. The information provided by the variable  $(\text{W}_{\text{it}}/\text{E}_{\text{it}})/\text{CPI}_{\text{t}}$  to workers in industry (i,j) allows them to formulate expectations, in the permanent wage framework, on their relative state of well-being. Their method of improving their relative position, as measured by their real wage,

<sup>12.</sup> Lucas, R. and L.A. Rapping, "Real Wages, Employment and Inflation", in Microeconomic Foundations of Employment and Inflation Theory, ed. by E.J. Phelps, Norton, 1970, p. 266. Wachter, M.L., "Relating Wage Equations for U.S. Manufacturing Industries 1947-1967", Review of Economics and Statistics, November 1970, pp. 405-410.

is to modify their participation as workers in this industry. A priori,  $\gamma$  is expected to be positive.

 $(W_{it}/E_{it})/CPI_t$  can equally be considered as a provisional indicator of the same type as profits in the case of production.  $^{13}$ 

As far as the variable  $[(X_{ijt}/E_{ijt})/(X_{it}/E_{it})]$  is concerned, it is used to measure forces which influence labour demand. Estimated as a ratio of productivities, it plays, according to employers, a role analogous to that of relative wages in the minds of the workers. An improvement in this measure, for a regional industry, urges employers to increase their labour demand in order to realize locally the potential increase in profits derived from higher productivity. Thus the coefficient  $\beta$  (in equation (1)) must be positive to conform with a priori expectations.

Returning to equation (1) the ratio of productivities can be considered, in the manner of  $\operatorname{Kuh}^{14}$ , ceteris paribus, as a ratio of profits of regional industry (i)

<sup>13.</sup> According to the terminology of the National Bureau of Economic Research on this subject, see B.M. Friedman and M.L. Wachter, "Unemployment: Okun's Law Labor Force and Productivity", Review of Economics and Statistics, May 1974.

<sup>14.</sup> Kuh, E., op. cit.

<sup>15.</sup> In order to establish a parallel between these two concepts, it is necessary that the two variables be measured in value terms. The measure of the ratio of productivity in value should not differ from our case as we have assumed that the implicit regional industrial production price indexes corresponded to the national prices. (Cont'd next page)...

to those of the same national industry. Thus, an increase in relative productivity would lead to increased profits of industry (i,j) which would then expect an increased demand for its product. 16

With reference to the variable U<sub>ijt</sub> it was noted earlier that it plays mainly a technical role in the estimation of a state of disequilibrium. Upon further examination of equation (1), one sees that the aggregation of the regional relationships takes the form:

$$W_{it} = f(U_t, CPI_t)$$
 (3)

or

$$dW_{it}/dt = (\partial W_{it}/\partial U_t) \cdot (dU_t/dt) + (\partial W_{it}/\partial (CPI)_t)$$

$$\cdot (d(CPI)_t/dt) \qquad (4)$$

(15. cont'd from preceding page)...
If in [(P<sub>ijt</sub> · X<sub>ijt</sub>)/E<sub>ijt</sub>]/[(P<sub>it</sub>/X<sub>it</sub>)/E<sub>it</sub>] ,

P<sub>ijt</sub> = P<sub>it</sub> , then this ratio equals

16. For a discussion on the use of profits as a provisional indicator or as a variable at the base of anticipations in a neo-classical model, see respectively:

Evans, M.K., Macroeconomic Activity: Theory Forecasting and Control, Harper and Row, New York, 1969.

Gordon, D.F. and A. Hines, "On the Theory of Price Dynamics", in Microeconomic Foundations of Employment and Inflation Theory, F.S. Phelps et al. (eds.), Norton, New York, 1970.

Equation (4) resembles the Phillips curve equation used in CANDIDE 1.1, except that the unemployment rate appears in (4) as a rate of change. The presence of Uijt in equation (1) would thus add an important dimension to this model in allowing the regional unemployment rate to take account of structural changes in regional labour markets<sup>17</sup>, and of the fact that, according to Bowen and Berry<sup>18</sup>, the relationship between cyclical and structural unemployment varies through time. Therefore, since changes in the unemployment rate are more closely related to cyclical unemployment we should expect a significant correlation between the rate of change of wages and salaries and variations in the unemployment rate.

For long-period studies, changes in unemployment may well constitute a more reliable index of active labour market processes than the level of unemployment which is more susceptible to inter-period variation in the amount of structural unemployment.19

One last interesting theoretical aspect<sup>20</sup> relates to the mathematical formulation of equation (1), a formulation which leads us to believe that all formulae

<sup>17.</sup> Lipsey, R.G., op. cit.

<sup>18.</sup> Bowen, W.G. and R.A. Berry, "Unemployment Conditions and Movements of the Money Wage Level", Review of Economics and Statistics, 45, (2), 1963, pp. 163-172.

<sup>19. &</sup>lt;u>ibid</u>., p. 169

<sup>20.</sup> This comment is attributed to M. Michel Bergeron, of the Department of Economics, University of Sherbrooke.

for indexation of Canadian wages and salaries to a cost of living index would be ineffective in raising regional wages and salaries. In effect, indexation of  $W_{it}$  to the rate of increase of the consumer price index (CPI $_t$ ) would leave the variable  $W_{it}/\text{CPI}_t$  unchanged. The measure of productivity, in value terms, could not be the mechanism for transmission of the benefits of indexation to the regions, since an increase in regional prices following an increase in national prices would not modify the ratio of productivities, since it is in real terms.  $^{21}$ 

The Consumer Price Index of equation (1) gets around this problem.

Since sufficient time series of regional price indexes do not exist, it was necessary to use the national Consumer Price Index in their place.

It is worth noting once more that this model was first built as a peripheral model to the national CANDIDE 1.1 model. The original equations of CANDIDE supply estimates of national averages of wages, required for the estimation of regional wages. At a second stage wages and salaries are determined by region, taking into account cyclical and structural differences between regions in the labour market. In a last stage, the regional

<sup>21.</sup> Even if the productivity ratios were expressed in value terms

an increase in regional prices following an increase in the national price would not modify the ratio of productivities.

salary estimates are adjusted in order that their sum correspond to that obtained at the beginning, from the national equations. The adjustment mechanism splits the difference between the two national wage bill estimates according to the relative importance of each of the regions:

$$\mathbf{W}_{ijt}^{*} = \hat{\mathbf{W}}_{ijt} + [\hat{\mathbf{W}}_{ijt}/\Sigma \hat{\mathbf{W}}_{ijt}] [\hat{\mathbf{W}}_{it} - \Sigma \hat{\mathbf{W}}_{ijt}]$$
 (5)

where ^ = estimated

\* = corrected

In principle a different procedure would have been desirable for this third stage. Originally it was intended to reaggregate the regional wage estimates from the second stage and to substitute these values for the national estimates obtained from the original CANDIDE 1.1 equations. This aggregation would carry the effects of regional disparities throughout the model where wages are used as explanatory variables. This characteristic would allow a study of the impact of regional disparities of wages and salaries on the national average and on the economy as a whole. There are, of course, a range of intermediate alternatives possible, if it is decided to correct the regional estimates from the second stage by dividing the difference between the estimates defined as follows:

[X % \* national estimate from the first stage for an industry
+ (100 - X)% sum of regional estimates of that industry]
- [sum of regional estimates of that industry].

This would result in the following corrected regional values:

$$W_{ijt}^{*} = \hat{W}_{ijt}^{*} + \hat{W}_{ijt}^{*} / \sum_{j=1}^{N} \hat{W}_{ijt}^{*} + ((1 - \beta) * \sum_{j=1}^{N} \hat{W}_{ijt}^{*}) + ((1 - \beta) * \sum_{j=1}^{N} \hat{W}_{ijt}^{*}) + ((1 - \beta) * \sum_{j=1}^{N} \hat{W}_{ijt}^{*})$$

$$- \sum_{j=1}^{N} \hat{W}_{ijt}^{*}$$

$$(6)$$

As with consumption in CANDIDE 1.1, the choice for the value of  $\beta$  would be left to the user. ple the choice would be made after a series of experiments with the value of  $\beta$ , selecting that value which would minimize the variance of the consumption estimates, or in this case of  $\Sigma W_{iit}$ . However, such experiments run a high resulting in the maintenance of status risk of quo, or  $\beta = 1$ . Theoretically one should benefit by maintaining a portion of the information from the estimated aggregates of the first stage and a portion of the information from the disaggregated estimates of the second stage, since the greater detail at the regional level adds information not obtainable at the national level. cularly because the information provided by the national wage estimates is required in estimating regional wages, the performance of regional equations in the model is partially tied to the quality of the national estimates. errors in the national equations are thus automatically transmitted to the regions - as well as errors from the specification as such of the explanatory variables and from the poor quality of much of the regional data. the extent that there is a desire to improve national equations, a large improvement would be made in favour of adopting the original aggregation approach, since the estimation results at the regional level are very encouraqinq.

## 3.0 Empirical Results

The empirical results for the equations described above follow. The quality of the estimation may be judged on the basis of the significance of the coefficients, the  $\overline{R}^2$  and the signs of the coefficients (i.e. whether they conform with a priori expectations). Little importance is attached to the Durbin Watson test in this case due to the limited number of observations. It is worth noting that the coefficients may be interpreted as elasticities, and that the difference in elasticities between regions well displays the regional characteristics.

A set of graphs of manufacturing wages follows the empirical results. These graphs are expressed in level form rather than in logarithmic ratios as they were estimated.

#### Salaries and Wages

#### Agriculture

Wages, Salaries and Other Labour Income Agriculture Atlantic

(13.23) 
$$WSAGE = \exp \cdot [7.9691 - 1.2041 \ln (AGYE/AGYPE) [5.46] [4.86]$$

- + 1.0652 ln ((AGYE/AGETE)/(AGY/AGET))
  [5.63]
- + 1.5225 ln (WSAGT/AGET/CPI)]\* AGETE [5.51]

 $\vec{R}^2$  = 0.94 S.E.E. = 0.1070 D.W. = 1.78 (GLS, 1961-1971)

Wages, Salaries and Other Labour Income Agriculture Quebec

(13.35) 
$$WSAGQ = \exp \left[ 8.2623 - 0.8248 \ln \left( \frac{AGYQ}{AGYPQ} \right) \right]$$
 [8.85] [2.83]

- + 0.5171 ln ((AGYQ/AGETQ)/(AGY/AGET))
  [2.62]
- + 1.7187 ln (WSAGT/AGET/CPI)]\* AGETQ [10.24]

 $\overline{R}^2$  = 0.90 S.E.E. = 0.0765 D.W. = 1.29 (GLS, 1961 - 1971)

Wages, Salaries and Other Labour Income Agriculture Ontario

(13.47) 
$$WSAGO = \exp \begin{bmatrix} 7.8396 - 0.1541 \ln (AGYO/AGYPO) \\ [10.53] [0.36] \end{bmatrix}$$

- + 0.3429 ln ((AGYO/AGETO)/(AGY/AGET))
  [1.99]
- + 1.4845 *ln* (WSAGT/AGET/CPI)]\* AGETO [10.49]

```
= 1.04
         D.W.
          (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Agriculture Prairies
           WSAGW = \exp . [7.0642 - 0.3276 \ ln \ (AGYW/AGYPW)]
(13.59)
                         [16.30]
                                    [2.11]
                   + 0.4715 ln ((AGYW/AGETW)/(AGY/AGET))
                     [1.71]
                     1.5211 ln (WSAGT/AGET/CPI)]* AGETW
                     [20.99]
         \overline{R}^2
                  = 0.98
         S.E.E.
                  = 0.0305
                  = 1.42
          D.W.
          (GLS.
                 1961-1971)
Wages, Salaries and Other Labour Income Agriculture British Columbia
          WSAGC = \exp \left[7.4393 - 0.5000 \ln \left(AGYC/AGYPC\right)\right]
(13.7)
                         [5.57]
                                    [1.56]
                   + 0.6674 ((AGYC/AGETC)/(AGY/AGET))
                     [3.16]
                   + 1.4452 ln (WSAGT/AGET/CPI)]* AGETC
                     [5.65]
          \vec{R}^2
                   = 0.70
          S.E.E.
                  = 0.1195
          D.W.
                  = 1.26
         (GLS,
                 1961-1971)
Forestry
```

 $\vec{R}^2$ 

S.E.E.

= 0.91

= 0.0729

Wages, Salaries and Other Labour Income Forestry Atlantic

```
₹²
                   = 0.99
                   = 0.0110
          S.E.E.
                   = 2.40
          D.W.
          (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Forestry Quebec.
(13.36)
           WSFOQ
                             [0.7330 + 0.0454 ln (URATEQ)]
                      exp.
                             [1.72]
                                         [3.15]
                      1.0272 ln ((FOYQ/FOETQ)/(FOY/FOET))
                      [54.76]
                      1.1148 ln (WSFOT/FOET/CPI)
                      [21.13]
                      0.9061 ln (CPI)]* FOETQ
                      [14.45]
         \overline{R}^2
                       0.99
         S.E.E.
                   =
                       0.0082
         D.W.
                       2.59
         (GLS,
                   1961-1971)
Wages, Salaries and Other Labour Income Forestry Ontario.
(13.48)
          WSF00
                              [0.0133 +
                       exp.
                                          0.0119 ln (URATEO)
                              [0.03]
                                          [1.11]
                       1.0102 ln ((FOYO/FOETO)/(FOY/FOET))
                   ÷
                       [66.73]
                       1.0006 ln (WSFOT/FOET/CPI)
                       [20.31]
                       0.9982 ln (CPI)]* FOETO
                       [17.32]
          \bar{R}^2
                   = 0.99
          S.E.E.
                   = 0.0078
          D.W.
                   = 2.73
          (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Forestry Prairies
                       exp. [-0.1839 + 0.0162 ln (URATEW)]
(13.60)
          WSFOW
                               [0.14]
                                         [0.48]
                       0.9062 ((FOYW/FOETW)/(FOY/FOET))
                       [41.22]
                       0.8362 ln (WSFOT/FOET/CPI)
                       [4.95]
```

0.9446 ln (CPI)]\* FOETW

[5.14]

```
\bar{R}^2 = 0.99
           S.E.E. = 0.0241
           D.W. = 1.59
           (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Forestry British Columbia
           WSFOC = \exp \cdot [-0.7218 - 0.0212 \ ln \ (URATEC)]
(13.72)
                          [1.91]
                                    [1.85]
                  + 1.0445 ln ((FOYC/FOETC)/(FOY/FOET))
                    [36.13]
                   + 0.9629 ln (WSFOT/FOET/CPI)
                    [22.89]
                   + 1.1258 ln (CPI)]* FOETC
                     [19.55]
         · ar{R}^2
                  = 0.99
           S.E.E. = 0.0077
           D.W.
                  = 2.88
           (GLS, (1961-1971)
Fishing and Trapping
Wages, Salaries and Other Labour Income Fishing Atlantic
          WSFSE = \exp [0.9093 - 0.0689 \ ln \ (URATEE)]
{13.25}
                          [0.93]
                                   [1.58]
                   + 0.8990 \ ln \ ((FSYE/FSETE)/(FSY/FSET))
                   [12.94]
                   + 0.9894 ln (WSFST/FSET/CPI)
                     [11.73]
                     0.8382 ln
                                 (CPI)]* FSETE
                     [5.92]
           \bar{R}^2 = 0.99
           S.E.E. = 0.0216
           D.W. = 2.64
           (GLS, (1961-1971)
Wages, Salaries and Other Labour Income Fishing Quebec
(13.37)
           WSFSQ = \exp [-6.5291 - 0.0551 \ln (URATEQ)]
                          L0.87 J L0.18 J
                   + 0.4858 ln ((FSYQ/FSETQ)/) FSY/FSET))
                     [2.60]
```

+ 1.0668 *ln* (WSFST/FSET/CPI)

+ 2.4735 ln (CPI)]\* FSETQ

[1.32]

[2.56]

```
\bar{R} = 0.83

S.E.E. = 0.2000
D.W. = 2.32
(GLS, 1961-1971)
```

Wages, Salaries and Other Labour Income Fishing British Columbia

## Mines, Quarries and Oil Wells

(GLS, 1961-1971)

Wages, Salaries and Other Labour Income Mining Atlantic

Wages, Salaries and Other Labour Income Mining Quebec

```
[9.18]
           ₹2
                   = 0.98
           S.E.E. = 0.0279
           D.W. = 1.39
            (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Mining Ontario
                   = \exp \cdot [-1.0499 - 0.0798 \ ln \ (URATEO)]
(13.50)
            WSMIO
                           [2.96]
                                       [2.80]
                   + 0.9331 ln ((MIYO/MIETO)/(MIY/MIET))
                      [14.72]
                   + 1.0057 ln (WSMIT/MIET/CPI)
                      [11.69]
                    + 1.2798 ln (CPI)]* MIETO
                      [18.26]
            \overline{R}^2
                   = 0.97
            S.E.E. = 0.0205
                   = 1.91
            D.W.
            (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Mining Prairies
                   = \exp \cdot [-1.4246 + 0.0455 \ ln \ (URATEW)]
(13.62)
            WSMIW
                                       [1.94]
                           [4.84]
                    + 1.1840 ln ((MIYW/MIETW)/(MIY/MIET))
                      [16.72]
                    + 1.0327 ln (WSMIT/MIET/CPI)
                      [13.62]
                    + 1.2271 ln (CPI)]* MIETW
                      [26.68]
            \overline{R}^{2}
                   = 0.99
            S.E.E. = 0.0147
            D.W. = 2.95
            (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Mining British Columbia
(13.73)
                    = \exp \cdot [-0.3183 + 0.1464 \ ln \ (URATEC)]
            WSMIC
                           [0.61]
                                        [3.39]
                    + 0.9647 ln ((MIYC/MIETC)/(MIY/MIET))
                      [28.14]
```

+ 0.5891 ln (WSMIT/MIET/CPI)

[4.79]

+ 0.9840 ln (CPI)]\* MIETQ

#### Manufacturing

Wages, Salaries and Other Labour Income Manufacturing Atlantic

Wages, Salaries and Other Labour Income Manufacturing Quebec

Wages, Salaries and Other Labour Income Manufacturing Ontario

```
+ 0.9784 ln (CPI)]* MAETO
                     [49.23]
           \bar{R}^2
                   = 0.99
           S.E.E. = 0.0010
           D.W.
                  = 3.17
            (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Manufacturing Prairies
(13.63)
           WSMAW
                   = \exp \cdot [-0.1339 - 0.0081 \ ln \ (URATEW)]
                          [0.21]
                                      [1.70]
                   + 1.0410 ln ((MAYW/MAETW)/(MAY/MAET))
                     [35.62]
                   + 1.0003 ln (WSMAT/MAET/CPI)
                     [11.64]
                   + 1.0251 ln (CPI)]* MAETW
                     [12.37]
           \overline{R}^2
                   = 0.99
           S.E.E. = 0.0037
           D.W. = 2.72
           (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Manufacturing British Columbia
(13.75)
           WSMAC = \exp . [1.1796 + 0.0007 ln (URATEC)]
                          [0.91]
                                    [0.06]
                   + 0.9463 ln ((MAYC/MAETC)/(MAY/MAET))
                     [16.87]
                   + 1.1467 ln (WSMAT/MAET/CPI)
                     [6.64]
                   + 0.8375 ln (CPI)]* MAETC
                     [5.06]
           \overline{R}^2
               = 0.99
           S.E.E. = 0.0087
           D.W.
                  = 1.55
           (GLS, 1961-1971)
Construction
Wages, Salaries and Other Labour Income Construction Atlantic
(13.28)
           WSCOE
                   = exp. [-1.6855 - 0.0118 ln (URATEE)]
                           [1.50]
                                      [0.08]
                   + 0.9494 ln ((COYE/COETE)/(COY/COET))
```

[28.06]

[6.38]

+ 0.7630 ln (WSCOT/COET/CPI)

```
+ 1.2146 ln (CPI)]* COETE
                 [7.4]
        ₹²
               = 0.99
        S.E.E. = 0.0094
        D.W.
              = 1.93
        (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Construction Quebec
(13.40) WSCOQ =
                     exp. [2.0186 +
                                      0.0199 ln (URATEQ)
                          [2.10]
                                      [1.84]
                     0.9753 ln ((COYQ/COETQ)/(COY/COET))
                     [39.26]
                     1.1461 ln (WSCOT/COET/CPI)
                     [11.49]
                + 0.6635 ln (CPI)]* COETQ
                    [4.63]
        \bar{R}^2
                 =
                     0.99
        S.E.E.
                 = 0.0070
        D.W.
                 = 1.18
                 1961-1971)
        (GLS.
Wages, Salaries and Other Labour Income Construction Ontario
           WSCOO = \exp \left[-0.9427 + 0.0123 \right]^{n} (URATEO)
(13.52)
                                      [2.17]
                          [1.29]
                     0.9523 ln ((COYO/COETO)/(COY/COET))
                     [26.18]
                     0.9377 ln (WSCOT/COET/CPI)
                     [12.24]
                     1.1559 ln (CIP)]* COETO
                     [10.74]
           \bar{R}^2 = 0.99
           S.E.E. = 0.0063
                 = 2.54
           D.W.
           (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Construction Prairies
(13.64)
            WSCOW = \exp \left[-2.7103 - 0.0255 \right]^{n} (URATEW)
                          [2.76]
                                      [2.85]
                   + 1.0159 ln ((COYW/COETW)/(COY/COET))
                     [31.27]
                   + 0.7735 ln (WSCOT/COET/CPI)
```

[7.62]

```
1.4314 ln (CPI)]* COETW [9.86]
```

$$\bar{R}^2$$
 = 0.99  
 $S.E.E. = 0.0057$   
 $D.W. = 2.49$   
(GLS, 1961-1971)

Wages, Salaries and Other Labour Income Construction British Columbia

(13.76) 
$$WSCOC = \exp \left[ 3.4743 - 0.0299 \ ln \ (URATEC) \right]$$

- + 1.1448 ln ((COYC/COETC)/(COY/COET))
  [83.29]
- + 1.4168 ln (WSCOT/COET/CPI)
  [18.33]
- + 0.5299 ln (CPI) \* COETC [4.92]

 $\bar{R}^2 = 0.99$  S.E.E. = 0.0066 D.W. = 2.45 (GLS, 1961-1971)

## Transportation, Storage, Communications and Public Utilities

Wages, Salaries and Other Labour Income Transportation Atlantic

(13.29) 
$$WSTSE = \exp \left[0.3369 - 0.0147 \ln (URATEE)\right]$$
[0.50] [0.94]

- + 1.2018 ln ((TSYE/TSETE)/((TSY+UTY)/TSUTET))
  [23.71]
- + 0.8471 ln ((WSTST+WSUTT)/TSUTET/CPI)
  [8.09]
- + 0.8434 ln (CPI)]\* TSETE [10.17]

 $\bar{R}^2$  = 0.99 S.E.E. = 0.0065 D.W. = 2.46 (GLS, 1961-1971)

Wages, Salaries and Other Labour Income Transportation Quebec

(13.41) 
$$WSTSQ = \exp \begin{bmatrix} 0.1266 + 0.0254 \ln (URATEQ) \\ [0.40] [5.22] \end{bmatrix}$$

+ 0.9792 ln ((TSYQ/TSETQ)/(TSY+UTY)/(TSUTET))
[32.53]

```
+ 1.0741 ln ((WSTST+WSUTT)/TSUTET/CPI)
                     [23.87]
                   + 1.0120 ln (CPI)]* TSETE
                     [25.47]
           ₽²
                   = 0.99
           S.E.E. = 0.0027
                  = 1.51
           D.W.
           (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Transportation Ontario
(13.53)
          WSTSO
                  = \exp \cdot [-0.0979 - 0.0096 ln (URATEO)]
                          [0.35]
                                     [3.33]
                   + 0.9724 ln ((TSYO/TSETO)/((TSY+UTY)/TSUTET))
                     [30.75]
                   + 0.9864 ln ((WSTST+WSUTT)/TSUTET/CPI)
                     [25.38]
                   + 1.0216 ln (CPI)]* TSETO
                     [28.39]
         ₹2
                  = 0.99
         S.E.E.
                  = 0.0029
                   = 1.83
         D.W.
         (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Transportation Prairies
(13.65)
          WSTSW
                   = exp. [0.0269 - 0.0104 ln (URATEW)]
                          [0.05]
                                    [1.78]
                   + 0.9573 ln ((TSYW/TSETW)/((TSY+UTY)/TSUTET))
                     [26.78]
                   + 1.0009 ln ((WSTST+WSUTT)/TSUTET/CPI)
                     [13.51]
                   + 0.9891 ln (CPI)]* TSETW
                     [14.95]
          ₹2
                  = 0.99
          S.E.E.
                  = 0.0052
          D.W.
                  = 1.78
                 1961-1971)
          (GLS,
Wages, Salaries and Other Labour Income Transportation British Columbia
(13.77)
          WSTSC
                   = \exp \cdot [-0.0222 - 0.0027 \ ln \ (URATEC)]
                          [0.03]
                                     [0.27]
                   + 1.0425 ln ((TSYC/TSETC)/(TSY+UTY)/TSUTET))
                     [17.37]
                   + 0.9602 ln ((WSTST +WSUTT)/TSUTET/CPI)
```

[8.86]

$$\bar{R}^2$$
 = 0.99  
 $S.E.E.$  = 0.0067  
 $D.W.$  = 0.90  
(GLS, 1961-1971)

#### Trade

Wages, Salaries and Other Labour Income Trade Atlantic

(13.30) WSTRE = exp. [-1.3193 - 0.0278 
$$ln$$
 (URATEE) [1.45] [2.56]   
+ 1.0180  $ln$  ((TRYE/TRETE)/(TRY/TRET)) [26.08]   
+ 0.7806  $ln$  (WSTRT/TRET/CPI) [6.96]   
+ 1.1298  $ln$  (CPI)]\* TRETE [10.02]

$$R^2$$
 = 0.99  
 $S.E.E.$  = 0.0052  
 $D.W.$  = 2.52  
(GLS, 1961-1971)

Wages, Salaries and Other Labour Income Trade Quebec

Wages, Salaries and Other Labour Income Trade Ontario

```
+ 1.0881 ln (WSTRT/TRET/CPI)
[20.79]
```

$$\bar{R}^2 = 0.99$$
 $S.E.E. = 0.0023$ 
 $D.W. = 3.01$ 
 $(GLS, (1961-1971)$ 

### Wages, Salaries and Other Labour Income Trade Prairies

(13.66) 
$$WSTRW = \exp \left[0.5135 - 0.0155 \ln (URATEW)\right]$$
[1.29] [4.86]

$$\overline{R}^2$$
 = 0.99  
 $S.E.E.$  = 0.0026  
 $D.W.$  = 1.91  
(GLS, 1961-1971)

#### Wages, Salaries and Other Labour Income Trade British Columbia

(13.78) 
$$WSTRC = \exp \left[0.4744 + 0.0054 \ln (URATEC)\right]$$
 [0.62] [0.79]

$$\bar{R}^2 = 0.99$$
 $S.E.E. = 0.0044$ 
 $D.W. = 2.59$ 
 $(GLS, 1961-1971)$ 

## Finance, Insurance and Real Estate

Wages, Salaries and Other Labour Income Finance Atlantic

(13.31) 
$$WSFIE = \exp \left[-2.4927 - 0.0027 \ln (URATEE)\right]$$
  
[0.74] [0.05]

```
+ 1.0896 ln ((FIYE/FIETE)/(FIY+HGY)/FIET))
                     [4.85]
                    + 0.5742 ln (WSFIT/FIET/CPI)
                      [1.00]
                    + 1.1935 ln (CPI)]* FIETE
                      [3.23]
            ₹²
                   = 0.99
            S.E.E. = 0.0194
            D.W.
                  = 2.32
            (GLS, 1961-1971)
Wages, Salaries and Other Labour Income Finance Quebec
                      exp. [-1.0663 - 0.0580 \ ln \ (URATEQ)]
(13.43)
          WSFIQ
                           [1.11]
                                      [3.17]
                    + 0.7931 ln ((FIYQ/FIETQ)/((FIY+HGY)/FIET))
                      [10.49]
                    + 1.2793 ln (CPI)]* FIETQ
                      [11.52]
          \overline{R}^2
                    = 0.99
          S.E.E.
                    = 0.0074
          D.W.
                    = 2.48
           (GLS,
                  (1961 - 1971)
Wages, Salaries and Other Labour Income Finance Ontario
                     exp. [0.7444 - 0.0348 ln (URATEO)]
(13.55)
          WSFIO
                          [0.98]
                                   [3.47]
                     1.0792
                             ln ((FIYO/FIETO)/((FIY+HGY)/FIET))
                     [16.91]
                     0.9725
                              ln (WSFIT/FIET/CPI)
                     [8.15]
                     0.8470 ln (CPI)]* FIETO
                     [9.56]
           ₹2
                     0.99
                  =
                     0.0070
           S.E.E. =
           D.W.
                     2.37
                 =
           (GLS, (1961-1971)
Wages, Salaries and Other Labour Income Finance Prairies
(13.67)
           WSFIW
                    exp. [-3.5518 +
                                       0.0620 ln (URATEW)
```

[1.17]

[2.15]

```
+ 0.7792 ln ((FIYW/FIETW)/((FIY+HGY)/FIET))
         [8.43]
       + 0.5514 ln (WSFIT/FIET/CPI)
         [1.17]
       + 1.3975 ln (CPI)]* FIETW
         [3.95]
      = 0.96
S.E.E. = 0.0326
D.W. = 2.24
(GLS, (1961-1971)
```

Wages, Salaries and Other Labour Income Finance British Columbia

### Public Administration and Defence

₹ 2

Wages, Salaries and Other Labour Income Public Administration Atlantic

```
Wages, Salaries and Other Labour Income Public Administration
       Ouebec
                    exp. [-2.4325 + 0.0044 ln (URATEQ)]
(13.44)
          WSADQ =
                          [1.16]
                                      [0.46]
                    0.5811 \ ln \ ((ADYQ/ADETQ)/(ADY/ADET))
                     [17.69]
                    0.6411 ln (WSADT/ADET/CPI)
                    [2.66]
                    1.2856 ln (CPI)]* ADETQ
                     [4.41]
           R^2
                   = 0.99
           S.E.E. = 0.0096
           D.W.
                  = 2.29
           (GLS, 1961.-1971)
Wages, Salaries and Other Labour Income Public Administration
        Ontario
           WSADO =
                     exp. [-1.0556 - 0.0233 \ ln \ (URATEO)]
 (13.56)
                           [0.99]
                                      [2.96]
                     0.8793 ln ((ADYO/ADETO)/(ADY/ADET))
                     [15.06]
                     0.8517 ln (WSADT/ADET/CPI)
                     [6.82]
                     1.1542 ln (CPI)]* ADETO
                     [7.92]
           ₹2
                =
                     0.99
           S.E.E.=
                     0.0047
           D.W. =
                     2.11
           (GLS, (1961-1971
Wages, Salaries and Other Labour Income Public Administration
        Prairies
 (13.68)
           WSADW =
                     exp. [1.5731 + 0.0281 \ln (URATEW)]
                           [0.77]
                                    [2.32]
                     0.8104 ln ((ADYC/ADETC)/(ADY/ADET))
                     [7.96]
                     1.1423 ln (WSADT/ADET/CPI)
                     [4.91]
                     0.7385 ln (CPI)]* ADETW
                     [2.60]
           ₹²
                     0.99
```

S.E.E.=

D.W. =

0.0089

2.86

(GLS, (1961-1971)

# Wages, Salaries and Other Labour Income Public Administration British Columbia

#### Public and Private Services

Wages, Salaries and Other Labour Income Services Atlantic

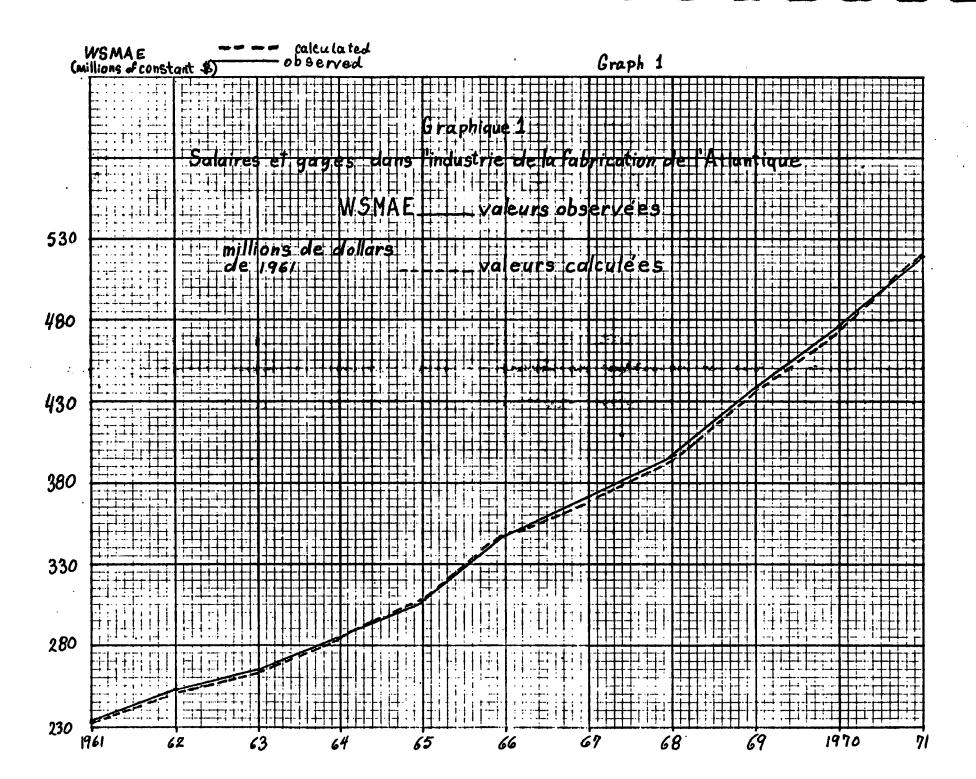
Wages, Salaries and Other Labour Income Services Quebec

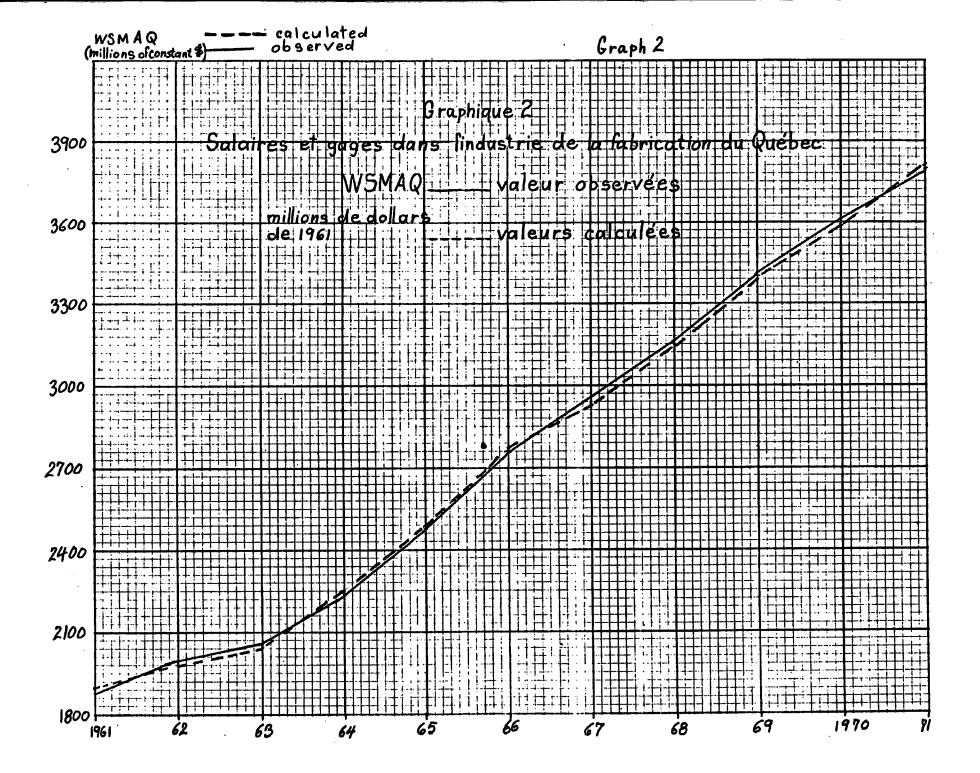
 $\bar{R}^2 = 0.99$  S.E.E. = 0.0036 D.W. = 2.91 (GLS, (1961-1971)

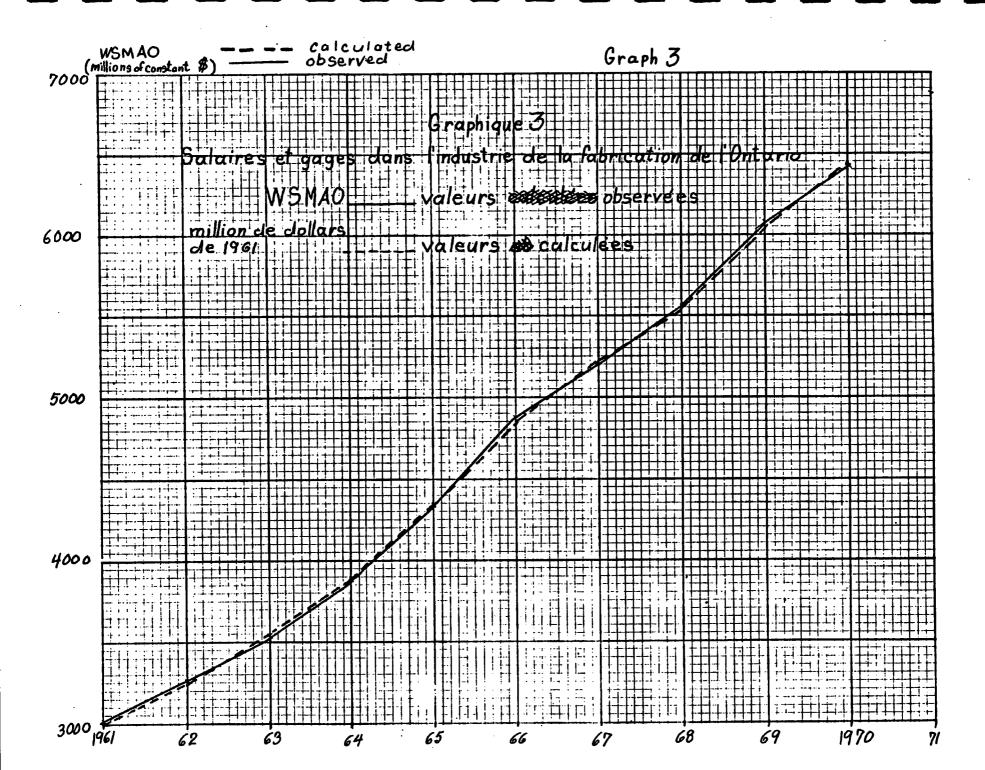
## Wages, Salaries and Other Labour Income Services Ontario

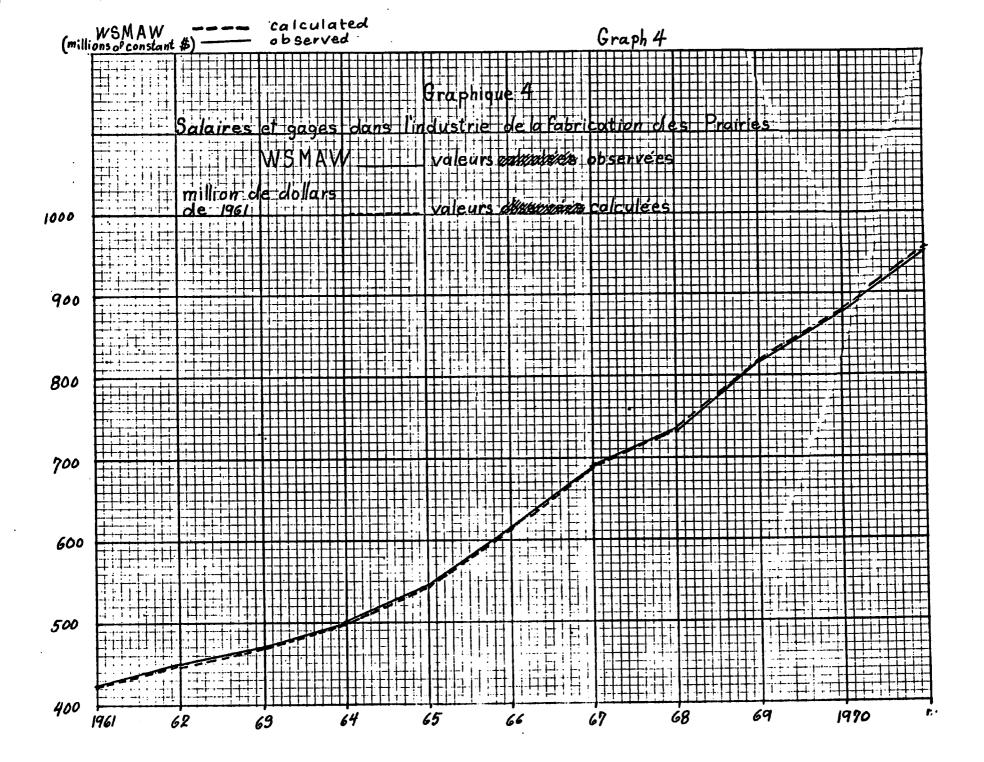
#### Wages, Salaries and Other Labour Income Services Prairies

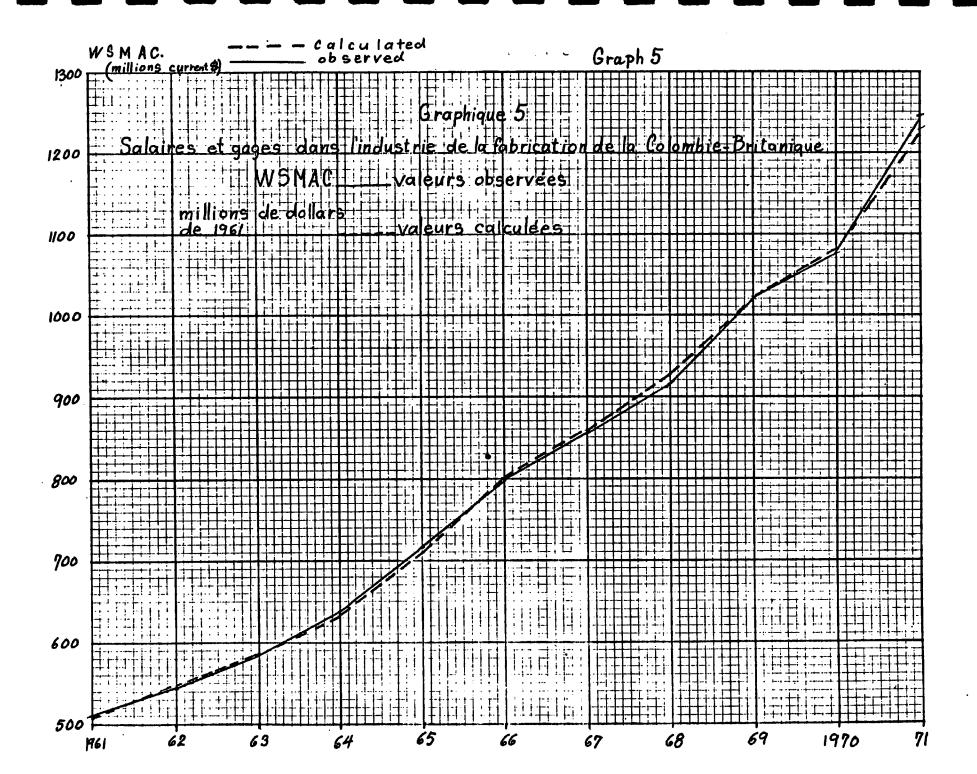
## Wages, Salaries and Other Labour Income Services British Columbia











## 4.0 Conclusion

One of the characteristic features of the regionalization of salaries and wages is the uniformity of specification in our model. We feel that the strength of reasoning here rests precisely on this uniformity which constitutes an important criterion for the implicit acceptance of the hypotheses.

In point of fact the originality of this approach has led to new research in the field of regionalization. One of the more interesting aspects, which may be the subject of future research, is the theory of relative regional salaries and the concept of permanent income. It goes without saying that special attention should be given to the collection of improved regional data.

#### APPENDIX A

#### Estimation Techniques

The difficulties in estimating equation (1) are not related solely to data problems. The data were taken from regional national accounts in the form of annual time series for the period 1961 - 1971. The data for wages and salaries and output were compiled by Statistics Canada and the Department of Regional Economic Expansion respectively. The following eleven industries are included:

Agriculture
Forestry
Fishing and Trapping
Mines, Quarries, and Oil Wells
Manufacturing
Construction
Transportation, Storage, Communications
and Public Utilities
Trade
Finance, Insurance and Real Estate
Public Administration and Defence
Public and Private Services

As we wanted to estimate equation (1) by region and by industry the first difficulty arose from the limited number of observations. The second difficulty arose from the form of the data which aimed to describe chronologically a spatial phenomenon. The most crucial econometric problem concerned the inter-regional dependence of error terms.

As a solution to these problems we have estimated equation (1) in logarithmic form by industry and simultaneously for all regions using Zellner's Generalized Least Squares approach.

Assume:

$$X_j = matrix (1, ln U_{ij}, ln P_{ij}, ln R_{ij})$$

Under these conditions we can rewrite equation (1) in the following form:

$$Y = X_{j}\gamma_{j} + \mu_{j}$$
 (5)

and the system of equations as:

$$\begin{bmatrix} Y_1 \\ Y_2 \\ Y_3 \\ Y_4 \\ Y_5 \end{bmatrix} = \begin{bmatrix} X_1 & 0 & 0 & 0 & 0 \\ 0 & X_2 & 0 & 0 & 0 \\ 0 & 0 & X_3 & 0 & 0 \\ 0 & 0 & 0 & X_4 & 0 \\ 0 & 0 & 0 & 0 & X_5 \end{bmatrix} \begin{bmatrix} Y_1 \\ Y_2 \\ Y_3 \\ Y_4 \\ Y_5 \end{bmatrix} + \begin{bmatrix} \mu_1 \\ \mu_2 \\ \mu_3 \\ \mu_4 \\ \mu_5 \end{bmatrix}$$
(6)

or as:

$$Y = X\gamma + \mu \tag{7}$$

Zellner, A. "An Efficient Method of Estimating Seemingly Unrelated Regressions and Tests for Aggregation Bias", <u>Journal of the American Statistical Association</u>, Vol. 57, June 1962, pp. 348-368.

This technique of combining the equations allows us to assume that they are apparently independent but offers the possibility of taking into account interdependencies of the errors terms, while increasing the precision of estimation. Considering the restricted number of observations per region, in this case, such advantages should not be overlooked.

The assumptions which we make regarding covariance are the following:

assuming j = 
$$\begin{bmatrix} \mu_1 \\ \mu_2 \\ \mu_{tj} \end{bmatrix}$$
 where T is the number of observations per region.

i) 
$$E(\mu_{tj} \mu_{tj}) = 0, V_j$$

ii) 
$$E(\mu_{tj} \mu_{sr}) = 0$$
 
$$V_{jr} = 1, 2, ... 5 \quad (8)$$
iii)  $E(\mu_{tj} \mu_{tr}) = \sigma_{jr}$  or  $E(\mu_{i} \mu_{r}^{1}) = \sigma_{ir}$  I

where

where I is a (TxT) matrix.

The symbol & signifies a Kronecker matrix product.

Applying Generalized Least Squares to equation (7), we obtain:

$$\hat{\gamma} = (x^1 \ \Sigma^{-1} \ x)^{-1} \ x^1 \ \Sigma^{-1} \ y$$
 (10)

as the best linear estimator with the variance - covariance matrix.

$$\operatorname{Var}(\widehat{\gamma}) = (x^{1} \Sigma^{-1} x)^{-1} \tag{11}$$

and where S is used as an estimate of  $\Sigma$ 

$$\begin{bmatrix} e_1^1 \\ e_2^1 \\ \vdots \\ \vdots \\ \vdots \\ e_5^1 \end{bmatrix}$$

$$\begin{bmatrix} e_1 & e_2 & \dots & e_5 \end{bmatrix}$$

where  $e_j = Y_j - X_j$  is calculated in the first stage, through Ordinary Least Squares.

MNEMONIC LIST -- BLOCK 13 -- LISIE DES MNEMONIGUES

```
AULT
                    TOTAL EMPLOYMENT -PUBLIC AUMINISTRATION
       tl 120/3 2
                                                              HRIT. CULUMBIA
ADEIL
       £n 12059 2
                    TUTAL EMPLOYMENT -PUHLIC AUMINISTRATION
AULIE
       fo 12011 3
                    TUTAL EMPLOYMENT -PUBLIC ADMINISTRATION
                                                              ATLANTIC
ADETH
       FP 15080 1
                    TUTAL MANHUURS
                                     -PUBLIC ADMINISTRATION
AULTU
                    TOTAL EMPLOYMENT -PUMLIC ADMINISTRATION
      En 12035 2
                                                              UNTARIU
ADELU
       En 12023 3
                    TOTAL EMPLOYMENT -PUBLIC ADMINISTRATION WHEBEC
AULIA
       to 12047 3
                    TUTAL EMPLOYMENT -PUBLIC AUMINISTRATION PRAIRIES
       E1 13095 1
ADAA
                    WAGES, SALARIES & UINER LAHOUR INC. PUBLIC ADMIN.
       tl 15004 1
AL SHU
                    CUMPENS. PER MANMUUH FORKED, PUB, ADM, ODEF, EX, AF. &CURR.
AD VHOLE
       tl 13005 5
                    INDEX OF AUPHU
4 () Y
                    PUBLIC AUMINISTRATION -REAL DUNESTIC PROD. SMILL-1961
       th 23012 3
                    GRUSS DUMESTIC PRODUCT SCIS PUBLIC ADMIN. BRIT. COLUMBIA
AUYC
       En 50056 3
                    GRUSS DOMESTIC PRODUCT SCIS PUBLIC ADMIN.
ANYE
       En 50010 3
                                                                    AILANTIC
AUYU
                    GRUSS DOMESTIC PRODUCT SCIS PUBLIC ADMIN.
                                                                   ULHATHU
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